## Interleukin-6, C-reactive protein and interleukin-10 after antidepressant treatment in people with depression: a meta-analysis

### S. A. Hiles<sup>1\*</sup>, A. L. Baker<sup>1</sup>, T. de Malmanche<sup>2</sup> and J. Attia<sup>3,4</sup>

<sup>1</sup> Centre for Brain and Mental Health Research, Faculty of Health, University of Newcastle, New South Wales, Australia

<sup>2</sup> Immunology, Hunter Area Pathology Service, John Hunter Hospital, New South Wales, Australia

<sup>8</sup> Centre for Clinical Epidemiology and Biostatistics, Faculty of Health, University of Newcastle, New South Wales, Australia

<sup>4</sup> Hunter Medical Research Institute, John Hunter Hospital, New South Wales, Australia

**Background.** Cross-sectional studies support an association between depression and inflammatory markers. However, little is known of their relationship in the context of antidepressant treatment. Our aim was to explore via meta-analysis whether antidepressant treatment is associated with a reduction in three inflammatory markers associated with depression.

**Method.** A computerized search of EMBASE, Medline, PsycINFO and Cochrane Library databases was completed using subject headings for depression and either interleukin-6, C-reactive protein or interleukin-10, selecting studies which reported circulating levels of inflammatory markers before and after antidepressant treatment for people with depression. Outcome and moderator variables were coded for analysis, including inflammatory marker change, depression severity change, age, gender ratio, assay brand, treatment response and weight change.

**Results.** Pooled effect sizes showed a significant decrease in interleukin-6 (n = 14, d = -0.42, p = 0.02), marginally significant decrease in C-reactive protein (n = 8, d = -0.57, p = 0.05) and a non-significant decrease in interleukin-10 (n = 3, d = -0.45, p = 0.14) after treatment. High levels of heterogeneity were observed, which may be associated with clinical variations between the studies such as weight gain, anxiety, incomplete remission and other individual differences and co-morbidities.

**Conclusions.** The findings of this meta-analysis indicate that there may be a normalization of overactive inflammatory processes following antidepressant treatment.

Received 25 August 2011; Revised 9 January 2012; Accepted 17 January 2012; First published online 16 February 2012

Key words: C-reactive protein, depression, inflammation, interleukin, meta-analysis, systematic review.

### Introduction

The co-morbidity of depression and inflammationrelated physical illnesses, particularly cardiovascular disease, has raised the possibility of a shared, underlying inflammatory pathway involving central and systemic responses (Irwin & Miller, 2007). Proposed bi-directional mechanisms for how inflammatory processes may induce depressive mood, and vice versa, incorporate evidence that prolonged exposure to inflammatory mediators can impair the regulation of neuroendocrine stress, influence the availability of monoamine neurotransmitters, and decrease

\* Address for correspondence : Ms. S. A. Hiles, Centre for Brain and Mental Health Research, University of Newcastle, University Drive, Callaghan, NSW, 2308, Australia.

(Email: sarah.hiles@uon.edu.au)

neurogenesis and neurotropic support (Maes *et al.* 2009; Miller *et al.* 2009).

Much of the evidence supporting inflammatory theories of depression is cross-sectional with many studies demonstrating that compared to people without depressive symptoms, people with high depressive symptoms show elevated levels of inflammatory markers (including cytokines, chemokines and acutephase proteins) in peripheral serum and cerebrospinal fluid (Zorrilla et al. 2001; Lindqvist et al. 2009; Dowlati et al. 2010). These differences have been observed in both medically healthy people with depressive symptoms and those with co-occurring medical comorbidities such as cardiovascular disease, renal disease and cancer (Howren et al. 2009; Bossola et al. 2010). Fewer studies examine the association between depression and inflammatory markers over time, although this can help develop the case for whether

<sup>(</sup>Energinal) (1000) (2000) (10000000)

elevated levels of inflammatory markers may be a cause or simply a consequence of depression.

This meta-analysis is one of the first to estimate change in inflammatory markers before and after antidepressant treatment. Through narrative review and meta-analytical moderator analysis, we also explore sources of heterogeneity related to individual differences and co-morbidity. The cytokine interleukin (IL)-6 and an acute-phase protein, C-reactive protein (CRP), were selected as prototypical pro-inflammatory markers, robustly associated with systemic inflammatory response, and IL-10 was selected as an important anti-inflammatory marker. These are repeatedly shown to be elevated in people with depression, compared to people without depression (Howren *et al.* 2009; Dowlati et al. 2010) and assuming there is an association, we would expect these to decrease as depressive symptoms decrease in association with antidepressant treatment.

### Method

### Systematic search

Studies were included in the meta-analyses if they met the following criteria: (i) participants were adults either diagnosed with major depression/dysthymia or endorsing high depressive symptoms on a standardized inventory; (ii) patients were not undergoing cytokine treatment (e.g. interferon or IL-2); (iii) the study explicitly reported antidepressant treatment in the methods; (iv) mean or median resting levels of IL-6, CRP or IL-10 in circulating plasma or serum were reported before and at least once after starting antidepressant treatment; (v) either a pre-test/post-test design or a randomized controlled trial (RCT) design was used; (vi) publication was in English in a peerreviewed journal; and (vii) enough information to calculate an effect size was reported. Studies of people with depression and a co-morbid general medical condition were included providing the methods stated that the disease and medications were stable for the duration of the study.

A computerized search of EMBASE, Medline, PsycINFO and Cochrane Library databases was completed in March 2011 using two different search strategies: (i) the key terms mapped to subject headings for depression (depression, depressive disorder, major depressive disorder, dysthymic disorder) and either 'interleukin-6', 'C-reactive protein' or 'interleukin-10'; and (ii) depression subject headings and 'cytokine/interleukin' or 'acute phase protein' and 'antidepressants'. Both searches were limited to human and English-language literature. The abstract of each article identified in the search was screened for relevance. If the abstracts mentioned antidepressant use and measurement of inflammatory markers, or the abstracts lacked detail, full text articles were extracted and compared against selection criteria for eligibility. The reference lists of included studies and relevant review articles were screened for additional articles.

Eligible studies were coded and blindly checked by one author (S.H.) for the sample size and the outcome variable of mean (with standard deviation/standard error) circulating IL-6, CRP or IL-10 before and after treatment. Means and standard errors of inflammatory markers reported only graphically were converted to numerical values using Data Thief III, version 1.5 (Tummers *et al.* 2008). Moderator and mediator variables were also coded, including: diagnosis, mean group depressive symptom rating at baseline and follow-up, mean age, proportion of males, treatment duration, inpatient or outpatient status, proportion of treatment responders and type of antidepressant. Where not reported, these variables were coded as missing values.

### Statistical methods

Statistical analyses were completed using Comprehensive Meta-Analysis II (Biostat Inc., USA). For the antidepressant treatment arm in each study, individual study effect sizes for change in inflammatory markers and depressive symptoms were calculated as repeated-measures Cohen's d standardized mean difference using the pre- and post-test means, standard deviation of the difference calculated from the pre- and post-test standard deviations and an estimation of the pre- and post-test correlation (Borenstein et al. 2009). If the mean was not reported, the median was used as an estimate because sample sizes were sufficiently large (Pudar Hozo et al. 2005). We used a conservative correlation of 0.5 as an estimate for the correlation between pre- and post-treatment inflammatory markers but also re-tested the meta-analysis model using other correlations (r = 0.4, 0.6, 0.7), with none producing a large difference as a result. For one study which failed to report standard deviations/ errors but otherwise met selection criteria (Dawood et al. 2007), the effect size was calculated using the difference in means, sample size and paired *p* value. For the few studies that included more than one follow-up measurement of IL-6, CRP or IL-10 (Kubera et al. 2000; Hernandez et al. 2008; Mackay et al. 2009), we calculated a study effect size for the follow-up point closest to the mean duration of the remaining studies for that inflammatory marker, in order to reduce the variability that too large a range of treatment duration might cause. For studies reporting data for subgroups only, whole sample data points were



Fig. 1. Summary of article extraction process after two systematic search strategies either based on search for 'depression' and either the specific mediator interleukin (IL)-6, C-reactive protein (CRP) or IL-10, or 'antidepressants', 'cytokine/interleukin' and 'acute phase protein'. Reasons for exclusion are listed.

imputed by collapsing subgroup means (Sluzewska *et al.* 1995; Leo *et al.* 2006; Yoshimura *et al.* 2009). A positive effect size indicates that there was an increase in the inflammatory marker over time.

Individual study effect sizes were synthesized to generate an overall effect size using a random-effects model, weighted by the inverse of variance. We also completed a sensitivity analysis to identify potential outliers by removing each study one by one to examine the influence of each individual study on the overall effect size. Egger's test of funnel plot asymmetry (Egger et al. 1997) and failsafe N (Rosenthal, 1979) were calculated to assess publication bias. Heterogeneity in the meta-analysis was assessed using Cochrane's Q and  $I^2$  which calculates a proportion of variation attributed to heterogeneity (Higgins et al. 2003). Moderator analysis was undertaken to explore sources of heterogeneity. Subgroup analysis was completed by comparing pooled inflammatory marker effect size in subgroups of categorical moderator/mediator variables, while method-of-moments meta-regression was completed to explore the relationship between effect size and continuous variables.

## Results

There were 22 studies relevant to these meta-analyses (n = 14 for IL-6, n = 8 for CRP, n = 3 for IL-10; methodsand relevant findings are summarized in online Supplementary Table S1, article extraction process is summarized in Fig. 1). Most study designs were of a single group of people with depression, measured before and after (or during) antidepressant treatment (n = 15). Seven studies were RCTs comparing an antidepressant treatment condition and either a nonantidepressant treatment or antidepressant in conjunction with non-antidepressant pharmacological treatment. Many studies included a control condition of people without depression who were measured for inflammatory markers at a single time point (n = 17) to infer whether levels of inflammatory markers in the depression group changed to a level comparable to a non-depressed control group. Every study reporting depressive symptoms showed a significant group reduction in depressive symptoms after antidepressant treatment.

Most of the included studies had similar exclusion criteria: excluding people with DSM-IV Axis I

disorders besides major depressive episodes or dysthymic disorder (n=12; except Sluzewska *et al.* 1997, who additionally included people with a lifetime diagnosis of bipolar disorder in a current depressive episode), people on psychotropic medication in the 2 weeks before blood collection (n=22), people with general medical illness (n=15), people with recent allergic reactions (n=7) and pregnant women (n=7). The study where patients had stable coronary heart disease were not acutely ill, and had been stable on medication (Pizzi *et al.* 2009). Most studies collected blood in the morning.

## IL-6 and CRP

Although in two studies administration of the selective serotonin reuptake inhibitor (SSRI) fluoxetine resulted in no significant change in IL-6 (Jazayeri *et al.* 2010) or CRP over multiple time-points (Mackay *et al.* 2009), most of the remaining studies that administered SSRIs reported significant reductions in IL-6 or CRP, including other studies of fluoxetine (significant reduction only observed in people with elevated IL-6 at baseline; Sluzewska *et al.* 1995), sertraline (Leo *et al.* 2006; Pizzi *et al.* 2009) and citalopram (Leo *et al.* 2006).

Few studies examined non-SSRI antidepressants. No significant changes in IL-6 were observed in studies administering a serotonin-norepinephrine reuptake inhibitor (duloxetine; Fornaro et al. 2011) or a serotonin antagonist and reuptake inhibitor (trazodone) either alone or augmented with pindodol (Maes et al. 1997). Significant reductions in CRP were observed after administration of the tricyclic antidepressant, amitriptyline, in people with and without 50% reductions in depressive symptoms (treatment responders non-responders, and respectively; Lanquillon et al. 2000). People with higher levels of CRP after treatment also recorded significantly higher levels of trait anxiety than people with lower levels of CRP. Finally, Chen et al. (2010) reported a significant increase in IL-6, but no change in CRP, in young, normal-weight males who were maintained on a strict diet with restricted alcohol and coffee intake while in an inpatient facility. While participants were randomized to four different SSRI and non-SSRI antidepressant conditions, the authors only reported summary statistics for the entire study sample, as the increase in IL-6 was comparable across conditions. The authors reported an increase in weight in two of the four conditions which while not statistically significant, may still be clinically meaningful in contributing to elevated IL-6.

In the remaining studies, a clinician decided caseby-case which antidepressant was administered. In the two studies where clinicians were unrestricted in

their choice of antidepressant, neither demonstrated a significant change in IL-6 (Kubera et al. 2000; Kagaya et al. 2001). No significant change in IL-6 was observed in two studies where participants were administered either a tricyclic antidepressant or a particular SSRI (fluoxetine: Maes et al. 1995; paroxetine: Mikova et al. 2001). However, significant decreases were observed in IL-6 and CRP after administration of various other antidepressants from SSRI classes (Tuglu et al. 2003; Basterzi et al. 2005; O'Brien et al. 2006). In contrast, Dawood et al. (2007) demonstrated a significant increase in CRP, with only 5/24 patients recording a decrease in CRP after treatment. Two studies that selected from SSRI and serotonin-norepinephrine reuptake inhibitors also reported significant reductions in IL-6 (Yao et al. 2004; Yoshimura et al. 2009), but not for treatment non-responders (Yoshimura et al. 2009). Sluzewska et al. (1997) also reported a significant reduction in CRP for both treatment responders and non-responders in participants with a refractory major depressive episode receiving various antidepressants and lithium carbonate. Neither change in depressive symptoms nor change in CRP differed significantly between people with lifetime bipolar disorder or major depressive disorder admitted to the study during a depressive episode.

## IL-10

No significant changes in IL-10 were observed after administration of fluoxetine, although levels were initially lower than controls (Song *et al.* 2009). The remaining studies reported that initially elevated levels of IL-10 reduced over treatment to a level below the controls at baseline (Kubera *et al.* 2000; Hernandez *et al.* 2008). Additionally, Hernandez *et al.* (2008) reported that while IL-10 measurements at 5, 20, 36 and 52 weeks showed a decreasing trend, participants reached clinical remission for depression by 20 weeks.

# Antidepressants v. non-antidepressant treatment conditions

There were only two RCTs with placebo treatment arms. Pizzi *et al.* (2009) found no change in IL-6 or CRP after 20 weeks of placebo treatment. Furthermore, treatment with sertraline led to significant reductions in IL-6 and CRP compared to placebo at follow-up (between-subjects IL-6: d = -0.74, 95% CI -1.16 to -0.32; CRP: d = -1.00, 95% CI -1.43 to -0.56). Song *et al.* (2009) also showed no significant decline in IL-10 after placebo treatment; however, there was no significant difference in IL-10 between placebo and active treatment groups at follow-up, as IL-10 did not



**Fig. 2.** Forest plot for change in interleukin (IL)-6 after antidepressant treatment including study name identifier, total number of participants, the standardized paired difference in IL-6 (d, 95% CI) and the relative weight that each study contributes to the overall pooled estimate of effect. The diamond at the bottom of the effect size plot represents the overall pooled effect size for standardized change in IL-6 (d, 95% CI) and the  $I^2$  measure of heterogeneity. Negative effect sizes represent a decrease in IL-6 following antidepressant treatment.

significantly decline in the active treatment groups either (fluoxetine and sham electroacupuncture or antidepression sequence of electroacupuncture and placebo capsules). The null result may possibly be due to non-specific treatment effects of sham electroacupuncture. In both studies, there was no significant decline in depressive symptoms in the placebo treatments, but significant declines in the active treatments.

Two studies randomized participants to antidepressants and other antidepressive treatments. Neither study showed that IL-6 or CRP significantly declined within the active treatment groups over time nor significantly differed between the active treatment groups at follow-up (IL-6 after treatment with fluoxetine, eicosapentaenoic acid, and fluoxetine and eicosapentaenoic acid combined for Jazayeri *et al.* 2010; and CRP after treatment with fluoxetine with tri-iodothyronine and counselling groups for Mackay *et al.* 2009). One final study treated a depressed and non-depressed control group with duloxetine, and found that IL-6 did not change in either group, although depressive symptoms declined in the depressed but not the control group (Fornaro *et al.* 2011).

### Meta-analysis

First, we examined whether depression significantly declined over the course of antidepressant treatment. Several studies were excluded from this analysis as they did not report sufficient information (Sluzewska *et al.* 1995, 1997; Mikova *et al.* 2001; O'Brien *et al.* 2006; Yoshimura *et al.* 2009; Jazayeri *et al.* 2010). Meta-analysis of the depression severity scores (chiefly the Hamilton Rating Scale for Depression) revealed a significant decrease in the studies which investigated IL-6 (n=10, d=-1.82, 95% CI -2.36 to -1.28,  $l^2$ =85.4), CRP (n=7, d=-1.93, 95% CI -2.59 to -1.27,  $l^2$ =90.1%) and IL-10 (n=3, d=-4.77, 95% CI -8.13 to -1.41,  $l^2$ =92.0%).

Second, we analysed changes in inflammatory markers after antidepressant treatment. Meta-analysis revealed a significant decrease in IL-6 (n=14, d=-0.42, 95% CI -0.78 to -0.06, Z=2.30, p=0.02;



Fig. 3. Forest plot for change in CRP after antidepressant treatment, details as for Fig. 2.



Fig. 4. Forest plot for change in IL-10 after antidepressant treatment, details as for Fig. 2.

Fig. 2). There was a marginally significant decrease in CRP (n=8, d=-0.57, 95% CI -1.140 to 0.005, Z=1.94, p=0.052; Fig. 3). In both instances, there was high heterogeneity (IL-6:  $l^2=88.2\%$ ,  $Q_{13}=110.23$ , p<0.001; CRP:  $l^2=93.2\%$ ,  $Q_7=103.66$ , p<0.001). There was no evidence of publication bias via Egger's test (IL-6:  $t_{13}=0.42$ , p=0.68; CRP:  $t_7=1.53$ , p=0.24).

The failsafe *N* was 148 for IL-6, which means 148 extra studies would be required for the *p* value to be > 0.05. Failsafe *N* was 82 for CRP.

Based on the very few studies and a failsafe N calculation of 5, the meta-analysis for IL-10 should be considered exploratory and considered a description of currently available data, rather than a

Subgroup	k	Effect size ( <i>d</i> )	95 % CI	I <sup>2</sup> (%)
Overall	14	-0.42*	-0.78 to $-0.06$	88.2
Formal diagnosis				
Yes	13	$-0.41^{*}$	-0.79 to $-0.04$	87.1
No	1	$-0.90^{***}$	-1.24 to $-0.56$	
Treatment type				
Clinical decision	8	-0.50*	-0.96 to $-0.04$	85.9
Uniform administration	6	-0.32	-0.91 to 0.27	90.8
Patient type				
Inpatient	7	-0.32	-0.93 to 0.29	91.5
Outpatient	7	$-0.53^{**}$	-0.90 to $-0.16$	78.3
CRP				
Overall	8	-0.57	-1.14 to 0.01	93.2
Formal diagnosis				
Yes	7	-0.52	-1.56 to 0.12	93.3
No	1	-0.92***	-1.26 to $-0.58$	
Treatment type				
Clinical decision	2	-0.66	-1.38 to 0.06	78.2
Uniform administration	5	-0.78	-1.58 to 0.01	94.7
Patient type				

 $-0.93^{*}$ 

-0.53

**Table 1.** Cohen's d (95% CI) pooled effect sizes under a random-effects model for the change in CRP or IL-6 following antidepressant treatment for subgroups of studies within the overall meta-analysis, with k number of studies and  $l^2$  measure of heterogeneity

CRP, C-reactive protein; IL-6, interleukin-6.

\* p < 0.05, \*\* p < 0.01, \*\*\* p < 0.001 (based on a single study).

4

true representation of the effect. There was a nonsignificant decrease in IL-10 following antidepressant treatment (n=3, d=-0.45, 95% CI -1.03 to 0.14, Z=1.49, p=0.14; Fig. 4). Significantly high heterogeneity was also observed ( $I^2=77.3\%$ ,  $Q_2=8.81$ , p=0.01). There was no significant publication bias according to Egger's test although this test lacks power when there are few studies ( $t_2=0.76$ , p=0.59). Sensitivity analyses revealed no extreme influence of any single study in the IL-6, CRP or IL-10 analyses.

#### Moderator analysis

Inpatient

Outpatient

Moderator analysis was completed on the IL-6 and CRP studies (no analysis was undertaken for IL-10 because there were too few studies). Table 1 shows the results of random-effects analysis on smaller subgroups of the included studies (whether or not a formal diagnosis was completed, whether it was a uniform or tailored antidepressant administration, whether participants were inpatients or outpatients, and IL-6 assay brand). Only the subgroup of outpatients for IL-6 resulted in a strongly significant pooled effect size (p < 0.01), recording a larger decrease in effect size than in the overall meta-analysis.

There was no substantial reduction in  $I^2$  in each subgroup, suggesting that none of these factors alone were responsible for the heterogeneity.

Meta-regression compared the standardized changes in IL-6 or CRP and continuous variables of standardized change in depression severity score, weeks in treatment, age, percentage males, and percentage of responders, for studies in which these variables were reported. Meta-regression revealed significant, positive associations between standardized change in IL-6 and percentage of males (n = 17, n) $\beta = -0.02$ , s.e. = 0.006, b = -1.12, p = 0.03). No other associations were significant (p > 0.05). Finally, metaregression of the standardized change in depressive symptoms against the baseline levels of IL-6 and CRP showed a negative, although non-significant, relationship (IL-6: n = 10,  $\beta = -0.14$ , s.e. = 0.15, b = -1.44, p = 0.39; CRP: n = 6,  $\beta = -0.26$ , s.e. = 0.25, b = -1.58, p = 0.35).

### Discussion

-1.85 to -0.003

-1.44 to 0.38

94.3

93.1

This study is the first to provide pooled estimates of the change in resting, circulating IL-6, CRP and IL-10 in people with depression after antidepressant treatment. Meta-analysis indicated that after various antidepressant treatment durations there was a significant pooled reduction in IL-6, marginally significant reduction in CRP, and non-significant decrease in IL-10. There was good support for the reliability of the pooled reduction in IL-6 and CRP, with no evidence of publication bias and a high failsafe N, meaning it would take many unpublished, or yet to be published, studies to negate this pooled estimate. The few studies and low failsafe N for the IL-10 meta-analysis means that this observed pooled reduction should be interpreted with caution and considered preliminary. There was high heterogeneity in all the pooled estimates which could not be explained by metaregression and subgroup analysis. Narrative review identified several issues which may help clarify unexplained sources of heterogeneity regarding comorbidity and individual differences.

In cross-sectional studies, levels of IL-6, CRP and IL-10 are elevated in people with high compared to low depressive symptoms (Howren et al. 2009; Dowlati et al. 2010). It is fitting that these inflammatory markers decrease as depressive symptoms decrease in association with antidepressant treatment. There may be a difference in the sensitivity of CRP and IL-6 for recording a change after treatment, with some studies demonstrating a change in IL-6 but not CRP (Chen et al. 2010) and the pooled decrease only marginally significant for CRP. Nevertheless, the association between antidepressant treatment and reduction in inflammatory markers was demonstrated with different antidepressants, in a depressive episode irrespective of lifetime diagnoses of major depressive disorder or bipolar disorder (Sluzewska et al. 1997), and to a limited extent, over time (Kubera et al. 2000; Hernandez et al. 2008; Mackay et al. 2009).

IL-6, CRP and IL-10 play a regulatory role in the acute phase of inflammation, with IL-6 and CRP being primarily pro-inflammatory and IL-10 being inhibitory (Gabay & Kushner, 1999; Moore et al. 2001), yet the role of each of these markers in depression is unclear. Indeed, there may be a problem in balancing pro- and anti-inflammatory agents, and the relative concentrations of the pro- and anti-inflammatory markers may be more important than the absolute levels. At a broad level, elevations in pro- and antiinflammatory markers during depressive illness may represent a generalized over-activation of the inflammatory system during the acute emotional state, which normalizes following alleviation of depressive symptoms. This is supported in studies which showed no significant difference between levels of inflammatory markers in the depression group post-treatment and in the healthy control group (Yao et al. 2004). However,

the relationship between depression and inflammatory markers over time is not necessarily linear. For example, the longest study with multiple measurements demonstrated constant reductions in IL-10 at weeks 5, 20, 36 and 52, although depressive symptoms remained relatively consistent from week 20 (Hernandez *et al.* 2008). These results require replication due to the methodological limitations of this study, including that only 35% of participants were retained by 52 weeks.

Although no outliers were detected and there was no evidence of publication bias for the IL-6 and CRP meta-analyses, there are limitations to the interpretation of these pooled effect sizes. First, it is likely that the magnitude of the decrease in these inflammatory markers is slightly inflated due to regression to the mean (Bland & Altman, 1994). This problem is characteristic of single-group pre-test/post-test designs, highlighting the need for further placebocontrolled RCTs. The available placebo-capsule RCT demonstrated that sertraline use was superior to placebo at reducing inflammatory markers (Pizzi et al. 2009). Furthermore, the high levels of heterogeneity observed in each meta-analysis make interpretation of the overall pooled estimates difficult. This level of heterogeneity was expected because liberal inclusion criteria were employed to maximally canvass the literature. The liberal inclusion criteria are both a weakness and strength of this meta-analysis. While studies of low methodological quality and studies with diverse samples were included, it has also identified potential sources of heterogeneity to be considered in future research.

The main drivers of the high heterogeneity in the IL-6 and CRP data were not identified using subgroup analysis or meta-regression. This implies that there is probably a cumulative effect of patient characteristics on the degree of change in pro-inflammatory markers and factors not reported in the primary studies may more precisely account for unexplained heterogeneity. For instance, many studies lacked potentially useful prognostic information, such as compliance information, to consider whether declines are associated with the biochemical effects of antidepressants compared to decline in depressive symptoms over time. In the subgroup analyses, the only highly significant effect size was for outpatients in IL-6, which only slightly decreased the level of heterogeneity and generated a slightly larger pooled decrease in IL-6. The only significant positive linear association in the metaregression analyses was that having fewer males in a study was associated with a larger standardized decrease in IL-6. The reason for this is unclear but may be because females tend to have higher response rates to particular antidepressant treatments (particularly SSRIs) in clinical trials (Khan et al. 2005).

Narrative review identified several other potential sources of heterogeneity in co-morbidity and individual differences. Changes in weight variables were rarely reported despite the co-morbidity between depression and obesity, and the relationship between antidepressant use and weight gain (Evans et al. 2005; Serretti & Mandelli, 2010). Studies which recorded an increase in IL-6 indicated that there were increases in fat distribution (Chen et al. 2010), and adiposity is significantly associated with inflammatory changes, particularly increases in IL-6 (Park et al. 2005). Furthermore, people with higher levels of CRP after treatment had significantly higher levels of trait anxiety than people with lower levels of CRP (Lanquillon et al. 2000). This finding may explain continued elevations in CRP or IL-6 in other studies which often fail to measure anxiety, despite its high comorbidity with depression (Rush et al. 2005) and independent association with inflammatory markers (O'Donovan et al. 2010).

Major depressive disorder is a clinically heterogeneous disorder and these differences may extend to differences in inflammation. Sluzewska et al. (1995) highlighted that only certain people with depression may exhibit elevated levels of IL-6, and only these people decreased their levels of IL-6. Disregarding issues of regression to the mean, this study highlights that certain people with depression may be more susceptible to elevations in pro-inflammatory markers. For instance, some research suggests that people with depression with melancholic features may exhibit a different inflammatory profile to those without melancholic features (Rothermundt et al. 2001). Depression with melancholic features is associated with a decrease in CRP in serum following treatment (O'Brien et al. 2006) and there are indications that levels of stimulated cytokines decline in melancholic but not other depression (Rothermundt et al. 2001), implying that perhaps the inflammatory state is more closely associated with organic rather than cognitive symptoms of depression.

Furthermore, there is a possibility that baseline levels of inflammatory markers may identify those who may respond to treatment. Meta-regression in the current review showed no significant association between baseline IL-6 or CRP and change in depressive symptoms, although the pattern across studies was that higher baseline IL-6 and CRP were related to larger decreases in depressive symptoms. In individual studies, there is evidence in support of this, with higher baseline levels of CRP in treatment responders (Sluzewska *et al.* 1997), and evidence to the contrary that lower stimulated levels of IL-6 at baseline (but not baseline serum CRP) in treatment responders compared to non-responders (Lanquillon *et al.* 2000). Additionally, meta-regression showed no significant relationship between percentage of individuals who responded to treatment and inflammatory marker change, perhaps because it was infrequently reported and thus the meta-regression was restricted in range. At the individual study level, there was evidence of decreases in IL-6 for treatment responders, but not treatment non-responders (Yoshimura et al. 2009), although two studies demonstrated a significant decrease in inflammatory markers in both responders and non-responders (people with and without a 50% reduction in depression, respectively; Sluzewska et al. 1997; Lanquillon et al. 2000). It is possible that even with small decreases in depressive symptoms, substantial changes to inflammatory markers may occur. Future studies should verify whether differences exist in inflammatory marker change and baseline inflammatory markers for treatment responders and nonresponders, to support whether improvement in depressive symptoms is associated with a normalization of inflammatory markers and provide evidence for whether inflammatory markers can act as a biomarker of treatment response.

On the whole, the evidence presented in this metaanalysis is consistent with the inflammation theory of depression; with a reduction in depressive symptoms, there is a co-occurring reduction in inflammatory markers. At an illustrative level, these data support the idea that the causal chain is 'depression driving inflammation', because treatment for depression also has the capacity to change inflammatory markers. Yet at the same time, antidepressants may have a direct anti-inflammatory effect, thus potentially causing the reductions in depressive mood. In vitro studies demonstrate that administration of antidepressants, particularly SSRIs, produces anti-inflammatory effects in the blood of both people with depression and healthy volunteers, decreasing pro-inflammatory markers including IL-6, IL-8 and tumour necrosis factor-alpha and increasing anti-inflammatory markers including IL-10 (Kenis & Maes, 2002; Janssen et al. 2010). The observed anti-inflammatory effects may occur through antidepressants increasing glucocorticoid receptormediated negative feedback of the hypothalamicpituitary-adrenal axis or increasing intracellular cyclic adenosine monophosphate (for reviews see Maes, 2001; Carvalho & Pariante, 2008; Carvalho et al. 2010; Janssen et al. 2010). Alternatively, studies into the effect of anti-inflammatory medications on depressive symptoms would provide evidence for the 'inflammation driving depression' causal chain, and if supporting the idea of an underlying common cause would be persuasive. To further explore the directionality, more longitudinal and prospective measurement of depression and inflammatory 2024 S. A. Hiles et al.

markers are necessary. It would also be of benefit to investigate changes in other immunomarkers after antidepressant treatment to provide the context of the changes in the few inflammatory markers reported in this study, as cross-sectional evidence suggests that depression is associated with many markers of cellmediated immune activation (Zorrilla *et al.* 2001).

## Note

Supplementary material accompanies this paper on the Journal's website (http://journals.cambridge.org/psm).

### Acknowledgements

Sarah Hiles is supported by an Australian Postgraduate Award and the Vice Chancellor's Award for Outstanding Candidates. Amanda Baker is supported by an Australian National Health and Medical Research Council Senior Research Fellowship.

## **Declaration of Interest**

None.

### References

Basterzi AD, Aydemir C, Kisa C, Aksaray S, Tuzer V, Yazici K, Goka E (2005). IL-6 levels decrease with SSRI treatment in patients with major depression. *Human Psychopharmacology* 20, 473–476.

Bland JM, Altman DG (1994). Statistics notes: some examples of regression towards the mean. *British Medical Journal* **309**, 780.

Borenstein M, Hedges LV, Higgins JPT, Rothstein HR (2009). Introduction to Meta-analysis. John Wiley & Sons: Chichester, UK.

Bossola M, Ciciarelli C, Di Stasio E, Conte GL, Vulpio C, Luciani G, Tazza L (2010). Correlates of symptoms of depression and anxiety in chronic hemodialysis patients. *General Hospital Psychiatry* **32**, 125–131.

Carvalho LA, Garner BA, Dew T, Fazakerley H, Pariante CM (2010). Antidepressants, but not antipsychotics, modulate GR function in human whole blood : an insight into molecular mechanisms. *European Neuropsychopharmacology* **20**, 379–387.

**Carvalho LA, Pariante CM** (2008). In vitro modulation of the glucocorticoid receptor by antidepressants. *Stress: The International Journal on the Biology of Stress* **11**, 411–424.

Chen Y-C, Lin W-W, Chen Y-J, Mao W-C, Hung Y-J (2010). Antidepressant effects on insulin sensitivity and proinflammatory cytokines in the depressed males. *Mediators of Inflammation*. Published online: 18 May 2010. doi:10.1155/2010/573594.

Dawood T, Lambert EA, Barton DA, Laude D, Elghozi JL, Esler MD, Haikerwal D, Kaye DM, Hotchkin EJ, Lambert GW (2007). Specific serotonin reuptake inhibition in major depressive disorder adversely affects novel markers of cardiac risk. *Hypertension Research* **30**, 285–293.

Dowlati Y, Herrmann N, Swardfager W, Liu H, Sham L, Reim EK, Lanctôt KL (2010). A meta-analysis of cytokines in major depression. *Biological Psychiatry* **67**, 446–457.

**Egger M, Davey Smith G, Schneider M, Minder C** (1997). Bias in meta-analysis detected by a simple, graphical test. *British Medical Journal* **315**, 629–634.

Evans DL, Charney DS, Lewis L, Golden RN, Gorman JM, Krishnan KRR, Nemeroff CB, Bremner JD, Carney RM, Coyne JC, Delong MR, Frasure-Smith N, Glassman AH, Gold PW, Grant I, Gwyther L, Ironson G, Johnson RL, Kanner AM, Katon WJ, Kaufmann PG, Keefe FJ, Ketter T, Laughren TP, Leserman J, Lyketsos CG, McDonald WM, McEwen BS, Miller AH, Musselman D, O'Connor C, Petitto JM, Pollock BG, Robinson RG, Roose SP, Rowland J, Sheline Y, Sheps DS, Simon G, Spiegel D, Stunkard A, Sunderland T, Tibbits JP, Valvo WJ (2005). Mood disorders in the medically ill: scientific review and recommendations. *Biological Psychiatry* 58, 175–189.

Fornaro M, Martino M, Battaglia F, Colicchio S, Perugi G (2011). Increase in IL-6 levels among major depressive disorder patients after a 6-week treatment with duloxetine 60 mg/day: a preliminary observation. *Neuropsychiatric Disease and Treatment* 7, 51–56.

Gabay C, Kushner I (1999). Acute-phase proteins and other systemic responses to inflammation. *New England Journal of Medicine* 340, 448–454.

Hernandez ME, Mendieta D, Martinez-Fong D, Loria F, Moreno J, Estrada I, Bojalil R, Pavon L (2008). Variations in circulating cytokine levels during 52 week course of treatment with SSRI for major depressive disorder. *European Neuropsychopharmacology* **18**, 917–924.

Higgins JPT, Thompson SG, Deeks JJ, Altman DG (2003). Measuring inconsistency in meta-analyses. *British Medical Journal* **327**, 557–560.

- Howren MB, Lamkin DM, Suls J (2009). Associations of depression with C-reactive protein, IL-1, and IL-6: a metaanalysis. *Psychosomatic Medicine* 71, 171–186.
- Irwin MR, Miller AH (2007). Depressive disorders and immunity: 20 years of progress and discovery. *Brain*, *Behavior*, and *Immunity* 21, 374–383.

Janssen DGA, Caniato RN, Verster JC, Baune BT (2010). A psychoneuroimmunological review on cytokines involved in antidepressant treatment response. *Human Psychopharmacology* **25**, 201–215.

Jazayeri S, Keshavarz SA, Tehrani-Doost M, Djalali M, Hosseini M, Amini H, Chamari M, Djazayery A (2010). Effects of eicosapentaenoic acid and fluoxetine on plasma cortisol, serum interleukin-1beta and interleukin-6 concentrations in patients with major depressive disorder. *Psychiatry Research* **178**, 112–115.

Kagaya A, Kugaya A, Takebayashi M, Fukue-Saeki M, Saeki T, Yamawaki S, Uchitomi Y (2001). Plasma concentrations of interleukin-1beta, interleukin-6, soluble interleukin-2 receptor and tumor necrosis factor alpha of depressed patients in Japan. *Neuropsychobiology* **43**, 59–62.

Kenis G, Maes M (2002). Effects of antidepressants on the production of cytokines. *International Journal of Neuropsychopharmacology* 5, 401–142.

Khan AMD, Brodhead AEMS, Schwartz KAMS, Kolts RLP, Brown WAMD (2005). Sex differences in antidepressant response in recent antidepressant clinical trials. *Journal of Clinical Psychopharmacology* **25**, 318–324.

Kubera M, Kenis G, Bosmans E, Zieba A, Dudek D, Nowak G, Maes M (2000). Plasma levels of interleukin-6, interleukin-10, and interleukin-1 receptor antagonist in depression: comparison between the acute state and after remission. *Polish Journal of Pharmacology* **52**, 237–241.

Lanquillon S, Krieg JC, Bening-Abu-Shach U, Vedder H (2000). Cytokine production and treatment response in major depressive disorder. *Neuropsychopharmacology* 22, 370–379.

Leo R, Di Lorenzo G, Tesauro M, Razzini C, Forleo GB, Chiricolo G, Cola C, Zanasi M, Troisi A, Siracusano A, Lauro R, Romeo F (2006). Association between enhanced soluble CD40 ligand and proinflammatory and prothrombotic states in major depressive disorder: pilot observations on the effects of selective serotonin reuptake inhibitor therapy. *Journal of Clinical Psychiatry* 67, 1760–1766.

Lindqvist D, Janelidze S, Hagell P, Erhardt S, Samuelsson M, Minthon L, Hansson O, Bjorkqvist M, Traskman-Bendz L, Brundin L (2009). Interleukin-6 is elevated in the cerebrospinal fluid of suicide attempters and related to symptom severity. *Biological Psychiatry* **66**, 287–292.

Mackay GM, Forrest CM, Christofides J, Bridel MA, Mitchell S, Cowlard R, Stone TW, Darlington LG (2009). Kynurenine metabolites and inflammation markers in depressed patients treated with fluoxetine or counselling. *Clinical and Experimental Pharmacology and Physiology* **36**, 425–435.

Maes M (2001). The immunoregulatory effects of antidepressants. *Human Psychopharmacology: Clinical and Experimental* 16, 95–103.

Maes M, Bosmans E, De Jongh R, Kenis G, Vandoolaeghe E, Neels H (1997). Increased serum IL-6 and IL-1 receptor antagonist concentrations in major depression and treatment resistant depression. *Cytokine* **9**, 853–858.

Maes M, Meltzer HY, Bosmans E, Bergmans R, Vandoolaeghe E, Ranjan R, Desnyder R (1995). Increased plasma concentrations of interleukin-6, soluble interleukin-6, soluble interleukin-2 and transferrin receptor in major depression. *Journal of Affective Disorders* **34**, 301–309.

Maes M, Yirmyia R, Noraberg J, Brene S, Hibbeln J, Perini G, Kubera M, Bob P, Lerer B, Maj M (2009). The inflammatory and neurodegenerative hypothesis of depression: leads for future research and new drug developments in depression. *Metabolic Brain Disease* 24, 27–53.

Mikova O, Yakimova R, Bosmans E, Kenis G, Maes M (2001). Increased serum tumor necrosis factor alpha concentrations in major depression and multiple sclerosis. *European Neuropsychopharmacology* **11**, 203–208. Miller AH, Maletic V, Raison CL (2009). Inflammation and its discontents: the role of cytokines in the pathophysiology of major depression. *Biological Psychiatry* 65, 732–741.

Moore KW, de Waal Malefyt R, Coffman RL, O'Garra A (2001). Interleukin-10 and the interleukin-10 receptor. *Annual Review of Immunology* **19**, 683–765.

O'Brien SM, Scott LV, Dinan TG (2006). Antidepressant therapy and C-reactive protein levels. *British Journal of Psychiatry* 188, 449–452.

O'Donovan A, Hughes BM, Slavich GM, Lynch L, Cronin M-T, O'Farrelly C, Malone KM (2010). Clinical anxiety, cortisol and interleukin-6: evidence for specificity in emotion-biology relationships. *Brain, Behavior, and Immunity* 24, 1074–1077.

Park HS, Park JY, Yu R (2005). Relationship of obesity and visceral adiposity with serum concentrations of CRP, TNF-α and IL-6. *Diabetes Research and Clinical Practice* 69, 29–35.

Pizzi C, Mancini S, Angeloni L, Fontana F, Manzoli L, Costa GM (2009). Effects of selective serotonin reuptake inhibitor therapy on endothelial function and inflammatory markers in patients with coronary heart disease. *Clinical Pharmacology & Therapeutics* 86, 527–32.

**Pudar Hozo S, Djulbegovic B, Hozo I** (2005). Estimating the mean and variance from the median, range, and the size of a sample. *BMC Medical Research Methodology* **5**(13).

**Rosenthal R** (1979). The file drawer problem and tolerance for null results. *Psychological Bulletin* **86**, 638–641.

Rothermundt M, Arolt V, Fenker J, Gutbrodt H, Peters M, Kirchner H (2001). Different immune patterns in melancholic and non-melancholic major depression. *European Archives of Psychiatry and Clinical Neuroscience* 251, 90–97.

Rush AJ, Zimmerman M, Wisniewski SR, Fava M, Hollon SD, Warden D, Biggs MM, Shores-Wilson K, Shelton RC, Luther JF, Thomas B, Trivedi MH (2005). Comorbid psychiatric disorders in depressed outpatients: demographic and clinical features. *Journal of Affective* Disorders 87, 43–55.

Serretti A, Mandelli L (2010). Antidepressants and body weight: a comprehensive review and meta-analysis. *Journal of Clinical Psychiatry* **71**, 1259–1272.

Sluzewska A, Rybakowski JK, Laciak M, Mackiewicz A, Sobieska M, Wiktorowicz K (1995). Interleukin-6 serum levels in depressed patients before and after treatment with fluoxetine. *Annals of the New York Academy of Sciences* 762, 474–476.

Sluzewska A, Sobieska M, Rybakowski JK (1997). Changes in acute-phase proteins during lithium potentiation of antidepressants in refractory depression. *Neuropsychobiology* 35, 123–127.

Song C, Halbreich U, Han C, Leonard BE, Luo H (2009). Imbalance between pro- and anti-inflammatory cytokines, and between Th1 and Th2 cytokines in depressed patients: the effect of electroacupuncture or fluoxetine treatment. *Pharmacopsychiatry* **42**, 182–188.

**Tuglu C, Kara SH, Caliyurt O, Vardar E, Abay E** (2003). Increased serum tumor necrosis factor-alpha levels and treatment response in major depressive disorder. *Psychopharmacology* **170**, 429–433.

Tummers B, van der Laan J, Huyser K (2008). Data Thief III, v. 1.5, computer software (http://datathief.org/). Accessed 8 September 2010.

Yao H, Tao L-G, Zhang X-B, Sha W-W, Hou G, Zhang Z-J (2004). Variety of cell-mediated immunity and the metabolic products of monoamine neural transmitter in depression patients before and after therapy. *Chinese Journal of Clinical Rehabilitation* **8**, 2978–2979. Yoshimura R, Hori H, Ikenouchi-Sugita A, Umene-Nakano W, Ueda N, Nakamura J (2009). Higher plasma interleukin-6 (IL-6) level is associated with SSRI- or SNRI-refractory depression. *Progress in Neuro-Psychopharmacology and Biological Psychiatry* 33, 722–726.

Zorrilla EP, Luborsky L, McKay JR, Rosenthal R, Houldin A, Tax A, McCorkle R, Seligman DA, Schmidt K (2001). The relationship of depression and stressors to immunological assays: a meta-analytic review. *Brain, Behavior and Immunity* **15**, 199–226.