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RESEARCH ARTICLE

Sociocultural and economic determinants of stunting and thinness among adolescent boys and girls in Nepal

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Abstract

Despite the increasing interest in the determinants of adolescent undernutrition in low- and middle-income countries, a comprehensive multi-level overview at the country level is missing. Using the nationally representative 2014 Nepal Adolescent Nutrition Survey, this study aimed to provide a comprehensive overview of the sociocultural and economic determinants of stunting and thinness of adolescent boys and girls in Nepal. Multivariate logistic regression models were used to estimate the associations between multiple individual, household and community determinants and stunting and thinness among 3773 adolescents aged 10–19 years (1888 boys and 1885 girls). The prevalence rates of stunting and thinness indicated suboptimal nutritional status and an inadequate growth environment in Nepali adolescents. The results highlighted an association of paternal occupation and education, household income, number of earning household members, geographical place of residence, caste/ethnicity and nutritional knowledge with stunting, with higher odds for males and older adolescents. Paternal occupation, education, household income, geographical region, caste/ethnicity and nutrition knowledge were associated with thinness, with higher odds for males and younger adolescents. The findings underscore the importance of involving adolescents, their parents and their communities in interventions. Such interventions should not only be aimed at improving adolescent nutrition but also at optimizing adolescents' growth environment for better health and development. Future research should focus on context-specific causal pathways and mechanisms through which sociocultural and economic determinants influence nutritional outcomes within broader societal, cultural and political settings. A longitudinal approach, including a range of dietary and nutrition indicators would allow understanding how and when the relative importance of these factors change during adolescence.

Keywords: Child undernutrition; Population health; Auxology

Introduction

Adolescence is a crucial phase in human development when rapid social, biological and psychological development takes place (WHO, 2014). During this phase, adolescents (aged 10 to 19, as defined by the WHO) adopt more-defined social roles and lay the foundation for their futures (Sawyer *et al.*, 2012, 2018). Adequate nutrition during this stage is essential for optimal growth and development and may affect the health of future generations (Black *et al.*, 2013). Neglecting the health of adolescents might impair earlier investments in maternal and child health (Resnick *et al.*, 2012). Although adolescent nutrition has recently gained more attention (Akseer *et al.*, 2017), research has mostly focused on 15- to 19-year old girls (Salam *et al.*, 2016). Studies including boys and 10- to 14-year old adolescents are limited (The Lancet, 2015). Moreover, although in higher-income countries the interest in the determinants of adolescent nutrition has increased

over recent decades, relatively little is known about the determinants of adolescent nutrition in low- and middle-income countries (LMICs). Furthermore, studies on adolescent nutrition have mostly focused on the effects of single determinants of nutritional status, such as food supplementation, media exposure, nutrition education and parental education (Singh *et al.*, 2014; Lassi *et al.*, 2017; Chau *et al.*, 2018), while only a few studies have tried to obtain a comprehensive overview of the sociocultural and economic determinants. Such a comprehensive overview is fundamental because human health and development, including nutritional status, are not only influenced by biological traits and the immediate setting, but also by underlying factors, broader environments and the interconnections between environments (Dahlgren & Whitehead, 1991). Suboptimal growth due to, among other things but not limited to, infant and child malnutrition and poor health might lead to linear growth retardation during childhood and adolescence and consequently short stature or stunting. Stunting is therefore a marker, rather than an outcome, of undernutrition and has been associated with impaired development, disease and loss of economic productivity (Leroy & Frongillo, 2019). Underweight, or thinness, more directly indicates nutrition deprivation (Best *et al.*, 2010; Black *et al.*, 2013) and is associated with reduced bone density and muscle strength and delayed pubertal maturation (Sellen, 1998). Identifying the relative contributions of individual-, household- and community-level determinants of adolescent stunting and thinness would allow policymakers to design integrated public health interventions targeting several levels. This is all the more necessary since the positive effect of one level can be cancelled out by determinants at another level (Dahlgren & Whitehead, 1991).

This study was inspired by a framework resulting from a systematic review in which the evidence on the multilevel determinants of adolescent undernutrition and micronutrient deficiencies in LMICs were charted (Madjdian *et al.*, 2018). This framework includes a broad range of sociocultural and economic determinants at the individual, household and community level. The present study aimed to test this model empirically and address the above-identified knowledge gaps using data from the 2014 Nepal Adolescent Nutrition Survey (NANS). The main research question was: which individual-, household- and community-level sociocultural and economic factors influence stunting and thinness of adolescent boys and girls aged 10–19 years in Nepal? To answer this research question, NANS data were used, including a wide range of individual, household, and community characteristics of adolescent boys and girls aged 10–19 years (Aryal *et al.*, 2016).

Nepal has a geographically diverse landscape with three agro-ecological zones: Mountains (*Himal*), Hills (*Pahad*) and Plains (*Terai*). Only 7% of the population resides in the Mountains area, where access to transport and communication facilities is limited (NPCS, 2012). In comparison, about 43% of the population lives in the Hills (NPCS, 2012). The southern part of Nepal, the *Terai*, is relatively flat and while it only comprises 23% of the total land area of Nepal, 50% of the population lives here (NPCS, 2012). At the time of the survey, Nepal was administratively divided into five development regions, fourteen zones and 75 districts restructured into seven provinces and 753 urban and rural municipalities after the promulgation of the new constitution in 2015.

Nepal is a multi-ethnic and multi-lingual country. In total, there are 126 different castes and ethnic groups and 123 spoken languages, with Nepali being the official language and mother tongue of 44.6% of the population. Most of the population (81.3%) is Hindu, followed by Buddhist (9%) and Muslim (4.4%) (NPCS, 2012). Like other South Asian countries, Nepal's society largely remains patriarchal with prevailing inequalities in household decision-making processes according to gender, generation and age. Son preference and arranged and early marriages (below 18) remain commonplace (MoH *et al.*, 2017; Yeung *et al.*, 2018).

Nepal is considered to be at stage four of the nutrition transition (Subedi *et al.*, 2017). This transition is related to an increase in the national income level, population growth, migration, urbanization, reduced fertility and mortality rates and increasing life expectancy. Furthermore, the prevalence of underweight among children has decreased over the last 40 years, while rates
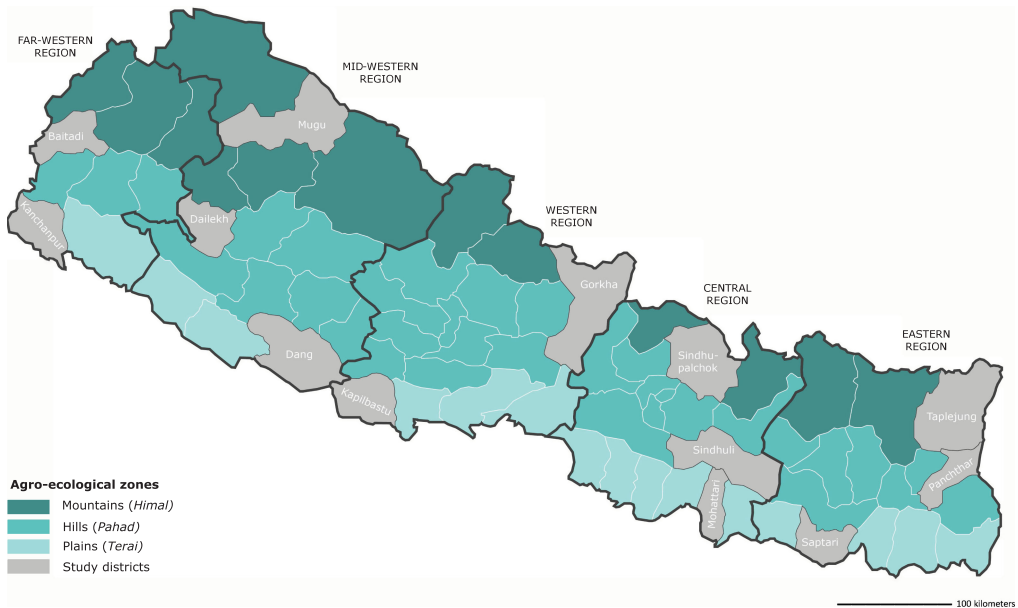
Figure 1. Study districts selected for the 2014 Nepal Adolescent Nutrition Survey depicted in grey.

of overweight, obesity and non-communicable diseases among women are increasing (Subedi *et al.*, 2017). Stunting (32.0% and 32.1%) and thinness (23.3% and 14.0%) among boys and girls aged 10–19 are, however, still a larger concern than overweight and obesity (4.8% and 4.3%) (MoHP *et al.*, 2018). Large differences exist in life expectancy, nutritional status, access to health care and education between ethnicities and ecological zones (MoHP *et al.*, 2018), making Nepal an interesting country to consider.

Methods

Data

Data from the 2014 Nepal Adolescent Nutrition Survey (Aryal *et al.*, 2016) were used. This survey aimed to explore the nutritional status of a nationally representative sample of 3773 adolescents (1888 boys and 1885 girls) aged 10–19 years in Nepal. The cross-sectional survey included demographic, socioeconomic and health characteristics, as well as anthropometric measurements. It was administered to the adolescents by trained enumerators. The sample was obtained by multi-stage cluster sampling in which agro-ecological (Mountains, Hills and *Terai*) and administrative (Eastern, Central, Western, Mid-Western and Far-Western) stratification were taken into account. In the first sampling stage, the thirteen study districts (Fig. 1) of the 2011 Nepal Demographic and Health Survey (NDHS) were used to obtain a representative sample of the agro-ecological and administrative strata. In the second stage, the Village Development Committees (VDCs) in rural areas or municipalities in urban areas, home to each selected district, were pooled and thirty VDCs were sampled by probability proportionate to size (PPS) from each ecological stratum. In the third stage, 21 male and 21 female adolescents were randomly selected from each sampled VDC or municipality. The sample size was weighted by the appropriate sample weights to obtain nationally representative results comparable to the 2011 NDHS to correct the data for the cluster survey design.

Measures

Dependent variables

The dependent adolescent nutrition variables were the dichotomous variables short stature or stunting, indicated by height-for-age z-scores (HAZ), and underweight or thinness, indicated by BMI-for-age z-scores (BAZ). The HAZ indicates linear growth and is a broader marker of 'the inadequacy of the environment to which children have been exposed' (Leroy & Frongillo, 2019). Stunting (the failure to reach linear growth) is indicated by HAZ less than −2 SD from the population median. Hence, stunting is used as an indicator of a deficient growth environment, which includes, but is not limited to, nutrition (Leroy & Frongillo, 2019). Thinness, as indicated by BAZ less than −2 SD from the population median, reflects nutritional status and is therefore a more direct indicator of (acute) undernutrition (Onis *et al.*, 1997). HAZ and BAZ were calculated based on the 2007 WHO growth references using WHO Anthro Plus software (Onis *et al.*, 2007). BAZ less than −5 or BAZ higher than 5 and HAZ less than −6 or HAZ higher than 6 were considered as inconsistent and unlikely and therefore coded as missing values (Onis *et al.*, 2007). Due to missing data or inconsistencies in height or weight, 52 adolescents (1.4% of the total sample) were excluded from the analysis on stunting and 61 adolescents (1.6% of total sample) from the analysis on thinness.

Independent variables

The independent variables were selected based on previous research on the determinants of nutritional status in LMICs and in Nepal specifically. They were derived from the child nutritional status literature and the review by Madjdian *et al.* (2018). The variables were divided into three groups: individual, household and community determinants (Table 1). A small majority (53.1%) of the sample was female. The age of the adolescents ranged from 10 to 19 years and the mean age of the sample was 13.92 years. Adolescents were split into young (10–14 years, 60.0%) and older (15–19 years, 40.0%) age groups. Only 1.8% of the sampled adolescent boys and 3.6% of the adolescent girls were married. Of these married adolescents, 21.0% resided in the Central Mountains area, 19.1% in the Mid-Western Hills and 12.1% in the Western Plains (*Terai*). Religion was categorized as Hindu or non-Hindu. The majority (83.2%) of adolescents were Hindu. The caste/ethnicity categorization (six groups) was based on the STEPS Survey Nepal 2013 categorization (Aryal *et al.*, 2014): 39.8% of adolescents were *Brahmin/Chhetri* (Upper Caste), followed by disadvantaged *Janajati* (31.7%), *Dalit* (12.5%), disadvantaged non-*Dalit Terai* castes (9.7%), relatively advantaged *Janajati* (5.7%) and religious minorities (1.6%). The last group mainly included Muslim adolescents (95%). Footwear was coded '1' if adolescents reported wearing shoes when going outside and '0' otherwise. Wearing shoes was common among the majority of adolescents (96.9%). Almost half of the adolescents had nutrition knowledge, which was determined by asking them whether they knew about nutrition (1) or not (0).

Household determinants included household size (numbers of members, including the respondent), ranging from two to 32 with an average of 6.41. Paternal and maternal education indicated the highest level of education completed by the father or mother. A distinction was made between no formal education, pre-primary or primary (grades 1–5) education and higher than primary education, including secondary education (grade 6–12) and beyond (university). Overall, 23.8% of the fathers and 67.5% of mothers had no formal education. Paternal occupation was divided into four sectors: government or private sector (including non-governmental organizations), agriculture, foreign employment (meaning that the father worked outside Nepal) and other paid employment, including business, daily wage (indicating that income was dependent on day-to-day and often seasonal work), retirement and 'other', comprising all occupations that did not fit into the aforementioned classification. More than two-fifths of the fathers were engaged in agriculture, followed by other paid employment (including business, daily wage, retirement and other) and foreign employment (28.5% and 14.3% respectively). Of the mothers, 67.7% were engaged in

Table 1. Distribution of Nepalese adolescents by background characteristics, 2014 Nepal Adolescent Nutrition Survey

	All adolescents				Boys				Girls			
	Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*	
	n (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)
Individual determinants												
Sex												
Male	1769 (46.9)	572 (32.4)	243 (13.7)									
Female	2003 (53.1)	555 (27.7)	142 (7.1)									
Age (years)				13.92 (2.30)				13.97 (2.29)				13.86 (2.31)
15–19	1508 (40.0)	492 (32.6)	111 (7.4)		1059 (59.9)	261 (36.7)	78 (10.9)		1204 (60.1)	232 (29.0)	34 (4.2)	
10–14	2263 (60.0)	635 (28.1)	274 (12.1)		709 (40.1)	312 (29.4)	165 (15.6)		798 (39.9)	324 (26.9)	108 (9.0)	
Marital status												
Married	104 (2.8)	39 (37.6)	2 (1.9)		32 (1.8)	9 (26.2)	0 (0.6)		72 (3.6)	31 (42.8)	2 (2.5)	
Unmarried	3667 (97.2)	1088 (29.7)	383 (10.4)		1736 (98.2)	564 (32.5)	243 (14.0)		1931 (96.4)	524 (27.2)	140 (7.3)	
Religion												
Non-Hindu	634 (16.8)	197 (31.1)	50 (7.9)		307 (17.4)	111 (36.1)	27 (8.8)		326 (16.3)	86 (26.5)	23 (7.1)	
Hindu	3138 (83.2)	930 (29.6)	335 (10.7)		1462 (82.6)	461 (31.6)	216 (14.8)		1676 (83.7)	469 (28.0)	119 (7.1)	
Ethnicity												
Dalit	473 (12.5)	162 (34.2)	51 (10.8)		215 (12.2)	76 (35.1)	29 (13.3)		258 (12.9)	86 (33.4)	23 (8.8)	
Disadvantaged *Janajati*	1195 (31.7)	311 (26.0)	79 (6.6)		549 (31.1)	170 (31.0)	52 (9.5)		646 (32.3)	140 (21.7)	27 (4.2)	
Disadvantaged non-*Dalit Terai*	364 (9.7)	125 (34.2)	78 (21.4)		191 (10.8)	59 (30.9)	58 (30.4)		174 (8.7)	66 (37.8)	20 (11.6)	
Religious minority	59 (1.6)	28 (48.5)	15 (26.3)		36 (2.1)	18 (49.3)	7 (20.1)		22 (1.1)	10 (47.1)	8 (36.5)	
Relatively advantaged *Janajati*	215 (5.7)	60 (27.8)	3 (1.4)		115 (6.5)	30 (26.2)	3 (2.5)		100 (5.0)	30 (29.7)	0 (0.0)	
Upper Caste	1465 (39.8)	442 (30.2)	158 (10.8)		662 (37.4)	219 (33.1)	94 (14.2)		803 (40.1)	223 (27.7)	64 (8.0)	

(Continued)

Table 1. (*Continued*)

	All adolescents				Boys				Girls			
	Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*	
	n (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)
Footwear outside												
Yes	3653 (96.9)	1085 (29.7)	365 (10.0)		1725 (97.5)	559 (32.4)	238 (13.8)		1928 (96.3)	525 (27.3)	126 (6.6)	
No	118 (3.1)	43 (36.1)	20 (17.1)		43 (2.5)	13 (30.0)	5 (10.5)		75 (3.7)	30 (39.6)	16 (20.9)	
Nutrition knowledge												
Yes	1858 (49.3)	481 (25.9)	135 (7.3)		890 (49.7)	239 (26.9)	79 (8.9)		968 (48.3)	242 (25.0)	56 (5.8)	
No	1913 (50.7)	646 (33.8)	250 (13.1)		890 (50.3)	333 (37.9)	164 (18.6)		1035 (51.7)	313 (30.3)	86 (8.3)	
Household determinants												
Household size				6.41 (2.54)				6.23 (2.52)				6.57 (2.54)
Paternal education												
No formal education	896 (23.8)	297 (33.2)	90 (10.0)		473 (26.8)	144 (30.4)	62 (13.1)		423 (21.1)	153 (36.2)	28 (6.6)	
Pre-primary or primary	1149 (30.5)	382 (33.3)	114 (9.9)		577 (32.6)	232 (40.9)	83 (14.7)		572 (28.6)	150 (26.2)	31 (5.4)	
Higher than primary	1726 (45.8)	448 (25.9)	181 (10.5)		718 (40.6)	196 (27.5)	98 (13.8)		1008 (50.3)	252 (25.0)	84 (8.3)	
Maternal education												
No formal education	2544 (67.5)	803 (31.6)	246 (9.7)		1245 (70.4)	421 (33.8)	168 (13.5)		1299 (64.9)	382 (29.4)	78 (6.0)	
Pre-primary or primary	618 (16.4)	187 (30.2)	67 (10.8)		295 (16.7)	97 (33.0)	41 (13.8)		323 (16.1)	89 (27.5)	26 (8.1)	
Higher than primary	609 (15.1)	138 (22.6)	72 (11.9)		229 (12.9)	54 (23.4)	34 (14.9)		380 (19.0)	84 (22.2)	38 (10.1)	
Paternal occupation												
Government or private sector	521 (13.8)	145 (27.8)	43 (8.3)		246 (13.9)	78 (31.8)	30 (12.1)		276 (13.8)	67 (24.2)	14 (5.0)	
Foreign employment	538 (14.3)	152 (28.2)	77 (14.3)		197 (11.2)	63 (31.9)	41 (20.9)		341 (17.0)	89 (26.1)	36 (10.5)	

Table 1. (*Continued*)

	All adolescents				Boys				Girls			
	Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*	
	n (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)
Business, daily wage or retirement	1075 (28.5)	363 (33.8)	104 (9.5)		539 (30.5)	185 (34.3)	67 (12.5)		536 (26.8)	178 (33.3)	36 (6.8)	
Agriculture	1637 (43.4)	468 (28.6)	161 (9.9)		787 (44.5)	247 (31.3)	105 (13.3)		850 (42.5)	221 (26.0)	56 (6.6)	
Maternal occupation												
Agriculture	859 (22.8)	248 (28.9)	65 (7.6)		358 (20.2)	115 (32.0)	40 (11.3)		501 (25.0)	133 (26.6)	25 (5.0)	
Other paid employment	360 (9.5)	103 (28.5)	25 (6.9)		165 (9.3)	56 (34.1)	12 (7.4)		195 (9.7)	46 (23.7)	13 (6.4)	
Homemaker	2552 (67.7)	777 (30.4)	295 (11.6)		1246 (70.4)	401 (32.2)	191 (15.3)		1307 (65.2)	375 (28.7)	105 (8.0)	
No. earning household members				1.42 (0.60)				1.37 (0.58)				1.46 (0.62)
Monthly household income (NPR)												
≤5000	1052 (27.9)	358 (34.1)	109 (10.3)		545 (30.8)	186 (34.1)	77 (14.2)		507 (25.3)	173 (34.0)	31 (6.1)	
5001–10,000	1356 (36.0)	408 (30.1)	166 (12.3)		618 (35.0)	211 (34.2)	96 (15.5)		738 (36.9)	197 (26.7)	70 (9.5)	
≥10,001	1362 (36.1)	361 (26.5)	110 (8.1)		605 (34.2)	175 (28.9)	70 (11.5)		757 (37.8)	186 (24.6)	41 (5.4)	
Source of drinking water												
Pump, tank, river/pond or well	1582 (41.9)	489 (30.9)	224 (14.1)		777 (43.9)	258 (33.2)	144 (18.5)		804 (40.2)	231 (28.7)	80 (9.9)	
Piped or public tap	2190 (58.1)	638 (29.1)	161 (7.4)		991 (56.1)	314 (31.7)	99 (10.0)		1198 (59.8)	324 (27.1)	62 (5.2)	
Type of latrine												
No flush toilet	1194 (31.7)	394 (33.0)	155 (13.0)		589 (33.3)	215 (36.6)	99 (16.8)		606 (30.2)	178 (29.5)	56 (9.2)	
Flush toilet	2577 (68.3)	733 (28.5)	230 (8.9)		1180 (66.7)	357 (30.2)	144 (12.2)		1397 (69.8)	377 (27.0)	86 (6.2)	

(*Continued*)

Table 1. (*Continued*)

	All adolescents				Boys				Girls			
	Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*		Total	*Stunted*	*Thin*	
	n (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)	*n* (%)	*n* (%)	*n* (%)	Mean (SD)
Community determinants												
Region												
Eastern	771 (20.4)	180 (23.3)	73 (9.5)		349 (19.7)	91 (26.0)	44 (12.6)		422 (21.1)	89 (21.1)	29 (6.9)	
Central	1288 (34.2)	406 (31.5)	130 (10.1)		612 (34.6)	205 (33.5)	82 (13.4)		676 (33.8)	201 (29.7)	48 (7.1)	
Western	484 (12.8)	157 (32.4)	69 (14.2)		221 (12.5)	80 (36.2)	41 (18.7)		263 (13.1)	77 (29.2)	28 (10.5)	
Mid-Western	710 (18.8)	250 (35.2)	63 (8.9)		351 (19.8)	130 (37.0)	44 (12.7)		359 (17.9)	120 (33.4)	19 (5.3)	
Far-Western	518 (13.7)	135 (26.0)	49 (9.5)		236 (13.3)	205 (33.5)	31 (13.2)		282 (14.1)	68 (24.1)	18 (6.4)	
Agro-ecological zone												
Mountains (*Himal*)	1211 (32.1)	414 (34.2)	86 (7.1)		581 (32.9)	218 (37.5)	51 (8.8)		630 (31.4)	196 (31.1)	35 (5.6)	
Hills (*Pahad*)	1228 (32.6)	321 (26.1)	92 (7.5)		540 (30.5)	165 (30.6)	63 (11.6)		688 (34.4)	156 (22.6)	29 (4.2)	
Plains (*Terai*)	1332 (35.3)	393 (29.5)	207 (15.6)		648 (36.6)	189 (29.2)	129 (20.0)		685 (34.2)	204 (29.7)	78 (11.4)	

household labour (homemakers), followed by agriculture (22.8%) and other paid employment (9.5%). Monthly household income was used as a proxy for resources available to the household. The following categorization, as defined in the survey, was used: ≤5000 NPR, 5001–10,000 NPR and ≥10,001 NPR. The adolescents were almost evenly distributed among these groups. The average number of earning members in the household, indicating the number of members who contributed to the monthly household income, was 1.42. The main source of drinking water was piped water (either from a public tap or at home), which was available in almost three-fifths of the households. The remaining two-fifth sourced their water from a river/pond, well or pump. About 68.3% had fixed flush toilets at home.

Community determinants included development (administrative) region, with Central taken as the reference category, accounting for the highest percentage of respondents in the sample. Due to the regionally stratified sampling design, adolescents were quite evenly distributed among the agro-ecological zones (Mountains, Hills and *Terai*). The *Terai* was taken as reference category, being home to a small majority of the adolescents.

Statistical analyses

Data were analysed using IBM SPSS Statistics 23. Dummy variables were created for all categorical variables. Multivariate logistic regressions were performed to determine the strength of the association between the potential predictors and stunting and thinness outcomes (Model 1). Due to substantial differences in the prevalence rates of stunting and thinness between boys and girls and young and older adolescents and expected differences between determinants, separate regression models were estimated for boys (Model 2), girls (Model 3), younger (10–14 years, Model 4) and older (15–19 years, Model 5) adolescents. Predictors were checked for multicollinearity using Variance Inflation Factors. The results of the multivariate logistic regression are presented using odds ratios (ORs) and 95% confidence intervals (CIs). Different levels of statistical significance $p < 0.05$, $p < 0.01$ and $p < 0.001$ are discerned and goodness-of-fit tests are presented.

Results

Prevalence of stunting and thinness

Figure 2 presents distributions of HAZ and BAZ by sex and age group. Stunting (HAZ<−2SD) was prevalent in almost a third of the study population (Table 2). The prevalence among adolescent boys was 32.7% (mean HAZ −1.46 ± 1.16) compared with 28.1% (mean HAZ −1.44 ± 1.05) in girls. Stunting was more prevalent in older (33.1%, mean HAZ −1.62 ± 0.97) than younger (28.3%, mean HAZ −1.33 ± 1.15) adolescents (Fig. 2a). Thinness (BAZ< ± 2SD) was less pronounced and affected only a tenth of the total study population. Thinness was prevalent in 14.0% of boys compared with 7.2% of girls (Table 2). Figure 2b shows that the mean BAZ of girls was closer to 0 than the mean BAZ of boys. In contrast to stunting, thinness was more frequent in the younger age group (12.3%, BAZ −0.74 ± 1.12) than in the older age group (7.5%, BAZ −0.64 ± 0.97) (Table 2).

Determinants of stunting

Binary multivariate logistic regression models were estimated to assess the effects of the selected determinants on stunting over normal stature. The presented ORs show the odds of being stunted compared with not being stunted for each specific determinant while keeping the other determinants constant. The results presented in Table 3 show that in the total population, boys and older adolescents had higher odds of stunting (OR = 1.17; $p < 0.05$ and OR = 1.58; $p < 0.001$, respectively). The increased odds of stunting in the male population were only found in the older age group (OR = 1.35; $p < 0.05$) and the increased odds in the older age group were found among
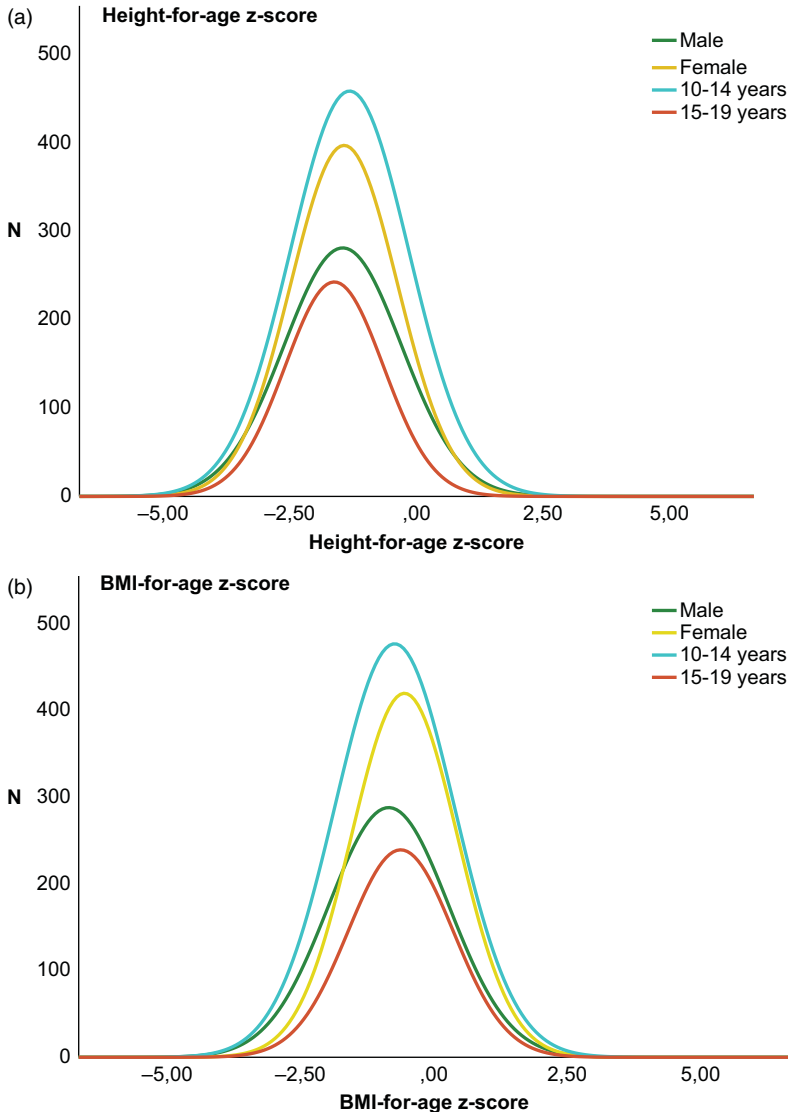
Figure 2. Sex- and age-specific distribution of a) stunting and b) thinness.

both boys (OR = 2.05; $p < 0.001$) and girls (OR = 1.30; $p < 0.05$). In the total population, adolescents from the disadvantaged *Janajati* group had lower odds of stunting (OR = 0.70; $p < 0.01$). This association was also found for girls and younger adolescents (OR = 0.62; $p < 0.01$ and OR = 0.66; $p < 0.01$ respectively). The disadvantaged non-*Dalit Terai* group showed increased odds of stunting in the total, female and older adolescent populations (OR = 1.50; $p < 0.05$, OR = 2.00; $p < 0.01$ and OR = 1.76; $p < 0.05$, respectively). Among religious minorities, increased odds of stunting were found in the total and younger age groups (OR = 2.42; $p < 0.01$ and OR = 3.80; $p < 0.01$ respectively). Nutrition knowledge decreased the odds of stunting in all groups. The largest decrease, of 48% ($p < 0.001$), was found among boys.

At the household level, every additional household member increased the odds of stunting among boys by 7% ($p < 0.01$). Paternal pre-primary or primary education, compared with no

Table 2. Weighted prevalence of stunting and thinness among adolescents

Indicator	All adolescents				Boys				Girls				10–14 years				15–19 years			
	n	%	Mean	SD	n	%	Mean	SD	n	%	Mean	SD	n	%	Mean	SD	n	%	Mean	SD
HAZ			−1.45	1.09			−1.46	1.14			−1.44	1.05			−1.33	1.15			−1.62	0.97
Stunted*	1127	29.9			572	32.7			555	28.1			635	28.3			492	33.1		
Not stunted	2599	68.9			1176	67.3			1423	71.9			1605	71.7			993	66.9		
Total	3726	100			1748	100.0			1978	100.0			2240	100			1485	100		
BAZ			−0.70	1.06			−0.86	1.11			−0.56	1.00			−0.74	1.12			−0.64	0.97
Thinness**	385	10.2			243	14.0			142	7.2			274	12.3			111	7.5		
No thinness	3332	88.4			1498	86.0			1834	92.8			1959	87.7			1373	92.5		
Total	3717	100			1741	100.0			1976	100.0			2233	100.0			1484	100.0		

*$p < 0.01$; **$p < 0.001$.

Table 3. Adjusted odds ratios (aOR) and 95% confidence intervals (CIs) of stunting by individual, household and community determinants

	Model 1 (Total)	Model 2 (Boys)	Model 3 (Girls)	Model 4 (10–14 years)	Model 5 (15–19 years)
	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)
Individual determinants					
Sex (Ref. = Female)					
Male	1.17 (1.01; 1.35)*			1.04 (0.86; 1.27)	1.35 (1.07; 1.70)*
Age (Ref. = 10–14 years)					
15–19 years	1.58 (1.34; 1.86)***	2.05 (1.62; 2.60)***	1.30 (1.03; 1.64)*		
Marital status (Ref. = Unmarried)					
Married	1.13 (0.74; 1.74)	0.57 (0.25; 1.30)	1.51 (0.90; 2.53)	0.61 (0.22; 1.70)	1.36 (0.82; 2.26)
Religion (Ref. = Hindu)					
Non-Hindu	1.18 (0.91; 1.53)	1.10 (0.75; 1.62)	1.17 (0.81; 1.70)	1.08 (0.77; 1.51)	1.4 (0.92; 2.16)
Caste/ethnicity (Ref. = Upper Caste)					
Dalit	1.10 (0.86; 1.41)	1.03 (0.71; 1.50)	1.19 (0.84; 1.68)	1.06 (0.76; 1.47)	1.16 (0.78; 1.73)
Disadvantaged *Janajati*	0.70 (0.56; 0.88)**	0.79 (0.56; 1.10)	0.62 (0.45; 0.86)**	0.66 (0.49; 0.90)**	0.77 (0.53; 1.11)
Disadvantaged non-*Dalit Terai*	1.50 (1.07; 2.10)*	1.09 (0.66; 1.80)	2.00 (1.24; 3.23)**	1.50 (0.95; 2.38)	1.76 (1.04; 2.97)*
Religious minority	2.42 (1.26; 4.66)**	2.23 (0.91; 5.46)	2.63 (0.94; 7.31)	3.80 (1.66; 8.73)**	1.40 (0.46; 4.23)
Relatively advantaged *Janajati*	0.84 (0.60; 1.17)	0.70 (0.43; 1.12)	1.01 (0.62; 1.64)	0.78 (0.50; 1.24)	0.93 (0.56; 1.54)
Footwear outside (Ref. = No)					
Yes	0.75 (0.5; 1.12)	0.98 (0.49; 1.97)	0.74 (0.44; 1.24)	0.70 (0.44; 1.12)	0.90 (0.38; 2.1)
Nutrition knowledge (Ref. = No)					
Yes	0.65 (0.55; 0.76)***	0.52 (0.41; 0.66)***	0.77 (0.61; 0.97)*	0.60 (0.48; 0.74)***	0.74 (0.57; 0.95)*

(*Continued*)

Table 3. (*Continued*)

	Model 1 (Total)	Model 2 (Boys)	Model 3 (Girls)	Model 4 (10–14 years)	Model 5 (15–19 years)
	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)
Household determinants					
Household size	1.02 (0.99; 1.06)	1.07 (1.02; 1.12)**	0.99 (0.95; 1.04)	1.03 (0.98; 1.07)	1.02 (0.97; 1.07)
Paternal education (Ref. = No formal education)					
Pre-primary or primary	1.08 (0.88; 1.32)	1.56 (1.18; 2.07)**	0.72 (0.53; 0.96)*	0.93 (0.71; 1.22)	1.25 (0.92; 1.70)
Higher than primary	0.81 (0.66; 1.00)*	0.91 (0.67; 1.24)	0.66 (0.49; 0.88)**	0.73 (0.55; 0.97)*	0.93 (0.68; 1.28)
Maternal education (Ref. = No formal education)					
Pre-primary or primary	1.08 (0.88; 1.33)	1.03 (0.76; 1.38)	1.09 (0.81; 1.47)	1.13 (0.86; 1.48)	0.96 (0.68; 1.35)
Higher than primary	0.90 (0.71; 1.15)	0.84 (0.57; 1.24)	0.89 (0.65; 1.24)	0.88 (0.65; 1.20)	0.88 (0.58; 1.34)
Paternal occupation (Ref. = Agriculture)					
Government or private sector	1.13 (0.88; 1.46)	1.28 (0.89; 1.83)	1.06 (0.74; 1.52)	1.00 (0.71; 1.41)	1.30 (0.89; 1.90)
Foreign employment	1.16 (0.91; 1.50)	1.20 (0.81; 1.79)	1.20 (0.86; 1.68)	1.00 (0.72; 1.38)	1.42 (0.95; 2.14)
Business, daily wage or retirement	1.35 (1.12; 1.63)**	1.18 (0.90; 1.54)	1.46 (1.11; 1.92)**	1.15 (0.90; 1.47)	1.72 (1.28; 2.32)***
Maternal occupation (Ref. = Homemaker)					
Agriculture	1.11 (0.91; 1.36)	1.19 (0.88; 1.59)	1.14 (0.86; 1.52)	1.15 (0.88; 1.51)	1.16 (0.85; 1.58)
Other paid employment	1.15 (0.87; 1.53)	1.40 (0.93; 2.09)	0.99 (0.66; 1.49)	1.31 (0.91; 1.87)	1.16 (0.71; 1.90)
No. earning household members	0.81 (0.70; 0.94)**	0.62 (0.50; 0.78)***	1.01 (0.82; 1.26)	0.67 (0.54; 0.84)***	0.96 (0.77; 1.19)
Monthly household income (Ref. = ≥10,001 NPR)					
≤5000 NPR	1.24 (1.00; 1.53)	1.05 (0.77; 1.44)	1.31 (0.97; 1.77)	1.05 (0.79; 1.41)	1.52 (1.11; 2.09)**
5001–10,000 NPR	1.11 (0.92; 1.33)	1.01 (0.76; 1.33)	1.08 (0.84; 1.40)	1.16 (0.90; 1.49)	1.05 (0.79; 1.39)

(*Continued*)

Table 3. (*Continued*)

	Model 1 (Total)	Model 2 (Boys)	Model 3 (Girls)	Model 4 (10–14 years)	Model 5 (15–19 years)
	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)
Source of drinking water (Ref = Piped or public tap)					
Pump, tank, river/pond or well	1.05 (0.82; 1.34)	1.27 (0.90; 1.81)	0.78 (0.54; 1.11)	1.11 (0.82; 1.52)	0.92 (0.61; 1.37)
Type of latrine in house (Ref. = Flush toilet)					
No flush toilet	1.14 (0.96; 1.36)	1.22 (0.95; 1.56)	1.01 (0.79; 1.30)	1.09 (0.87; 1.37)	1.20 (0.92; 1.57)
Community determinants					
Region (Ref. = Central)					
Eastern	0.67 (0.53; 0.84)***	0.79 (0.57; 1.10)	0.60 (0.43; 0.83)**	0.64 (0.47; 0.85)**	0.70 (0.48; 1.00)
Western	1.28 (0.98; 1.67)	1.33 (0.90; 1.98)	1.30 (0.89; 1.90)	0.90 (0.62; 1.28)	2.04 (1.34; 3.11)***
Mid-Western	1.62 (1.28; 2.04)***	1.55 (1.10; 2.19)*	1.61 (1.15; 2.25)**	1.34 (0.99; 1.82)	2.16 (1.48; 3.15)***
Far-Western	1.10 (0.83; 1.46)	0.98 (0.64; 1.49)	1.20 (0.81; 1.78)	1.04 (0.72; 1.50)	1.32 (0.84; 2.06)
Agro-ecological zone (Ref. = Plains)					
Mountains (*Himal*)	1.86 (1.36; 2.52)***	2.28 (1.45; 3.57)***	1.28 (0.82; 2.01)	2.39 (1.60; 3.57)***	1.29 (0.78; 2.15)
Hills (*Pahad*)	1.07 (0.81; 1.43)	1.58 (1.04; 2.39)*	0.63 (0.42; 0.96)*	1.38 (0.94; 2.03)	0.77 (0.49; 1.21)

*$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.
Model 1: Hosmer and Lemeshow χ^2 16.55 ($p = 0.08$); Nagelkerke R^2 0.08.
Model 2: Hosmer and Lemeshow χ^2 3.44 ($p = 0.90$); Nagelkerke R^2 0.12.
Model 3: Hosmer and Lemeshow χ^2 18.52 ($p = 0.02$); Nagelkerke R^2 0.09.
Model 4: Hosmer and Lemeshow χ^2 5.86 ($p = 0.66$); Nagelkerke R^2 0.10.
Model 5: Hosmer and Lemeshow χ^2 10.37 ($p = 0.24$); Nagelkerke R^2 0.10.

formal education, increased the odds of stunting among boys (OR = 1.56; $p < 0.01$), while it decreased the odds of stunting among girls (OR = 0.72; $p < 0.05$). Paternal education beyond primary decreased the risk of stunting among all adolescents, girls and the younger adolescent subgroups (OR = 0.81; $p < 0.05$, OR = 0.66; $p < 0.01$ and OR = 0.73; $p < 0.05$). As for paternal education, paternal occupation in business, daily wage or retirement increased the odds in the total, female and older populations (OR = 1.35; $p < 0.01$, OR = 1.46; $p < 0.01$ and OR = 1.72; $p < 0.001$, respectively). An increasing number of earning members in the household was associated with a decreased likelihood of stunting in the total, male and younger age groups. The largest decrease was found in adolescent boys (OR = 0.62; $p < 0.001$). Furthermore, a monthly household income below 5000 NPR compared with a monthly income above 10,001 NPR increased the odds of stunting in the older adolescent population (OR = 1.52; $p < 0.01$).

At the community level, residing in the Mid-Western development region was associated with higher stunting odds (OR = 1.62; $p < 0.001$), while residing in the Eastern region was associated with decreased stunting odds (OR = 0.67; $p < 0.001$). In the separate models, residing in the Mid-Western region increased the odds of stunting in adolescent boys (OR = 1.55; $p < 0.05$) and girls (OR = 1.61; $p < 0.01$), while residence in the Eastern region decreased the odds among younger adolescents (OR = 0.64; $p < 0.01$). Residence in both the Western and Mid-Western regions increased the odds in older adolescents (OR = 2.04; $p < 0.001$ and OR = 2.16; $p < 0.001$). Additionally, adolescent boys living in the Mountain and Hill areas were 2.28 ($p < 0.001$) and 1.58 ($p < 0.05$) times more likely to be stunted compared with adolescent boys living in the *Terai*, while adolescent girls living in the Hills, were less likely to be stunted (OR = 0.63; $p < 0.05$). Similarly, young adolescents residing in the Mountains showed increased odds of stunting (OR = 2.39; $p < 0.001$).

Determinants of thinness

Table 4 shows the relationship between sociocultural and economic determinants and thinness of adolescents. The binary logistic regression model shows that adolescent boys were 2.32 ($p < 0.001$) times more likely to be thin than girls. This association was confirmed in the separate age group models where younger and older boys were 2.07 ($p < 0.001$) and 3.12 ($p < 0.001$) times more likely to be thin than girls their age. In Model 1, the older age group had a lower likelihood of thinness (OR = 0.72; $p < 0.05$) compared with the younger age group. This association was only found for girls (OR = 0.54; $p < 0.01$). Being married decreased the mean chance of thinness (OR = 0.22; $p < 0.05$) in the total population. This association was not found in the separate models. Disadvantaged *Janajati*s showed lower thinness odds in all models. The largest decrease was found amongst older adolescents (OR = 0.30; $p < 0.01$). Adolescents belonging to a religious minority showed increased odds of thinness (OR = 1.28; $p < 0.001$). Model 2 showed lower odds for relatively advantaged *Janajati* boys (OR = 0.15; $p < 0.01$), while Model 3 showed decreased odds of thinness (OR = 0.38; $p < 0.001$) among disadvantaged *Janajati* girls. Among younger adolescents, lower odds were found for relatively advantaged *Janajati* (OR = 0.15; $p < 0.01$), while the older age group showed decreased odds for *Dalit* (OR = 0.42; $p < 0.05$). Wearing footwear outdoors was associated with lower thinness odds for girls only (OR = 0.37; $p < 0.01$). The likelihood of thinness was lower among adolescents with nutrition knowledge in all models, except for girls. The largest decrease was found in adolescent boys (OR = 0.45; $p < 0.001$).

At the household level, adolescents whose mothers attained beyond primary education were 1.52 times ($p < 0.05$) more likely to be thin that adolescents whose mothers had no formal education. This effect was also found in the younger adolescent subgroup (OR = 1.60; $p < 0.05$). Paternal foreign employment increased the odds of thinness by 47% ($p < 0.05$) in the total and 74% ($p < 0.05$) in the male population. Among young adolescents, paternal occupation in business, daily wage or retirement decreased the odds by 38% ($p < 0.01$) compared with agricultural occupation. A decrease of 37% in thinness was found in older adolescents with every

Table 4. Adjusted odds ratios (aOR) and 95% confidence intervals (CIs) of individual, household, and community determinants of thinness

	Model 1 (Total)	Model 2 (Boys)	Model 3 (Girls)	Model 4 (10–14 years)	Model 5 (15–19 years)
	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)
Individual determinants					
Sex (Ref. = Female)					
Male	2.32 (1.84; 2.94)***			2.07 (1.57; 2.74)***	3.12 (1.97; 4.95)***
Age (Ref. = 10–14 years)					
15–19 years	0.72 (0.56; 0.94)*	0.84 (0.61; 1.17)	0.54 (0.34; 0.84)**		
Marital status (Ref. = Unmarried)					
Married	0.22 (0.05; 0.90)*	0.05 (0.00; 3.97)	0.38 (0.08; 1.78)	0.30 (0.05; 1.88)	0.19 (0.02; 2.04)
Religion (Ref. = Hindu)					
Non-Hindu	1.02 (0.65; 1.60)	0.86 (0.46; 1.58)	1.28 (0.64; 2.56)	0.90 (0.55; 1.50)	1.18 (0.41; 3.43)
Caste/ethnicity (Ref. = Upper Caste)					
Dalit	0.74 (0.51; 1.08)	0.66 (0.40; 1.10)	0.89 (0.49; 1.61)	0.89 (0.57; 1.38)	0.42 (0.19; 0.92)*
Disadvantaged *Janajati*	0.46 (0.32; 0.65)***	0.51 (0.32; 0.81)**	0.38 (0.22; 0.67)***	0.53 (0.35; 0.79)**	0.30 (0.14; 0.62)**
Disadvantaged non-*Dalit Terai*	1.01 (0.66; 1.55)	1.36 (0.77; 2.40)	0.62 (0.30; 1.29)	1.14 (0.67; 1.93)	0.74 (0.35; 1.56)
Religious minority	1.28 (0.56; 2.94)***	0.85 (0.27; 2.72)	2.17 (0.60; 7.86)	0.86 (0.30; 2.48)	2.61 (0.53; 12.84)
Relatively advantaged *Janajati*	0.11 (0.03; 0.35)	0.15 (0.05; 0.51)**	0	0.15 (0.04; 0.49)**	0.01 (0.00; 5.22)
Footwear outside (Ref. = No)					
Yes	0.74 (0.44; 1.27)	2.12 (0.75; 5.99)	0.37 (0.19; 0.72)**	0.86 (0.46; 1.61)	0.39 (0.12; 1.23)
Nutrition knowledge (Ref. = No)					
Yes	0.61 (0.47; 0.78)***	0.45 (0.32; 0.62)***	0.96 (0.63; 1.47)	0.66 (0.48; 0.90)**	0.48 (0.30; 0.76)**

(*Continued*)

Table 4. (*Continued*)

	Model 1 (Total)	Model 2 (Boys)	Model 3 (Girls)	Model 4 (10–14 years)	Model 5 (15–19 years)
	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)	aOR (95%CI)
Household determinants					
Household size	1.04 (1.00; 1.09)	1.05 (0.98; 1.12)	1.04 (0.96; 1.12)	1.03 (0.98; 1.09)	1.09 (0.99; 1.19)
Paternal education (Ref. = No formal education)					
Pre-primary or primary	1.24 (0.90; 1.71)	1.41 (0.94; 2.11)	1.07 (0.60; 1.92)	1.42 (0.96; 2.09)	1.08 (0.57; 2.04)
Higher than primary	1.16 (0.84; 1.61)	1.11 (0.73; 1.71)	1.19 (0.69; 2.05)	1.03 (0.69; 1.54)	1.57 (0.87; 2.82)
Maternal education (Ref. = No formal education)					
Pre-primary or primary	1.22 (0.89; 1.68)	1.07 (0.71; 1.61)	1.41 (0.84; 2.37)	1.39 (0.96; 2.00)	0.79 (0.39; 1.60)
Higher than primary	1.52 (1.08; 2.14)*	1.42 (0.88; 2.31)	1.61 (0.96; 2.71)	1.60 (1.06; 2.39)*	1.58 (0.79; 3.15)
Paternal occupation (Ref. = Agriculture)					
Government or private sector	0.86 (0.57; 1.28)	1.14 (0.68; 1.92)	0.62 (0.31; 1.23)	0.70 (0.42; 1.14)	1.62 (0.76; 3.45)
Foreign employment	1.47 (1.03; 2.10)*	1.74 (1.07; 2.85)*	1.24 (0.72; 2.15)	1.33 (0.87; 2.03)	1.42 (0.70; 2.90)
Business, daily wage or retirement	0.87 (0.65; 1.16)	0.97 (0.67; 1.42)	0.71 (0.43; 1.17)	0.62 (0.43; 0.88)**	1.65 (0.94; 2.90)
Maternal occupation (Ref. = Homemaker)					
Agriculture	1.07 (0.77; 1.49)	1.16 (0.76; 1.77)	1.11 (0.63; 1.94)	1.05 (0.71; 1.56)	1.12 (0.57; 2.19)
Other paid employment	0.78 (0.48; 1.26)	0.53 (0.27; 1.05)	1.36 (0.67; 2.77)	0.72 (0.40; 1.27)	0.92 (0.34; 2.51)
No. earning household members	0.81 (0.64; 1.04)	0.83 (0.60; 1.14)	0.70 (0.46; 1.05)	0.92 (0.68; 1.24)	0.63 (0.41; 0.99)*
Monthly household income (Ref. = ≥10,001 NPR)					
≤5000 NPR	1.38 (0.98; 1.95)	1.52 (0.97; 2.37)	1.26 (0.70; 2.26)	1.29 (0.85; 1.95)	1.96 (1.02; 3.74)*
5001–10,000 NPR	1.71 (1.28; 2.27)***	1.51 (1.03; 2.21)*	2.03 (1.27; 3.25)**	1.52 (1.07; 2.15)*	2.45 (1.43; 4.22)**

(*Continued*)

additional earning household member. A monthly household income of 10,000 NPR or below increased the odds of thinness compared with a monthly income of 10,001 NPR or above. While a monthly family income below 5000 NPR increased the likelihood of thinness in the older population (OR = 1.96; $p < 0.05$), a monthly family income between 5001 and 10,000 NPR increased the risk of thinness in all models.

At the community level, residing in the Mid-Western and Far-Western development regions was associated with decreased odds of thinness (OR = 0.68; $p < 0.05$ and OR = 0.59; $p < 0.05$ respectively). This effect was only found for boys and for older adolescents residing in the Far-Western region (OR = 0.53; $p < 0.05$ and OR = 0.33; $p < 0.01$, respectively). Additionally, decreased odds of thinness (OR = 0.44; $p < 0.05$) were found for older adolescents residing in the Western development region compared with their counterparts in the Central region. Similarly, adolescents residing in the Mountains or Hills had a lower risk (OR = 0.43; $p < 0.001$ and OR = 0.48; $p < 0.01$, respectively) of thinness compared with their counterparts living in the *Terai*. This association was also found for boys living in the Mountains (OR = 0.40; $p < 0.01$), girls residing in the Hills (OR = 0.28; $p < 0.01$), young adolescents residing in the Mountains and Hills (OR = 0.47; $p < 0.01$ and OR = 0.31; $p < 0.001$ respectively) and older adolescents residing in the Mountains (OR = 0.22; $p < 0.01$).

Discussion

This study provides a multi-level overview of the wide range of sex- and age-specific individual-, household- and community-level sociocultural and economic determinants that influence the odds of stunting and thinness among adolescent boys and girls aged 10–19 years in Nepal. The findings show a widespread prevalence of both stunting and thinness among Nepali adolescents leading to the conclusion that many of them grow up in a deficient growth environment or might suffer from inadequate diets. Generally, boys were found to be more at risk of stunting and thinness than girls. Stunting increased with age, while thinness odds decreased with age. Male sex, older age, belonging to a religious minority, paternal occupation in business, daily wage, retirement or other and living in the Mid-Western and Mountain areas were among the main risk factors for stunting. Belonging to a disadvantaged *Janajati* caste/ethnicity, paternal education beyond primary education, higher number of household earning members, nutritional knowledge and living in the Eastern development region were protective against stunting. Male sex, belonging to a religious minority, a low household income, maternal education beyond primary education and paternal foreign employment were the main factors associated with increased odds of thinness. Older age, being married, identifying as disadvantaged *Janajati*, living in Mid-Western or Far-Western regions, and Hills and Mountains regions and having nutritional knowledge, decreased the risk of thinness in most populations.

In line with previous studies based on DHS data from South Asia, the male sex has been associated with higher odds of both stunting and thinness (Benedict *et al.*, 2018). Higher thinness odds might be explained by boys' increased energy expenditure due to more participation in labour activities, possibly in combination with school enrolment. Girls might be at lower risk due to increased involvement in cooking, during which they might be able to consume more food (Yamanaka & Ashworth, 2002; Madjdian *et al.*, 2018). However, stunting indicates more than nutrition deficiencies, but also links to a multitude of factors, including a low-quality diet, poor health status and environmental and societal challenges. Because these factors have not all been captured in this survey, it goes beyond the scope of this study to explain why adolescent boys, and in particular older boys, have higher odds of stunting than girls. Explanations provided in the literature range from biological to sociocultural, including different puberty growth spurt onsets, childhood feeding practices, boys' greater risk of morbidity and mortality in early life and societal influences due to gender inequality (Svefors *et al.*, 2020). Further research could thus benefit from including indicators such as energy expenditure and adolescent time use.

Age was found to be associated with stunting among both boys and girls in Nepal. This finding does not corroborate previous literature showing that a generally declining stunting trend during adolescence results from catch-up growth at later age, which compensates for earlier-age growth deficits (Gausman *et al.*, 2019; Astatkie, 2020). The present study also did not find a significantly higher prevalence of stunting among boys compared with girls, which has been previously explained by boys not having reached their full height due to a later onset of puberty than girls, who generally have reached their full height (Sawyer *et al.*, 2018). However, stunting is influenced by prenatal, infant and childhood experience and fluctuates during childhood and adolescence. Hence, both recovery and growth faltering or stunting persistence may take place (Dewey & Begum, 2011; Sawyer *et al.*, 2018). It is likely that adolescents who were already stunted during childhood have a higher probability of growing into stunted adults, while non-stunted adolescents have lower stunting odds in adolescence or adulthood (Gausman *et al.*, 2019; Astatkie, 2020). Another explanation could be that adolescents in the younger age group had lower stunting prevalence rates in early life and childhood than the older group. However, due to the cross-sectional design of this survey, it was not possible to determine causality or discern whether growth is inhibited during late adolescence. Additionally, it is important to note that the growth references commonly used to determine undernutrition face challenges related to ethnic and racial differences, as well as variations in growth spurts and potential and body composition and might thus overestimate stunting prevalence (Tumilowicz *et al.*, 2019). On the contrary to increased stunting with age, older age was associated with decreased odds of thinness in girls only. Although household 'serving orders', where for instance Hindu adolescent girls are served last during meals, still exist (Madjdian & Bras, 2016; Morrison *et al.*, 2018), it is also possible that older girls receive more food when being prepared for marriage or pregnancy (Harris-Fry *et al.*, 2017).

The positive association between nutrition knowledge and nutritional status found in all groups, except for girls for thinness (OR = 0.96; p = 0.84), could be explained by nutrition knowledge being a driver for healthier food choices. Moreover, it might create a spill-over effect by, for instance, educating parents about nutrition. The study findings are in line with a previous study among Bangladeshi adolescent girls aged 13–18 years, which found no association between the awareness of increased nutrient requirements during adolescence and stunting or thinness in girls (Alam *et al.*, 2010). The absence of an effect of nutrition knowledge on thinness among girls in this study could be explained by the fact that girls are more involved in food preparation and cooking from an early age, and underestimate their knowledge, or that girls have less agency than their male counterparts, who might be better able to act on their knowledge when it comes to food- and health-related decision-making. Assuming nutritional knowledge is partly gained at school it might also be related, particularly in the older age group, to lower (secondary) school completion rates. Out-of-school girls are 'less reached' by educational programmes (Alam *et al.*, 2010). Although the gender gap in primary and secondary education in Nepal is narrowing, many girls are at risk of dropping out of, or skipping, school due to financial or social reasons. Unfortunately, data on adolescents' educational attainment are missing. Nutrition knowledge could also be seen as a proxy for socioeconomic or educational status, which might explain its association with stunting (Akhter *et al.*, 2018). These results should be interpreted with caution as nutrition knowledge was self-reported and it does not explain what kind of knowledge this involves exactly.

Studies from Nepal have repeatedly reported health and nutrition disparities between caste/ethnic groups. Inequalities are probably caused by disadvantaged groups' lower socioeconomic position, resulting in decreased access to resources and higher vulnerability to poor health outcomes, but could also be attributed to different food consumption patterns (Adhikari, 2010). The present study showed mixed results and should be interpreted with caution. First, it could be questioned whether caste/ethnicity is a stand-alone predictor of undernutrition as Nepal's multitude of religious, ethnic and caste-bound practices overlap, making it difficult to identify with one specific caste/ethnic group. Second, it is likely that a combination of wealth, education and caste/ethnicity determine an individual's position and vulnerability to poor health outcomes (Adhikari, 2010).

Third, as discussed earlier, the growth reference used might not be fully able to discern ethnic and racial differences in body composition (Tumilowicz *et al.*, 2019). The study found some association between caste/ethnicity and stunting and thinness, for both adolescent boys and girls. Adolescents from the disadvantaged non-*Dalit Terai* caste/ethnic group had increased stunting odds, despite the finding that living in the Hills or Mountains increased stunting odds compared with living in the *Terai*. However, this corresponds with national data on lower use of health services and poorer health outcomes among *Terai* caste/ethnic groups (Ghimire *et al.*, 2019). Disadvantaged *Janajati* adolescents were less likely to be thin and stunted compared with Upper Caste adolescents. This is in line with previous research that reported lower odds of thinness among *Janajati* women aged 15–49 (Pandey *et al.*, 2013). The *Janajati* classification of Pandey *et al.* (2013) included the same ethnicities and castes as this study's disadvantaged *Janajati* classification, as well as *Thakali* and *Gurung* castes/ethnicities, which were included under relatively advantaged *Janajati*. Relatively advantaged *Janajati* adolescent boys and younger adolescents were less likely to be thin, which might be mediated by wealth. In the present sample, 42.5% of relatively advantaged *Janajati* boys had a monthly household income of 10,001 NPR or more, compared with 34.2% in the total male population. This group included the *Newar* caste/ethnicity, which represents a socioeconomically advantaged caste/ethnic group with a generally better health status (Bennett *et al.*, 2008; Adhikari, 2010). Religious minority adolescents showed higher risks of thinness, while younger adolescents belonging to a religious minority had higher stunting odds compared with the Upper Caste. In the study sample, 54 out of 59 adolescents identifying as a religious minority were Muslim. Although these results should be interpreted with caution given the small representation of religious minorities in the sample (1.6 %), in Nepal, Muslims are known to make less use of health services and have poorer health outcomes (MoH *et al.*, 2017). Additionally, the religious practice of fasting (*Roza*) from dawn to dusk during Ramadan, which leads to changes in meal schedules and meal types, might result (particularly for girls and combined with already existing unequal food allocation) in reduced food intake and weight loss leading to higher odds of thinness (Poh *et al.*, 1996). Another strand of literature suggests that pre-birth exposure to Ramadan (maternal fasting) negatively influences the stature and body mass of religious Muslims during childhood (Karimi, 2018) and adolescence (Kunto & Mandemakers, 2019).

The absence of an effect of latrine type or source of drinking water and stunting and thinness might be explained by the large effect of household income. Wearing shoes outside was associated with lower odds of thinness in girls, which is in line with evidence on the association between footwear and soil-transmitted helminth infections leading to undernutrition, or vice versa (Amare *et al.*, 2013). Similar results have been found in the context of anaemia prevalence amongst adolescents in Nepal (Chalise *et al.*, 2018). This finding could also be confounded by poverty. Hence, not (being able to) wear shoes could be a proxy indicator of economic status.

Household income was one of the main contributors to stunting and thinness in all groups. This is in line with previous research linking higher household income or wealth with access to high-quality food and health care services and less exposure to health risks (Victora *et al.*, 2003; Bashir & Schilizzi, 2013; Sreeramareddy *et al.*, 2015; Vollmer *et al.*, 2017b; Nepali *et al.*, 2019). A greater number of earning household members was associated with decreased odds of stunting in all but the female and older population. The study outcomes suggest a pro-male bias that could be explained by the 'buffer hypothesis', assuming that in the context of poverty and food insecurity, women and older household members buffer men and younger household members from the adverse effects of, for instance, food insecurity by eating less (Hadley *et al.*, 2008; Harris-Fry *et al.*, 2017; Moffitt & Ribar, 2018). An increasing number of earning members might lead to a higher household income and thus more resources to invest in food, health and care (Bashir & Schilizzi, 2013), thereby lowering the risk of stunting. However, no data were available on the work status of adolescents or adolescents' relation to the additional earning members. Such information would be helpful to understand whether there is a food allocation bias favouring

those who contribute more to household income, which in this context is probably boys, who are more often engaged in paid work and school than girls (Dercon & Singh, 2013; Harris-Fry *et al.*, 2017; Morrison *et al.*, 2018). Nevertheless, the clear association between economic indicators and stunting and thinness points to the need to improve household access to, and availability of, food and other more underlying but related factors and assets, such as health care.

While for both boys and girls poverty was an important determinant of stunting, the relative importance of parental education differed for boys and girls and by age group. Generally, higher paternal education was protective against stunting in girls and younger adolescents. Maternal education was to a lesser extent associated with stunting or thinness. For instance, maternal education beyond primary increased the odds of thinness in younger adolescents. Hence, while most studies have found a positive association between maternal education and especially girls' nutritional status through increased knowledge and use of health services, higher household income and increased decision-making power (Dancer & Rammohan, 2009; Aslam & Kingdon, 2012; Rahman *et al.*, 2015; Hasan *et al.*, 2016; Kunto & Bras, 2018), this study's findings suggest that paternal education is equally, or even more, important than maternal education. The absence of a protective effect of maternal education could be caused by women's work outside the household, thereby transferring feeding and caring tasks to others (Reed *et al.*, 1996). Some studies have pointed out that maternal occupation might result in less contact time including adverse child feeding and care practices (Kunto & Bras, 2018), leading to a negative effect on growth of children. Nevertheless, the present findings are in line with a recent study from Nepal that found that, while maternal secondary education or higher was a significant contributor to stunting in under-five children in 1996, it lost significance in 2016. This was explained by the fact that the gender gap in education decreases with many women being educated, thereby reducing the effect of maternal education (Angdembe *et al.*, 2019). The present findings are also in line with previous research among Indian adults that showed that male, but not female, education reduced the risk of being underweight (Siddiqui & Donato, 2017). In Nepal's predominantly patriarchal society, men are often decision-makers and are able to influence decisions related to food, health and care (Morrison *et al.*, 2018). A systematic review investigating differences in parental educational levels on child nutritional status concluded that paternal education is equally important for nutritional status of children as maternal education, although the mechanisms might differ (Vollmer *et al.*, 2017a). The association of boys' higher odds of stunting with paternal pre-primary or primary education might be due to mediation by household income. For instance, 40.5% of the boys' fathers who had completed pre-primary or primary education worked in daily wage and business or were retired, possibly resulting in an unstable income, compared with 30.5% in the overall male population. Further research is needed to understand through which pathways parental education affects the nutritional status of adolescents in order to tailor interventions aimed at improving nutritional status of adolescents involving both mothers and fathers.

Parental occupation was found to be associated with stunting and thinness, most likely because of associated incomes. Paternal occupation in business and daily wage or retirement, related to lower and more unstable incomes, increased girls' and older adolescents' odds of stunting. At the same time, it decreased the odds of thinness in younger adolescents. Again, these outcomes could be explained by the buffer hypothesis (Hadley *et al.*, 2008; Harris-Fry *et al.*, 2017; Moffitt & Ribar, 2018). Furthermore, paternal foreign employment was found to be associated with increased odds of thinness in boys and in the total population. This is surprising, because male out-migration is a strategy to improve household economic status (Ratha *et al.*, 2011). However, adolescents' work burden outside the house (i.e. in agriculture or other activities) might increase because of their father's absence. For instance, evidence from Mexico suggests that adolescent boys intensify their work outside the house when their fathers migrate, which may lead to increased energy expenditure, absence during meal times and eventually to thinness (Antman, 2011; McKenzie & Rapoport, 2011). Moreover, older adolescents seem

to be affected more by paternal occupation, possibly because they take up more responsibilities either in addition to, or instead of, going to school.

At the community level, residence in the *Terai* appears to increase thinness odds in adolescents. This might be explained by dependency on imported and purchased food, which is highly susceptible to seasonal availability of food and price fluctuations (Krishnamurthy *et al.*, 2013). When comparing stunting prevalence over the years, the *Terai* made the least progress in reducing stunting, potentially due to higher population density and lack of basic health and education services (Nepali *et al.*, 2019). Mountain residence as a risk factor for stunting in all but girls and older adolescents, might be explained by a lack of conditions required for healthy growth, including household food security, care and health from an early age (Krishnamurthy *et al.*, 2013). Similar to a recent study on stunting in Nepal, residence in the Mid-Western region, the least developed and geographically most isolated region, showed higher stunting odds among all but the younger adolescents, possibly associated with the lack of infrastructure, health services and sanitation, social exclusion and high rates of male out-migration (Nepali *et al.*, 2019).

The findings of this study could support integrated and context-specific public health interventions to improve adolescent health and present an opportunity for solid policymaking to facilitate a healthy transition into adulthood. Stunting and thinness may lead to the failure of realizing adolescents' social, economic and developmental potential, which is why it is essential to consider both immediate and underlying determinants across the lifespan and focus on environmental, societal and equity issues. Stunting in particular needs careful interpretation as it is not merely an indicator of undernutrition, and nutrition interventions alone will not improve the nutrition and health status, including stature, of present and future adolescents and generations (Leroy & Frongillo, 2019). Multi-component interventions should also address empowerment and gender equality and involve girls, boys, their parents, schools and communities, while considering household determinants and cultural variation. Moreover, this study's findings show clear differences in determinants of stunting and thinness between younger and older adolescents. Thus, it is vital that we acknowledge that adolescence is indeed a period of rapid development and that interventions should be tailored to specific age groups.

Methodological limitations of this study relate first of all to its cross-sectional design, which did not allow conclusions about causality to be drawn. Secondly, the extended number of independent variables and the relatively small population of around 3700 adolescents increased the likelihood of type I errors, incorrectly showing the differences between groups. Third, self-report might have led to response bias, particularly in the case of parental education and occupation questions. Fourth, despite the multi-cluster sampling design, the study population might not fully represent all Nepalese adolescents. For instance, the low percentage of married adolescents (1.8% boys and 3.6% girls) differs from national DHS 2016 data (6.4% boys and 27.1% girls aged 15–19) (MoH *et al.*, 2017). The difference might be explained by the fact that NANS intended to solely sample unmarried adolescents. However, some married adolescents were interviewed by mistake. Finally, although anthropometric data and the 2007 growth reference are commonly accepted and frequently used as indicators of undernutrition in the adolescent population, they have their limitations, particularly in the context of cross-sectional data, as recently highlighted by Tumilowicz *et al.* (2019).

There were also limitations related to challenges in determining the actual prevalence estimates in the study population due to ethnic and racial variations and differences in growth spurts and potential and body composition. To fully understand age- and sex-specific differences, data on biological maturation and the adolescent growth spurt, such as age of menarche and peak height velocity, are essential, as there might be variability between the study population and the reference population in terms of growth trajectories (Tumilowicz *et al.*, 2019). The WHO cut-offs might be problematic, especially for older adolescents as growth in both height and weight can continue into late adolescence, which may result in an overestimation of stunting among boys compared with girls. Moreover, the study's thinness prevalence of girls might be underestimated due to sex

differences in BAZ SD scores. Hence, in order to assess and combat adolescent undernutrition and to determine intervention points, additional nutritional indicators, such as dietary intake and quality, are important, but data are limited (Tumilowicz *et al.*, 2019). Moreover, proxies of nutritional status, including distal and proximate determinants such as environmental and macro-level factors related to agriculture, infrastructure, health services and food systems, as well as social and cultural norms and practices that favour girls or boys, younger or older adolescents, are vital in order to better understand and address regional and age- and sex-specific differences between stunting and thinness (WHO, 2018; Tumilowicz *et al.*, 2019). Such research would be ideally based on longitudinal data and adopt a life course perspective to explore the mechanisms and causal pathways through which sociocultural and economic determinants influence adolescent nutrition within their everyday-life setting.

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