

# THE STRONG LIMIT THEOREM FOR RELATIVE ENTROPY DENSITY RATES BETWEEN TWO ASYMPTOTICALLY CIRCULAR MARKOV CHAINS

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In this paper, we are going to study the strong limit theorem for the relative entropy density rates between two finite asymptotically circular Markov chains. Firstly, we prove some lemmas on which the main result is based. Then, we establish two strong limit theorems for non-homogeneous Markov chains. Finally, we obtain the main result of this paper. As corollaries, we get the strong limit theorem for the relative entropy density rates between two finite non-homogeneous Markov chains. We also prove that the relative entropy density rates between two finite non-homogeneous Markov chains are uniformly integrable under some conditions.

**Keywords:** asymptotic equipartition property, asymptotically circular Markov chains, relative entropy density rate, strong law of large numbers

## 1. INTRODUCTION

Let  $S = \{1, 2, \dots, N\}$  be a finite state space,  $\{X_n, n \geq 0\}$  be a sequence of  $S$ -valued random variables defined on measurable space  $(\Omega, \mathcal{F})$ . Let  $\mathbf{P}$  be a probability measure on  $(\Omega, \mathcal{F})$ , the joint distributions of  $\{X_n, n \geq 0\}$  under  $\mathbf{P}$  be

$$p(x_0, \dots, x_n) = \mathbf{P}\{X_0 = x_0, \dots, X_n = x_n\}, \quad x_i \in S, \quad 0 \leq i \leq n, \quad n \geq 0. \quad (1)$$

Let

$$f_n(\omega) = -\frac{1}{n} \log p(X_0, \dots, X_n), \quad (2)$$

where  $\log$  is the natural logarithm.  $f_n(\omega)$  is called the entropy density of  $\{X_i, 0 \leq i \leq n\}$ , and  $\omega$  is a sample point in  $\Omega$ .

Let  $\mathbf{Q}$  be another probability measure on  $(\Omega, \mathcal{F})$ , the joint distributions of  $\{X_n, n \geq 0\}$  under  $\mathbf{Q}$  be

$$q(x_0, \dots, x_n) = \mathbf{Q}\{X_0 = x_0, \dots, X_n = x_n\}, \quad x_i \in S, \quad 0 \leq i \leq n, \quad n \geq 0. \tag{3}$$

Let

$$L_n(\omega) = -\frac{1}{n} \log \frac{q(X_0, \dots, X_n)}{p(X_0, \dots, X_n)}. \tag{4}$$

$L_n(\omega)$  is called relative entropy density rate between  $\mathbf{P}_{x_0^n}$  and  $\mathbf{Q}_{x_0^n}$  (see [9, p. 93]). It is easy to see that

$$R_n \hat{=} E_{\mathbf{P}}(L_n(\omega)) = \frac{1}{n} \sum_{x_0^n} p(x_0^n) \log \frac{p(x_0^n)}{q(x_0^n)} = \frac{1}{n} D(\mathbf{P}_{x_0^n} \parallel \mathbf{Q}_{x_0^n}), \tag{5}$$

where  $X_0^n = (X_0, \dots, X_n)$ ,  $x_0^n$  is the realization of  $X_0^n$ .  $D(\mathbf{P}_{x_0^n} \parallel \mathbf{Q}_{x_0^n})$  is relative entropy between  $\mathbf{P}_{x_0^n}$  and  $\mathbf{Q}_{x_0^n}$ .  $R_n$  is called relative entropy rate between  $\mathbf{P}_{x_0^n}$  and  $\mathbf{Q}_{x_0^n}$ . If  $\lim_{n \rightarrow \infty} R_n = R$  exists,  $R$  is called the relative entropy rate between  $\mathbf{P}$  and  $\mathbf{Q}$  (see [9, p. 67]).

*Remark 1:* It is easy to see that if  $\{L_n(\omega), n \geq 0\}$  are  $\mathbf{P}$ -uniformly integrable and  $\lim_{n \rightarrow \infty} L_n(\omega) = L$ ,  $\mathbf{P}$ -a.e., then  $R = L$ .

Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain under  $\mathbf{P}$  with initial distribution

$$(p(1), p(2), \dots, p(N)), \tag{6}$$

and transition matrices

$$P_n = (p_n(i, j)), \quad i, j \in S, \quad n \geq 1, \tag{7}$$

where  $p_n(i, j) = \mathbf{P}(X_n = j | X_{n-1} = i)$ . Then

$$p(x_0, \dots, x_n) = p(x_0) \prod_{k=1}^n p_k(x_{k-1}, x_k). \tag{8}$$

In this case,

$$f_n(\omega) = -\frac{1}{n} \left[ \log p(X_0) + \sum_{k=1}^n \log p_k(X_{k-1}, X_k) \right]. \tag{9}$$

Let  $\{X_n, n \geq 0\}$  be also a non-homogeneous Markov chain under  $\mathbf{Q}$  with initial distribution (6) and transition matrices

$$Q_n = (q_n(i, j)), \quad i, j \in S, \quad n \geq 1, \tag{10}$$

where  $q_n(i, j) = \mathbf{Q}(X_n = j | X_{n-1} = i)$ . Then

$$q(x_0, \dots, x_n) = p(x_0) \prod_{k=1}^n q_k(x_{k-1}, x_k). \tag{11}$$

In this case,

$$L_n(\omega) = -\frac{1}{n} \sum_{k=1}^n \log \frac{q_k(X_{k-1}, X_k)}{p_k(X_{k-1}, X_k)}. \tag{12}$$

DEFINITION 1 (see [26]): Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (7). Let  $T_1, T_2, \dots, T_d$  be  $d$  transition matrices, where  $T_l = (t_l(i, j))$ ,  $l = 1, 2, \dots, d, i, j \in S$ . If

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=0}^n |p_{td+l}(i, j) - t_l(i, j)| = 0, \quad l = 1, 2, \dots, d, \quad \forall i, j \in S, \tag{13}$$

this Markov chain is called an asymptotically circular Markov chain. If, in particular,

$$P_{td+l} = T_l, \quad l = 1, 2, \dots, d, \quad t = 0, 1, 2, \dots, \tag{14}$$

this Markov chain is called a circular Markov chain (see [5]).

A question of importance in information theory is the convergence of  $f_n(\omega)$  to a constant in a sense ( $L_1$  convergence, convergence in probability, *a.e.* convergence), namely, the asymptotic equipartition property (AEP) of an information source. Shannon [20] first proved the AEP for convergence in probability for an ergodic homogeneous Markov information source. McMillan [18] and Breiman [4] obtained, for a stationary ergodic information source, the AEP in  $L_1$  and *a.e.* convergence, respectively. This is the famous Shannon–McMillan–Breiman theorem. The extensions of the Shannon–McMillan–Breiman theorem can be found, for example, in Barron [3], Chung [7], Kieffer [14], or Algoet and Cover [1]. Yang [22] established the AEP for a class of non-homogeneous Markov chains. Yang and Liu [23] showed the AEP for  $m$ th order non-homogeneous Markov chains. Zhong Yang and Liang [26] studied the AEP for asymptotically circular Markov chains.

Relative entropy was first defined by Kullback and Leibler [15]. It is known under a variety of names, including the Kullback–Leibler distance, cross-entropy, information divergence, and information for discrimination, and it was studied in detail by Csiszar [8] and Amari [2]. The relative entropy between two random variables was developed to two sequences of variables called the relative entropy rate [9, p. 67]) and it is used for comparing two stochastic processes. The relative entropy rate is a natural and useful measure of distance between two stochastic processes, and it also plays an important role in the statistical hypothesis testing theory and coding theory. The relative entropy rate has several natural interpretations. For example, it represents the discriminative power of one distribution over the other. Thus, if data are generated by the distribution  $\mathbf{P}$ , the relative entropy rate  $R$  represents the average difference in the likelihood score per-symbol between  $\mathbf{P}$  and  $\mathbf{Q}$ . The relative entropy rate is also the average loss-per-symbol when compressing the data, assuming (erroneously) it was generated by distribution  $\mathbf{Q}$ . Lai and Ford [16] studied the relative entropy rate-based Multiple Hidden Markov Model. Kesidis and Walrand [13] derived the relative entropy rate between two Markov transition matrices. Chazottes, Giardina, and Redig [6] applied it for comparing two Markov chains. Recently, Yari and Nikooravesh [25] obtained the relative entropy rate between a Markov chain and its corresponding hidden Markov chain.

There are also some works about relative entropy density rate. Gray [9, p. 94] studied a limit theorem for the relative entropy density rate between two distributions  $\mathbf{P}$  and  $\mathbf{Q}$ . Jia, Chen, and Lin [12] applied the relative entropy density rate to intrusion detection models. Yang et al. [24] studied the strong law of large numbers for the generalized relative entropy density rate of non-homogeneous Markov chains.

In this paper, we are going to prove that the relative entropy density rates between two asymptotically circular Markov chains converges to a constant  $\mathbf{P}$  – *a.e.* under some conditions. As corollary, we obtain that the relative entropy density rates between two

non-homogeneous Markov chains converges to a constant  $\mathbf{P}$  – a.e. We also prove that the relative density rates between two non-homogeneous Markov chains are uniformly  $\mathbf{P}$ -integrable under the some conditions. Observe that the relative entropy rate is found by integrating the relative entropy density rate, as we have pointed out as above, this constant is just the relative entropy rate between the distributions of two probability measures.

The paper is organized as follows. In Section 2, we prove two strong limit theorem for non-homogeneous Markov chains and some lemmas which are the basis of the main result. In Section 3, we obtain the main result of this paper – the relative entropy density rates between two asymptotically circular Markov chains converges to a constant  $\mathbf{P}$  – a.e. and in  $L_1$  under some conditions.

## 2. SOME LEMMAS

In this section, we will prove some lemmas which are the basis of the main result of this paper.

LEMMA 1: *Let  $D$  be a domain in the plane,  $\varphi(x, y)$  a bounded function defined on  $D$ ,  $(x_0, y_0)$  a interior point in  $D$ , and  $\varphi(x, y)$  continue at  $(x, y) = (x_0, y_0)$ . Let  $\{(x_k, y_k), k \geq 1\}$  be a sequence of points in  $D$ . If*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n |x_k - x_0| = 0, \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n |y_k - y_0| = 0, \tag{15}$$

then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n |\varphi(x_k, y_k) - \varphi(x_0, y_0)| = 0. \tag{16}$$

In reference [24], the authors of that paper state the generalized form of this lemma, but do not provide detailed proof. In the following, we will prove the detailed proof of this lemma.

PROOF OF LEMMA 1: It is easy to see that (15) is equivalent to the following equation:

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \|(x_k, y_k) - (x_0, y_0)\| \hat{=} \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \sqrt{(x_k - x_0)^2 + (y_k - y_0)^2} = 0. \tag{17}$$

Since  $\varphi(x, y)$  is continuous at  $(x_0, y_0)$ ,  $\forall \varepsilon > 0, \exists \delta > 0$  whenever  $\|(x, y) - (x_0, y_0)\| \leq \delta$ , we have  $|\varphi(x, y) - \varphi(x_0, y_0)| \leq \varepsilon$ . Let  $N_n(\delta)$  be the number of terms which are greater than  $\delta$  in first  $n$  terms of the sequence  $\{\|(x_k, y_k) - (x_0, y_0)\|, k \geq 1\}$ , and  $M_n(\varepsilon)$  be the number of terms which are greater than  $\varepsilon$  in first  $n$  terms of the sequence  $\{|\varphi(x_k, y_k) - \varphi(x_0, y_0)|, k \geq 1\}$ . It is easy to see that

$$M_n(\varepsilon) \leq N_n(\delta). \tag{18}$$

Let  $M$  be a upper bound of the function  $|\varphi(x, y)|$ , it is easy to prove that

$$\frac{\delta}{n} N_n(\delta) \leq \frac{1}{n} \sum_{k=1}^n \|(x_k, y_k) - (x_0, y_0)\|, \tag{19}$$

and

$$\frac{1}{n} \sum_{k=1}^n |\varphi(x_k, y_k) - \varphi(x_0, y_0)| \leq \varepsilon + \frac{2M}{n} M_n(\varepsilon). \tag{20}$$

This lemma follows from (18) – (20) and the arbitrariness of  $\varepsilon$ . ■

Let  $R$  be a finite transition matrix, if each row of  $R$  is the same, we call it a constant transition matrix.

A finite transition matrix  $P$  is said to be  $C$ -strongly ergodic [11, p. 194], if there exists a finite constant transition matrix  $R$  such that

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n P^k = R.$$

*Remark 2:* It is easy to see that for finite transition matrix, irreducibility implies  $C$ -strongly ergodic [19, p. 104], but the converse is false (see [17]).

LEMMA 2 (see [17,26]): Let  $T_1, T_2, \dots, T_d$  be  $d$  finite transition matrices as in Definition 1,  $R_1 = T_1 T_2 \dots T_d$ ,  $R_2 = T_2 T_3 \dots T_d T_1$ ,  $\dots$ ,  $R_d = T_d T_1 \dots T_{d-1}$ . If  $R_1$  is  $C$ -strongly ergodic, then  $R_2, R_3, \dots, R_d$  are also  $C$ -strongly ergodic, and let  $r_i^{(k)}(i, j)$  be the  $k$ -step transition probability determined by the transition matrix  $R_l$ , then

$$\lim_{m \rightarrow \infty} \frac{1}{m} \sum_{k=1}^m r_i^{(k)}(i, j) = \pi_j^l, \quad l = 1, 2, \dots, d, \tag{21}$$

where  $(\pi_1^l, \pi_2^l, \dots, \pi_N^l)$  is the unique stationary distribution determined by the transition matrix  $R_l$ .

Let  $\delta_j(\cdot)$  be the Kronecker delta function in  $S$ , that is,

$$\delta_j(i) = \begin{cases} 1, & \text{if } i = j, \\ 0, & \text{if } i \neq j, \end{cases} \quad j \in S$$

Let  $S_n^l(j, \omega)$  be the number of  $j$  in the  $td + (l - 1)$  items of the sequence  $X_0(\omega), X_1(\omega), \dots, X_{n-1}(\omega)$ , where  $l = 1, 2, \dots, d, t = 0, 1, 2, \dots$ . It is easy to see that

$$S_n^l(j, \omega) = \sum_{t=0}^{\lfloor n/d \rfloor - 1} \delta_j(X_{td+l-1}) \tag{22}$$

where  $\lfloor a \rfloor$  is the largest integer less than  $a$ .

LEMMA 3 (see [17,26]): Let  $\{X_n, n \geq 0\}$  be an asymptotically circular Markov chain under probability measure  $\mathbf{P}$  defined by Definition 1,  $R_1, R_2, \dots, R_d$  be as in Lemma 2. Let  $S_n^l(j, \omega)$  be the number of  $j$  in the  $td + (l - 1)$  items of the sequence  $X_0(\omega), X_1(\omega), \dots, X_{n-1}(\omega)$ , where  $l = 1, 2, \dots, d, t = 0, 1, 2, \dots$  defined by (22). If  $R_1$  is  $C$ -strongly ergodic, then

$$\lim_{n \rightarrow \infty} \frac{S_n^l(i, \omega)}{n} = \frac{\pi_i^l}{d}, \quad \mathbf{P} - \text{a.e.}, \tag{23}$$

where  $(\pi_1^l, \pi_2^l, \dots, \pi_N^l)$  is the unique stationary distribution determined by the transition matrix  $R_l$ .

LEMMA 4: Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (7) under probability measure  $\mathbf{P}$ , a non-homogeneous Markov chain with initial distribution (6) and transition matrices

$$Q_n = (q_n(i, j)), \quad q_n(i, j) \geq \tau, \quad 0 < \tau < 1, \quad i, j \in S, \quad n \geq 0, \tag{24}$$

under probability measure  $\mathbf{Q}$ . Then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \left\{ \log \frac{p_k(X_{k-1}, X_k)}{q_k(X_{k-1}, X_k)} - \sum_{j=1}^N p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} \right\} = 0, \quad \mathbf{P} - \text{a.e.} \tag{25}$$

PROOF: By the Theorem 2 of [22] and (9), we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \left\{ \log p_k(X_{k-1}, X_k) - \sum_{j=1}^N p_k(X_{k-1}, j) \log p_k(X_{k-1}, X_k) \right\} = 0, \quad \mathbf{P} - \text{a.e.} \tag{26}$$

Let  $f_k(s, t) = \log q_k(s, t)$  in Theorem 1 of [22]. Since

$$\begin{aligned} E_{\mathbf{P}} \left[ (\log q_k(X_{k-1}, X_k))^2 | X_{k-1} \right] &= \sum_{j=1}^N (\log q_k(X_{k-1}, j))^2 p_k(X_{k-1}, j) \\ &\leq \sum_{j=1}^N (\log q_k(X_{k-1}, j))^2 \leq N(\log \tau)^2. \end{aligned} \tag{27}$$

By Theorem 1 of [22], we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \left\{ \log q_k(X_{k-1}, X_k) - \sum_{j=1}^N p_k(X_{k-1}, j) \log q_k(X_{k-1}, X_k) \right\} = 0, \quad \mathbf{P} - \text{a.e.} \tag{28}$$

Then (25) follows from from (26) and (28). ■

LEMMA 5: Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (7) under probability measure  $\mathbf{P}$ , a non-homogeneous Markov chain with initial distribution (6) and transition matrices (10) under probability measure  $\mathbf{Q}$ . If

$$A = \sup_{n \geq 1, i, j \in S} \frac{p_n(i, j)}{q_n(i, j)} < \infty \quad \text{and} \quad B = \sup_{n \geq 1, i, j \in S} \frac{q_n(i, j)}{p_n(i, j)} < \infty, \tag{29}$$

then (25) holds.

PROOF: Let

$$Y_n = \log \frac{p_n(X_{n-1}, X_n)}{q_n(X_{n-1}, X_n)} - \sum_{j=1}^n p_n(X_{n-1}, j) \log \frac{p_n(X_{n-1}, j)}{q_n(X_{n-1}, j)}, \quad n \geq 1, \tag{30}$$

and  $Y_0 = 0$ . It is easy to see that  $\{Y_n, n \geq 0\}$  is a  $\mathbf{P}$ -martingale difference sequence. Since by (29), we have

$$\begin{aligned} E_{\mathbf{P}}[Y_n^2 | X_0, \dots, X_{n-1}] &\leq E_{\mathbf{P}}[(\log \frac{p_n(X_{n-1}, X_n)}{q_n(X_{n-1}, X_n)})^2 | X_{n-1}] \\ &\leq \max\{(\log A)^2, (\log B)^2\} \doteq C \end{aligned} \tag{31}$$

By Chow’s strong law of large number for martingale difference sequence (see [10, p. 35]), we have (25) holds. ■

*Example 1:* Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain taking values in  $\{1, 2\}$  with the transition matrices

$$P_n = \begin{pmatrix} \frac{1}{n} & 1 - \frac{1}{n} \\ \frac{1}{n} & 1 - \frac{1}{n} \end{pmatrix} \tag{32}$$

under  $\mathbf{P}$ , and let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain taking values in  $S = \{1, 2\}$  with the transition matrices

$$Q_n = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix} \tag{33}$$

under  $\mathbf{Q}$ . It is easy to see that the condition (24) is satisfied, but the condition (29) is not satisfied.

*Example 2:* Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain taking values in  $\{1, 2\}$  with the transition matrices (32) under  $\mathbf{P}$ , and let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain taking values in  $S = \{1, 2\}$  also with the transition matrices (32) under  $\mathbf{Q}$ . It is easy to see that the condition (29) is satisfied, but the condition (24) is not satisfied.

From above two examples, we know that Lemma 4 does not imply Lemma 5, and Lemma 5 does not imply Lemma 4. So above two lemmas are not overlapping.

LEMMA 6: *Under the conditions of Lemmas 4 or 5, let  $L_n(\omega)$  be defined by (12). Then  $L_n(\omega)$  are  $\mathbf{P}$ -uniformly integrable.*

PROOF: By (9) and (12), we have

$$L_n(\omega) = -f_n(\omega) - \frac{1}{n} \log p(X_0) - \frac{1}{n} \sum_{k=1}^n \log q_k(X_{k-1}, X_k). \tag{34}$$

If the condition of Lemma 4 holds, since  $f_n(\omega)$  are uniformly  $\mathbf{P}$ -integrable (see [10, Lemma 3.7] or [21, Lemma 3]), and  $\frac{1}{n} \log p(X_0) + \frac{1}{n} \sum_{k=1}^n \log q_k(X_{k-1}, X_k)$  are bounded, by (34),  $L_n(\omega)$  are  $\mathbf{P}$ -uniformly integrable. If the condition of Lemma 5 holds, It is easy to see that in this case  $L_n(\omega)$  are bounded, so they are also  $\mathbf{P}$ -uniformly integrable. ■

### 3. THE MAIN RESULT

In this section, we will prove the main result of this paper.

**THEOREM 1:** *Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (7) under probability measure  $\mathbf{P}$ . and  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chains with the initial distribution (6) and transition matrices (10) under probability measure  $\mathbf{Q}$ . Assume that one of the conditions of (24) or (29) hold. We further assume that  $\{X_n, n \geq 0\}$  is an asymptotically circular Markov chain under  $\mathbf{P}$  defined by Definition 1, that is (13) holds, and  $\{X_n, n \geq 0\}$  is also an asymptotically circular Markov chain under  $\mathbf{Q}$ , that is, there exist  $d$  transition matrices  $H_l = (h_l(i, j))$ ,  $l = 1, 2, \dots, d$ ,  $i, j \in S$  such that*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=0}^n |q_{td+l}(i, j) - h_l(i, j)| = 0, \quad l = 1, 2, \dots, d, \quad \forall i, j \in S. \tag{35}$$

Let  $R_l, l = 1, 2, \dots, d$  be defined as in Lemma 2. Assume that  $R_l$  is  $C$ -strongly ergodic. Let  $L_n(\omega)$  defined by (12). Then

$$\lim_{n \rightarrow \infty} L_n(\omega) = \sum_{l=1}^d \sum_{i=1}^N \sum_{j=1}^N \frac{\pi_i^l}{d} t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)}, \quad \mathbf{P} - \text{a.e. and } L_1, \tag{36}$$

where  $(\pi_1^l, \pi_2^l, \dots, \pi_N^l)$  is the unique stationary distribution determined by the transition matrices  $R_l, l = 1, 2, \dots, d$ .

*Remark 3:* From Lemma 6 we know that  $L_n(\omega)$  are  $\mathbf{P}$ -uniformly integrable under the conditions of Theorem 1, so the constant of right-hand side of (36) is also the relative entropy rate between two asymptotically circular Markov chains.

**PROOF OF THEOREM 1.1:** If one of the conditions (24) or (29) hold. By (12), Lemmas 4 or 5, we have

$$\lim_{n \rightarrow \infty} \left\{ L_n(\omega) - \frac{1}{n} \sum_{k=1}^n \sum_{j=1}^N p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} \right\} = 0, \quad \mathbf{P} - \text{a.e.}, \tag{37}$$

Now

$$\begin{aligned} &= \left| \frac{1}{n} \sum_{k=1}^n \sum_{j=1}^N p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} - \sum_{l=1}^d \sum_{i=1}^N \sum_{j=1}^N \frac{\pi_i^l}{d} t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \\ &= \left| \frac{1}{n} \sum_{j=1}^N \sum_{k=1}^{d \lfloor \frac{n}{d} \rfloor} p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} + \frac{1}{n} \sum_{j=1}^N \sum_{k=d \lfloor \frac{n}{d} \rfloor + 1}^n p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} \right. \\ &\quad \left. - \sum_{l=1}^d \sum_{i=1}^N \sum_{j=1}^N \frac{\pi_i^l}{d} t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \end{aligned}$$



$$\begin{aligned}
 &\leq \left| \frac{1}{n} \sum_{j=1}^N \sum_{l=1}^d \sum_{t=0}^{\lfloor \frac{n}{d} \rfloor - 1} \sum_{i=1}^N \delta_i(X_{td+l-1}) p_{td+l}(i, j) \log \frac{p_{td+l}(i, j)}{q_{td+l}(i, j)} \right. \\
 &\quad \left. - \frac{1}{n} \sum_{j=1}^N \sum_{l=1}^d \sum_{t=0}^{\lfloor \frac{n}{d} \rfloor - 1} \sum_{i=1}^N \delta_i(X_{td+l-1}) t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \\
 &\quad + \left| \frac{1}{n} \sum_{j=1}^N \sum_{l=1}^d \sum_{t=0}^{\lfloor \frac{n}{d} \rfloor - 1} \sum_{i=1}^N \delta_i(X_{td+l-1}) t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} - \sum_{l=1}^d \sum_{i=1}^N \sum_{j=1}^N \frac{\pi_i^l}{d} t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \\
 &\quad + \left| \frac{1}{n} \sum_{j=1}^N \sum_{k=d\lfloor \frac{n}{d} \rfloor + 1}^n p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} \right| \\
 &\leq \sum_{j=1}^N \sum_{i=1}^N \sum_{l=1}^d \frac{1}{n} \sum_{t=0}^n \left| p_{td+l}(i, j) \log \frac{p_{td+l}(i, j)}{q_{td+l}(i, j)} - t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \\
 &\quad + \sum_{j=1}^N \sum_{l=1}^d \sum_{i=1}^N \left| \frac{S_n^l(i, \omega)}{n} - \frac{\pi_i^l}{d} \right| \cdot \left| t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| \\
 &\quad + \left| \frac{1}{n} \sum_{j=1}^N \sum_{k=d\lfloor \frac{n}{d} \rfloor + 1}^n p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} \right|. \tag{38}
 \end{aligned}$$

If (24) holds, by (24) and (35), it is easy to prove that  $h_l(i, j) \geq \tau, \forall i, j \in S$ . Letting  $\varphi(x, y) = x \log \frac{x}{y}$  ( $\varphi(0, y) = 0$ ) and  $D_1 = [0, 1] \times [\tau, 1]$  in Lemma 1, it is easy to see that  $\varphi(x, y)$  is bounded on  $D$  and continuous at  $(t_l(i, j), h_l(i, j))$ . By (13), (35) and Lemma 1, we have  $\forall i, j \in S$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=0}^n \left| p_{td+l}(i, j) \log \frac{p_{td+l}(i, j)}{q_{td+l}(i, j)} - t_l(i, j) \log \frac{t_l(i, j)}{h_l(i, j)} \right| = 0. \tag{39}$$

If (29) holds, by (13) and (35), it is easy to prove that for  $l = 1, 2, \dots, d$  and  $\forall i, j \in S$

$$\frac{t_l(i, j)}{h_l(i, j)} \leq A < \infty \text{ and } \frac{h_l(i, j)}{t_l(i, j)} \leq B < \infty.$$

It is clear that  $\varphi(x, y) = x \log \frac{x}{y}$  is bounded and continuous on the set  $D_2 = \{(x, y) : \frac{1}{B} \leq \frac{x}{y} \leq A, 0 \leq x \leq 1\}$  (It is easy to see that  $B > 0$ ). By (13), (35) and Lemma 1, we also have (39) holds. If one of the conditions (24) or (29) hold, it is easy to prove that  $p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)}$  are bounded. So we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^N \sum_{k=d\lfloor \frac{n}{d} \rfloor + 1}^n p_k(X_{k-1}, j) \log \frac{p_k(X_{k-1}, j)}{q_k(X_{k-1}, j)} = 0. \tag{40}$$

By Lemma 3, (37)–(40), (36) holds  $\mathbf{P}$  – a.e. By Lemma 6,  $L_n(\omega)$  are uniformly integrable, (36) also holds in  $L_1$  convergence. ■

COROLLARY 1 see [24]: Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (7) under probability measure  $\mathbf{P}$ . Let

$P = (p(i, j)), i, j \in S$  be another transition matrix and assume that  $P$  be irreducible. Assume that

$$\lim_{n \rightarrow \infty} \frac{1}{n+1} \sum_{k=0}^n |p_k(i, j) - p(i, j)| = 0, \quad \forall i, j \in S. \quad (41)$$

Let  $\{X_n, n \geq 0\}$  be a non-homogeneous Markov chain with initial distribution (6) and transition matrices (10) under measure  $\mathbf{Q}$  such that (24) hold. Let  $Q = (q(i, j)), i, j \in S$  be another transition matrix. Assume that

$$\lim_{n \rightarrow \infty} \frac{1}{n+1} \sum_{k=0}^n |q_k(i, j) - q(i, j)| = 0, \quad \forall i, j \in S. \quad (42)$$

Then

$$\lim_{n \rightarrow \infty} L_n(\omega) = \sum_{i=1}^m \sum_{j=1}^m \pi_i p(i, j) \log \frac{p(i, j)}{q(i, j)}, \quad \mathbf{P} - \text{a.e. and } L_1, \quad (43)$$

where  $(\pi_1, \pi_2, \dots, \pi_N)$  is the unique stationary distribution determined by the transition matrix  $P$ .

PROOF: Since irreducibility implies  $C$ -strongly ergodic (see [19, p. 104]). Letting  $d = 1$  in Theorem 1, this corollary follows.  $\blacksquare$

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