

MEASUREMENT AND TRANSFER OF CATASTROPHIC RISKS. A SIMULATION ANALYSIS

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ABSTRACT

When analyzing catastrophic risk, traditional measures for evaluating risk, such as the probable maximum loss (PML), value at risk (VaR), tail-VaR, and others, can become practically impossible to obtain analytically in certain types of insurance, such as earthquake, and certain types of reinsurance arrangements, specially non-proportional with reinstatements. Given the available information, it can be very difficult for an insurer to measure its risk exposure. The transfer of risk in this type of insurance is usually done through reinsurance schemes combining diverse types of contracts that can greatly reduce the extreme tail of the cedant's loss distribution. This effect can be assessed mathematically. The PML is defined in terms of a very extreme quantile. Also, under standard operating conditions, insurers use several "layers" of non proportional reinsurance that may or may not be combined with some type of proportional reinsurance. The resulting reinsurance structures will then be very complicated to analyze and to evaluate their mitigation or transfer effects analytically, so it may be necessary to use alternative approaches, such as Monte Carlo simulation methods. This is what we do in this paper in order to measure the effect of a complex reinsurance treaty on the risk profile of an insurance company. We compute the pure risk premium, PML as well as a host of results: impact on the insured portfolio, risk transfer effect of reinsurance programs, proportion of times reinsurance is exhausted, percentage of years it was necessary to use the contractual reinstatements, etc. Since the estimators of quantiles are known to be biased, we explore the alternative of using an Extreme Value approach to complement the analysis.

KEYWORDS

Quantile, Extreme Value, Monte Carlo Methods, PML, VAR, Reinsurance.

1. INTRODUCTION

The measurement and transfer of risk are at the essence of the insurance business. This has prompted the development of quantitative techniques to achieve both. They are important for all the stakeholders in the industry: the

direct insurer, a potential reinsurer, regulators, rating agencies, and consumers. In the case of catastrophic risks (defined for the purpose of this paper as those with low frequency and high severity), they become particularly relevant due to the magnitude of potential losses, Woo (1999). A large earthquake or hurricane (or sequence of them) will impact losses in an extreme fashion, such that if not adequately reserved and capitalized, or covered by reinsurance or retrocession, it can cause the ruin of either the insurer or the reinsurer, with 'catastrophic' consequences for stockholders and society. Hence it is important to measure this kind of risk and evaluate how it is transferred.

Mexico is a country with a large number of earthquakes per year. On average, there are 80 of magnitude larger than 4.3 every year. The available information and models allow us to analyze the information on earthquake intensities over the last 100 years. The National Seismological Service of Mexico, SSN (1999), has published the magnitudes of earthquakes larger than 6.5 on the Richter scale during the 20th century; see Figure 1.

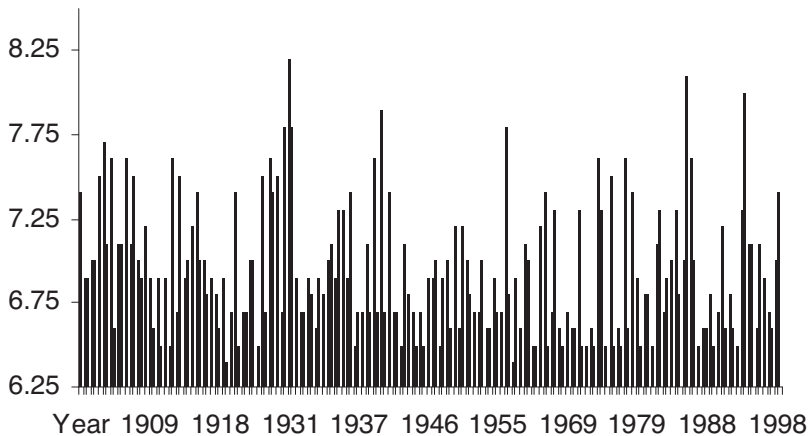


FIGURE 1: Magnitude of Large Earthquakes in Mexico.

It is clear that the correct evaluation of the potential catastrophic losses due to large earthquakes is of great importance. There have been large earthquakes in 1932 (8.2), 1942 (7.9), 1957 (7.8), 1985 (8.1) and 1995 (8.0). The largest losses were due to the 1985 earthquake, as a result of the location of its epicenter off the coast of Acapulco, more than 300 kilometers away from Mexico City, but with catastrophic local effects in Mexico City. By some accounts it was the strongest earthquake to hit Mexico in the twentieth century. It is clearly of great importance to have adequate models to evaluate and measure this kind of risk. Even though Mexico is also exposed to other natural disasters, such as hurricane, we focus on earthquake losses.

Catastrophic Risk Measurement

In many countries earthquake catastrophic risk is measured in terms on the probable maximum loss (PML), an extreme quantile of the corresponding loss distribution. Quantile-based risk measures such as PML (another name for a specific VaR) and tail-VaR have been in use for a long time now to measure risk in insurance contexts, Dowd and Blake (2006), and the PML is the 'official' measure for catastrophic risk in Mexico. Given the available information it can be very difficult for an insurer to measure this risk. Furthermore, since the distribution of losses due to earthquakes for a large portfolio of risks will usually be unknown, the only way to quantify risk may be through simulation. That is the process we follow here.

The methodology for estimating probable maximum loss (PML) for natural catastrophes has slowly evolved over the past few decades from being deterministic to one based on stochastic models, due to the increased use of complex risk transfer agreements, the need to describe in greater detail the vulnerability of an insurance entity's strategy on the commercial, underwriting and risk transfer practices: Regarding its commercial practices it must take into account how increased sales in a given zone increases risk exposure to catastrophes and the impact on capital and risk transfer adequacy; with respect to underwriting it must evaluate how accepting a class of risks, or rejecting another, changes the risk profile; and with regard to its risk transfer it must determine the kind of single or multiple events that will produce the higher retained loss.

The specification of deterministic earthquake models involved civil engineers and geologists. Civil engineers dealt with inherent uncertainty in construction by incorporating safety factors in building design. There was no stochastic element involved. Geologists would identify, for a given zone, faults posing the greatest threat, then they would estimate (guess) the maximum possible earthquake magnitude from each fault considered, the consequential ground shaking severity and, hence, the severity of ground shaking at the site. Once this had been done they would select the highest ground motion value. With these elements the PML could be obtained deterministically in several ways, one might be by choosing the insured losses corresponding to an earthquake of a large magnitude. No probabilistic or stochastic models were involved. For a clear and detailed review of the development process see Woo (2002).

Stochastic earthquake models used for measuring insurer's risk on a gross basis and the magnitude of its transfer vehicles to produce its net risk measurement (after reinsurance), incorporate probabilistic elements into the models of each one of the aspects considered in the loss generating process. Thus there may be a separate stochastic model for the process of earthquake generation, another one for shock wave diffusion from the source of the earthquake to the site of the insured property, a model for the severity of ground shaking at the site where the insured buildings are located, and one for the ensuing damage to buildings with the corresponding loss, and others. Some of these models will be interrelated and their relations must be identified. Once all the models

have been constructed the whole structure can be used to analyze losses from an insured portfolio due to earthquake activity. The analysis might be done analytically but in general it must be done using Monte Carlo simulation methods, due to the complexity of the models, ISAWP (2004).

In all cases, however, a guiding principle has been that any technique for computing PML should reflect the dual, hazard and vulnerability, aspects of loss. The hazard aspect considers, for a given phenomenon, elements such as its source, intensity and trajectory, while vulnerability assesses physical and economic attributes of the covered objects, such as a building and its contents. An essential flaw in deterministic models is that they exclude numerous earthquake sources that collectively contribute to the earthquake risk profile of the portfolio. Not just one, but a considerable number of sites may each produce earthquakes capable of causing major losses to a given portfolio. A multiplicity of hazard sources should be taken into account. The preceding limitation is especially important for hurricanes, where multiple events can hit different covered areas during the term of an insurance contracts portfolio and its reinsurance program. The approach described in this paper is a probabilistic one.

One reason for the slow evolution is due to the fact that the processes of earthquake generation, shock wave diffusion, damage to buildings, etc. are very complex and their interaction makes their analysis even more so. Also, insurance and reinsurance contracts contain diverse conditions to limit the losses of the involved company to those that its capital can cover, so it is necessary to bring together the expertise of geophysicists, structure engineers, actuaries, financial experts and others in order to construct a model that represents the overall process reasonably well. It is interesting to note that, as pointed out by Woo (2002) “*The rate of progress was slowed by the reluctance of regulators...*”.

In statistical terms the PML is the q -th quantile, ξ_q , of the loss distribution, where q is usually large, say 0.998. The exceedance probability is defined as $\Pr\{loss \geq \xi_q\}$ and the corresponding return period is $1/\Pr\{loss \geq \xi_q\}$. Hence if in a loss process the PML is defined with $q = 0.998$ it will correspond to a loss that occurs every 500 years. In catastrophe insurance return periods are usually required to be large so that the corresponding quantiles are very high, Woo (2002), and difficult to compute.

Clark (2002) explains how computer models can be used in estimating catastrophe losses. She points out the various components of such models and stages in their construction and application. Those components are present in the modeling process followed by the firm *ERN Ingenieros Consultores, S.C.* (ERN) commissioned by the Mexican regulator to develop the models used here. However, in addition to the problems inherent in the modeling process and estimation we must also face the fact that the usual risk measures are not coherent McNeil et al. (2005). In particular this is true of the PML and VaR even though they are extensively utilized. It is also known that quantile based measures are biased, Inui et al. (2005) and Kim and Hardy (2007). We will complement the simulation analysis with some coherent risk measures: the Conditional Tail Expectation or tail-VaR, as advocated by Artzner et al. (1999).

McNeil et al. (2005) distinguish several approaches to measuring risk. Here we concentrate on risk measures based on the loss distribution since this is how earthquake catastrophic risk is generally measured. They indicate that one of the main problems in working with distributions is that “even under a stationary environment it is difficult to estimate the loss distribution accurately” (McNeil et al. 2005, page 36). In particular, when analyzing catastrophic risk, traditional measures for evaluating risk, such as the probable maximum loss (PML), value at risk (VaR) (both are quantiles of the distribution), tail-VaR (also known as Conditional Tail Expectation, or CTE) and others can become nearly impossible to obtain analytically in certain types of insurance, such as earthquake where normally losses are due to a main event and a replica. Another element that complicates the analysis, and which is of key importance for the industry, is that a measure of the risk is required for a whole year of coverage, and especially with multiple areas of exposure.

Given the recent trends on risk based management on regulation, methods like the one described in this paper to measure catastrophic risks are much more relevant. To comply with the requirements of regulatory and accounting frameworks, such as Solvency II, that in article 43 on Risk management states the following:

“1. *Insurance and reinsurance undertakings shall have in place an effective risk management system comprising strategies, processes and reporting procedures necessary to identify, measure, monitor, manage and report, on a continuous basis the risks, on an individual and aggregated level, to which they are or could be exposed, and their interdependencies.*

That risk management system shall be effective and well integrated into the organisational structure and in the decision making processes of the insurance or reinsurance undertaking with proper consideration of the persons who effectively run the undertaking or have other key functions.

2. *The risk management system shall cover the risks to be included in the calculation of the Solvency Capital Requirement as set out in Article 101(4) as well as the risks which are not or not fully included in the calculation thereof.*

It shall cover at least the following areas:

(a) underwriting and reserving;

(b) ...

(e) reinsurance and other risk mitigation techniques.

The written policy on risk management referred to in Article 41(3) shall comprise policies relating to points (a) to (e) of the second subparagraph of this paragraph.”

Clearly, it is necessary not only to have the correct measurement, but to have it available for auditing, adequately documented, and with enough and clear elements for communication with, and disclosure to, interested parties

(company management and board, reinsurers, authorities and rating agencies). As we will see, this approach is especially useful on the communication side.

In terms of risk management, correct measurement of the risk of a succession of catastrophic events and not only a single one is a must, specially in countries exposed to both seismological and hydro meteorological dangers, such as the USA, Japan, México, and others.

Risk Transfer

Because of the magnitude of potential losses, risk mitigation in this type of insurance is usually done through diverse types of traditional (proportional and non-proportional reinsurance) and alternative (cat-bonds and the like) risk transfer schemes. To clearly understand the exposure of the portfolio and achieve effective risk mitigation from financial and regulatory points of view, proper measurements of the magnitude of the losses, with and without the risk transfer chosen, are needed.

Again, processes like the one described in this paper are necessary to comply with of regulatory framework and accounting requirements, such as Solvency II that, in article Article 101 on the Calculation of the Solvency Capital Requirement states that:

“1. *The Solvency Capital Requirement shall be calculated in accordance with paragraphs 2 to 5:*

2. *The Solvency Capital Requirement shall be calculated on the presumption that the undertaking will carry on its business as a going concern.*
3. *The Solvency Capital Requirement shall be calibrated so as to ensure that all quantifiable risks to which an insurance or reinsurance undertaking is exposed are taken into account. It shall cover existing business, as well as the new business expected to be written over the next twelve months. With respect to existing business, it shall cover unexpected losses only.*

It shall correspond to the Value-at-Risk of the basic own funds of an insurance or reinsurance undertaking subject to a confidence level of 99,5% over a one-year period.

4. *The Solvency Capital Requirement shall cover at least the following risks:*
 - (a) *non-life underwriting risk;*
 - (b) *life underwriting risk;*
 - (c) *health underwriting risk;*
 - (d) *market risk;*
 - (e) *credit risk;*
 - (f) *operational risk.*

Operational risk as referred to in point (f) of the first subparagraph shall include legal risks, and exclude risks arising from strategic decisions, as well as reputation risks.

5. *When calculating the Solvency Capital Requirement, insurance and reinsurance undertakings shall take account of the effect of risk mitigation techniques, provided that credit risk and other risks arising from the use of such techniques are properly reflected in the Solvency Capital Requirement.”*

Valuation of recoverables from reinsurance contracts and special purpose vehicles on the same basis as the valuation of the contractual obligations, and both capturing the risk profile of the portfolio, can be done as required by Solvency II with procedures like the one described in the present paper.

In terms of risk transfer or mitigation, it is well known that in proportional reinsurance (quota share), the insurer takes a proportion of every loss, so that if X is the random variable that represents gross losses in a given time period, then the loss net of reinsurance is $Y = \alpha X$ where α is the retention rate. Alternatively non-proportional reinsurance (e.g. excess-loss) states that for every loss exceeding a specified threshold or priority (P), the reinsurer will pay the loss up to a certain limit (L), so that for each gross loss occurrence the direct insurer will pay only $\text{Max}\{0, \min(X - P, L)\}$, Booth et al. (1999). This has the effect of truncating the loss distribution. Usually, excess-loss treaties include provisions for coverage reinstatement, after the initial coverage has been used up, in one or more events.

Non-proportional reinsurance can greatly reduce the extreme tail of the cedant's loss distribution. This effect can be assessed mathematically. If the PML is being defined in terms of a very extreme quantile we argue that in simple cases, and if there is a limit to the non-proportional reinsurance, the reduction in the PML can be very unstable, depending on the relation between the limit and the PML. Also, under standard operating conditions, insurers use several “layers” of non proportional reinsurance that will be combined with some type of proportional reinsurance. The resulting reinsurance structures will then be very complicated to analyze, Verlaak and Beirlant (2003). This is further complicated if the probability distribution of losses is not known analytically. In fact most of the literature on optimal reinsurance assumes it is known. Recently, Silvestrov et al. (2006) developed criteria for evaluating alternative reinsurance contracts that are large and mathematically complex. They use a Monte Carlo based approach.

Regulatory Considerations

It has been argued that it is impossible to measure the mitigation effect, or transfer of risk, of non proportional reinsurance and so it should not be given recognition for solvency assessment. The following reasons, among others, have been given to support this:

- a) The difficulty of estimating their effect;
- b) The inclusion of aggregate limits in some of the non proportional reinsurance schemes;

- c) Given the fact that we are concerned not only with exposure to one event, but with exposure to a series of them, it will be absolutely necessary to have “reinstatement” clauses and a proper measure of their adequacy.

Nevertheless, proper recognition of reinsurance is necessary in order to assess the risk reduction for the ceding company. This has implications for capital requirements to ensure effective solvency supervision. The Insurer Solvency Assessment Working Party (ISAWP) of the International Actuarial Association (IAA), ISAWP (2004), states that

“While proper treatments and recognition of reinsurance arrangements are necessary to assess the impact of the of a ceding company’s risk profile, this is a difficult task for a number of reasons.

The first complexity comes from the tremendous diversity in the types of reinsurance contracts:

- *Typical reinsurance arrangement comprise both proportional and non-proportional covers*
- *Some contracts have variable rating terms, ... for a proportional reinsurance treaty, and reinstatements or contingent commissions for an excess-of-loss treaty*
- *Some contracts cover just one line of business, others cover multiple lines of business ...*
- *Some contracts are on an aggregate basis, with aggregate deductibles and aggregate limits*
- *Some financial type reinsurance contracts cover a hybrid of underwriting and financial risks.*

The second complexity comes from the fact that many reinsurance contracts do not bear a linear relationship with the underlying risks.”

The ISAWP further indicates that *“the proper evaluation of the risk reducing impact of non-proportional reinsurance contracts is still not possible without either relatively complex mathematical transformations, which are typically beyond the of supervisory control mechanisms, or the use of simulations, which are standard routines for more complex risk modelling in internal models.”*

In addition the ISAWP also indicates that if applied properly to evaluate the solvency of a direct insurer, reinsurance is a very efficient means of reducing risk (particularly if measured by tail-VaR) and hence can be a useful alternative for capital.

Hence when reinsurance schemes are very sophisticated, it becomes very complicated, if not impossible, to evaluate their mitigation or transfer effects analytically then it may be necessary to use alternative approaches, such as Monte Carlo simulation methods, Silvestrov et al. (2006). That is what we do in this paper in order to measure the effect of a complex reinsurance treaty on the risk profile of an insurance company.

Something that also should be taken into account is that simulations generally produce results that help make better management decisions, improving

communication to different stakeholders, such as underwriters, reinsurers, the company's board and rating agencies.

The Model

In general, models developed to estimate catastrophic losses are based on the physical laws of nature that govern the specific phenomena, in our case earthquake occurrence, and on the equations that embody them. Thus by combining mathematical representations of the natural occurrence patterns and characteristics of earthquakes, with complementary information on property values, construction types, and other characteristics, as well as information on insurance and reinsurance contracts, these models can provide extensive information to companies concerning the potential for large losses before they actually occur.

In Mexico, the insurance regulatory body (Comision Nacional de Seguros y Fianzas, CNSF) commissioned to the firm ERN Ingenieros Consultores, S.C. the construction of an earthquake loss model that must be used to compute the pure risk premium as well as the PML. These results are used to verify compliance with corresponding regulation and compute statutory reserves.

Even better, the software produces additional output that can be used for simulation. These simulation exercises can provide a rich output that can be used for many different applications. Probability distributions of losses and their complement, exceedance probabilities, can be estimated for potential levels of annual aggregate and per-occurrence losses that a company may experience given its portfolio of property exposures, Clark (2002). There are several commercial simulation models (AIR¹, EQECAT², RMS³) that do this. Here we intend to show how a similar model can be used by the insurance companies.

Hence, based on the arguments put forth by international associations, such as the International Association of Insurer Supervisors (IAIS) and the International Actuarial Association, the latter through the ISAWP, that encourage the use of mathematical models and simulation methods, we have used the

¹ AIR Worldwide Corporation (AIR) models the risk from natural catastrophes (among them earthquake) and terrorism in more than 50 countries, including Mexico. AIR introduced its U.S. Earthquake Model to the insurance industry in 1990. Uncertainty is quantified and incorporated throughout the modeling process, creating fully probabilistic estimates of financial loss. SOURCE: <http://www.air-worldwide.com/About-AIR.aspx>.

² EQECAT provides state-of-the-art products and services for managing natural risks. It has developed models for key catastrophe perils that allow quantification and mitigation of their financial consequences by the insurance and reinsurance industries. EQECAT has earthquake and hurricane models for Mexico. SOURCE: http://www.eqecat.com/WC_unbiasedRiskClarity.html

³ RMS (Risk management Solutions) pioneered the development of catastrophe models for insurance markets in the late 1980s, focusing on event-specific probabilistic modeling to quantify risk for individual locations and for portfolios of aggregate risk. RMS catastrophe models are built upon detailed databases describing highly localized variations in hazard characteristics, as well as insurance exposures. SOURCE: <http://www.rms.com/Catastrophe/Models/>

mentioned ERN model output and constructed a program that allows the actuary to generate the distributions of gross and net yearly losses for an insurance portfolio. The algorithm includes the possibility of simulating

- a) the occurrence of one or several earthquakes in a year
- b) their impact on the insured portfolio
- c) the risk transfer effect of reinsurance programs that mix different types or reinsurance
- d) descriptive statistics for gross losses, and losses net of reinsurance
- e) the proportion of times the reinsurance is exhausted
- f) average cost per year of reinstatements
- g) distribution of loss by reinsurance layer according to their magnitude
- h) percentage of years it was necessary to contract additional reinstatements

The model consists of a series of sub-models corresponding to different aspects of the earthquake loss generation, shock wave transfer and impact processes. The initial component is earthquake occurrence. This is modeled as a spatial Poisson distribution for each of a number of potential seismic sites, i.e. space has been discretized in 3600 points. Then there is the distribution of earthquake magnitudes at each one of the sites. Let M be the random variable representing the magnitude, and $\lambda_i(M)$ is the magnitude exceedance rate at source i , and it represents the number of earthquakes of magnitude greater than M at source i . It will typically be modeled as

$$\lambda_i(M) = \lambda_{0i} \frac{e^{-\beta_i M} - e^{-\beta_i M_{ui}}}{e^{-\beta_i M_0} - e^{-\beta_i M_{ui}}} \quad \text{for } M_0 \leq M \leq M_{ui} \tag{1}$$

where λ_{0i} is the number of earthquakes occurring at site i per year, M_0 is the minimum relevant magnitude, and M_{ui} the maximum magnitude that can be observed at the i -th seismic site; the parameters λ_{0i} and β_i need to be estimated; for a detailed explanation of the modeling process see Ordaz et al. (2000), and references therein.

The exceedance rate for the i -th site, $v_i(y)$, is the average number of events, per unit time, that produce losses larger than y from an earthquake at seismic source i . It is specified as:

$$v_i(y) = \int_{M_0}^{M_{ui}} -\frac{d\lambda_i(M)}{dM} \Pr(Y_i > y | M) dM = -\int_{M_0}^{M_{ui}} \Pr(Y_i > y | M) d\lambda_i(M) \tag{2}$$

where Y_i represents the losses to the whole portfolio due to an earthquake at the i -th site. Then the total exceedance rate (The average number of events that produce losses that will exceed a given value y) for the whole portfolio is:

$$v(y) = \sum_{i=1}^{Nf} \int_{M_0}^{M_{ui}} -\frac{d\lambda_i(M)}{dM} \Pr(Y_i > y | M) dM \tag{3}$$

where M_0 and M_{ii} are as defined above and N_f is the number of seismic sites. Given equation (3) we have $\nu(0)$ = average number of events by unit time, that produce losses greater than 0, the probability distribution for the losses of the whole portfolio is

$$F(y) = 1 - \frac{\nu(y)}{\nu(0)}. \quad (4)$$

In equation (3) $\Pr(Y_i > y | M)$ is obtained in the program as follows: given an earthquake of magnitude M , at site i is

$$\Pr(Y_i > y | M) = \Pr(S_{\text{exp}} \beta_i > y | M) = \Pr\left(\beta_i > \frac{y}{S_{\text{exp}}} \middle| M\right)$$

where β_i = relative loss, as a proportion of the total amount exposed S_{exp} , i.e. $\beta_i = \frac{Y_i}{S_{\text{exp}}}$, for $i = 1, 2, \dots, N_f$. The distributions for the β_i are produced by the software, ERN (2002), by aggregating the corresponding distributions for the proportions of losses at each insured building. The corresponding densities for the portfolio are specified as:

$$f_B(\beta_i) = P_0 \delta(\beta_i) + (1 - P_0 - P_1) B(\beta_i; a, b | M) + P_1 \delta(\beta_i - 1) \quad 0 \leq \beta_i \leq 1. \quad (5)$$

In equation (5) P_0 is the probability of zero losses, P_1 the probability of total loss; $B(y; a, b | M)$ represents a beta density with parameters a, b , conditional on the magnitude, and δ is the Dirac delta. These distributions are obtained using information on the construction characteristics for each insured building combined with shock wave diffusion and local effects from earthquakes at the given 'site-magnitude'; they lead to a ratio damage distribution for each building. The individual distributions for the loss proportions in each building are then aggregated over the entire portfolio to obtain (5); for a detailed description see ERN (2002). Each one of these component models was validated at every stage of its development by the scientists and engineers who developed them, by comparing model results with actual data from historical events and specific portfolios of property exposures, de Alba and Zúñiga-San Martín (2006).

The programs necessary to run the simulations presented in this paper are two; see Appendix 1. In addition to the pure premium and the PML, that is required for computing catastrophe reserves, the ERN program produces output that includes the probability of an earthquake of a given magnitude from site i , plus the parameters of the distribution in (5) for each combination; see Appendix 2. This information allows the user to carry out simulations.

In very broad terms the simulation algorithm is as follows:

- a) Choose an earthquake site at random

- b) Given the site, generate a magnitude at random from the corresponding distribution
- c) Use the distribution of proportion of losses for the site-magnitude combination to generate a random loss proportion (damage) for each insured building, equation (5) and combine to obtain a global proportion for the portfolio.
- d) Multiply the proportion resulting in c) by the total value insured for the portfolio and obtain a loss amount.
- e) Apply any reinsurance and risk transfer vehicle that are in force.

This process is applied as many times as there are earthquakes in a year to derive a figure of total yearly losses. As many yearly replications are generated as are needed according to the required precision.

Simulation Results

We apply the algorithm to the portfolio from a real Mexican insurance company that has been disguised by multiplying loss amounts by a constant. The portfolio consists of 25,000 buildings. The non-proportional reinsurance scheme (in thousands of dollars) is as shown in Table 1. The insurance company also has a quota share with 10% retention for losses below 7,500, the priority of the first layer. In this table, the heading 'Rol' stands for the Rate on Line cost of the reinsurance premium for the specific layer. In turn 'Reins' indicates the number of reinstatements that the reinsurance contract establishes for each layer. The column labeled 'Deductible' corresponds to the deductible or attachment point of each layer.

TABLE 1

Layers	Deductible	Cover	Reinstatement Premium	Rol	Reins
1	\$7,500	\$7,500	\$1,586	21.15%	2
2	\$15,000	\$15,000	\$1,890	12.60%	2
3	\$30,000	\$30,000	\$2,268	7.56%	1
4	\$60,000	\$40,000	\$1,548	3.87%	1
5	\$100,000	\$130,000	\$2,574	1.98%	1
Superior	\$230,000	None	NA	NA	NA

Some of the business questions that can be answered with the simulation analysis are the following: Is total reinsurance coverage adequate for the company? Does the program have enough reinstatements? How much risk relief is achieved with the (reinsurance) program, in monetary terms? How much capital does the company require to guarantee coverage of the insured portfolio? Which are the higher contributors to PML? What are the most cost effective risk transfer programs?

Applying the algorithm described above and through simple statistical analysis we evaluate the mitigation effect of the reinsurance contract. In Table 2 we show some statistics for the gross losses (without deducting any reinsurance) and for losses net of all reinsurance. The statistics given for net loss in the second row of that table include reinstatement costs. The third row shows net losses without these costs. The resulting retention level for the whole portfolio is given in the last row.

TABLE 2

	Mean	ST. D.	Minimum	Q1	Median	Q3	Maximum
Gross Loss	\$4,908	\$18,636	\$9	\$752	\$1,558	\$3,653	\$1,241,000
Net Loss	\$696	\$9,280	\$1	\$75	\$156	\$365	\$1,019,000
Net Loss W.O. Reins	\$480	\$9,008	\$1	\$75	\$156	\$365	\$1,010,000
Retention	9.78%	1.51%	10.00%	10.00%	10.00%	10.00%	81.39%

Table 3 shows gross and net losses for several return periods. This is relevant for complying with the regulatory authority with respect to solvency, that requires a return period of 1500 years in the computation of the PML which in turn is used in calculating the corresponding earthquake catastrophe reserves.

TABLE 3

Gross Losses	Net Losses	Net Losses without Reinstatement Premiums	% Reduction	Fn	Return Period
\$304,623	\$84,801	\$74,084	72.16%	0.999333333	1500
\$230,102	\$12,509	\$2,756	94.56%	0.999	1000
\$155,290	\$9,613	\$1,947	93.81%	0.998	500
\$81,936	\$7,813	\$1,648	90.46%	0.995	200
\$53,503	\$6,167	\$1,493	88.47%	0.99	100

The table tells us that the relief obtained from the program in this case goes from 74% to 95%, so the program is effective for the company, because with at most \$84,801 it can cover a portfolio with a PML of up to \$304,623, all in thousands of dollars.

Solvency II states that the Solvency Capital Requirement shall “correspond to the Value-at-Risk of the basic own funds of an insurance or reinsurance undertaking subject to a confidence level of 99.5% over a one-year period.” So in this case, the amount should be \$7,813 and the relief produced by reinsurance will be \$74,123 ($81,936 - 7,813$), complying with the other Solvency II

statement “insurance and reinsurance undertakings shall take account of the effect of risk mitigation techniques, provided that credit risk and other risks arising from the use of such techniques are properly reflected in the Solvency Capital Requirement.” Credit risk adjustments are not illustrated here.

Other regulations apply a higher confidence level, such as the Mexican, which for earthquake is 99.93%. In this case the amount should be \$84,801 and the relief produced by reinsurance \$230,539.

Further analysis yields the results in Table 4, where we can see if the reinsurance strategy is what the company needs. This can be deduced from some statistics concerning the layers used in the reinsurance scheme described in Table 1. From column (2) one can see that reinsurance was insufficient in only an extremely low percentage of the total number of years simulated (150,000). Column (5) presents the distribution of events by layer.

TABLE 4

Layers	% years all reinstatements were used	% years only one reinstatement was used	% years the second reinstatement was used	Distribution of events by layer
(1)	(2)	(3)	(4)	(5)
Priority	NA	NA	NA	99.8635%
1	0.01%	9.54%	0.54%	0.0711%
2	0.00%	4.85%	0.13%	0.0393%
3	0.01%	2.02%	NA	0.0168%
4	0.00%	0.73%	NA	0.0047%
5	0.00%	0.36%	NA	0.0034%
Sup	NA	NA	NA	0.0013%

We also use the Monte Carlo results to show that for a large portfolio and a complicated reinsurance contract, non-proportional reinsurance risk relief can clearly be compared with that achieved with a proportional reinsurance contract, Insurer Solvency Assessment Working Party (2004). For this purpose Figure 2 shows gross losses together with net losses resulting from a proportional reinsurance scheme, as well as from a non-proportional one, for a certain range of loss amounts. At the lower levels shown in the graph there are more losses net of non-proportional insurance than those net of the proportional reinsurance schemes. However as the amount of the loss increases, the frequency of losses net of non-proportional reinsurance practically disappears, which is to say that this kind of reinsurance provides more protection to the insurer, at least for this portfolio and this reinsurance contract.

As mentioned, Mexican regulation specifies the use of the 0.99933% quantile to compute the PML and the corresponding reserves. Considering the fact that VaR is an incoherent measure of risk, McNeil et al. (2005), we will now complement the previous analysis using the Tail-VaR, as well as its variance,

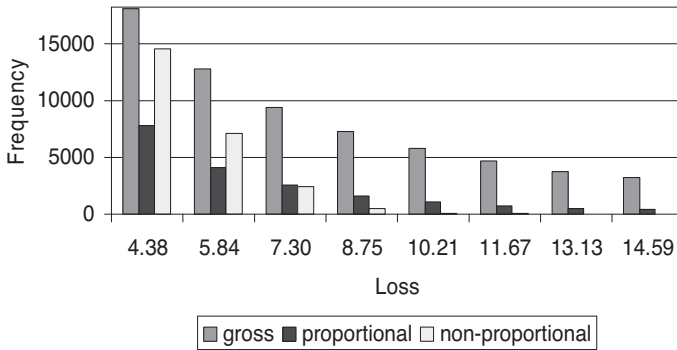


FIGURE 2: Frequency of Losses.

Manistre and Hancock (2005). In Table 5 we provide the corresponding Tail-VaR for gross and net losses, as well as the variance for the gross losses.

TABLE 5
TAIL-VAR

alpha	Return	k	Gross Losses				Net Losses			
			CTE (alpha)	FSE (CTE)	LL	UL	CTE (alpha)	FSE (CTE)	LL	UL
0.01	100	1500	124,320	3423	117,610	131,030	16,339	1478	13,443	19,236
0.005	200	750	185,873	6051	174,012	197,733	26,399	2909	20,697	32,102
0.002	500	300	309,852	11877	286,573	333,130	53,237	6991	39,535	66,940
0.001	1000	150	442,367	17975	407,136	477,599	96,297	13059	70,703	121,892
0.000667	1500	100	530,843	22049	487,626	574,059	138,066	18178	102,436	173,695

Since for the computations in Table 5, we are using the data from the simulations, the Tail-VaR's are computed as the average of the values exceeding the corresponding quantile. Table 6 shows the results of VaR and tail-VaR for gross and net losses using different return periods.

TABLE 6

Gross Losses		Net Losses		Return Period
VaR	Tail VaR	VaR	Tail VaR	
\$300,710	\$530,843	\$14,278	\$138,066	1500
\$231,938	\$442,367	\$10,968	\$96,297	1000
\$143,763	\$309,852	\$9,313	\$53,237	500
\$90,040	\$185,873	\$8,038	\$26,399	200
\$78,011	\$124,320	\$7,684	\$16,339	100

The results are as expected: the Tail-VaR at each return period is larger than the corresponding VaR. There is no reason to expect the Tail-VaR to be equal to the quantile based PML measure. If these are seen as too large the Tail-VaR will be much more so. It would be necessary to define new criteria for PML calculation if a coherent measure of risk is to be applied. In addition the percentage reduction between the gross and net losses is smaller than when using VaR, but they are still considerably large. In practice, if a quantile based PML is established for a given return period, then a smaller value of the return period might be appropriate for the Tail-VaR.

Although the paper is written in the context of earthquake catastrophe insurance, it is applicable to others, such as hurricane, provided the hurricane model is available. It is shown how a relatively simple simulation model can provide a wealth of information not obtainable by analytic procedures. In fact, the Regulatory Authority has also commissioned a model to evaluate hurricane risk, along the lines of the earthquake model, which is already in use by the market, although without the use of simulations as described above. We have begun to carry out simulation exercises similar to those presented here.

Extreme Value Analysis

The previous analysis provides a large amount of information to the insurer. Yet, there are some technical details that can be explored further. For example, it is known that quantile estimates obtained by simulation are biased, Inui et al. (2005); and that the bias tends to zero as the sample size increases. It can also be corrected by bootstrapping, Kim and Hardy (2007). We have used large sample sizes so that the bias should be small. Nevertheless we will carry out additional analyses in order to fine tune the results. In particular the quantile corresponding to the 1500 year return period, which must be used to compute the PML and hence required reserves for earthquake catastrophe risk. We will also complement the analysis with the computation of confidence intervals for the quantiles.

Since PML estimation is essentially an exercise in estimating a large quantile it falls in the field of extreme values. We use the approach of fitting excesses over a threshold and proceed along the lines set out in Embrechts et al. (1997, page 352). We assume that the losses from earthquakes are X_1, X_2, \dots, X_n , i.i.d. with distribution $F(x)$, a Generalized Extreme Value Distribution (GEV). Then we choose a high threshold u , and so the corresponding excesses are denoted by $Y_i = X_i - u, i = 1, \dots, n$; and N_u is the number of exceedances of u by X_1, \dots, X_n . The exceedances follow a generalized Pareto distribution (GPD). This GPD, denoted by $G_{\xi, \beta}$, with parameters $\xi \in \mathbb{R}$ and $\beta > 0$ has distribution tail

$$\bar{G}_{\xi, \beta}(x) = \begin{cases} \left(1 + \xi \frac{x}{\beta}\right)^{-1/\xi} & \text{if } \xi \neq 0, \\ e^{-x/\beta} & \text{if } \xi = 0, \end{cases} \quad x \in D(\xi, \beta), \tag{6}$$

where

$$D(\xi, \beta) = \begin{cases} [0, \infty) & \text{if } \xi \geq 0, \\ [0, -\beta/\xi] & \text{if } \xi < 0. \end{cases}$$

Embrechts et al. (1997). We must choose a high threshold u , although there are no clear criteria. It should not be too small so as to not produce biased estimators and it should not be too high because it will produce high variance estimators. To decide what value to use for u , we plot the Mean Excess for a range of values from 0 to 800, Figure 3. The graph is roughly linear up to $u = 200$ and so we decided to use $u = 66.22$ (in millions of dollars), which yields $N_u = 1000$. This value has the added advantage that (empirically) $\text{Prob}\{X > 66.22\} = 1/150$, and can easily be used to compute our 1500 year return period quantile, as will be shown below. We used the program “ExtRemes” to carry out the analysis, Gilleland and Katz (2005). We then use the 1000 exceedances that were for fitting the tail of the GPD. Figure 4 shows the resulting histogram, which indicates a good fit.

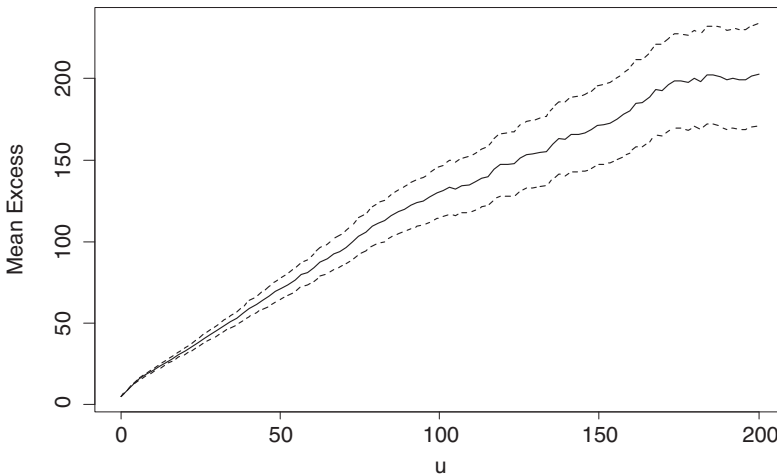


FIGURE 3.

Figure 5 shows the corresponding fit diagnostics. Since we are fitting to the exceedances the two graphs that involve return periods must be viewed with care. In order to compare these results with those obtained directly by simulation we compute the required quantiles from the GPD using the estimator

$$\hat{x}_p = u + \frac{\hat{\beta}}{\hat{\xi}} \left(\left(\frac{n}{N_u} (1 - p) \right)^{-\hat{\xi}} - 1 \right), \tag{7}$$

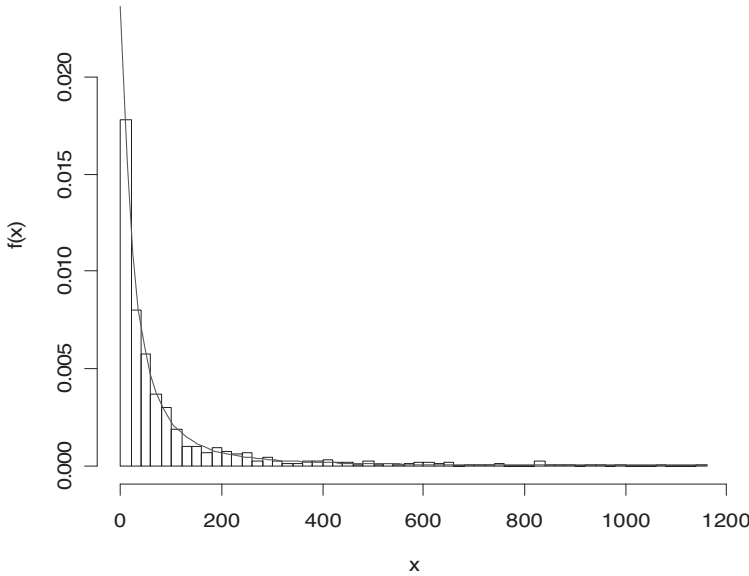


FIGURE 4: Density Plot.

Embrechts et al. (1997), where $\hat{\beta}$ and $\hat{\xi}$ are the parameter estimates. The value of p must be determined such that

$$\bar{F}(u + y) = \bar{F}(u) \cdot \bar{F}_u(y) = 1 - p,$$

where $\bar{F}(x) = 1 - F(x)$, and $\bar{F}(u)$ is estimated by $N_u/n = 1/150$. For example, to get the quantile for the 1500 year return we take $p = .999333$ and we use an approximation based on (7) to estimate $\bar{F}_u(y)$ and hence the quantile. But from these specifications we end up with

$$0.000667 = 1 - 0.999333 = (1/150) \cdot \bar{F}_u(y).$$

Now we use (7) to obtain a p' -quantile for the exceedances (the Y 's) and where $p' = 1 - 0.000667 * 150 = 1 - 0.10 = 0.9$. This is done with the “extRemes Toolkit”, Gilleland and Katz (2005), by obtaining the 10 year return level. Table 7 shows the return levels, along with their confidence intervals and the simulation results for the return periods 500, 1000 and 1500. These would be the corresponding PML's for this insurance company using the Mexican earthquake data. Note they are fairly close and in all cases the simulation results are within the 95% confidence intervals. Interpretation of the confidence interval must be done with care; these are not to be confused with the much wider confidence intervals which would be obtained if one should account adequately for the epistemic parameter uncertainty in the earthquake loss model. This is

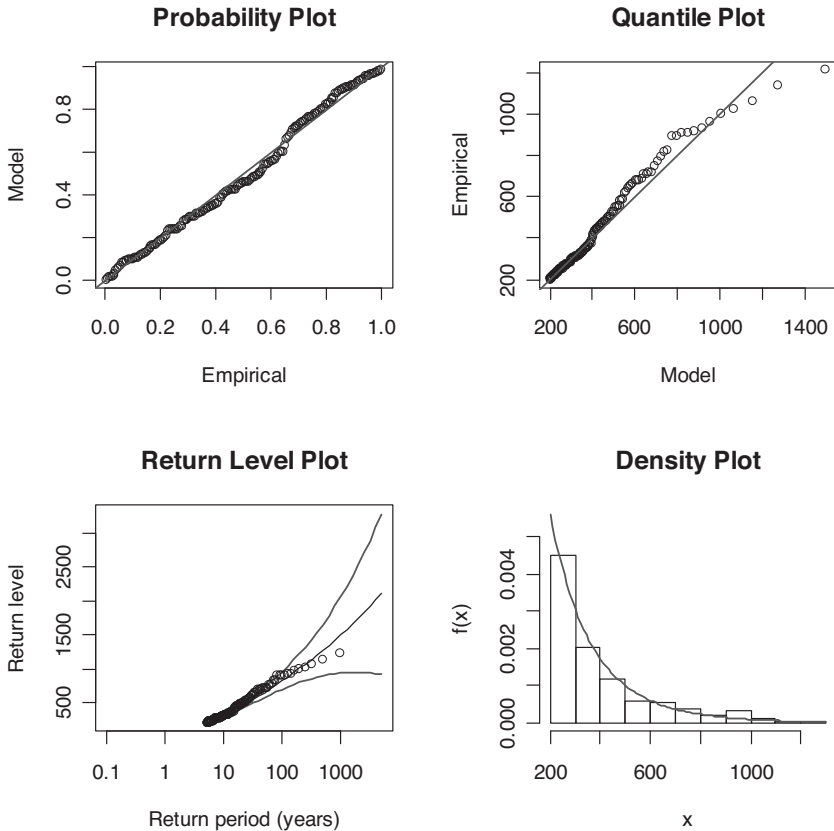


FIGURE 5.

beyond the scope of this paper, so that essentially we are assuming the model used is correct.

The result obtained from the ERN program, which is supposed to be exact, turns out to be $PML = \$284,718.06$. This is clearly very close to the one obtained using an Extreme Value approach. However the simulation approach yields a wealth of information not obtainable via the ERN program.

The results in Table 7 show that there is consistency between the different concepts and approaches. The computations obtained via Extreme Value Theory provide additional information to that obtained directly from the simulations and with a formal justification, from the statistical point of view. The estimates of the quantiles obtained from the simulations are near the upper bound of the confidence intervals in agreement with the known fact that they are known to be biased. Apparently, in this case they have a positive bias.

Hence they provide conservative estimates for the PML. The ERN results are roughly in the middle of the interval, but ERN is less flexible for evaluating complex reinsurance schemes.

TABLE 7

Return periods	Simulation	Extreme value	Lower limit	Upper limit	ERN
1500	\$300,710	\$280,076	\$256,524	\$309,758	\$284,718
1000	\$231,938	\$217,937	\$203,170	\$235,519	\$224,775
500	\$143,763	\$141,439	\$134,747	\$148,824	\$144,087

We have shown how Monte Carlo methods can be used to analyze the effect of complicated reinsurance treaties on a heterogeneous portfolio. Simulation also allows the evaluation of large quantiles although the results may be biased if we do not have a large sample. However this may be further explored via extreme values.

CONCLUSIONS

The methodology for estimating probable maximum loss (PML) for natural catastrophes has evolved over the past few decades from being deterministic to one based on stochastic models, due to:

1. The increased use of complex risk transfer agreements,
2. The need to describe in greater detail the vulnerability of an insurance entity's strategy on its commercial, underwriting and risk transfer practices.

It is necessary to bring together the expertise of geophysicists, structure engineers, actuaries, financial experts and others in order to construct a model that represents the overall process reasonably well. All are present in the modeling process followed by the firm *ERN Ingenieros Consultores, S.C.* (ERN) commissioned by the Mexican regulator to develop the models used here.

Given the recent trends on risk based management on regulation, methods like the one described in this paper to measure catastrophic risks are much more relevant. To comply with the requirements of regulatory and accounting frameworks, such as Solvency II:

1. In terms of risk management, the correct measurement of the risk of a succession of catastrophic events and not only a single one is a must, specially in countries exposed to both seismological and hydro meteorological dangers.
2. To clearly understand the exposure of an insured portfolio and achieve effective risk mitigation from financial and regulatory points of view, proper measurements of the magnitude of the losses, with and without the risk transfer chosen, are needed.
3. Proper recognition of reinsurance is necessary in order to assess risk reduction for the ceding company. This has implications for capital requirements to ensure effective solvency supervision.

4. It is necessary not only to have the right measurement, but to have it available for auditing and adequately documented
5. Models can only be accepted by insurance regulators if there is evidence that their use is “embedded” in the governance of the organizations, so it is necessary to have enough and clear elements for communication with and disclosure to interested parties (company management and board, reinsurers, authorities and rating agencies). As we saw, this approach is specially useful on the communication side, given the access to different views of the results.

Among the business questions that can be answered with the analysis described in this paper are the following:

1. Is total reinsurance coverage adequate for the company?
2. Does the program have enough reinstatements?
3. How much risk relief is achieved with the program, in monetary terms?
4. How much capital does the company require to guarantee coverage of the insured portfolio?
5. Which are the higher contributors to PML?
6. What are the most cost effective risk transfer programs?

We have shown how Monte Carlo methods can be used to analyze the effect of complicated reinsurance treaties on a heterogeneous portfolio. Simulation also allows the estimation of large quantiles although the results may be biased if we do not have a large sample. However this may be further explored via the theory of extreme values.

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