

On the Length of a Random Minimum Spanning Tree

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We study the expected value of the length L_n of the minimum spanning tree of the complete graph K_n when each edge e is given an independent uniform $[0, 1]$ edge weight. We sharpen the result of Frieze [6] that $\lim_{n \rightarrow \infty} \mathbb{E}(L_n) = \zeta(3)$ and show that

$$\mathbb{E}(L_n) = \zeta(3) + \frac{c_1}{n} + \frac{c_2 + o(1)}{n^{4/3}},$$

where c_1, c_2 are explicitly defined constants.

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1. Introduction

We study the expected value of the length L_n of the minimum spanning tree of the complete graph K_n when each edge e is given an independent uniform $[0, 1]$ edge weight X_e . Frieze [6] showed that

$$\lim_{n \rightarrow \infty} \mathbb{E}(L_n) = \zeta(3) = \sum_{k=1}^{\infty} \frac{1}{k^3} = 1.202\dots \quad (1.1)$$

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Since then there have been several generalizations and improvements. Steele [27] extended the applicability of (1.1) distribution-wise. Janson [11] proved a central limit theorem for L_n . Penrose [23], Frieze and McDiarmid [7], Beveridge, Frieze and McDiarmid [2] and Frieze, Ruszinkó and Thoma [8] analysed L_n for graphs other than the complete graph. Steele [28] and Fill and Steele [4] used the Tutte polynomial to compute $\mathbb{E}(L_n)$ exactly for small values. Nishikawa, Otto and Starr [21] studied the coefficients of a polynomial derived from the formula in [28]. Gamarnik [9] computed $\mathbb{E}_{\text{exp}}(L_n)$ exactly up to $n \leq 45$ using a more efficient algorithm, where $\mathbb{E}_{\text{exp}}(L_n)$ is the expectation when the distribution of the X_e is exponential with mean one. Li and Zhang [18] consider more general distributions and prove in particular that

$$\mathbb{E}_{\text{exp}}(L_n) - \mathbb{E}(L_n) = \frac{\zeta(3)}{n} + O\left(\frac{\log^2 n}{n^2}\right). \tag{1.2}$$

Flaxman [5] gives an upper bound on the lower tail of L_n .

Equation (1.1) says that $\mathbb{E}(L_n) = \zeta(3) + o(1)$ as $n \rightarrow \infty$. Ideally, one would like to have an exact expansion for $\mathbb{E}(L_n)$ as there is for the assignment problem; see Wästlund [29] and the references therein. Such an expansion has proved elusive. In this work we improve the asymptotics of $\mathbb{E}[L_n]$ by giving the secondary and tertiary terms.

Theorem 1.1.

$$\mathbb{E}(L_n) = \zeta(3) + \frac{c_1}{n} + \frac{c_2 + o(1)}{n^{4/3}},$$

where

$$c_1 = -1 - \zeta(3) - \frac{1}{2} \int_{x=0}^{\infty} \log(1 - (1+x)e^{-x}) dx$$

and

$$\begin{aligned} c_2 &= \int_{x=0}^{\infty} \left(x^{-3} \psi(x^{3/2}) e^{-x^3/24} - x^{-3} - \sqrt{\frac{\pi}{8}} x^{-3/2} - \frac{1}{2} \right) dx \\ &= \frac{2}{3} \int_{y=0}^{\infty} \left(y^{-2} \psi(y) e^{-y^2/24} - y^{-2} - \sqrt{\frac{\pi}{8}} y^{-1} - \frac{1}{2} \right) y^{-1/3} dy, \end{aligned}$$

with ψ defined in (1.3) below.

The two integral expressions defining c_2 are equal by the change of variable $x = y^{2/3}$.

A numerical integration (with Maple) yields $c_1 = 0.0384956\dots$. This shows that the rate of convergence to $\zeta(3)$ is order $1/n$ and is from above. Further numerical computations show that $c_2 \approx -1.7295$, and these are explained in the Appendix.

To define ψ , we let the random variable $\mathcal{B}_{\text{ex}} = \int_{s=0}^1 \mathcal{B}_{\text{ex}}(s) ds$ be the area under a normalized Brownian excursion; we then let

$$\psi(t) = \mathbb{E} e^{t\mathcal{B}_{\text{ex}}}, \tag{1.3}$$

the moment generating function ψ of \mathcal{B}_{ex} . The Brownian excursion area \mathcal{B}_{ex} and its moments $\mathbb{E} \mathcal{B}_{\text{ex}}^k$ and moment generating function ψ have been studied by several authors;

see Louchard [19, 20], for example, and the survey by Janson [12], where further references are given. From these results, we derive an expression (see (1.7)) that will show that c_2 is well-defined. Note that $\psi(t)$ is finite for all $t > 0$ (and thus (1.3) holds for all complex t). Indeed, it is well known that

$$\mathbb{E} \mathcal{B}'_{\text{ex}} \sim \sqrt{18} \ell (12e)^{-\ell/2} \ell^{\ell/2} \quad \text{as } \ell \rightarrow \infty \tag{1.4}$$

(see [12, (53)] and the references there), and thus [13, Lemma 4.1(ii)] implies that

$$\psi(t) \sim \frac{1}{2} t^2 e^{t^2/24} \quad \text{as } t \rightarrow +\infty \tag{1.5}$$

(cf. [13, Remarks 3.1 and 4.9], where $\xi = 2\mathcal{B}_{\text{ex}}$). More precisely, Janson and Louchard [15] show that the density f_{ex} of \mathcal{B}_{ex} satisfies

$$f_{\text{ex}}(x) = \frac{72\sqrt{6}}{\sqrt{\pi}} x^2 e^{-6x^2} (1 + O(x^{-2})), \quad x > 0, \tag{1.6}$$

from which routine calculations show that

$$\psi(t) = \int_{x=0}^{\infty} e^{tx} f_{\text{ex}}(x) dx = \frac{t^2}{2} e^{t^2/24} (1 + O(t^{-2})), \quad t > 0. \tag{1.7}$$

Hence the integrand in the second integral defining c_2 in Theorem 1.1 is $O(y^{-4/3})$ as $y \rightarrow \infty$. Moreover, $\psi(0) = 1$ and $\psi'(0) = \mathbb{E} \mathcal{B}_{\text{ex}} = \sqrt{\pi/8}$, and thus a Taylor expansion shows that the integrand is $O(y^{-1/3})$ as $y \rightarrow 0$. (Similarly, the integrand in the first integral is $O(x^{-3/2})$ and $O(1)$.) Consequently, the integrals defining c_2 converge absolutely.

2. Proof of Theorem 1.1

We prove the theorem by using the expression

$$\mathbb{E}(L_n) = \int_{p=0}^1 \mathbb{E}(\kappa(G_{n,p})) dp - 1 \tag{2.1}$$

(see Janson [11] and a related expression in Frieze and McDiarmid [7, equation (7)]). Here $\kappa(G_{n,p})$ is the (random) number of components in the random graph $G_{n,p}$.

To evaluate (2.1), we let $\kappa(k, j, p) = \kappa_n(k, j, p)$ denote the number of components of $G_{n,p}$ with k vertices and $k + j$ edges in $G_{n,p}$. The components split neatly into three categories: trees ($j = -1$), unicyclic ($j = 0$), and complex ($j \geq 1$) components. These are evaluated separately.

Lemma 2.1.

(a)

$$\begin{aligned} \int_{p=0}^1 \sum_{k \geq 1} \mathbb{E}(\kappa(k, -1, p)) dp &= \zeta(3) + \frac{3(\zeta(2) - \zeta(3))}{2n} \\ &\quad - \frac{1}{n^{4/3}} \int_{x=0}^{\infty} x^{-3} (1 - e^{-x^3/24}) dx + o(n^{-4/3}). \end{aligned}$$

(b)

$$\int_{p=0}^1 \sum_{k \geq 3} \mathbb{E}(\kappa(k, 0, p)) dp = \frac{1}{2n} \left(\zeta(3) - 3\zeta(2) - \int_{x=0}^\infty \log(1 - (1+x)e^{-x}) dx \right) - \frac{\sqrt{\pi/8}}{n^{4/3}} \int_{x=0}^\infty x^{-3/2}(1 - e^{-x^3/24}) dx + o(n^{-4/3}).$$

(c) With $\psi_2(x) = \psi(x) - 1 - \sqrt{\pi/2}x$,

$$\int_{p=0}^1 \sum_{k \geq 1} \sum_{j \geq 1} \mathbb{E}(\kappa(k, j, p)) dp = 1 - \frac{1}{n} + \frac{1}{n^{4/3}} \int_{x=0}^\infty \left(x^{-3} \psi_2(x^{3/2}) e^{-x^3/24} - \frac{1}{2} \right) dx + o(n^{-4/3}).$$

Remark 1. Tree components contribute the main $\zeta(3)$ term. Unicyclic components contribute a secondary $O(1/n)$ addend. Roughly speaking there are no complex components for $p \leq 1/n$ and precisely one complex component (the famous ‘giant component’) for $p \geq 1/n$. Were this to be precisely the case, the contribution of complex components would be $1 - 1/n$. The additional $\Theta(n^{-4/3})$ term in Lemma 2.1(c) comes from the behaviour of complex components in the critical window $p = 1/n + \lambda n^{-4/3}$.

Remark 2. The coefficients of $n^{-4/3}$ in Lemma 2.1(a,b) are easily evaluated as $-\frac{1}{8}3^{-2/3}\Gamma(1/3)$ and $-\frac{1}{2}3^{-1/6}\sqrt{\pi}\Gamma(5/6)$, respectively; see the Appendix. The coefficient in Lemma 2.1(c) is expressed as an infinite sum and evaluated numerically in the Appendix.

Proof. In the proof we assume tacitly that n is large enough when necessary. We let C_1, \dots denote some unimportant universal constants.

Let $C(k, \ell)$ be the number of connected graphs on a vertex set $[k]$ with ℓ edges. We begin by noting the standard formula

$$\mathbb{E} \kappa(k, j, p) = \binom{n}{k} C(k, k+j) p^{k+j} (1-p)^{k(n-k) + \binom{k}{2} - k-j}. \tag{2.2}$$

By Cayley’s formula, $C(k, k-1) = k^{k-2}$. Moreover, Wright [30] proved that, for every fixed $j \geq -1$,

$$C(k, k+j) \sim w_{j+1} k^{k+3j/2-1/2}, \quad \text{as } k \rightarrow \infty, \tag{2.3}$$

for some constants $w_\ell > 0$. (See also [14, §8] and the references there. In the notation of [30], $w_{j+1} = \rho_j$.) We have $w_0 = 1$ and $w_1 = \sqrt{\pi/8}$. Spencer [26] showed that

$$w_\ell = \frac{\mathbb{E} \mathcal{B}_{\text{ex}}^\ell}{\ell!}, \quad \ell \geq 0, \tag{2.4}$$

where \mathcal{B}_{ex} is the Brownian excursion area defined above. See further Janson [12]. Hence,

$$\psi(t) = \mathbb{E} e^{t\mathcal{B}_{\text{ex}}} = \sum_{\ell=0}^\infty w_\ell t^\ell. \tag{2.5}$$

Let

$$\begin{aligned}
 A(k, k + j) &= \int_{p=0}^1 \mathbb{E}(\kappa(k, j, p)) dp \\
 &= \binom{n}{k} C(k, k + j) \int_{p=0}^1 p^{k+j} (1-p)^{k(n-k) + \binom{k}{2} - k - j} dp \\
 &= \binom{n}{k} C(k, k + j) \frac{(k + j)! (k(n - k) + \binom{k}{2}) - k - j)!}{(k(n - k) + \binom{k}{2} + 1)!} \\
 &= \frac{C(k, k + j) (k + j)!}{k!} \times B(k, k + j)
 \end{aligned} \tag{2.6}$$

where, provided $k \leq n$ and $k + j \leq \binom{k}{2}$ (as in our case),

$$\begin{aligned}
 B(k, k + j) &= \frac{n!}{(n - k)!} \cdot \frac{(k(n - k) + \binom{k}{2}) - k - j)!}{(k(n - k) + \binom{k}{2} + 1)!} \\
 &= \frac{1}{n^{j+1} k^{k+j+1}} \frac{\prod_{i=0}^{k-1} (1 - \frac{i}{n})}{\prod_{i=0}^{k+j} (1 - \frac{k+1}{2n} - \frac{i-1}{kn})} \\
 &= \frac{1}{n^{j+1} k^{k+j+1}} \exp \left\{ \sum_{m=1}^{\infty} \frac{1}{m n^m} \left(\sum_{i=0}^{k+j} \left(\frac{k+1}{2} + \frac{i-1}{k} \right)^m - \sum_{i=0}^{k-1} i^m \right) \right\} \\
 &= \frac{1}{n^{j+1} k^{k+j+1}} \exp \left\{ \sum_{m=1}^{\infty} \frac{t_m(k, j)}{m n^m} \right\}.
 \end{aligned} \tag{2.7}$$

Observe that as

$$\sum_{i=1}^a i^m \geq \int_0^a x^m dx,$$

for $\ell = k + j$ we have

$$t_m(k, j) = \sum_{i=0}^{\ell} \left(\frac{k+1}{2} + \frac{i-1}{k} \right)^m - \sum_{i=0}^{k-1} i^m \leq (\ell + 1) \left(\frac{k+1}{2} + \frac{\ell-1}{k} \right)^m - \frac{(k-1)^{m+1}}{m+1}. \tag{2.8}$$

This implies that, as is easily verified,

$$t_m(k, j) \leq 0 \text{ if } m \geq 2 \text{ and } j \in \{0, -1\} \text{ and } k \geq 100. \tag{2.9}$$

Case (a): $1 \leq k \leq n, j = -1$ (tree components). Now we have by (2.7)

$$\begin{aligned}
 B(k, k - 1) &= \frac{1}{k^k} \exp \left\{ \frac{1}{n} \sum_{i=0}^{k-1} \left(\frac{k+1}{2} + \frac{i-1}{k} \right) - \frac{1}{n} \sum_{i=0}^{k-1} i \right. \\
 &\quad \left. + \frac{1}{2n^2} \sum_{i=0}^{k-1} \left(\frac{k+1}{2} + \frac{i-1}{k} \right)^2 - \frac{1}{2n^2} \sum_{i=0}^{k-1} i^2 + \zeta \right\}
 \end{aligned}$$

where, using (2.9),

$$|\xi| \leq \sum_{m=3}^{\infty} \frac{10^{2m+1}}{mn^m} = O(n^{-3}), \quad 1 \leq k \leq 100, \tag{2.10}$$

$$0 \geq \xi \geq -\sum_{m=3}^{\infty} \frac{k^{m+1}}{m(m+1)n^m} \geq -\frac{k^4}{n^3}, \quad k > 100, \tag{2.11}$$

and hence, for all $k \leq n$,

$$\xi = O(k^4/n^3). \tag{2.12}$$

This implies, after some calculation, that for $1 \leq k \leq n$,

$$B(k, k - 1) = \frac{1}{k^k} \exp\left\{ \frac{3(k - 1)}{2n} - \frac{k^3}{24n^2} + O\left(\frac{k^2}{n^2} + \frac{k^4}{n^3}\right) \right\}$$

and then, by (2.6),

$$\begin{aligned} \sum_{k=1}^{n^{0.7}} A(k, k - 1) &= \sum_{k=1}^{n^{0.7}} \frac{k^{k-2}}{k} \cdot B(k, k - 1) \\ &= \sum_{k=1}^{n^{0.7}} \frac{1}{k^3} \exp\left\{ \frac{3(k - 1)}{2n} - \frac{k^3}{24n^2} + O\left(\frac{k^2}{n^2} + \frac{k^4}{n^3}\right) \right\} \\ &= \sum_{k=1}^{n^{0.7}} \frac{e^{-k^3/24n^2}}{k^3} \left(1 + \frac{3(k - 1)}{2n} + O\left(\frac{k^2}{n^2} + \frac{k^4}{n^3}\right) \right). \end{aligned}$$

Now, by simple estimates,

$$\sum_{k=1}^{n^{0.7}} \frac{e^{-k^3/24n^2}}{k^3} \times O\left(\frac{k^2}{n^2} + \frac{k^4}{n^3}\right) = O(n^{-5/3}) \tag{2.13}$$

and

$$\begin{aligned} \sum_{k=1}^{n^{0.7}} \frac{(1 - e^{-k^3/24n^2})}{k^3} \left(1 + \frac{3(k - 1)}{2n} \right) &= o(n^{-4/3}) + \sum_{k=n^{2/3}/\ln n}^{n^{2/3} \ln n} \frac{(1 - e^{-k^3/24n^2})}{k^3} \\ &= o(n^{-4/3}) + \frac{1}{n^{4/3}} \int_{x=0}^{\infty} x^{-3} (1 - e^{-x^3/24}) dx. \end{aligned} \tag{2.14}$$

Thus

$$\begin{aligned} \sum_{k=1}^{n^{0.7}} A(k, k - 1) &= \sum_{k=1}^{n^{0.7}} \frac{1}{k^3} + \frac{1}{n} \sum_{k=1}^{n^{0.7}} \frac{3(k - 1)}{2k^3} - \frac{1}{n^{4/3}} \int_{x=0}^{\infty} x^{-3} (1 - e^{-x^3/24}) dx + o(n^{-4/3}) \end{aligned}$$

$$\begin{aligned}
 &= \zeta(3) + O(n^{-1.4}) + \frac{3(\zeta(2) - \zeta(3))}{2n} + O(n^{-1.7}) \\
 &\quad - \frac{1}{n^{4/3}} \int_{x=0}^{\infty} x^{-3}(1 - e^{-x^3/24}) dx + o(n^{-4/3}) \\
 &= \zeta(3) + \frac{3(\zeta(2) - \zeta(3))}{2n} - \frac{1}{n^{4/3}} \int_{x=0}^{\infty} x^{-3}(1 - e^{-x^3/24}) dx + o(n^{-4/3}). \tag{2.15}
 \end{aligned}$$

When $k \geq n^{0.7}$, we have from (2.7) and (2.9) that

$$B(k, k - 1) \leq \frac{1}{k^k} \exp\left(\frac{1}{n} \sum_{i=0}^{k-1} \left(\frac{k+1}{2} + \frac{i-1}{k}\right) - \frac{1}{n} \sum_{i=0}^{k-1} i\right) = \frac{1}{k^k} \exp\left\{\frac{3(k-1)}{2n}\right\} \leq \frac{e^{3/2}}{k^k}.$$

This implies that $A(k, k - 1) \leq k^{-3}e^{3/2}$. This gives

$$\sum_{k > n^{0.7}} A(k, k - 1) \leq \sum_{k > n^{0.7}} \frac{e^{3/2}}{k^3} = O(n^{-1.4}) = o(n^{-4/3}).$$

Together with (2.15), this verifies (a).

Case (b): $1 \leq k \leq n, j = 0$ (unicyclic components). Rényi [25] proved (see, e.g., Bollobás [3, Theorem 5.18]) that

$$C(k, k) = \frac{(k-1)!}{2} \sum_{l=0}^{k-3} \frac{k^l}{l!} \sim \sqrt{\frac{\pi}{8}} k^{k-1/2} \tag{2.16}$$

(cf. the more general (2.3) above). Now, for $1 \leq k \leq n$ we have by (2.7)

$$\begin{aligned}
 B(k, k) &= \frac{1}{nk^{k+1}} \exp\left\{\frac{1}{n} \sum_{i=0}^k \left(\frac{k+1}{2} + \frac{i-1}{k}\right) - \frac{1}{n} \sum_{i=0}^{k-1} i\right. \\
 &\quad \left. + \frac{1}{2n^2} \sum_{i=0}^k \left(\frac{k+1}{2} + \frac{i-1}{k}\right)^2 - \frac{1}{2n^2} \sum_{i=0}^{k-1} i^2 + \xi\right\},
 \end{aligned}$$

where ξ satisfies (2.10)–(2.12). Thus, after some calculation,

$$B(k, k) = \frac{1}{k^{k+1}n} \exp\left\{\frac{2k}{n} - \frac{1}{kn} - \frac{k^3}{24n^2} + O\left(\frac{k^2}{n^2} + \frac{k^4}{n^3}\right)\right\}$$

and then

$$\begin{aligned}
 \sum_{k=3}^{n^{0.7}} A(k, k) &= \frac{1}{n} \sum_{k=3}^{n^{0.7}} \frac{C(k, k)}{k^{k+1}} \exp\left\{-\frac{k^3}{24n^2} + O\left(\frac{k}{n} + \frac{k^4}{n^3}\right)\right\} \\
 &= \frac{1}{n} \sum_{k=3}^{n^{0.7}} \frac{C(k, k)e^{-k^3/24n^2}}{k^{k+1}} \left\{1 + O\left(\frac{k}{n} + \frac{k^4}{n^3}\right)\right\}. \tag{2.17}
 \end{aligned}$$

Now (2.16) implies

$$\frac{1}{n} \sum_{k=3}^{n^{0.7}} \frac{C(k, k)e^{-k^3/24n^2}}{k^{k+1}} \times O\left(\frac{k}{n} + \frac{k^4}{n^3}\right) = O(n^{-5/3}) \tag{2.18}$$

and

$$\begin{aligned} \frac{1}{n} \sum_{k=3}^{n^{0.7}} \frac{C(k,k)(1 - e^{-k^3/24n^2})}{k^{k+1}} &= o(n^{-4/3}) + \frac{1}{n} \sum_{k=n^{2/3}/\ln n}^{n^{2/3} \ln n} \frac{C(k,k)(1 - e^{-k^3/24n^2})}{k^{k+1}} \\ &= o(n^{-4/3}) + \frac{\sqrt{\pi/8}}{n^{4/3}} \int_{x=0}^{\infty} x^{-3/2}(1 - e^{-x^3/24}) dx. \end{aligned} \tag{2.19}$$

It follows from (2.17), (2.18) and (2.19) that

$$\sum_{k=3}^{n^{0.7}} A(k,k) = \frac{1}{n} \sum_{k=3}^{\infty} \frac{C(k,k)}{k^{k+1}} - \frac{\sqrt{\pi/8}}{n^{4/3}} \int_{x=0}^{\infty} x^{-3/2}(1 - e^{-x^3/24}) dx + o(n^{-4/3}). \tag{2.20}$$

For $k > n^{0.7}$ we observe that $t_1(k, 0) \leq 2k$ in (2.8) and $t_m(k, 0) \leq 0$ for $m \geq 2$, so

$$B(k,k) \leq \frac{e^2}{k^{k+1}n}$$

and thus

$$A(k,k) \leq e^2 \frac{C(k,k)}{k^{k+1}n} = O\left(\frac{1}{k^{3/2}n}\right).$$

It follows from this that

$$\sum_{k=n^{0.7}}^n A(k,k) = O(n^{-1.35}) = o(n^{-4/3}). \tag{2.21}$$

We are almost done: we need to simplify the sum

$$\sum_{k=3}^{\infty} \frac{C(k,k)}{k^{k+1}}.$$

Now, by (2.16),

$$\sum_{k=3}^{\infty} \frac{2C(k,k)}{k^{k+1}} = \sum_{k=3}^{\infty} \frac{(k-1)!}{k^{k+1}} \sum_{i=0}^{k-3} \frac{k^i}{i!} = \sum_{i=0}^{\infty} \sum_{k=i+3}^{\infty} \frac{k^i}{k^{k+1}} \frac{(k-1)!}{i!}. \tag{2.22}$$

In the last double sum, let us also add the terms with $k = i + 2$, $k = i + 1$ and $k = i \geq 1$. The terms with $k = i + 2$ add up to

$$\sum_{k=2}^{\infty} \frac{k^{k-2}}{k^{k+1}} \frac{(k-1)!}{(k-2)!} = \sum_{k=2}^{\infty} \frac{k-1}{k^3} = \sum_{k=1}^{\infty} \frac{k-1}{k^3} = \zeta(2) - \zeta(3).$$

The terms with $k = i + 1$ add up to

$$\sum_{k=1}^{\infty} \frac{k^{k-1}}{k^{k+1}} \frac{(k-1)!}{(k-1)!} = \sum_{k=1}^{\infty} \frac{1}{k^2} = \zeta(2).$$

The terms with $k = i \geq 1$ add up to

$$\sum_{k=1}^{\infty} \frac{k^k}{k^{k+1}} \frac{(k-1)!}{k!} = \sum_{k=1}^{\infty} \frac{1}{k^2} = \zeta(2).$$

Consequently, (2.22) yields

$$\sum_{k=3}^{\infty} \frac{2C(k, k)}{k^{k+1}} = \zeta(3) - 3\zeta(2) + \sum_{k=1}^{\infty} \sum_{i=0}^k \frac{k^i}{k^{k+1}} \frac{(k-1)!}{i!} = \zeta(3) - 3\zeta(2) + \sum_{k=1}^{\infty} \sum_{i=0}^k \frac{k!}{i!} k^{i-k-2}. \tag{2.23}$$

We transform the sum further:

$$\begin{aligned} \sum_{k=1}^{\infty} \sum_{i=0}^k \frac{k!}{i!} k^{i-k-2} &= \sum_{k=1}^{\infty} \sum_{i=0}^k \binom{k}{i} (k-i)! k^{i-k-2} \\ &= \sum_{k=1}^{\infty} \sum_{i=0}^k \binom{k}{i} k^{-1} \int_{x=0}^{\infty} x^{k-i} e^{-kx} dx \\ &= \int_{x=0}^{\infty} \sum_{k=1}^{\infty} \sum_{i=0}^k k^{-1} \binom{k}{i} x^{k-i} e^{-kx} dx \\ &= \int_{x=0}^{\infty} \sum_{k=1}^{\infty} k^{-1} (1+x)^k e^{-kx} dx \\ &= \int_{x=0}^{\infty} -\log(1 - (1+x)e^{-x}) dx. \end{aligned}$$

Consequently, (2.23) yields

$$2 \sum_{k=3}^{\infty} \frac{C(k, k)}{k^{k+1}} = \zeta(3) - 3\zeta(2) - \int_{x=0}^{\infty} \log(1 - (1+x)e^{-x}) dx. \tag{2.24}$$

Together with (2.20) and (2.21), this verifies (b).

Case (c): $1 \leq k \leq n, j \geq 1$ (complex components). Let

$$\kappa_c(p) = \kappa_{c,n}(p) := \sum_{k=1}^{\infty} \sum_{j=1}^{\infty} \kappa(k, j, p), \tag{2.25}$$

that is, the number of complex components in $G_{n,p}$, and

$$f_n(p) = \mathbb{E} \kappa_c(p) = \sum_{k \geq 1} \sum_{j \geq 1} \mathbb{E} \kappa(k, j, p), \tag{2.26}$$

the expected number of complex components in $G_{n,p}$. The contribution to (2.1) from the complex components is thus $\int_{p=0}^1 f_n(p) dp$. We make a change of variables and let

$$p = n^{-1} + \lambda n^{-4/3}, \tag{2.27}$$

which means that we focus on the critical window. We will assume this relation between p and λ in the rest of the proof. We thus define $\bar{f}_n(\lambda) = f_n(p) = f_n(n^{-1} + \lambda n^{-4/3})$, and obtain the contribution, letting $\mathbf{1}\{\dots\}$ denote the indicator of an event,

$$\begin{aligned} \int_{p=0}^1 f_n(p) dp &= 1 - \frac{1}{n} + \int_{p=0}^1 (f_n(p) - \mathbf{1}\{p > 1/n\}) dp \\ &= 1 - \frac{1}{n} + n^{-4/3} \int_{\lambda=-n^{1/3}}^{n^{4/3}-n^{1/3}} (\bar{f}_n(\lambda) - \mathbf{1}\{\lambda > 0\}) d\lambda. \end{aligned} \tag{2.28}$$

We begin by showing that the integrand in the final integral converges pointwise. We define

$$\psi_2(t) = \sum_{\ell=2}^{\infty} w_{\ell} t^{\ell} = \psi(t) - 1 - \sqrt{\pi/8} t \tag{2.29}$$

(cf. (2.5)), and

$$F(x, \lambda) = \frac{1}{6}x^3 - \frac{1}{2}x^2\lambda + \frac{1}{2}x\lambda^2 = \frac{x}{2} \left(\lambda - \frac{x}{2} \right)^2 + \frac{1}{24}x^3. \tag{2.30}$$

Lemma 3.1 below proves that for any fixed $\lambda \in (-\infty, \infty)$, as $n \rightarrow \infty$,

$$\bar{f}_n(\lambda) \rightarrow f(\lambda) = \frac{1}{\sqrt{2\pi}} \int_{x=0}^{\infty} \psi_2(x^{3/2}) e^{-F(x,\lambda)} x^{-5/2} dx. \tag{2.31}$$

The next step is to use dominated convergence in (2.28). For this we use the following estimates. For convenience, we let $\kappa_c(n^{-1} + \lambda n^{-4/3})$ and its expectation $\bar{f}_n(\lambda)$ be defined for all real λ , by trivially defining $\kappa_c(p) = \kappa_c(0) = 0$ for $p < 0$ and $\kappa_c(p) = \kappa_c(1) = 1$ for $p > 1$.

There exist integrable functions $g_1(\lambda), g_2(\lambda), g_3(\lambda)$, not depending on n , such that

$$\bar{f}_n(\lambda) = \mathbb{E} \kappa_c(n^{-1} + \lambda n^{-4/3}) \leq g_1(\lambda), \quad \lambda \leq 0, \tag{2.32}$$

$$\mathbb{P}(\kappa_c(n^{-1} + \lambda n^{-4/3}) = 0) \leq g_2(\lambda), \quad \lambda \geq 0, \tag{2.33}$$

$$\bar{f}_n(\lambda) - 1 = \mathbb{E} \kappa_c(n^{-1} + \lambda n^{-4/3}) - 1 \leq g_3(\lambda), \quad \lambda \geq 0. \tag{2.34}$$

Equations (2.32), (2.33) and (2.34) are proved in Lemma 3.2 below.

Equation (2.33) implies that $1 - \bar{f}_n(\lambda) \leq g_2(\lambda)$ for $\lambda \geq 0$, and so

$$|\bar{f}_n(\lambda) - \mathbf{1}\{\lambda > 0\}| \leq \begin{cases} g_1(\lambda) & \lambda \leq 0, \\ g_2(\lambda) + g_3(\lambda) & \lambda > 0. \end{cases}$$

This justifies using dominated convergence in the integral in (2.28), and equation (3.1) implies

$$\int_{\lambda=-n^{1/3}}^{n^{4/3}-n^{1/3}} (\bar{f}_n(\lambda) - \mathbf{1}\{\lambda > 0\}) d\lambda \rightarrow c_{2c} = \int_{\lambda=-\infty}^{\infty} (f(\lambda) - \mathbf{1}\{\lambda > 0\}) d\lambda. \tag{2.35}$$

Hence (2.28) yields

$$\int_{p=0}^1 f_n(p) dp = 1 - \frac{1}{n} + c_{2c} n^{-4/3} + o(n^{-4/3}), \tag{2.36}$$

which is the desired result except for the expression for c_{2c} .

We transform the expression for c_{2c} in (2.35) by first writing it as

$$\begin{aligned} c_{2c} &= \lim_{A \rightarrow \infty} \left(-A + \int_{\lambda=-\infty}^A f(\lambda) d\lambda \right) \\ &= \lim_{A \rightarrow \infty} \left(-A + \frac{1}{\sqrt{2\pi}} \int_{\lambda=-\infty}^A \int_{x=0}^{\infty} \psi_2(x^{3/2}) e^{-F(x,\lambda)} x^{-5/2} dx d\lambda \right). \end{aligned} \tag{2.37}$$

By (2.29) we have $\psi_2(t) = O(t^2)$ for small t , which together with (1.7) shows that

$$\psi_2(t) = O(t^2 e^{t^2/24}), \quad t \geq 0,$$

and thus by (2.30), for all $x > 0$ and $\lambda \in (-\infty, \infty)$,

$$\psi_2(x^{3/2})e^{-F(x,\lambda)} \leq C_1 x^3 e^{-x(\lambda-x/2)^2/2}.$$

Hence, for $A > 0$, with the substitutions $x = 2A + s$ and $\lambda = A - t$,

$$\begin{aligned} \int_{x>2A} \int_{\lambda<A} \psi_2(x^{3/2})e^{-F(x,\lambda)} x^{-5/2} dx d\lambda &\leq C_1 \int_{x>2A} \int_{\lambda<A} e^{-x(\lambda-x/2)^2/2} x^{1/2} dx d\lambda \\ &= C_1 \int_{s>0} \int_{t>0} e^{-(2A+s)(t+s/2)^2/2} (2A+s)^{1/2} dt ds \\ &\leq C_1 \int_{s>0} \int_{t>0} e^{-(2A+s)(t^2/2+s^2/8)} (2A+s)^{1/2} dt ds \\ &= C_2 \int_{s>0} e^{-(2A+s)s^2/8} ds \leq C_3 A^{-1/2}. \end{aligned}$$

Similar estimates show also that

$$\int_{x<2A} \int_{\lambda>A} \psi_2(x^{3/2})e^{-F(x,\lambda)} x^{-5/2} dx d\lambda \leq C_4 \int_{s=0}^{2A} e^{-(2A-s)s^2/8} ds \leq C_5 A^{-1/2}.$$

Consequently, we can subtract and add these integrals to (2.37), yielding

$$c_{2c} = \lim_{A \rightarrow \infty} \left(-A + \frac{1}{\sqrt{2\pi}} \int_{\lambda=-\infty}^{\infty} \int_{x=0}^{2A} \psi_2(x^{3/2})e^{-F(x,\lambda)} x^{-5/2} dx d\lambda \right). \tag{2.38}$$

It follows from (2.30) that

$$\int_{\lambda=-\infty}^{\infty} e^{-F(x,\lambda)} d\lambda = e^{-x^3/24} \int_{\lambda=-\infty}^{\infty} e^{-x(\lambda-x/2)^2/2} d\lambda = e^{-x^3/24} \sqrt{2\pi/x}. \tag{2.39}$$

Hence, by Fubini, (2.38) yields

$$c_{2c} = \lim_{A \rightarrow \infty} \left(-A + \int_{x=0}^{2A} \psi_2(x^{3/2})e^{-x^3/24} x^{-3} dx \right) = \int_{x=0}^{\infty} \left(x^{-3} \psi_2(x^{3/2})e^{-x^3/24} - \frac{1}{2} \right) dx. \tag{2.40}$$

This completes the proof of Lemma 2.1 and the proof of Theorem 1.1. □

3. Auxiliary lemmas

Lemma 3.1. For any fixed $\lambda \in (-\infty, \infty)$, as $n \rightarrow \infty$,

$$\bar{f}_n(\lambda) \rightarrow f(\lambda) = \frac{1}{\sqrt{2\pi}} \int_{x=0}^{\infty} \psi_2(x^{3/2})e^{-F(x,\lambda)} x^{-5/2} dx. \tag{3.1}$$

Proof. We note first that the integral in (3.1) is convergent; for small x we have $\psi_2(x) = O(x^2)$ and for large x we have $\psi_2(x) = O(x^2 e^{x^2/24})$ by (1.5) and

$$e^{-F(x,\lambda)} \leq e^{-x^3/6 + \lambda x^2/2} = O(e^{-x^3/7})$$

by (2.30); remember that λ is fixed in the integral.

We convert the sum over k in (2.26) to an integral by setting $k = \lceil xn^{2/3} \rceil$. Thus

$$\bar{f}_n(\lambda) = f_n(p) = \int_{x=0}^{\infty} \sum_{j \geq 1} \mathbb{E} \kappa(\lceil xn^{2/3} \rceil, j, p) n^{2/3} dx. \tag{3.2}$$

For any fixed λ and fixed $x > 0$, $j \geq 1$, and $p = n^{-1} + \lambda n^{-4/3}$ and $k = \lceil xn^{2/3} \rceil$ as above, we have as $n \rightarrow \infty$, by (2.2) and (2.3) and standard calculations,

$$\begin{aligned} \mathbb{E} \kappa(k, j, p) &\sim \frac{n^k}{k!} \exp\left(-\frac{k^2}{2n} - \frac{k^3}{6n^2}\right) C(k, k + j) n^{-k-j} (1 + \lambda n^{-1/3})^k \exp(-p(nk - k^2/2)) \\ &\sim n^{-j} \frac{C(k, k + j)}{k!} \exp(-k - F(kn^{-2/3}, \lambda)) \\ &\sim (2\pi)^{-1/2} w_{j+1} k^{-1} \left(\frac{k^{3/2}}{n}\right)^j e^{-F(kn^{-2/3}, \lambda)} \\ &\sim n^{-2/3} (2\pi)^{-1/2} w_{j+1} x^{3j/2-1} e^{-F(x, \lambda)} \end{aligned}$$

(for further details see, e.g., [17, Section 4] or [1, Section 11.10]). Thus, as $n \rightarrow \infty$,

$$n^{2/3} \mathbb{E} \kappa(\lceil xn^{2/3} \rceil, j, p) \rightarrow (2\pi)^{-1/2} w_{j+1} x^{3j/2-1} e^{-F(x, \lambda)}. \tag{3.3}$$

Moreover, Bollobás [3, Theorem 5.20] has shown the uniform bound

$$C(k, k + j) \leq \left(\frac{C_6}{j}\right)^{j/2} k^{k+(3j-1)/2} \tag{3.4}$$

for some constant C_6 and all $k, j \geq 1$. Let $A \geq 1$ be a constant, and first consider only components of size $k \leq An^{2/3}$. For such k , all $j \geq 1$ and $p = n^{-1} + \lambda n^{-4/3}$, (2.2) and (3.4) yield, by calculations similar to those above,

$$\begin{aligned} \mathbb{E} \kappa(k, j, p) &\leq C_7 \frac{n^k}{k!} \exp\left(-\frac{k^2}{2n}\right) C(k, k + j) n^{-k-j} (1 + \lambda n^{-1/3})^{k+j} \exp(-p(nk - k^2/2 - j)) \\ &\leq C_8 n^{-j} \frac{C(k, k + j)}{k!} e^{-k+j \times o(1)} \\ &\leq C_8 n^{-j} \left(\frac{2C_6}{j}\right)^{j/2} k^{3j/2-1} \end{aligned}$$

(with C_3 possibly depending on A), and thus

$$n^{2/3} \mathbb{E} \kappa(k, j, p) \leq C_8 \left(\frac{C_9 A^{3/2}}{j}\right)^{j/2}.$$

The sum over j of the right-hand side converges, and thus (3.3) and dominated convergence yield (recalling (2.29))

$$\int_{x=0}^A \sum_{j \geq 1} \mathbb{E} \kappa(\lceil xn^{2/3} \rceil, j, p) n^{2/3} dx \rightarrow \frac{1}{\sqrt{2\pi}} \int_{x=0}^A \psi_2(x^{3/2}) e^{-F(x, \lambda)} x^{-5/2} dx. \tag{3.5}$$

For $k > An^{2/3}$ we use the fact shown in [17, (6.6)] that the expected number of vertices in tree components of size at most $n^{2/3}$ is $n - O(n^{2/3})$; consequently, the expected number of vertices in all components (complex or not) of size larger than $n^{2/3}$ is $O(n^{2/3})$, and the expected number of components larger than $An^{2/3}$ is $\leq C_{10}/A$. The left-hand side of (3.5) thus converges uniformly to the right-hand side of (3.2) as $n \rightarrow \infty$, and the result (3.1) follows from (3.5) by letting $A \rightarrow \infty$. □

Lemma 3.2. *There exist integrable functions $g_1(\lambda), g_2(\lambda), g_3(\lambda)$, not depending on n , such that*

(i)

$$\bar{f}_n(\lambda) = \mathbb{E} \kappa_c(n^{-1} + \lambda n^{-4/3}) \leq g_1(\lambda), \quad \lambda \leq 0,$$

(ii)

$$\mathbb{P}(\kappa_c(n^{-1} + \lambda n^{-4/3}) = 0) \leq g_2(\lambda), \quad \lambda \geq 0,$$

(iii)

$$\bar{f}_n(\lambda) - 1 = \mathbb{E} \kappa_c(n^{-1} + \lambda n^{-4/3}) - 1 \leq g_3(\lambda), \quad \lambda \geq 0.$$

Proof. We use the method in Janson [10]. We consider $G(n, p)$, $p \in [0, 1]$, as a random graph process in the usual way: we regard p as time, edges are added as p grows from 0 to 1, and an edge e is added at a time T_e with a uniform distribution on $[0, 1]$, with all T_e independent.

As $G(n, p)$ evolves, there are at first only tree components, but later unicyclic components and complex components appear as edges are added to the graph. If we consider only the complex components, a new complex component is created if a new edge is added to a unicyclic component, or if it joins two unicyclic components. (Note that these are the only possibilities: we do not regard the growth of an already existing complex component as creating a new complex component. Creation of a new complex component may occur one or several times. It is shown in [14] that it occurs only once with probability converging to $5\pi/18$, but we will not need this.) As evolution continues, the complex components may grow by merging with trees or unicyclic components, and they may merge with each other, until at the end only one complex component remains, containing all vertices.

Let $\varphi_n(k, p)$ be the intensity of creation on new complex components of size k , that is, the probability of creating a new complex component of size k in the interval $[p, p + dp]$ is $\varphi_n(k, p) dp$. (For $p < 0$, $p > 1$ or $k > n$, we set $\varphi_n(k, p) = 0$.) Further, let

$$\Phi_n(p) = \sum_{k \geq 1} \varphi_n(k, p),$$

the intensity of creation of complex components regardless of size. We change variables as above and also define

$$\begin{aligned} \psi_n(x, \lambda) &= n^{-2/3} \varphi_n(\lceil xn^{2/3} \rceil, n^{-1} + \lambda n^{-4/3}), \\ \Psi_n(\lambda) &= n^{-4/3} \Phi_n(n^{-1} + \lambda n^{-4/3}) = \int_{x=0}^{\infty} \psi_n(x, \lambda) dx. \end{aligned}$$

(The notation is not exactly as in [10], where the two ways of creating a complex component are treated separately, but the estimates are the same.)

We have

$$\varphi_n(k, p) = \binom{n}{k} \hat{C}(k) p^k (1 - p)^{(n-k)k + \binom{k}{2} - k - 1},$$

where $\hat{C}(k)$ is the number of ways to create a multicyclic component by either adding an edge to a unicyclic component on $[k]$ or adding an edge joining two unicyclic components whose vertex sets are complementary subsets of $[k]$. The first case contributes

$$C(k, k) \binom{k}{2} - k = O(k^{k+3/2})$$

to $\hat{C}(k)$ and the second

$$\begin{aligned} \frac{1}{2} \sum_{i=3}^{k-3} \binom{k}{i} C(i, i) C(k-i, k-i) i(k-i) &\leq C_{11} \sum_{i=3}^{k-3} \binom{k}{i} e^i i! e^{k-i} (k-i)! \\ &\leq C_{11} k e^k k! = O(k^{k+3/2}); \end{aligned}$$

hence

$$\hat{C}(k) = O(k^{k+3/2}) = O(k e^k k!).$$

(Compare the more precise [10, (2.30)].) The intensity $\psi_n(x, \lambda)$ is bounded in [10, (2.12)–(2.19)] by calculations similar to those in the proof of Lemma 3.1. (In these bounds, and our versions below, δ, δ_1, \dots are some positive constants.)

We use the results of [10] with some small modifications. Equation (2.12) of [10] shows (together with the comments after it) that

$$\psi_n(x, \lambda) \leq C_{12} x e^{-\delta x^3 - \delta x \lambda^2} \quad \text{for } k \leq \delta_1 n \quad \text{and} \quad -n^{1/3} \leq \lambda \leq \delta_2 n^{1/3}.$$

Then one line before (2.15) of [10] proves that

$$\psi_n(x, \lambda) \leq C_{13} x e^{-\delta x^3 - \delta_3 x \lambda n^{1/3}/3} \quad \text{for } k \leq \delta_3 n \text{ and } \lambda \geq \delta_2 n^{1/3}.$$

Because $\lambda \leq n^{4/3}$ always, it is legitimate to replace $-\delta_3 x \lambda n^{1/3}/3$ by $-\delta_3 x \lambda^{5/4}$ to give

$$\psi_n(x, \lambda) \leq C_{13} x e^{-\delta x^3 - \delta_3 x \lambda^{5/4}/3} \quad \text{for } k \leq \delta_3 n \text{ and } \lambda \geq \delta_2 n^{1/3}.$$

Then (2.17) of [10] proves that

$$\psi_n(x, \lambda) \leq C_{14} n e^{-2\delta_5 n} \quad \text{for } \min\{\delta_1, \delta_3\} n \leq k \leq n.$$

Using $\min\{\delta_1, \delta_3\} n^{1/3} \leq x \leq n^{1/3}$ and $\lambda \leq n^{4/3}$, we replace this by

$$\psi_n(x, \lambda) \leq C_{15} x e^{-\delta_5 x^3} (1 + \lambda^4)^{-1}.$$

We therefore have, for all x and λ ,

$$\begin{aligned} 0 &\leq \psi_n(x, \lambda) \leq g(x, \lambda) \\ &= C_{12} x e^{-\delta x^3 - \delta x \lambda^2} + C_{13} x e^{-\delta x^3 - \delta_3 x |\lambda|^{5/4}/3} + C_{15} x e^{-\delta_5 x^3} (1 + \lambda^4)^{-1} \end{aligned} \tag{3.6}$$

(recalling that $\psi_n(x, \lambda) = 0$ if $x > n^{1/3}$, $\lambda < -n^{1/3}$ or $\lambda > n^{4/3}$).

Integrating, we find

$$\Psi(\lambda) \leq \int_{x=0}^{\infty} g(x, \lambda) dx \leq \frac{C_{16}}{1 + |\lambda|^{5/2}}. \tag{3.7}$$

The number of complex components at any time is at most the number of complex components that have been created so far. Taking expectations, we thus obtain, using (3.7),

$$\bar{f}_n(\lambda) = \mathbb{E} \kappa_c(n^{-1} + \lambda n^{-4/3}) \leq \int_{\mu=-\infty}^{\lambda} \Psi(\mu) d\mu \leq \int_{\mu=-\infty}^{\lambda} \frac{C_{16}}{1 + |\mu|^{5/2}} d\mu. \tag{3.8}$$

This verifies (i), with $g_1(\lambda) = C_{17}(1 + |\lambda|^{3/2})^{-1}$ for $\lambda \leq 0$.

Similarly, if at some time there is no complex component, at least one complex component has to be created later. Thus,

$$\mathbb{P}(\kappa_c(n^{-1} + \lambda n^{-4/3}) = 0) \leq \int_{\mu=\lambda}^{\infty} \Psi(\mu) d\mu \leq \int_{\mu=\lambda}^{\infty} \frac{C_{16}}{1 + |\mu|^{5/2}} d\mu, \tag{3.9}$$

which verifies (ii) with $g_2(\lambda) = C_{18}(1 + \lambda^{3/2})^{-1}$ for $\lambda \geq 0$.

For (iii), let $Y(p) = \binom{\kappa_c(p)}{2}$ be the number of pairs of complex components in $G_{n,p}$. Since $\kappa_c(p) - 1 \leq Y(p)$, it suffices to estimate $\mathbb{E} Y(p)$.

If there is a pair of complex components in $G_{n,p}$, then these components have been created at times p_1 and p_2 with $p_1 \leq p_2 \leq p$. The intensity of this happening, with sizes $k_1 = \lceil x_1 n^{2/3} \rceil$ and $k_2 = \lceil x_2 n^{2/3} \rceil$ of the components at the moments of their creations, is bounded in [10, (2.24)–(2.26)] by

$$C_{19}g(x_1, \lambda_1)g(x_2, \lambda_2) d\lambda_1 d\lambda_2 dx_1 dx_2$$

(using modifications as above, and g is defined in (3.6)). Moreover, if the two components are still distinct components in $G_{n,p}$, then, at least (ignoring further conditions from the growth of the components), the original vertex sets of sizes k_1 and k_2 are not connected by any edge in the time interval $[p_2, p]$; the (conditional) probability of this is

$$\left(1 - \frac{p - p_2}{1 - p_2}\right)^{k_1 k_2} \leq (1 - (p - p_2))^{k_1 k_2} \leq e^{-k_1 k_2 (p - p_2)} \leq e^{-x_1 x_2 (\lambda - \lambda_2)}.$$

Consequently,

$$\begin{aligned} &\bar{f}_n(\lambda) - 1 \\ &\leq \mathbb{E} Y(n^{-1} + \lambda n^{-4/3}) \leq g_3(\lambda) \\ &= \int_{\lambda_1=-\infty}^{\lambda} \int_{\lambda_2=\lambda_1}^{\lambda} \int_{x_1=0}^{\infty} \int_{x_2=0}^{\infty} C_{19}g(x_1, \lambda_1)g(x_2, \lambda_2)e^{-x_1 x_2 (\lambda - \lambda_2)} d\lambda_1 d\lambda_2 dx_1 dx_2. \end{aligned}$$

This yields (iii), but it remains to verify that $\int_{\lambda=0}^{\infty} g_3(\lambda) d\lambda < \infty$. Indeed, by Fubini and (3.6),

$$\begin{aligned} & \int_{\lambda=-\infty}^{\infty} g_3(\lambda) d\lambda \\ &= \int_{\lambda_1=-\infty}^{\infty} \int_{\lambda_2=\lambda_1}^{\infty} \int_{x_1=0}^{\infty} \int_{x_2=0}^{\infty} C_{19} g(x_1, \lambda_1) g(x_2, \lambda_2) \\ & \quad \times \int_{\lambda=\lambda_2}^{\infty} e^{-x_1 x_2 (\lambda - \lambda_2)} d\lambda d\lambda_1 d\lambda_2 dx_1 dx_2 \\ &= \int_{\lambda_1=-\infty}^{\infty} \int_{\lambda_2=\lambda_1}^{\infty} \int_{x_1=0}^{\infty} \int_{x_2=0}^{\infty} C_{19} \frac{g(x_1, \lambda_1) g(x_2, \lambda_2)}{x_1 x_2} d\lambda_1 d\lambda_2 dx_1 dx_2 \\ & \leq C_{19} \left(\int_{\lambda=-\infty}^{\infty} \int_{x=0}^{\infty} \frac{g(x, \lambda)}{x} d\lambda dx \right)^2 < \infty. \end{aligned} \quad \square$$

4. Final remarks

Remark 3. We have shown that when the X_e are uniform $[0, 1]$ then $\mathbb{E}(L_n)$ converges to $\zeta(3)$ with an error term of order $1/n$. The constant c_1 is positive, and so for large n we have $\mathbb{E}(L_n) > \zeta(3)$. Fill and Steele [4] computed $\mathbb{E}(L_n)$ for $n \leq 8$. $\mathbb{E}(L_n)$ increased monotonically, and it was natural to conjecture from this that $\mathbb{E}(L_n)$ increases monotonically for all n . However, since $\mathbb{E}(L_n)$ converges to $\zeta(3)$ from above, we now see that this turns out not to be true. Note, however, that $c_2 < 0$, and that $|c_2|$ is much larger than c_1 . Thus we expect that $\mathbb{E} L_n > \zeta(3)$ only for very large n .

If our numerical estimates are correct, we have $|c_2|/c_1 \approx 45$, so a naive guess, ignoring higher-order terms, would be that $\mathbb{E} L_n > \zeta(3)$ for $n > 45^3 \approx 10^5$. We do not want to conjecture this, as we have no idea about the next term.

Remark 4. By (1.2), we obtain for $\mathbb{E}_{\text{exp}}(L_n)$ the same result as in Theorem 1.1 except that c_1 is increased by $\zeta(3)$ (while c_2 remains the same). This gives a somewhat simpler c_1 , which suggests that this version might be slightly simpler to analyse. Note that formula (2.1) holds for $\mathbb{E}_{\text{exp}}(L_n)$ if we replace $G_{n,p}$ by the multigraph where each pair of vertices is connected by a $\text{Po}(t)$ number of edges, and integrate for $t \in (0, \infty)$. This suggests that it might be profitable to make a version of the argument below using these multigraphs, but we have not pursued this. (Compare the use of multigraphs in [14].)

Acknowledgement

In an earlier version of this paper, we showed that

$$\mathbb{E}(L_n) = \zeta(3) + \frac{c_1 + o(1)}{n}.$$

Nick Read directed us to his article [24], which suggested that the $o(1/n)$ term could be replaced by $(c + o(1))/n^{4/3}$. This encouraged us to go the extra mile and find the next term, and prove Nick’s conjecture.

Appendix: Estimation of c_2

The constant c_2 in Theorem 1.1 is the sum of the three coefficients for $n^{-4/3}$ in Lemma 2.1(a–c), which we denote by c_{2a} , c_{2b} and c_{2c} . By the change of variable $t = x^3/24$, and integration by parts (cf. [22, § 5.9.5]), we obtain, as stated in Remark 2,

$$c_{2a} = \frac{24^{-2/3}}{3} \int_{t=0}^{\infty} t^{-5/3}(e^{-t} - 1) dt = -\frac{1}{8}3^{-2/3}\Gamma\left(\frac{1}{3}\right) = -0.16098\dots, \tag{A.1}$$

$$c_{2b} = \sqrt{\frac{\pi}{8}} \frac{24^{-1/6}}{3} \int_{t=0}^{\infty} t^{-7/6}(e^{-t} - 1) dt = -\frac{1}{2}3^{-1/6}\sqrt{\pi}\Gamma\left(\frac{5}{6}\right) = -0.83298\dots \tag{A.2}$$

The coefficient c_{2c} is given by an integral in Lemma 2.1; see also (2.40). To evaluate c_{2c} , we change variables by $x = y^{1/3}$ and use the definition (2.29) of ψ_2 to obtain

$$\begin{aligned} c_{2c} &= \frac{1}{3} \int_{y=0}^{\infty} \left(y^{-1}\psi_2(y^{1/2}) - \frac{1}{2}e^{y/24} \right) e^{-y/24}y^{-2/3} dy \\ &= \frac{1}{3} \int_{y=0}^{\infty} \sum_{k=1}^{\infty} \left(w_{2k}y^{k-1} + w_{2k+1}y^{k-1/2} - \frac{y^{k-1}}{2 \cdot 24^{k-1}(k-1)!} \right) e^{-y/24}y^{-2/3} dy. \end{aligned} \tag{A.3}$$

We interchange the order of integration and summation, which is justified below, and obtain

$$\begin{aligned} c_{2c} &= \frac{1}{3} \sum_{k=1}^{\infty} \int_{y=0}^{\infty} \left(w_{2k}y^{k-1} + w_{2k+1}y^{k-1/2} - \frac{y^{k-1}}{2 \cdot 24^{k-1}(k-1)!} \right) e^{-y/24}y^{-2/3} dy \\ &= \frac{24^{1/3}}{3} \sum_{k=1}^{\infty} \left(w_{2k}24^{k-1}\Gamma(k-2/3) + w_{2k+1}24^{k-1/2}\Gamma(k-1/6) - \frac{\Gamma(k-2/3)}{2\Gamma(k)} \right). \end{aligned} \tag{A.4}$$

We note that (2.4) and (1.4) yield, together with Stirling’s formula, $w_\ell \sim 6 \cdot 24^{-\ell/2}/\Gamma(\ell/2)$, which implies that

$$w_{2k}24^{k-1}\Gamma(k-2/3) \sim w_{2k+1}24^{k-1/2}\Gamma(k-1/6) \sim \frac{1}{4}k^{-2/3} \quad \text{as } k \rightarrow \infty,$$

so the three terms in the sum in (A.4) are all of order $k^{-2/3}$, showing that we cannot sum them separately. However, their leading terms cancel. A more precise calculation using (1.6) yields

$$\mathbb{E} \mathcal{B}_{\text{ex}}^r = \sqrt{18}r \left(\frac{r}{12e} \right)^{r/2} (1 + O(r^{-1})), \quad r > 0, \tag{A.5}$$

and thus by (2.4) and Stirling’s formula,

$$w_\ell = \frac{3\sqrt{\ell}}{\sqrt{\pi}} \left(\frac{e}{12\ell} \right)^{\ell/2} (1 + O(\ell^{-1})) = \frac{6 \cdot 24^{-\ell/2}}{\Gamma(\ell/2)} (1 + O(\ell^{-1})), \quad \ell \geq 1. \tag{A.6}$$

Hence,

$$w_{2k}24^{k-1}\Gamma(k-2/3) = \frac{1}{4}k^{-2/3}(1 + O(k^{-1})), \quad \text{as } k \rightarrow \infty, \tag{A.7}$$

and the same estimate holds for $w_{2k+1}24^{k-1/2}\Gamma(k-1/6)$, while

$$\Gamma(k-2/3)/\Gamma(k) = k^{-2/3}(1 + O(k^{-1})).$$

Consequently, the summand in (A.4) is $O(k^{-5/3})$.

The constants w_k can be computed by a recursion formula (see [30] and [12]), and a numerical summation of the first 1000 terms in (A.4) yields -0.7331 . It can be shown, using (1.6) with the further second-order term given in [15] (which replaces $O(x^{-2})$ by $-\frac{1}{9}x^{-2} + O(x^{-4})$), that the terms in the sum in (A.4) are $\sim -\frac{1}{6}k^{-5/3}$, and using this to estimate the sum of the terms with $k > 1000$ yields the estimate $c_{2c} \approx -0.7355$, which together with (A.1)–(A.2) yields

$$c_2 \approx -1.7295. \tag{A.8}$$

The tail estimate is not rigorous. Replacing $O(x^{-4})$ by $\leq Cx^{-4}$ for some estimate C is what is needed to make the tail estimate rigorous. Nevertheless, it seems unlikely that the estimate in (A.8) is very far off.

To justify the interchange of summation and integration above, by Fubini’s theorem it is sufficient to verify that

$$\sum_{k=1}^{\infty} \int_{y=0}^{\infty} \left| w_{2k} y^{k-1} + w_{2k+1} y^{k-1/2} - \frac{y^{k-1}}{2 \cdot 24^{k-1} (k-1)!} \right| e^{-y/24} y^{-2/3} dy < \infty. \tag{A.9}$$

Indeed, we claim that the integral in (A.9) is $O(k^{-7/6})$. Using (A.7), its analogue for $2k + 1$, and $\Gamma(k - 2/3)/\Gamma(k) = k^{-2/3}(1 + O(k^{-1}))$, it follows easily that the integral is, after another change of variable $t = y/24$,

$$\frac{24^{1/3}}{4} k^{-2/3} \int_{t=0}^{\infty} \left| \frac{t^{k-7/6}}{\Gamma(k-1/6)} - \frac{t^{k-5/3}}{\Gamma(k-2/3)} \right| e^{-t} dt + O(k^{-5/3}). \tag{A.10}$$

Let I_k denote the integral in (A.10). By the Cauchy–Schwarz inequality,

$$\begin{aligned} I_k^2 &\leq \int_{t=0}^{\infty} t^{k-1} e^{-t} dt \cdot \int_{t=0}^{\infty} \left(\frac{t^{k-7/6}}{\Gamma(k-1/6)} - \frac{t^{k-5/3}}{\Gamma(k-2/3)} \right)^2 t^{1-k} e^{-t} dt \\ &= \Gamma(k) \left(\frac{\Gamma(k-2/6)}{\Gamma(k-1/6)^2} - 2 \frac{\Gamma(k-5/6)}{\Gamma(k-1/6)\Gamma(k-4/6)} + \frac{\Gamma(k-8/6)}{\Gamma(k-4/6)^2} \right) \\ &= O(k^{-1}). \end{aligned}$$

Consequently, $I_k = O(k^{-1/2})$, which shows that (A.10) is $O(k^{-7/6})$, and thus (A.9) holds as claimed above.

References

- [1] Alon, N. and Spencer, J. H. (2008) *The Probabilistic Method*, third edition, Wiley.
- [2] Beveridge, A., Frieze, A. M. and McDiarmid, C. J. H. (1998) Minimum length spanning trees in regular graphs. *Combinatorica* **18** 311–333.
- [3] Bollobás, B. (2001) *Random Graphs*, second edition, Cambridge University Press.
- [4] Fill, J. A. and Steele, J. M. (2005) Exact expectations of minimal spanning trees for graphs with random edge weights. In *Stein’s Method and Applications*, Singapore University Press, pp. 169–180.
- [5] Flaxman, A. (2007) The lower tail of the random minimum spanning tree. *Electron. J. Combin.* **14** N3.

- [6] Frieze, A. M. (1985) On the value of a random minimum spanning tree problem. *Discrete Appl. Math.* **10** 47–56.
- [7] Frieze, A. M. and McDiarmid, C. J. H. (1989) On random minimum length spanning trees. *Combinatorica* **9** 363–374.
- [8] Frieze, A. M., Ruzinkó, M. and Thoma, L. (2000) A note on random minimum length spanning trees. *Electron. J. Combin.* **7** R41.
- [9] Gamarnik, D. (2005) The expected value of random minimal length spanning tree of a complete graph. In *Proc. Sixteenth Annual ACM–SIAM Symposium on Discrete Algorithms: SODA 2005*, ACM, pp. 700–704.
- [10] Janson, S. (1993) Multicyclic components in a random graph process. *Random Struct. Alg.* **4** 71–84.
- [11] Janson, S. (1995) The minimal spanning tree in a complete graph and a functional limit theorem for trees in a random graph. *Random Struct. Alg.* **7** 337–355.
- [12] Janson, S. (2007) Brownian excursion area, Wright’s constants in graph enumeration, and other Brownian areas. *Probab. Surveys* **3** 80–145.
- [13] Janson, S. and Chassaing, P. (2004) The center of mass of the ISE and the Wiener index of trees. *Electron. Comm. Probab.* **9** 178–187.
- [14] Janson, S., Knuth, D. E., Łuczak, T. and Pittel, B. (1993) The birth of the giant component. *Random Struct. Alg.* **3** 233–358.
- [15] Janson, S. and Louchard, G. (2007) Tail estimates for the Brownian excursion area and other Brownian areas. *Electron. J. Probab.* **12** 1600–1632.
- [16] Janson, S., Łuczak, T. and Ruciński, A. (2000) *Random Graphs*, Wiley.
- [17] Janson, S. and Spencer, J. (2007) A point process describing the component sizes in the critical window of the random graph evolution. *Combin. Probab. Comput.* **16** 631–658.
- [18] Li, W. and Zhang, X. (2009) On the difference of expected lengths of minimum spanning trees. *Combin. Probab. Comput.* **18** 423–434.
- [19] Louchard, G. (1984) Kac’s formula, Lévy’s local time and Brownian excursion. *J. Appl. Probab.* **21** 479–499.
- [20] Louchard, G. (1984) The Brownian excursion area: A numerical analysis. *Comput. Math. Appl.* **10** 413–417. Erratum: *Comput. Math. Appl.* A **12** (1986) 375.
- [21] Nishikawa, J., Otto, P. T. and Starr, C. (2012) Polynomial representation for the expected length of minimal spanning trees. *Pi Mu Epsilon J.* **13** 357–365.
- [22] NIST Digital Library of Mathematical Functions. <http://dlmf.nist.gov/>
- [23] Penrose, M. (1998) Random minimum spanning tree and percolation on the n -cube. *Random Struct. Alg.* **12** 63–82.
- [24] Read, N. (2005) Minimum spanning trees and random resistor networks in d dimensions. *Phys. Rev. E* **72** 036114.
- [25] Rényi, A. (1959) Some remarks on the theory of trees. *Publ. Math. Inst. Hungar. Acad. Sci.* **4** 73–85.
- [26] Spencer, J. (1997) Enumerating graphs and Brownian motion. *Comm. Pure Appl. Math.* **50** 291–294.
- [27] Steele, J. M. (1987) On Frieze’s $\zeta(3)$ limit for lengths of minimal spanning trees. *Discrete Appl. Math.* **18** 99–103.
- [28] Steele, J. M. (2002) Minimum spanning trees for graphs with random edge lengths. In *Mathematics and Computer Science II: Algorithms, Trees, Combinatorics and Probabilities* (B. Chauvin *et al.*, eds), Springer, pp. 223–245.
- [29] Wästlund, J. (2009) An easy proof of the $\zeta(2)$ limit in the random assignment problem. *Electron. Comm. Probab.* **14** 261–269.
- [30] Wright, E. M. (1977) The number of connected sparsely edged graphs. *J. Graph Theory* **1** 317–330.