

Organizational Citizenship Behavior in Schools: Validation of a Questionnaire

Paula C. Neves¹, Rui Paixão², Madalena Alarcão² and A. Duarte Gomes²

¹ Instituto Politécnico de Coimbra (Portugal)

² Universidade de Coimbra (Portugal)

Abstract. The present study examines the psychometric properties (including factorial validity) of an organizational citizenship behavior (OCB) scale in a school context. A total of 321 middle and high school teachers from 59 schools in urban and rural areas of central Portugal completed the OCB scale at their schools. The confirmatory factor analysis validated a hierarchical model with four latent factors on the first level (altruism, conscientiousness, civic participation and courtesy) and a second order factor (OCB). The revised model fit with the data, $\chi^2/df = 1.97$; CFI = .962; GFI = .952, RMSEA = .05. The proposed scale (comportamentos de cidadania organizacional em escolas- Revista CCOE-R)- is a valid instrument to assess teacher's perceptions of OCB in their schools, allowing investigation at the organizational level of analysis.

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The Organizational Citizenship Behavior concept (OCB), defined as “individual behavior that is discretionary, not directly or explicitly recognized by the formal reward system, and that in the aggregate promotes the effective functioning of the organization” (Organ, 1988, p.4), has been studied at an entrepreneurial level (Podsakoff, MacKenzie, Paine, & Bachrach, 2000; Podsakoff, Whiting, Podsakoff, & Blume, 2009) and, more recently, at an educational level (DiPaola, Tarter, & Hoy, 2005; DiPaola & Tschannen-Moran, 2001; Oplatka, 2006; Oplatka & Stundi, 2011; Somech & Drach-Zahavy, 2000; Yilmaz & Taşdan, 2009).

Initially proposed by Organ (1988), the concept refers to the individual behavior and actions of an employee that objectively benefit organizational effectiveness but cannot be expected from the employee, due to not being part of their formal job description, nor fostered, due to the lack of a formal reward system for it (Organ, 1990). It is positive employee behavior which strongly contributes to organizational performance and efficiency (Kidwell, Mossholder, & Bennett, 1997; Organ, 1988, 1990; Podsakoff et al., 2000; Podsakoff et al., 2009).

Studies in this field have focused more on the identification of antecedent and consequence factors of OCB rather than on the clarification of the nature and dimension of the construct. With regards to the antecedent factors, variables have been identified which

are related to personality characteristics (Borman, Penner, Allen, & Motowidlo, 2001), personal motivation (Dávila & Finkelstein, 2010; Rioux & Penner, 2001), job attitudes (Organ & Ryan, 1995), task characteristics (Cardona, Lawrence, & Bentler, 2004; Chiu & Chen, 2005; Passos & Caetano, 2000), context characteristics (Jex, Adams, Bachrach, & Sorenson 2003), and leadership types (Podsakoff et al., 2000), among others. As regards the consequence factors, studies have shown the relationship between OCB and organizational performance measures, such as productivity, efficiency, cost reduction, client satisfaction and turnover (Podsakoff et al., 2009).

As of yet, there is no consensus on the dimensionality of OCB, nor is the relationship between the construct and its dimensions clearly defined, although it is consensual that it is a multidimensional construct (Law, Wong, & Chen, 2005). Various configurations and multiple instruments for evaluation of the OCB structural dimension have been proposed (Graham, 1991; Organ, 1990; Podsakoff & MacKenzie, 1994; van Dyne, Graham, & Dienesch, 1994; van Dyne & LePine, 1998; van Scotter & Motowidlo, 1996; Williams & Anderson, 1991). Of those proposed two main types stand out: those that associate behavior according to its nature (Organ, 1988; Podsakoff, MacKenzie, Moorman, & Fetter, 1990) and those that associate behavior according to its objective (Skarlicki & Lathan, 1995; Williams & Anderson, 1991).

In the first instance, although in literature specific to the field more than 30 possible types of OCB are identified (Podsakoff et al., 2000), the five dimensional structure initially proposed by Organ (1988) is the most

Correspondence concerning this article should be addressed to Paula Costa Neves. Escola Superior de Educação. Rua Dom João, III. Solum. 3030-329. Coimbra (Portugal).
E-mail: pneves@esec.pt

consensual (Rego, 2002). These dimensions incorporate behavior of different natures and include: altruism (behavior with the intention of helping a specific work colleague); conscientiousness (behavior that goes beyond the minimum role requirements of the organization, e.g. break management); courtesy (behavior aimed at preventing work-related conflicts); sportsmanship (tolerance to less-than-ideal circumstances without complaining) and civic virtue (behavior indicating concern and active interest in the life of the organization).

In the second instance, Williams and Anderson (1991) present OCB as a bi-dimensional construct where each dimension is structured according to whom the behavior was directed at: 1) OCBI or behaviors that immediately benefit other specific individuals; 2) OCBO or behaviors that benefit the organization as a whole.

With regards to educational institutions, literature suggests that in schools that function in an adequate and effective fashion, the teaching staff do not restrict their activities to those strictly necessary but foster others that, not being of a compulsory nature, make a difference because they have an impact in student performance (DiPaola & Hoy, 2005). As a matter of fact, in order to achieve their objectives, schools depend on the availability of the teachers to make significant efforts, beyond the formal requirements of their function (Bogler & Somech, 2004; Somech & Bogler, 2002). Studies have shown the positive consequences OCB have had on schools, namely as regards: discipline and the fostering of a positive response from superiors, students and colleagues (Oplatka, 2009); the success of students (DiPaola & Hoy, 2005; Oplatka, 2009); civic behavior and help of the students (Oplatka, 2009); coordination of activities between staff members and their contribution for the creations of a more pleasant workplace (Podsakoff et al., 2000); a positive relationship with the school image and a cooperative and collegial school environment (Oplatka, 2009).

Although OCB are considered to be extra-role behaviors, the difficulty in defining what is or is not part of one's function is notorious when it comes to teaching career (DiPaola & Neves, 2009). In educational organizations, the role of a teacher encompasses a wide spectrum of tasks and responsibilities (in and out of the classroom) and it is therefore difficult to objectively define the boundaries of the profession, depending mostly on individual representations regarding the essence of the teaching career, or in other words professionalism (Gimeno, 1995). Teaching is a field of activity with a strong emphasis on helping other individuals; therefore some aspects of OCB may eventually overlap with teaching responsibilities (DiPaola & Neves, 2009; DiPaola & Tschannen-Moran, 2001; Jimmieson, Hannam, & Yeo, 2010).

In any case, studies focusing on OCB in an educational context are scarce (Jimmieson et al., 2010; Oplatka & Stundi, 2011) which makes the nature of the construct in an educational environment neither very clear nor consensually operationalized. Many of the studies have attempted to identify OCB antecedents in relation with other variables (Jimmieson et al., 2010; Somech & Ron, 2007; Yilmaz & Taşdan, 2009). Among the studies attempting to clarify and measure the concept in an educational environment, two tendencies stand out: one that considers OCB as a three dimensional construct and another that presents OCB as one dimensional. In the first instance, the three dimensions are organized according to whom the behavior is directed at (Christ, Van Dick, Wagner, & Stellmacher, 2003; Oplatka, 2006; Somech & Drach-Zahavy, 2000) and include: (1) behavior directed at the students; (2) behavior directed at colleagues; and (3) behavior directed at the organization as a whole. Based on this model, Somech and Drach-Zahavy (2000) developed the first instrument designed specifically for application in an educational environment (Extra Role Behavior-Schools).

The one dimensional perspective (DiPaola & Tschannen-Moran, 2001) is based on the nature of the educational institutions. A school is an organization structured to help by which the distinction between helping the student or the school is artificial, seeing that all behaviors have the same objective: helping the students in being successful. With this assumption in mind, the authors developed an unifactorial scale on OCB for schools (OCBSS) of which the more parsimonious version (OCB-Scale: DiPaola et al., 2005) was translated into various languages and adapted for use in various countries (DiPaola & Neves, 2009; Yilmaz & Taşdan, 2009).

In addition, and due to the lack of consensus regarding the construct in an educational environment, that the OCB may have a different operationalization due to cultural differences (Farh, Zhong, & Organ, 2004), makes even more pertinent its inter-cultural clarification as well as validation of the instruments used in those evaluations.

The objective of the work presented here is to contribute towards the clarification of the OCB concept in Portuguese schools, and in so doing validates an instrument of measure specifically designed to collect data regarding this subject.

Method

Participants

The sample under study includes 321 teachers of students between 10-15 years old at 59 schools in urban and rural areas of Central Portugal. The sample

includes teachers who teach in 26 of the 28 discipline groups. Not represented are 3rd cycle basic education teachers in Agricultural Science and Music. The average age of the participants is 40.3 ($SD = 8.3$), of which 70 (21.8%) are male and 251 (78.2%) female. From this sample 79.5% of the teachers have permanent work contracts and 20.2% have temporary work contracts. The average duration of employment is 14.93 years and 82.2% have more than seven years of experience in the profession.

Measures

The Organizational Citizenship Behavior Questionnaire for Schools (OCBS) (Neves, 2010, see Annex 1) is a questionnaire that is completed by the individual, with 16 items arranged in Likert type (classification of 1 to 6) ordinal scale, which registers the perception each teacher has about the OCB at his/her school and allows it, through data aggregation to be studied at an organizational level. This instrument was developed based on items from *OCB-Scale* (DiPaola et al., 2005) and *ERB Schools* (Somech & Drach-Zahavy, 2000) and rests on the conception of OCB as a four dimensional construct referenced to the nature of behaviors, to note: (1) altruism, which is a component of helping peers who directly or indirectly improve their performance (items 1, 2, 10, 14, and 16); (2) conscientiousness, which encompasses behaviors relating to work effort, such as, for example, effective time management in order to maximize the effectiveness of tasks (items 3, 5, 11, and 13); (3) civic participation, which encompasses behaviors relating to participation in tasks with the objective of promoting organizational interests as a whole (items 7, 8, 12 and 15); courtesy, which includes behavior aimed at preventing problems by planning ahead and devising solutions for predictable changes (items 4, 6 and 9).

Procedures

The data was collected at different schools by teachers studying for post-graduation and master degrees at the Escola Superior de Educação de Coimbra, once they had been briefed on how to uniformly administer the instrument.

Sensitivity of the items was studied taking into consideration asymmetry and kurtosis coefficients. For the reliability study Cronbach's alpha index was applied to each of the factors and, for the total scale, a standardized alpha since it was considered to be a better estimator of true reliability (Kamata, Turhan, & Darandari, 2003). Factorial validation and the second-order order structural hierarchy were studied using AMOS 20 software. In this analysis, the following adjustment indexes were used χ^2/df ; Comparative Fit

Index (CFI); Goodness of Fit Index (GFI), Root Mean Square Error of Approximation (RMSEA). The quality of the modified models and of the second-order factor structure was also evaluated and compared using Akaike Information Criterion (AIC), Bays Information Criterion (BIC) and Browne-Cudeck criterion (BCC) information criteria. The criterion used were: χ^2/df between 1 and 2; CFI and GFI greater than .9 RMSEA lower than .5; it was assumed in the information criteria lower values indicated a better adjustment to the model (Byrne, 2010).

Results

Temporal stability

Temporal stability was assessed by means of retesting of a sub-sample of 187 teachers within 3 weeks of the first test. The data showed the following correlations between dimensions: Altruism .62 ($p < .01$); Conscientiousness .68 ($p < .01$); Civic Participation .58 ($p < .01$); Courtesy .57 ($p < .01$) and .69 ($p < .01$) on the global scale score, by which the temporal stability is considered to be acceptable (Robinson, Shaver, & Wrightsman, 1991).

Sensitivity

The sensitivity of the 16 items was evaluated using asymmetry and kurtosis coefficients and respective critical ratios (Table 1). Asymmetry and kurtosis coefficients with absolute values greater than 1 were not identified, therefore, it is considered that there are no sensitivity or significant deviations from normal (Maroco, 2003). Because all manifest variables presents asymmetry and kurtosis near zero the assumption of multivariate normality is plausible (Maroco, 2010).

Factorial validity

The factorial validity of the instrument was tested on four models: model 1 (unifactorial), model 2 (four first-order factors), model 3 (four modified first-order factors) and model 4 (four modified first-order factors with 1 second-order factor). The sample size and plausibility of the normality justify the use of the estimate method of maximum likelihood (ML) because its produces results that are more efficient and consistent and is more robust to violation of the assumption of normality if the asymmetry and flattening of the manifest variables are not too large (Maroco, 2010).

The confirmatory factor analysis of model 1 showed adjusted fit indexes of poor or unacceptable factorial validity indicating the existence of more than one factor (Table 2).

The confirmatory factor analysis of the second model (four first-order factors) showed questionable

Table 1. Sensitivity of the 16-item CCOE-R item

Item	Mdn	M	SE	Sk	Sk/SESk	Ku	Ku/SEKu	Range
1	4	3.93	.072	-.39	-2.85	-.55	-2.01	1-6
2	3	3.18	.060	.13	0.98	-.11	-0.40	1-6
3	5	4.42	.065	-.55	-4.04	-.55	-2.01	2-6
4	3	3.13	.079	.25	1.82	-.76	-2.80	1-6
5	5	4.46	.060	-.57	-4.22	.09	0.32	1-6
6	4	3.68	.074	-.08	-0.62	-.78	-2.88	1-6
7	4	3.97	.063	-.32	-2.38	-.41	-1.49	1-6
8	3	3.51	.073	.03	0.19	-.68	-2.49	1-6
9	3	3.20	.078	.25	1.85	-.77	-2.84	1-6
10	3	3.15	.069	.25	1.85	-.48	-1.77	1-6
11	4	4.29	.062	-.34	-2.46	-.41	-1.51	1-6
12	3	3.38	.067	-.04	-0.32	-.51	-1.90	1-6
13	5	4.57	.057	-.44	-3.24	-.15	-0.56	1-6
14	4	3.99	.073	-.43	-3.17	-.43	-1.59	1-6
15	3	2.81	.073	.56	4.15	-.29	-1.06	1-6
16	4	4.21	.066	-.34	-2.53	-.51	-1.88	1-6

Note: Mdn=Median; M=Mean; SE= Standard Error; Sk=Asymmetry; Ku= kurtosis; Sk/SESk = Skcritical ratio; Ku/SEKu = Kucritical ratio.

Table 2. Adjustment indexes for each proposed model

Factorial models	χ^2	gl	χ^2/gl	CFI	GFI	RMSEA	AIC	BIC	BCC
1.One factor	454.08	104	4.37	0.82	0.83	0.103	518	638	521
2.Four first-order factors	268.23	98	2.73	0.91	0.90	0.074	344	487.5	348.5
3.Four first-order modified factors	96.48	48	2.01	0.96	0.95	0.056	156.5	269.6	159
4.Four first-order modified factors with 1 second-order factor	98.72	50	1.97	0.96	0.95	0.055	154.7	260.3	157.1

adjustment of fit indexes (Table 2), although the standardized factor values of all the items are greater than .5. The reliability of the first-order factors was assessed using Cronbach's alpha internal consistency coefficient and presents the following values: altruism .83; conscientiousness .77; civic participation .75; courtesy .76. The total scale presents a standardized alpha of .90.

The identification of a questionable factorial validity value justified the adjustment of the original model in accordance with the modification indices supplied by the program. In the third model (four modified first-order factors), items 10 and 16 relating to altruism, item 5 relating to conscientiousness and item 12 relating to civic participation, were removed because the modification indices suggested the saturation of these items on factors other than those suggested initially. The final modified scale presents the same four factors but with three items in each one. The adjustment of fit indexes for model 3 range between good and very good (Table 2). The factors are all positively correlated: altruism is positively correlated with conscientiousness

($r = .85$; $p < .001$), with civic participation ($r = .69$; $p < .001$) and with courtesy ($r = .78$; $p < .001$); conscientiousness correlates with civic participation ($r = .68$; $p < .001$) and with courtesy ($r = .66$; $p < .001$); and finally, civic participation correlates with courtesy ($r = .60$; $p < .001$).

The theoretical framework considered (Organ, 1988) and the high correlation between factors pointed to the existence of a second-order factor, by which the hypothesis was tested. For this new model (model 4: four modified first-order factors with 1 second-order factor), organized hierarchically, the OCB are depicted as a second-order factor (Figure 1) with adjustment of fit indexes ranging between good and very good (Table 2) indicating a high factorial validity for this revised scale.

In this hierarchical model the second-order factors altruism, conscientiousness, civic participation and courtesy (each with three items) present alpha values of .73, .71, .70 and .76 respectively. The global scale presents a standardized alpha of .87.

In comparison with the OCBS proposed in the second model tested this new structure, which we will call *Escala de Comportamentos de Cidadania Organizacional em Escolas – Revista (CCOE-R)* (Organizational Citizenship Behavior Scale in Schools – Revised) (Figure 1), presents a significantly better and parsimonious adjustment [$\Delta\chi^2(48) = 169.5; p < .001$] since the AIC, BIC and BCC criteria present lower values than in the original scale ($\Delta AIC = 190; \Delta BIC = 227.2; \Delta BCC = 191.4$).

Thus, the OCB can be considered a second-order factor with estimated numerical indexes based on the factorial scores. The standardized values of the OCB levels are calculated based on the following expression (1):

$$OCB = .037It_9 + .05It_6 + .075It_4 + .021It_{15} + .047It_8 + .082It_7 + .112It_{13} + .098It_{11} + .054It_3 + .138It_{14} + .116It_2 + .098It_1. \quad (1)$$

The scores for the four sub-scales are obtained by expressions (2), (3), (4) and (5):

$$\text{Altruism} = .130It_1 + .153It_2 + .182It_{14} \quad (2)$$

$$\text{Conscientiousness} = .96It_3 + .173It_{11} + .198It_{13} \quad (3)$$

$$\text{Civic Participation} = .358It_7 + .206It_8 + .09It_{15} \quad (4)$$

$$\text{Courtesy} = .326It_4 + .218It_6 + .162It_9 \quad (5)$$

Discussion

The nature of the OCB construct and the instruments used to study it in an educational environment are not, yet, consensual and this may be related with its low level of study (Jimmieson et al., 2010). In this article

we studied OCBS characteristics on the Portuguese teacher population and propose an instrument of measure to evaluate the perceptions of teachers regarding the application of OCB in schools.

The four factor structure, introduced in the initial version of the instrument (Neves, 2010), held up and was expanded to a second-order hierarchical structure, which fitted the data better.

The multi-factorial four factor structure fits the taxonomy subgroup that organizes the OCBs based on the nature of behavior and identifies itself more closely with Organ’s initial proposal (1988) for the business environment. The global similarity between the two structures shows that, although the teaching profession has specific characteristics, what may be described as extra-role behavior may be of an equivalent nature to any other profession. In other words, “activities that support the social and psychological environment in which task performance takes place” (Organ, 1997, p.95) may be carried out by teachers using behavior that, although specific to the teaching profession, are equivalent, in their nature, to those of any other profession.

For any scale of measure all the items should measure the same construct, with a determined margin of error. In a multi-dimensional construct, the dimensions represent their different materializations in which case

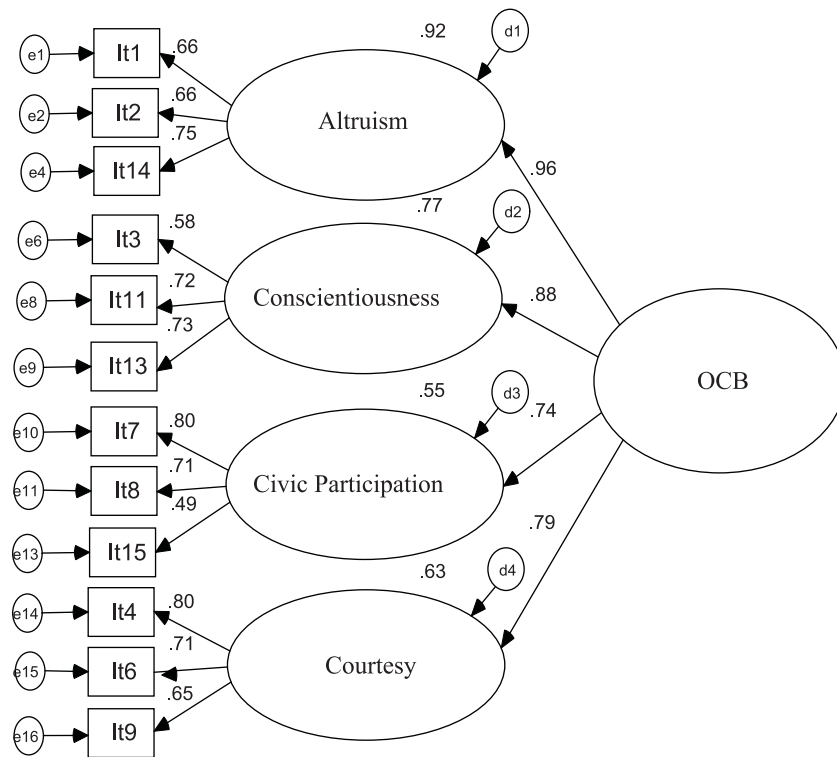


Figure 1. Four modified first-order factors with one second-order factor, including the standard factor influences for each of the items and R² values for each of the first-order factors.

the construct is defined by the amount of correlated variance shared by the dimensions (latent non-observable variables), underlying a set of observable indicators (Jarvis, MacKenzie, & Podsakoff, 2003, 2012). The model presented shows the OCB conception as a latent construct with four factors (altruism, conscientiousness, civic participation and courtesy), resulting from different forms of expression of the same construct. In this manner, it can be assumed that we stand before a set of first level latent factors (dimensions), with observable indicators (variables / items) which, in turn, are themselves the observable indicators for the underlying second level construct (OCB). Although the construct may also be viewed as an aggregate model (Law et al., 2005), the conceptualization of the OCB as a latent model is the most used, establishing itself as a good way to conceptualize and describe the relations between the dimensions of the OCB construct.

Finally, the CCOE-R presents itself as an instrument that allows for the collection of information directly at teacher level, not at their superior's level, regarding their perception of OCB application in the schools where they work. Since it does not collect specific teacher information, it allows a valid index to be obtained that expresses the shared perceptions of the teachers regarding the levels of OCB of schools, making possible studies at an organizational level.

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Appendix 1

Items of the scale comportamentos de cidadania organizacional em escolas - CCOE (Neves, 2010)

Instructions for filling in: Indicate to what extent you disagree or agree with the following statements about your current school.

1. Os professores ajudam voluntariamente os novos professores
2. Os professores oferecem-se para integrar novos grupos de trabalho
3. Os professores chegam pontualmente ao trabalho e às reuniões
4. Quando necessitam de ser substituídos (faltas, atestado...) os professores tomam a iniciativa de se apresentarem aos professores substitutos e de os ajudarem
5. Os professores iniciam prontamente os trabalhos e fazem uma utilização eficaz do tempo de aulas
6. Os professores avisam com antecedência os colegas sobre alterações de horário ou calendário (faltas...)
7. Os professores apresentam sugestões inovadoras para melhorar a qualidade global da escola
8. Os professores organizam atividades sociais para a escola
9. Os professores preparam os materiais a utilizar pelos professores que os vão substituir
10. Os professores ajudam os colegas que estão mais sobrecarregados
11. Os professores participam ativamente nas reuniões
12. Os professores oferecem-se para desempenhar papéis e tarefas não obrigatórias
13. Os professores esforçam-se por adquirir novas competências em matérias que contribuem para o seu trabalho
14. Os professores oferecem aos seus colegas materiais que prepararam para as suas aulas
15. Os professores apoiam a direção nos seus tempos livres
16. Os professores trabalham em colaboração com os colegas (planificação, projetos...)