CRITICAL REVIEW

Screening Tools for Cognitive Impairment in Adults with Substance Use Disorders: A Systematic Review

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(RECEIVED JANUARY 6, 2021; FINAL REVISION MAY 2, 2021; ACCEPTED JULY 6, 2021; FIRST PUBLISHED ONLINE AUGUST 26, 2021)

ABSTRACT

Objectives: Cognitive impairment is common in individuals with substance use disorders (SUDs), yet no evidencebased guidelines exist regarding the most appropriate screening measure for use in this population. This systematic review aimed to (1) describe different cognitive screening measures used in adults with SUDs, (2) identify substance use populations and contexts these tools are utilised in, (3) review diagnostic accuracy of these screening measures versus an accepted objective reference standard, and (4) evaluate methodology of included studies for risk of bias. Methods: Online databases (PsycINFO, MEDLINE, Embase, and CINAHL) were searched for relevant studies according to pre-determined criteria, and risk of bias and applicability was assessed using the Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2). At each review phase, dual screening, extraction, and quality ratings were performed. Results: Fourteen studies met inclusion, identifying 10 unique cognitive screening tools. The Montreal Cognitive Assessment (MoCA) was the most common, and two novel screening tools (Brief Evaluation of Alcohol-Related Neuropsychological Impairments [BEARNI] and Brief Executive Function Assessment Tool [BEAT]) were specifically developed for use within SUD populations. Twelve studies reported on classification accuracy and relevant psychometric parameters (e.g., sensitivity and specificity). While several tools yielded acceptable to outstanding classification accuracy, there was poor adherence to the Standards for Reporting Diagnostic Accuracy Studies (STARD) across all studies, with high or unclear risk of methodological bias. Conclusions: While some screening tools exhibit promise for use within SUD populations, further evaluation with stronger methodological design and reporting is required. Clinical recommendations and future directions for research are discussed.

Keywords: Cognition, Screening, Adults, Substance-related disorders, Neuropsychology

Substance use disorder (SUD) is a pathological cluster of cognitive, behavioural, and physiological symptoms stemming from compulsive use of a substance despite adverse effects on psychosocial functioning (American Psychological Association, 2013). Between 30% and 80% of individuals presenting for the treatment of SUDs experience some form of cognitive impairment (CI; Copersino et al., 2009), with neurophysiological changes occurring as a result of intoxication, withdrawal, or chronic use. Substance users frequently demonstrate deficits in attention, memory, executive functions, and decision-making (Ramey & Regier, 2020), with nature and chronicity of impairment differing between types of substances, quantity, and frequency of use (Bruijnen et al., 2019; Chen, Strain, Crum, & Mojtabai, 2013).

CI may also result from other comorbid factors. Psychiatric illness is highly prevalent in individuals with SUD (Gould, 2010; Kelly & Daley, 2013) and can increase severity of health issues and complicate recovery (Daley & Moss, 2002; Grella, Hser, Joshi, & Rounds-Bryant, 2001). Other common comorbidities that can impact cognition include physical illness, such as vascular conditions (Volkow, Baler, Compton, & Weiss, 2014), respiratory

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problems (Owen, Sutter & Albertson, 2014), and liver damage (Brodersen et al., 2014). Between 37% and 66% of individuals with an acquired brain injury (ABI, e.g., due to falls, head trauma, overdose, or chronic substance use; Ridley et al., 2018) misuse alcohol, and 10%-44% use illicit substances (Parry-Jones, Vaughan, & Miles Cox, 2006), suggesting potential cumulative cognitive dysfunction. Developmental disorders (e.g., intellectual or learning disability, attentiondeficit hyperactivity disorder) also co-occur with SUDs, with rates varying (between 1% and 45%; Levin, 2007; Lin et al., 2016; Salavert et al., 2018). Many developmental disorders may remain undiagnosed at treatment (Gooden et al., 2020), and it is possible that similar factors that form barriers to diagnosis and support for individuals with developmental disorders (i.e., social and environmental disadvantages; Carroll Chapman & Wu, 2012) are also risk factors for developing SUDs.

Regardless of aetiology, CI in individuals with SUDs can severely impede everyday functioning. The degree of executive dysfunction could determine how effectively individuals with SUDs are able to cope with day-to-day demands (see Verdejo-Garcia, Garcia-Fernandez, & Dom, 2019). Impaired inhibitory control can contribute to behavioural difficulties, including increased impulsivity and poor decisionmaking (Ramey & Regier, 2020). CI also has impacts on the ability to engage successfully with SUD treatment (Bates, Bowden, & Barry, 2002), with individuals less likely to benefit from the therapeutic process or acquire effective strategies, related to diminished insight and motivation (Aharonovich, Amrhein, Bisaga, Nunes, & Hasin, 2008). CI is also associated with poorer treatment compliance (Copersino et al., 2012). However, individuals with CI can benefit from treatment, particularly when treatment is appropriately matched to cognitive level (Bates, Buckman, & Nguyen, 2013). For instance, cognitive-behavioural approaches are geared to enhance higher-order cognitive processes related to self-awareness and decision-making. These cognitive skills may be fundamentally reduced in individuals with CI (Hagen et al., 2016), who may benefit more from interactional programmes with reduced cognitive demands (Secades-Villa, García-Rodríguez, & Fernández-Hermida, 2015) or increased involvement of social supports (Buckman, Bates, & Cisler, 2007; Buckman, Bates, & Morgenstern, 2008).

Given the impact of CI on SUD treatment, accurate and timely detection of CI is essential. The wide range of factors contributing to CI in individuals presenting for SUD treatment, however, results in heterogeneous cognitive profiles that vary in their clinical presentation, making CI difficult to detect (Horner, Harvey, & Denier, 1999). Consequently, individuals with CI presenting for SUD treatment might not receive the most appropriate support (Braatveit, Torsheim, & Hove, 2018).

While neuropsychological assessment is the gold standard in evaluating cognitive functioning (Roebuck-Spencer et al., 2017), this is not always feasible as it is time-consuming (taking up to several hours) and requires extensive training in administration and interpretation (Ridley et al., 2018; Roebuck-Spencer et al., 2017). According to the 2014 American Psychological Association Working Group on Screening and Assessment (APA–WGSA) guidelines, cognitive screening tools are used for early identification of individuals at high risk for a specific disorder, are generally brief, monitor treatment progress or change in symptoms over time, can be administered by clinicians with appropriate training, and are not definitively diagnostic (APA–WGSA, 2014). Cognitive screening, therefore, represents a critical preliminary stage of triaging those in need of more comprehensive neuropsychological assessment (Roebuck-Spencer et al., 2017).

Cognitive screening is routinely practiced in primary and tertiary settings in the assessment of adults with mild CI, dementia, ABI, and psychiatric illness (Roebuck-Spencer et al., 2017); however, it is much less established as a standard clinical process for individuals with SUD. Cognitive screening tools should be chosen carefully, with reference to both clinical utility and psychometric properties for the population in which it will be used (Slater & Young, 2013). The measure should have high sensitivity (to correctly detect those with CI) and good specificity (to accurately rule out those without CI; Bunnage, 2017). If specificity is too low, clients may inaccurately be perceived as needing full assessment, with potential for resource overload as well as unnecessary alarm for the individual. Yet, if sensitivity is too low, clients that would benefit from full assessment may not receive this, leading to a potential mismatch in treatment provision.

There exists limited evidence to guide decision-making on which cognitive screening tool is appropriate for use with individuals presenting for the treatment of SUDs. The most widely used cognitive screening tools, the Mini-Mental State Examination (MMSE; Folstein, Folstein, & McHugh, 1975) and the Montreal Cognitive Assessment (MoCA; Nasreddine et al., 2005), have most consistently been used in older adult populations (e.g., for stroke and dementia screening). It remains unclear how well these measures can detect CI in individuals with SUDs, or whether there are more appropriate screening tools for use in this population. Given the impact of CI on SUD treatment outcomes, there is a clear need for evidence-based review and recommendations around the use of cognitive screening tools in adults engaging in SUD treatment.

Objective

The aims of this systematic review were to: (1) describe types of cognitive screening measures used in adults with SUDs; (2) identify specific substance use populations and settings these tools are utilised in; (3) review diagnostic accuracy of these screening measures in detecting CI versus an accepted objective reference standard; and (4) evaluate methodology of these studies for risk of bias. It was hoped this review would yield clinical recommendations for use of cognitive screening in individuals presenting for the treatment of SUDs in clinical settings, as well as identify areas for future research in this field.

METHOD

Search Strategy

The Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) statement guided the reporting standards of this review (Moher, Liberati, Tetzlaff, Altman, & The PRISMA Group, 2009). Results of systematic review registry searches (e.g., PROSPERO) did not reveal similar systematic reviews as of May 2020. The review protocol was registered with PROSPERO on 7 November 2020 (ID: CRD42020185902; www.crd.york.ac.uk/prospero).

A systematic search of the Medline, Embase, PsycINFO (OVID), and CINAHL (EBSCO) databases was undertaken in June 2019 and updated in May 2020. The search strategy was constructed around the aims of the review. Subject headings, keywords, and keyword phrases were developed for each of the search concepts, substance use disorder, cognition, and screening or assessment. Subject headings and keywords for each concept were combined using the Boolean operator 'OR' before the concepts were bought together using the 'AND' operator. The search strategy was developed by a senior research librarian (AS) and was designed in Medline before being translated to the other databases. Searches were limited to adult populations and studies in English. No date limits were applied. Comments, editorials, letters, news, book, and book chapters as well as conference abstracts were excluded (see Supplementary Material 1 for full search strategies). In addition, supplementary searching was undertaken in Google Scholar using a combination of search terms for the first 200 results. The reference lists of all papers that met the final inclusion criteria were screened for any additional relevant citations not captured in the database searches. Screening for articles prepared for publication but not yet in print was performed through a search of above search terms of the Psychology Preprint Archive (psyarxiv.com); an author known to be developing a measure in this field was also contacted for potential preprints. A comprehensive grey literature search was not conducted. Results were exported to EndNote and duplicates removed. Studies were then exported to Covidence (Veritas Health Innovation; www.covidence. org) for screening.

Inclusion Criteria

The following inclusion criteria were set: (1) original empirical article, published or prepared for publication (i.e., *under review*); (2) in English; (3) reported on a tool designed to screen for cognitive functioning, or the tool was used for this purpose (in isolation or embedded as part of a broader screening process); (4) examined or provided data that supported inferences regarding criterion validity of the screening tool compared to an accepted standard criterion of cognitive functioning (e.g., neuropsychological assessment); (5) participants were adults (aged 18 years and above); and (6) participants were reported to have some form of SUD (inclusive of alcohol-use or other specific disorder types). The standards described by the 2014 APA–WGSA statement guided screening tool eligibility, specifying they must: (1) be brief to administer; (2) collect objective test data with an established protocol for scoring; (3) require minimal, non-specialist training; and (4) not used for the purpose of definitive diagnosis. To maximise eligible studies, screening tools were not excluded if they also collected functional or self-report information, or focused on one cognitive domain. No restriction on type of criterion validity analysis was applied (e.g., correlations, regression, receiver operating curve).

Screening, Extraction, and Risk of Bias

Two authors (JK and KK) independently screened titles and abstracts from the database searches for articles meeting inclusion criteria (Phase 1). JK and KK independently reviewed full-text articles for eligibility using the same review criteria (Phase 2). If there was disagreement between authors on eligibility for an article, a third author (NR) was consulted and decisions were resolved by consensus. Additional studies retrieved from correspondence or Google Scholar searches were included if author consensus (JK, KK, and NR) indicated they met inclusion criteria.

Data extraction from articles meeting criteria was conducted by JK and KK, with a check by author NR. Variables of interest included target population, study design, recruitment method, participant demographics, index test, reference standard, and relevant psychometric data (e.g., sensitivity, specificity, likelihood ratios, and correlation coefficients).

Included studies were also assessed for risk of bias and applicability (JK and KK) using the revised Quality Assessment of Diagnostic Accuracy Studies (QUADAS–2) tool according to scoring guidelines presented in Whiting and colleagues (2011). Data extraction, risk of bias, and applicability ratings followed the same dual, independent examination and consensus procedures as described for article screening. NR was not involved in reviewing her respective paper (Ridley et al., 2018); instead, SB was consulted for consensus decisions.

Risk of bias assessment using QUADAS-2

The QUADAS–2 assessment tool (Whiting et al., 2011) evaluated risk of bias and applicability of diagnostic accuracy studies (www.quadas.org). Guided by the Standards for Reporting Diagnostic Accuracy Studies (STARD) criteria, this scale was developed to allow more transparency in reporting (Cohen et al., 2016; Whiting et al., 2011) and is recommended by the Cochrane Collaboration (www.cochrane. org) and National Institute for Health and Clinical Excellence (NICE; www.nice.org.uk).

The QUADAS-2 rating process has four phases: (1) summarise review question; (2) tailor the tool to the review; (3) construct a flow diagram for the study; and (4) assess risk of bias and applicability concerns. The tool consists of four key domains: (1) patient selection; (2) index test; (3) reference standard; and (4) flow and timing (i.e., flow of participants through the study, timing of administration of index test and reference standard). Each domain is assessed for risk of bias, the first three also for applicability concerns. Signalling questions guided decision-making on high or low risk of bias (Schueler, Schuetz, & Dewey, 2012), rated as 'Yes', 'No', or 'Unclear'. Risk of bias is deemed 'Low' if all relevant signalling questions are rated 'Yes'; however, bias may exist if these questions are rated as 'No' or 'Unclear'. If questions relating to concerns regarding applicability are rated as 'High' or 'Unclear', aspects of the study may not match the review question.

Statistical Interpretation

Sensitivity, the ability of a screening tool to correctly identify individuals with a condition of interest, and specificity, the ability of a screening tool to detect individuals who do not have a condition, are relative to a reference standard. A receiver operating characteristic (ROC) curve can be plotted by using sensitivity against the false-positive rate (1–specificity) for given cut-off points of a screening test. An area under ROC curve (AUC) value summarises overall test diagnostic accuracy (Trevethan, 2017), ranging from 0 to 1, where 0 and 1 represent perfectly inaccurate or accurate tests, respectively. As a guide, AUC values of 0.5 indicate no meaningful discrimination ability, 0.7 to 0.8 *acceptable*, 0.8 to 0.9 *excellent*, and >0.9 *outstanding* (Mandrekar, 2010).

Sensitivity and specificity are used to obtain predictive power and likelihood ratios. Positive predictive power describes the probability that an individual with a positive test result is a true positive; negative predictive power indicates the probability an individual who obtains a negative test result is a true negative. Both are dependent on the prevalence of the target condition in each population sample (Bunnage, 2017; Trevethan, 2017).

A positive likelihood ratio (LR+) refers to the ratio between the probability of correctly classifying an individual as having CI and the probability of incorrectly classifying an individual who does not have CI (i.e., sensitivity/100–specificity). A LR+ greater than 10 indicates CI is highly likely when a positive test result is yielded. A negative likelihood ratio (LR–) is the ratio between the probability of incorrectly classifying an individual as having CI and the probability of correctly classifying an individual who does not have CI (i.e., 100–sensitivity/specificity). A LR– below 0.1 indicates that impairment is highly unlikely when a negative test result is obtained (McGee, 2002).

For studies that examined correlations between cognitive screening tools and reference tests, coefficients of 0.1, 0.3,

and 0.5 indicated small, medium, and large relationships, respectively (Cohen, 1988).

RESULTS

Search Results

Figure 1 presents a flow diagram of search results. Searches of electronic databases of published articles identified 5,135 candidate studies. Removal of 1,672 duplicates resulted in 3,463 studies. An additional study identified from a Google Scholar search and another through correspondence with an author brought the total to 3,465 studies. A search of the Psychology Preprint Archive yielded no additional studies. After screening the titles and abstracts (Phase 1), the number of studies was reduced to 35. Examination of full-text versions of eligible studies identified 14 relevant papers. Full-text articles were excluded for not meeting the following criteria: not appropriate study design (e.g., did not evaluate screening tool against a standard criterion of cognitive functioning); did not meet 2014 APA-WGSA screening criteria; and did not include participants with SUDs. One duplicate was also removed during this phase.

Study Characteristics

Tables 1-3 present characteristics of the 14 included studies and an overview of the screening measures and reference standards. Studies were conducted in the United States of America (n = 3), France (n = 3), the Netherlands (n = 2), Spain (n = 1), Norway (n = 1), Brazil (n = 1), Belgium (n = 1), and Australia (n = 2). Across all 14 studies, the average clinical sample size was 88.33 (SD = 116.75); however, the range between studies varied widely from 20 to 501 participants. The majority of studies (n = 11; 78.57%) made reference to formal diagnostic criteria (e.g., Diagnostic and Statistical Manual of Mental Disorders) to operationalise SUD. Over a quarter (n = 4) included a sample with a single problematic substance (Ewert et al., 2018; Pelletier et al., 2018; Ritz et al., 2015; Wester et al., 2013). All studies included participants with alcohol use disorder, nearly twothirds also included those with cannabis use (n = 9); 64.29%), followed by cocaine use (n = 7; 50.00%), then opioids (n = 6; 42.86%). Only one study failed to report the specific substance of misuse (To et al., 2015). Over two-thirds (n = 8; 57.14%) required a period of abstinence before engaging in cognitive assessment, while three studies did not require any period of abstinence (Berry et al., 2021; Ridley et al., 2018; Rojo-Mota et al., 2017) and one study did not specify abstinence status (Wester et al., 2013). Nearly, half of studies (n = 6; 46.15%) were in an inpatient setting, two (15.38%) outpatient, four (30.77%) residential rehabilitation, and two (14.29%) did not report the setting.

As per inclusion criteria, each study that was included compared an index test (i.e., the cognitive screening measure) to a reference standard (i.e., neuropsychological assessment).



Fig. 1. PRISMA flow diagram.

In over a quarter of studies (28.57%), a particular syndrome of interest (e.g., Korsakoff's syndrome, intellectual disability) influenced test selection for the reference standard. The index tests and reference standards are presented in Tables 2 and 3.

The majority of studies (n = 12; 85.57%) examined only one screening tool. Across all 14 studies, 10 different index measures were used, with the most common being the MoCA (n = 6; 46.15%). Most (n = 8; 80.00%) were developed for use within other populations (e.g., dementia, intellectual disability, brain injury); however, two were developed specifically for use within a SUD population (Brief Evaluation Alcohol-Related Neuropsychological Impairments of [BEARNI] and Brief Executive Function Assessment Tool [BEAT]). Length of administration for cognitive screening measures varied between 10 and 45 minutes. Almost all index measures targeted multiple cognitive domains, including: orientation (n = 5; 35.71%), attention (n = 6; 42.86%), language (n = 6; 42.86%), visual-spatial abilities (n = 8;57.14%), memory (n = 7; 50%), motor ability (n = 3;21.43%), and executive functions (n = 6; 42.86%). The Frontal Assessment Battery (FAB) focused on one cognitive domain (executive functioning).

Across all 14 studies, reference standards varied in cognitive domains and specific neuropsychological tests. While the reference standard largely encompassed a range of cognitive domains, two studies (14.28%) utilised a reference standard focusing on one cognitive domain (i.e., executive functioning, memory; Berry et al., 2021; Wester et al., 2013). Another study (Cunha et al., 2010) restricted the reference standard to attention, working memory, and executive function. Other variations across studies included time between administration of index and reference standard (ranging from the same day to three months later), and operationalisation of 'impairment' for the reference standard.

Screening Measure Utility

Diagnostic accuracy

Table 4 displays psychometric data provided by 12 studies that reported classification accuracy (Berry et al., 2021; Braatveit et al., 2018; Bruijnen et al., 2019; Copersino et al., 2009; Ewert et al., 2018; Fals-Stewart, 1997; Grohman & Fals-Stewart, 2004; Pelletier et al., 2018; Ridley et al., 2018; Ritz et al., 2015; To et al., 2015; Wester et al., 2013). While all

Table 1. Study characteristics of articles included for review

| First author, year, country | Target population (clinical set- ting, if applicable) | Exclusion criteria | Study design | Study setting | Recruitment method | Clinical group (n) | Mean age, years (SD; range) | % Male | Education level, mean years (SD; range) |
|---|---|---|--|--|---|-----------------------|---|--------|---|
| Berry et al., 2021, Australia | Substance use disorder: alco- hol, cannabis, metham- phetamine, heroin, other stimulants, sedatives/hyp- notics/tranquilizers, other opioids | Abstinence: nil Exclusions: unable to provide informed consent, not proficient in English | Cross-sectional | Residential rehabilitation | Consecutive | 501 | 34.7 (11.1; NR) | 62 | 10.7 (2.1; NR) |
| Braatveit et al., 2018, Norway | Substance use disorder (ICD– 10): including alcohol, opioids, cannabis, cocaine, sedatives, hypnotics, other stimulants, hallucinogens, other psychoactive | Abstinence: 6 weeks Exclusions: native language other than Norwegian, WAIS–IV testing within 6 months, under acute influence of substances | Cross-sectional | Inpatient | Convenience | 84 | 33.31 (11.65, 19–64) | 67 | Mean NR Grade 1–6 4.8%, Grade 7–9 45.2%. Grade 10–12 33.3% Higher 0% |
| Bruijnen et al., 2019, The Netherlands | Substance use disorder: including alcohol, cannabis, stimulants, and opiates | Abstinence: nil Exclusions: severe cognitive impairment due to neurological injury, acute psychiatric disorder, severe lack of moti- vation, and insufficient Dutch language | Longitudinal (baseline, 8 weeks) | Outpatient | Convenience, part of larger study (N = 691) | 82 | 44.1 (13.8; NR) | 68 | Mean NR Low educated 14% Average educated 72% High educated 11% |
| Copersino et al., 2009, USA | Substance use disorder (DSM-IV): including alco- hol, opioids, cocaine, can- nabis, benzodiazepine, and amphetamines | Abstinence: 7 days Exclusions: acute intoxication/withdrawal, medical illness or psychiatric condition impeding capacity to consent | Cross-sectional | NR. Recent admis- sion to hospi- tal treatment or residential programme | NR | 60 | 38.3 (13.2; NR) | 52 | 15 (2.4; NR) |
| Cunha, 2010, Brazil | Cocaine dependence (DSM- IV-TR) with/without alco- hol or cannabis use | Abstinence: 2 weeks Exclusions: past/current psychiatric disor- ders, history of neurological disorders, diagnosis of learning disorder, IQ < 70 | Cross-sectional | Inpatient | NR | 30 | 27.17 (7.64; NR) | 100 | 9.93 (2.74; NR) |
| Ewert et al., 2018, France | Alcohol use disorder (DSM–5) | Abstinence: 1 week Exclusions: severe comorbid neurological or psychiatric diseases (dementia, psy- chosis, past history of stroke/coma/ encephalopathy, cardiac disease, HIV infection) | Cross-sectional | Inpatient | Consecutive | 56 | Reported by cognitive impairment (CI) group No CI: 49.3 (7.9; NR) CI: 49.6 (9.4; NR) | 71 | Mean NR No CI: 40% >12 years education CI: 9.7% >12 years edu- cation |
| Fals-Stewart, 1997, USA | Substance abuse or depend- ence (DSM-III-R): includ- ing alcohol, cocaine, heroin, cannabis, other | Abstinence: 2 weeks Exclusions: organic mental disorder, schizo- phrenia, delusional disorder, other psy- chotic disorder; physical limitations interfering with testing; significant birth or developmental problems; and psycho- tropic medications | Cross-sectional | Residential rehabilitation | Consecutive | 51 | Reported by CI group No CI: 27.9 (4.9; NR) CI: 31.4 (5.2; NR) | 75 | Mean NR Impaired group: 11.6 (1.1; NR) Intact group: 12.2 (1.4; NR) |
| Grohman and Fals- Stewart 2004, USA | Substance use disorder (DSM–IV): including alco- hol, cocaine, opiate, canna- bis, other | Abstinence: 2 weeks Exclusions: schizophrenia, delusional disor- der or other psychotic disorder; on meth- adone programme; significant head | Cross-sectional | Residential in- patient sub- stance use programme | Consecutive | 84 | 32.0 (6.2; NR) | 71 | 12.9 (1.7; NR) |

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| First author, year, country | Target population (clinical set- ting, if applicable) | Exclusion criteria | Study design | Study setting | Recruitment method | Clinical group (n) | Mean age, years (SD; range) | % Male | Education level, mean years (SD; range) |
|--|---|--|-----------------|--|--------------------|---|--|------------------|--|
| Pelletier et al., 2018, France | Alcohol dependence (DSM- IV) | trauma; medical disorder interfering with testing; and psychotropic medications Abstinence: 7–10 days post withdrawal Exclusions: inability to understand/speak French; severe neurological or psychiat- ric diseases such as dementia, psychosis, past history of stroke, coma, encepha- lopathy; and recent consumption of cocaine/cannabic/heroin | Cross-sectional | NR Rehab centre located in hospital | Convenience | 90 | 48.9 (9.6; NR) | 74 | Mean NR 12 or less: 79% >12 years: 21% |
| Ridley et al., 2018, Australia | Substance use or dependence (DSM–IV) including alco- hol, opioids, and cannabis | Abstinence: not required Exclusions: acute intoxication/withdrawal, limited English, significant sensory defi- cits | Cross-sectional | Outpatient | Consecutive | CL: 30 CO: 20 | CL: 52.3 (10.43; 32–76) CO: NR | CL: 70 CO: NR | CL: 11.07 (2.33; 7–11) CO: NR |
| Ritz et al., 2015, France | Alcohol dependence (DSM- IV) | Abstinence: undergoing withdrawal treat- ment Exclusions: lifetime history of neurological pathologies, endocrine or infectious dis- eases, mental illness, history of other substance misuse, and psychotropic med- ications | Cross-sectional | Inpatient | NR | 73 | 45.47 (8.85; 26–67) | 73 | 11.18 (1.69; 8–15) |
| Rojo-Mota et al., 2017, Spain | Substance use disorder (DSM–5), including alco- hol, heroin, cocaine, and cannabis | Abstinence: not under effects of substances Exclusions: brain damage, active psychosis, and inability to understand instructions | Cross-sectional | NR | Consecutive | 48 | 39.8 (12.6; 18–67) | 79 | Mean NR Primary: 25% Secondary: 33% Secondary (advanced): 35.4% Uni: 6.3% |
| To, 2015, Belgium | Substance abuse problem | Abstinence: at least two weeks Exclusions: non-native speakers of Dutch, no WAIS–III last 2 years | Cross-sectional | NR; Recruited from mental health ser- vices | Convenience | 90 | 32 (9.80; 18–64) | 83 | NR |
| Wester et al., 2013, The Netherlands | DSM-IV-TR criteria for alco- hol dependence or alcohol- induced persisting amnestic disorder (Korsakoff's Syndrome) | Abstinence: unclear, but undergoing in- patient treatment Exclusions: NR | Cross-sectional | Inpatient | Convenience | Alcohol depend- ence (AL): 26 Korsakoff's Syndrome (KS): 20 | AL: 54.5 (8.1; NR) KS: 57.6 (8.7; NR) | AL: 76 KS: 75 | AL: Mean 7–9 years (range 1–16) KS: Mean 7–8 years (range 6–16) (range 7–16) |

Note: Only clinical sample characteristics are presented here. As Ridley et al. (2018) and Wester et al. (2013) combined clinical and control data when calculating accuracy statistics, both groups are presented in this table. DSM–III–R, Diagnostic and Statistical Manual–Third Edition–Revised; DSM–IV, Diagnostic and Statistical Manual–Fourth Edition; DSM–IV–TR, Diagnostic and Statistical Manual–Fourth Edition–Text-Revised; DSM–5, Diagnostic and Statistical Manual–Fifth Edition; ICD–10, International Statistical Classification of Diseases and Related Health Problems–10th Revision; NR, not reported; SD, standard deviation; WAIS–III, Wechsler Adult Intelligence Scale–Third Edition; WAIS–IV, Wechsler Adult Intelligence Scale–Fourth Edition.

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Table 2. Overview of cognitive screening measures

| Screening measure | Domains evaluated | Total score | Standard impairment criteria | Administration time (minutes) | Status in adults with SUDs |
|---|---|--|--|----------------------------------|---|
| Addenbrooke's Cognitive Examination – Revised (ACE-R) | Five subtests assessing orientation, attention, ver- bal fluency, memory, language, and visuospatial function. The Mini-Mental State Examination (MMSE) is incorporated into this test | /100 Includes MMSE (/30 points) | <88 for increased sensi- tivity <82 for increased speci- ficity | 12–20 | Developed to screen for dementia (Mioshi et al., 2006). Not developed specifically for SUD. |
| Brief Evaluation of Alcohol-Related Neuropsychological Impairments (BEARNI) | Five subtests assessing visuospatial function, executive functions via alternating verbal flu- ency, verbal episodic memory, verbal working memory, and ataxia (through a balance task). | /30 Without ataxia /22 | Total (including ataxia): <i>Mild impairment</i> : ≤ 12 years education: ≤ 19 >12 years education: ≤ 21 <i>Mod. to sev. impairment</i> : ≤ 12 years education: ≤ 16 >12 years education: ≤ 17 Total (excluding ataxia): <i>Mild impairment</i> : ≤ 12 years education: ≤ 15 >12 years education: ≤ 16 <i>Mod. to sev. impairment</i> : ≤ 12 years education: ≤ 16 <i>Mod. to sev. impairment</i> : ≤ 12 years educ: ≤ 11.5 >12 years educ: ≤ 12.5 | 15–20 | Specifically developed for use with alcohol use disorder populations. Normative data for adults aged 20 to 66 years (Ritz et al., 2015). |
| Brief Executive Function Assessment Tool (BEAT) | Contains 20 items evaluating a range of cognitive domains, including self-reported executive func- tioning, visuoconstruction, learning and memory, attention and working memory, motor sequencing, executive functioning (abstraction, fluency, and delay discounting), reading, and naming. | /60 | Intact cognitive function- ing: 31–60 <i>Mild impairment:</i> 17–30 <i>Sev. impairment:</i> ≤16 | 20 | Specifically developed to assess cognitive impair- ment in substance use disorder populations (Berry et al., 2021). |
| Frontal Assessment Battery (FAB) | Six subtests evaluating executive functioning: conceptualisation, mental flexibility, motor pro- gramming, and resistance to interference, inhibi- tory control, and environmental autonomy. | /18 | Not specified. Higher score indicates better performance. | 10 | Developed to assess the presence and severity of a dysexecutive syndrome (Dubois et al., 2000). Not developed specifically for SUD. |
| Hayes Ability Screening Index (HASI) | Three subtests evaluating spelling, visuospatial, and visuoconstructional ability, as well as four questions on self-reported learning difficulties. | /96.4 | ≤85 | 5–10 | Developed to screen for intellectual disability within forensic justice setting (Hayes, 2000). Not developed specifically for individuals with substance use disorder. |
| Loewenstein Occupational Therapy Cognitive Assessment (LOTCA) | Contains 20 items categorised into 4 subscales: orientation, perception, visuomotor organisation, and thinking operations. | /80 | Not specified. Higher score indicates less cog- nitive impairment. | 30-45 | Developed to assess basic cognitive performance in brain injured patients (Katz et al., 1989). Not developed specifically for individuals with SUD. |

| Screening measure | Domains evaluated | Total score | Standard impairment criteria | Administration time (minutes) | Status in adults with SUDs |
|--|---|---|---|----------------------------------|--|
| Mini-Mental State Examination (MMSE) | Assesses orientation, registration of three words, attention, recall of three words, language, and visuoconstruction. | /30 | Score of <21 suggested by authors | 5–10 | Developed to screen for cognitive impairment in dementia syndromes, delirium, schizophrenia, or affective disorders (Folstein et al., 1975). Not developed specifically for individuals with SUD. |
| Montreal Cognitive Assessment (MoCA) | Assesses executive functioning, visuospatial abil- ities, attention, concentration and working memory, language, abstract reasoning, memory, and orientation. | /30 | <26 (including +1 point of <13 years education) | 10 | Developed to screen for mild cognitive impairment and Alzheimer's disease (Nasreddine et al., 2005). Not developed specifically for individ- uals with SUD. |
| Neuropsychological Assessment Battery–Screening Module (NAB– SM) | Assesses cognitive functioning across five domains: attention, language, memory, visuo- spatial, and executive. | Raw scores converted to standard score (M: 100, SD: 15) | <85 (standard score) | 35-45 | Developed to screen for cognitive impairment across a range of cognitive domains (Stern & White, 2003). Not developed specifically for individuals with SUD. |
| Neurobehavioural Cognitive Status Examination (NCSE/Cognistat) | Assesses orientation, attention, language (compre- hension, repetition, naming), construction, memory, calculations, and executive functions (reasoning, judgement). | 10 independent domain scores | Impairment defined by cut-off scores in each domain | 10–30 | Designed to assess cognition briefly within five major cognitive ability areas (Kiernan et al., 1987). Not developed specifically for individ- uals with SUD. |

M, mean; SD, standard deviation; SUD, substance use disorder.

Table 3. Overview of screening measures and reference standards for included studies

https://doi.org/10.1017/S135561772100103X Published online by Cambridge University Press

| First author (year) | Target population | Time between index and reference | Index test(s) | Index impairment criteria | Reference | Reference impairment criteria |
|---------------------------------------|--|---|---|--|---|---|
| Berry et al. (2021) | Substance use disorder | Same day | BEAT | Intact cognitive function- ing: 31–60 Mild impairment: 17–30 Sev. Impairment: ≤16 | Executive Functions: Alpha Span Task (Total Score), FPT (Unique Designs), Stroop (Interference Trial) | Mean <i>z</i> -score across three tasks: ≤ -1.5 SD below the mean was classified as impaired |
| Braatveit et al. (2018) | Intellectual disability in those with SUD | Same day (screen prior to NPA) | HASI, Norwegian version | N/A: range of cut-scores | Global Intellectual Ability: WAIS–IV | WAIS–IV FSIQ < 70; Vineland II score <70; Self- reported learning difficulties |
| Bruijnen et al. (2019) | Substance use disorder | Baseline MoCA at intake Repeat MoCA and NPA administered same day 6–8 weeks later (screen prior to NPA) | MoCA, Dutch version 7.1 (baseline), Dutch version 7.2 (follow-up) | N/A: range of scores | Orientation: CST-14; Processing Speed: SCWT word reading, colour naming; D-KEFS TMT motor speed; Attention: DSF and DSB; Visuospatial Abilities: WAIS-IV-NL Block Design, ROCF copy; Memory: RAVLT total correct and delayed recall, ROCF Immediate Recall); Abstract Reasoning: WAIS-IV-NL Similarities, DART; Executive Functions: D-KEFS TMT (letter sequencing, number sequencing, and letter-number switching), SCWT interference | At least 2 of 7 domain <i>z</i> –scores ≤–1.00 |
| Copersino et al. (2009) | Substance dependent | Same day (counterbal- anced) | MoCA, English version | N/A: range of scores | NAB–Screening Module (NAB–SM): Attention: Orientation, Digits Forward, Digits Backward, Numbers and Letters; Language: Picture Naming, Auditory Comprehension Spatial: Visual Discrimination and Design Construction; Memory: Story Learning (immediate and delayed recall), Shape Learning Executive Functions: Mazes, Word Generation | NAB–SM Total Screening Index score of <85 |
| Cunha et al. (2010) | Substance dependent | Same day (unclear order) | FAB, Portuguese version | N/A: range of scores | Attention and Working Memory: WAIS-IV DSF, DSB (Portuguese version) Executive: SCWT (Portuguese version), WCST | N/A |
| Ewert et al. (2018) | Alcohol use disorder | Screen conducted day prior to NPA | MoCA, French version 7.1 | N/A: range of scores | Attention and Working Memory: DSF, DSB, DSS; Visuospatial Skills: ROCF copy; Episodic Memory: CVLT learning, short delay free recall, short delay cued recall, delayed free recall, delayed cued recall, recognition; Executive Functions: TMT B, Stroop Interference-Naming, Fluency (Letter and Animal) | 1+ score within a domain either <1.65 SD or <5th percentile or 2+ for executive domain |
| Fals-Stewart (1997) | Substance abuse or dependence | Same day (counterbal- anced) | NCSE | Classified as impaired if they scored in the impaired range on any of the 10 subscales. | Premorbid IQ: NART Neuropsychological Screening Battery; Psychomotor Speed: SDMT Sequencing Efficiency: TMT A and B; Visual Attention: Numerical attention test; Verbal Learning: WMS Story; Non-Verbal Learning: ROCF: Delayed Recall: WMS Story, ROCF; Visuoconstruction: ROCF; Language: MAE (fluency, auditory sequencing, naming, repetition, speech articulation) | For each subtest score, <1 SD = mildly impaired = 1 score; <5th percentile = severely impaired = 2 points; a cumulative score of 6 points or more indicated overall impairment |
| Grohman and Fals-Stewart (2004) | Substance use disorder | One day apart (counter- balanced) | NAB-SM | Index score of 84 or less | Premorbid IQ: NART Neuropsychological Screening Battery; Motor Speed: SDMT Sequencing Efficiency: TMT A and B; Visual Attention: Numerical attention Test; Verbal Learning: WMS Learning Scores; Non- Verbal Learning: ROCF learning; Delayed Recall: WMS Memory, ROCF memory; Visuoconstructional Skills: ROCF copy; Language: MAE (fluency, auditory sequencing, confrontation nam- ing, repetition) | As above |
| Pelletier et al. (2018) | Alcohol use disorder | One day apart (random order) | MoCA, French version 7.1 | 25 or less, no educa- tional adjustment | French versions of validated tests Attention and Working Memory: Digit Span Forward, Backward; Visuospatial skills: ROCF copy; Executive Functions: TMT A and B, SCWT, fluency (letter, animals) Verbal Memory: FCSRT, cued recall trial 3, and delayed cued recall trial | 1+ subtest <1.65 SD or 5th percentile within any domain (2+ for executive domain) |

https://doi.org/10.1017/S135561772100103X Published online by Cambridge University Press

| First author (year) | Target population | Time between index and reference | Index test(s) | Index impairment criteria | Reference | Reference impairment criteria |
|----------------------------|--|--|---|---|--|--|
| Ridley et al. (2018) | Alcohol use dis- order | Same day (screens counterbalanced order with NPA) | MoCA (English version 7.1) ACE-R (Form A, Australian version) MMSE (in ACE-R) | N/A: range of cut-off scores | Attention: RBANS Digit Span, Coding; Language: RBANS Naming, Semantic Fluency; Visuospatial: RBANS Line Orientation, Figure Copy; Memory: RBANS List-Learning, Story-Learning, List Recall, List Recognition, Story Recall, Figure Recall; Executive Functions: Letter Fluency, Digits Backwards, TMT B, Hayling and Brixton Tests | <1.5 SD on domain index scores, 2+ subtests impairment in executive domain |
| Ritz et al. (2015) | Alcohol use disorder | NPA completed within 2 days following screen- ing measure | BEARNI MMSE | BEARNI Total: <i>Mild impairment</i> : ≤ 12 years ed: ≤ 19 >12 years ed: ≤ 21 <i>Modsev. impairment</i> : ≤ 12 years ed: ≤ 16 >12 years ed: ≤ 15 >12 years ed: ≤ 15 >12 years ed: ≤ 16 <i>Modsev. impairment</i> : ≤ 12 years ed: ≤ 11.5 > 12 years ed: ≤ 11.5 > 12 years ed: ≤ 12.5 MMSE Total: ≤ 24 | Working Memory: DSF, DSB, LNS; Visuospatial: ROCF; Verbal Memory: CVLT (French); Executive Functions: TMT, RFFT, SCWT, fluency (letter, animal); Ataxia: Walk-a-line Ataxia Battery | NR |
| Rojo-Mota et al. (2017) | Substance use disorder | All within one week (order unclear) | LOTCA | N/A, correlations only | Orientation: BCSE; Processing Speed: WAIS-IV Coding, Symbol Search; Working Memory: Digit Span, Symbol Span; WMS-IV Verbal Memory: Logical Memory; Visual Memory: ROCF; Executive Functions: FDT; Motor Ability: ACLS-5 | N/A |
| To (2015) | Substance abuse problems | Same day | HASI (Dutch version) | N/A: Range of cut-off scores | Global Intellectual Ability: WAIS–III (Dutch version) | <70 FSIQ |
| Wester et al. (2013) | Alcohol-related cognitive impairment, Korsakoff Syndrome | MoCA–D at intake, RBMT–3 (Dutch version) up to 3 months later | MoCA-D | N/A: Range of cut-off scores | RBMT-3 (Dutch version) Verbal memory: Names, Stories; Visual memory: Face and picture rec- ognition; Spatial memory: immediate and delayed recall of a route; Prospective Memory: remembering appointments, personal belong- ings, shopping items; Spatial and temporal orientation; New learn- ing: immediate and delayed recall of a novel complex nuzzle | Raw scores changes to standard scores, converted to GMI score; mean of 100, SD of 15 Grouped into: Sev. impairment: GMI of <70; Mild deficits: 70–84; Unimpaired: GMI of ≥85 |

ACE–R, Addenbrooke's Cognitive Examination–Revised; ACLS–5, Allen Cognitive Level Screen–Fifth Edition; BCSE, Brief Cognitive Status Examination; BEARNI, Brief Evaluation of Alcohol-Related Neuropsychological Impairments; BEAT, Brief Executive Function Assessment Tool; CI, cognitive impairment; CST–14, Cognitive Screening Test–14; CVLT, California Verbal Learning Test; DART, Dutch Adult Reading Test; D–KEFS TMT, Delis–Kaplan Executive Function System Trail Making Test; DSB, Digit Span Backward; DSF, Digit Span Forward; DSS, Digit Span Sequencing; FAB, Frontal Assessment Battery; FCSRT, Free and Cued Selective Reminding Test; FPT, Five-Point Test; FSIQ, Full Scale Intelligence Quotient; GMI, General Memory Index; HASI, Hayes Ability Screening Index; LNS, letter–number sequencing; LOTCA, Loewenstein Occupational Therapy Cognitive Assessment; MAE, Multilingual Aphasia Exam; MMSE, Mini-Mental State Examination; MoCA, Montreal Cognitive Assessment; NAB–SM, Neuropsychological Assessment Battery; RAVLT, Rey Auditory Verbal Learning Test; RBANS, Repeatable Battery for the Assessment of Neuropsychological assessment; NR, not reported; NSB, Neuropsychological Screening Battery; RAVLT, Rey Auditory Verbal Learning Test; RBANS, Repeatable Battery for the Assessment of Neuropsychological Status; RBMT–3, Rivermead Behavioural Memory Test–Third Edition; RFFT, Ruff Figural Fluency Test; ROCF, Rey–Osterrieth Complex Figure; SCWT, Stroop and Coluc Word Test; SD, substance use disorder; TMT, Trail Making Test; TMT A, Trail Making Test-Part A; TMT B, Trail Making Test-Part B; TMT B–A, Trail Making Test Part A minus Part B; Vineland II, Vineland Adaptive Behaviour Scales–Second Edition; WAIS–IV, Wechsler Adult Intelligence Scale–Fourth Edition; WAIS–IV, Wechsler Adult Intelligence Scale–Fourth Edition; WAIS–IV, Wechsler Adult Intelligence Scale–Fourth Edition; WAIS–IV, Wechsler Memory Scale, Fourth Edition; WAIS–IV, Wechsler Adult Intelligence Scale–Fourth Edition; WAIS–IV, Wechsler Memory Scale, WMS–IV, Wechsler Memory Scale–Fourt

of these studies provided sensitivity and specificity data for at least one defined cut-point, two did not provide overall AUC data (Fals-Stewart, 1997; Pelletier et al., 2018). Although several studies included control groups in the overall study, two included control groups in diagnostic accuracy data (Ridley et al., 2018; Wester et al., 2013).

Of the studies that provided overall AUC data, the MoCA ranged from acceptable to outstanding in the detection of CI (range: 0.75–0.96). The single study examining the Addenbrooke's Cognitive Examination-Revised (ACE-R; Ridley et al., 2018) revealed *excellent* discriminative ability (AUC = 0.85), as did the Hayes Ability Screening Index (HASI; AUC = 0.94-0.95) across two studies (Braatveit et al., 2018; To et al., 2015), and the Neuropsychological (NAB-SM; Battery-Screening Module Assessment AUC = 0.88; Fals-Stewart, 1997). The MMSE by comparison had lower accuracy that ranged from no discrimination ability to acceptable (AUC = 0.21-0.79) across two studies (Ridley et al., 2018; Ritz et al., 2015). For index measures specifically developed for SUD populations (BEAT and BEARNI), at least excellent accuracy in detecting CI was reported (Berry et al., 2021; Ritz et al., 2015).

Half of studies that reported diagnostic accuracy data (n = 6) included Positive Predictive Power and Negative Predictive Power; only two studies reported positive and negative likelihood ratios (Berry et al., 2021; Copersino et al., 2009). Berry and colleagues (2021) was the sole study that reported classification accuracy as well as predictive power and likelihood ratios. Using a LR+ of ≥ 10 and LR- of ≤ 0.1 as threshold (McGee, 2002), a score of ≤ 14 on the BEAT (Berry et al., 2021) was needed to rule in likely impairment, while a score of ≥ 31 was needed to likely rule out impairment. Copersino et al. (2009) indicated that a cut-score of ≤ 22 on the MoCA had strong evidence for ruling in CI and a score of ≥ 27 had strong evidence.

Correlational data

More than half of the 14 studies (n = 8) examined correlations between index tests and neuropsychological measures, with most (n = 6) presenting this in addition to classification accuracy data. Two studies (Cunha et al., 2010; Rojo-Mota et al., 2017) provided correlational data only. As outlined in Table 5, various correlations were performed, including: index total summary scores as compared to reference standard total summary score; index domain score as compared to corresponding reference standard domain scores; and total index score versus specific test outcome on the reference standard. One study (Ridley et al., 2018) combined control and clinical groups to produce correlational data, while all others used clinical samples only. The largest correlations reported were between total summary scores derived from a reference standard (i.e., the Repeatable Battery for the Assessment of Neuropsychological Status; RBANS) and total scores on the MoCA (r=0.75) and the ACE-R (r=0.81); Ridley et al., 2018). Similarly, large correlations were observed between

HASI total score and Wechsler Adult Intelligence Scale-Third Edition Full Scale IQ (r = 0.69-0.70; To et al., 2015). For the MoCA, medium to large correlations between MoCA total score and neuropsychological tasks assessing executive function (r = 0.32 - 0.46), learning and memory (r = 0.45 - 0.56), working memory (r = 0.27 - 0.29), and fluency (r = 0.37 - 0.53) were found (Ewert et al., 2018). MoCA domain scores also yielded medium to large correlations to corresponding neuropsychological domains (Bruijnen et al., 2019; Ridley et al., 2018), with the exception of the domains of language and attention/orientation items (Ridley et al., 2018). For the ACE-R, Ridley and colleagues (2018) reported medium to large correlations between reference domains and relevant ACE-R subscales, with the exception of language. For the FAB, Cuhna and colleagues (2010) reported medium to large correlations (r = 0.41 - 0.57) between FAB tasks and performance on specific executive function measures; however, not all correlations were significant (r = 0.03 - 0.27). For the Loewenstein Occupational Therapy Cognitive Assessment (LOTCA), medium to large correlations between index and reference tasks were found for measures of visual-perceptual abilities and psychomotor abilities (Rojo-Mota et al., 2017). However, reference measures assessing executive functioning and memory did not correspond consistently with LOTCA subscales. For the BEAT, there were medium to large correlations between overall score and performance on reference tasks measuring executive functioning (r = 0.43 - 0.57; Berry et al., 2021).

Quality

Ratings of risk of bias, applicability, and quality of study design are presented in Table 6. Almost all studies were judged to have a high or unclear risk of bias in terms of patient selection (QUADAS-2 Domain 1;92.86%) due to convenience or unclear sampling methods (57.14%; Braatveit et al., 2018; Bruijnen et al., 2019; Cunha et al., 2010; Copersino et al., 2009; Pelletier et al., 2018; Ritz et al., 2015, To et al., 2015; Wester et al., 2013); inappropriate or unclear exclusions (57.14%; Cunha et al., 2010; Ewert et al., 2018;Fals-Stewart, 1997; Grohman & Fals-Stewart, 2004; Pelletier et al., 2018; Ritz et al., 2015; Rojo-Mota et al., 2017, Wester et al., 2013); inclusion of a control group in analyses (14.29%; Ridley et al., 2018; Ritz et al., 2015); or a combination of these factors. Nearly half of the included studies (n = 6; 42.86%) had the main aim of validating the index measure within a SUD population (Berry et al., 2021; Braatveit et al., 2018; Copersino et al., 2009; Ritz et al., 2015; Rojo-Mota et al., 2017; To et al., 2015). Half of the studies had high or unclear risk of bias for the index test (QUADAS-2 Domain 2; 50.00%), while a greater number had a high or unclear risk of bias for the reference standard (QUADAS-2 Domain 3; 71.43%). The high risk was largely due to the non-blinding of index or reference test results; only four studies (Ewert et al., 2018; Pelletier et al., 2018; Rojo-Mota et al., 2017; Wester et al., 2013) used different https://doi.org/10.1017/S135561772100103X Published online by Cambridge University Press

| First author (year) | Impairment Rate (refer- ence standard) | Comparison parameters for impairment | Sensitivity | Specificity | Area Under curve | Positive predictive value | Negative predictive value | Positive likelihood ratio | Negative likelihood ratio | |
|---------------------------|---|--|-------------|-------------|------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|-------------|
| Berry et al. (2021) | 7.8% | Optimal cut-off score for BEAT of ≤24 against executive tasks | 87.2% | 71.4% | 0.86 | 20.5% | 98.5% | 3.04 | 0.18 | |
| Braatveit et al. (2018) | NR | HASI at standard cut-off score of ≤85 against diagnosis of intellectual disability (WAIS–IV ≤ 70) | 100% | 65.4% | 0.94 | NR | NR | NR | NR | |
| | | HASI performance at adjusted cut-off score of ≤80.7 against diagnosis of intellectual disability (including WAIS– IV ≤ 70) | 100% | 80.8% | NR | NR | NR | NR | NR | |
| Bruijnen et al. (2019) | 54.9% (sam- ple) | MoCA at baseline (Version 7.1) with cut-off score of 24 against NPA | 55.6% | 62.2% | 0.68 | 64.1% | 53.5% | NR | NR | |
| | | MoCA at follow-up (Version 7.2) with optimal cut-off score of 25 against NPA | 66.7% | 73.0% | 0.745 | 75.0% | 64.3% | NR | NR | |
| Copersino et al. (2009) | NR | Optimal cut-off score for MoCA of 25, against NAB–SM Total Screening Index impairment | 83.3% | 72.9% | 0.86 | NR | NR | 3.08 | 0.23 | |
| Ewert et al. (2018) | NR | MoCA score of \geq 25 uncorrected for education, against over- all impairment on NPA | 68% | 96% | 0.83 | NR | NR | NR | NR | |
| | | MoCA score of \geq 25 uncorrected for education, against one cognitive domain impaired on NPA | 62% | 96% | 0.82 | NR | NR | NR | NR | |
| | | MoCA score of \geq 25 uncorrected for education, against > one cognitive domain impaired on NPA | 73% | 96% | 0.92 | NR | NR | NR | NR | |
| | | MoCA score of \geq 26 uncorrected for education, against over- all impairment on NPA | 84% | 80% | 0.83 | NR | NR | NR | NR | |
| | | MoCA score of ≥ 26 uncorrected for education, against one cognitive domain impaired on NPA | 75% | 80% | 0.82 | NR | NR | NR | NR | |
| | | MoCA score of \geq 26 uncorrected for education, against > one cognitive domain impaired on NPA | 93% | 80% | 0.92 | NR | NR | NR | NR | |
| | | MoCA score of \geq 26 corrected for education, against overall impairment on NPA | 71% | 88% | 0.84 | NR | NR | NR | NR | |
| | | MoCA score of ≥ 26 corrected for education, against one cognitive domain impaired on NPA | 62% | 88% | 0.80 | NR | NR | NR | NR | |
| | | MoCA score of \geq 26 corrected for education, against > one cognitive domain impaired on NPA | 80% | 88% | 0.90 | NR | NR | NR | NR | |
| | | MoCA score of \geq 27 corrected for education, against overall impairment on NPA | 84% | 76% | 0.84 | NR | NR | NR | NR | |
| | | MoCA score of \geq 27 corrected for education, against one cognitive domain impaired on NPA | 75% | 76% | 0.80 | NR | NR | NR | NR | <i>K</i> .1 |
| | | MoCA score of \geq 27 corrected for education, against > one cognitive domain impaired on NPA | 93% | 76% | 0.90 | NR | NR | NR | NR | Y. Ko é |
| | | | | | | | | | (Continued) | ?t al. |

| Table 4. | (Continued) |
|----------|-------------|
|----------|-------------|

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| First author | Impairment Rate (refer- ence | | | | Area Under | Positive predictive | Negative predictive | Positive likelihood | Negative likelihood |
|---------------------------------------|------------------------------------|--|-------------|-------------|---------------|----------------------|----------------------|---------------------|------------------------|
| (year) | standard) | Comparison parameters for impairment | Sensitivity | Specificity | curve | value | value | ratio | ratio |
| Fals-Stewart (1997) | 43% (sam- ple) | NCSE impairment (in at least one domain) against impairment on NSA (summary score ≥ 6) | 36% | 86% | NR | 67% | 64% | NR | NR |
| Grohman and Fals-Stewart (2004) | NR | NAB–SM at established cut-off ≥ 6 compared with the NSB | 81% | 92% | 0.88 | 87% | 88% | NR | NR |
| Pelletier et al. (2018) | NR | MoCA at traditional cut-off <26 (uncorrected for education), against impairment based on NPA | 79% | 65% | NR | NR | NR | NR | NR |
| | | BEARNI total (including ataxia), against impairment based on NPA | 100% | 2% | NR | NR | NR | NR | NR |
| | | BEARNI cognition (excluding ataxia) against impairment based on NPA | 98% | 13% | NR | NR | NR | NR | NR |
| Ridley et al. (2018) | Two values: 67% and | MoCA at traditional cut-off score of <26, against impair- ment based on NPA | 70% | 83% | 0.84 | 67%: 89% 33%: 67% | 67%: 58% 33%: 85% | NR | NR |
| | 33% (general) | MoCA at optimal cut-off score of <27, against impairment based on NPA | 80% | 73% | 0.84 | 67%: 86% 33%: 59% | 67%: 64% 33%: 41% | NR | NR |
| | | ACE–R at traditional cut-off score of <88, against impairment based on NPA | 55% | 90% | 0.85 | 67%: 92% 33%: 73% | 67%: 50% 33%: 80% | NR | NR |
| | | ACE–R at optimal cut-off score of <93, against impairment based on NPA | 90 | 73% | 0.85 | 67%: 87% 33%: 62% | 67%: 78% 33%: 93% | NR | NR |
| | | MMSE at traditional cut-off score of <27, against impair- ment based on NPA | 15% | 100% | 0.79 | 67%:100% 33%:100% | 67%: 37% 33%: 70% | NR | NR |
| | | MMSE at optimal cut-off score of <29, against impairment based on NPA | 60% | 93% | 0.79 | 67%: 94% 33%: 81% | 67%: 53% 33%: 82% | NR | NR |
| Ritz et al. (2015) | Overall: Mild: 86% | BEARNI total optimal score (education corrected), against mild impairment on NPA | 98% | 50% | 0.92 | 93% | 83% | NR | NR |
| | Mod. to Sev.: 53% | BEARNI Total optimal score (education corrected) mod. to sev. impairment on NPA | 95% | 71% | 0.84 | 79% | 92% | NR | NR |
| | Without ataxia: | BEARNI cognitive optimal score (education corrected) against mild impairment on NPA | 93% | 50% | 0.86 | 90% | 60% | NR | NR |
| | Mild: 84% Mod. to | BEARNI cognitive optimal score (education corrected) against mod. to sev. impairment NPA | 82% | 85% | 0.84 | 82% | 85% | NR | NR |
| | Sev.: 45% | MMSE optimal score (education adjusted), against mild impairment on the NPA | 10% | 90% | 0.21 | 86% | 14% | NR | NR |
| | | MMSE optimal score (education adjusted), against mod. to sev. impairment on NPA | 8% | 88% | 0.24 | 43% | 46% | NR | NR |
| To et al. (2015) | NR | HASI at cut-off score of 84.95, against WAIS-III at IQ 70 | 91% | 80% | 0.95 | NR | NR | NR | NR |
| | | HASI at cut-off score of 85.55, against WAIS-III at IQ 70 | 91% | 79% | 0.95 | NR | NR | NR | NR |
| | | HASI at cut-off score of 86.40, against WAIS-III at IQ 70 | 100% | 91% | 0.95 | NR | NR | NR | NR |

(Continued) 769

| | Imnairment | | | | | | | | |
|---|--|---|--|--|---|---|--|--|---|
| First author | Rate (refer- ence | | | | Area Under | Positive predictive | Negative predictive | Positive likelihood | Negative likelihood |
| (year) | standard) | Comparison parameters for impairment | Sensitivity | Specificity | curve | ¹ value | value | ratio | ratio |
| | | HASI at cut-off score of 87.05, against WAIS-III at IQ 70 | 100% | %2000 | 0.95 | NR | NR | NR | NR |
| Wester et al. (2013) | NR | MoCA–D optimal score of ≤ 23 , against severe memory disorder (GMI < 70) | 91% | 88% | 0.96 | NR | NR | NR | NR |
| | | MoCA-D traditional cut-score of ≤ 25 , against severe memory disorder (GMI < 70) | 78% | 92% | 0.96 | NR | NR | NR | NR |
| | | MoCA-D optimal score of ≤ 24 , against mild memory disorder (GMI 70–84) | 88% | 71% | 0.82 | NR | NR | NR | NR |
| | | MoCA–D traditional cut-score of ≤ 25 , against mild memory disorder | 78% | 76% | 0.82 | NR | NR | NR | NR |
| | | MoCA-D optimal score of ≤ 20 , to classify mild versus severe memory disorder | 81% | %69 | 0.75 | NR | NR | NR | NR |
| ACE–R, Addenb Battery; GMI, G MoCA, Montrea Status Examinati | rooke's Cognitive eneral Memory Inc I Cognitive Assess on; NPA, europsy | Examination–Revised; BEARNI, Brief Evaluation of Alcohol-Related Neun lex; HASI, Hayes Ability Screening Index; IQ, intellectual quotient; LOT sment; MoCA–D, Montreal Cognitive Assessment (Dutch version); NAE chological assessment; NR, not reported; NSB, Neuropsychological Scre | ropsychologic ICA, Loewen B-SM, Neurc eening Batter | al Impairments stein Occupatic psychological y; WAIS-III, V | s; BEAT, Bri onal Therapy Assessment Wechsler Ac | ef Executive Fund Cognitive Asses Battery–Screenir Iult Intelligence S | tion Assessment ' sment; MMSE, M g Module; NCSE cale-Third Editic | Fool; FAB, Fron fini-Mental Statt G. Neurobehavio m; WAIS-IV, V | al Assessment Examination; ural Cognitive /echsler Adult |

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individuals to administer the index and reference measures, while in another seven the order of administration or prior knowledge of test results by the tester was unclear (Berry et al., 2021; Braatveit et al., 2018; Bruijnen et al., 2019; Cunha et al., 2010; Fals-Stewart, 1997; Grohman & Fals-Stewart, 2004; Ritz et al., 2015). For participant flow and timing (QUADAS-2 Domain 4), nearly half of included studies had either high or unclear risk of bias (42.86%) due to inappropriate or unclear time interval between the administration of the index test and the reference standard (n = 2), participants not all receiving the reference standard (n = 3), or not accounting for missing participant data, either clinically or statistically. Overall, quality rating results based on QUADAS-2 indicated generally poor adherence to STARD criteria across the included studies. Two studies (Berry et al., 2021; Bruijnen et al., 2019) were not rated as high risk across any domains; however, unclear ratings meant that risk was unable to be fully assessed.

By contrast, the majority of studies met applicability criteria for the current review. The exception was in relation to the reference standard (QUADAS-2 Domain 3) across two studies (Braatveit et al., 2018; To et al., 2015). These studies sought to evaluate presence of intellectual disability within SUD populations, and consequently, it is possible the reference standards could fail to detect CI specifically related to SUD, or CI more broadly in this population.

DISCUSSION

Intelligence Scale-Fourth Edition.

This review had four primary aims: (1) describe types of cognitive screening measures used in adults with SUDs; (2) identify specific SUD population and settings in which these tools are utilised; (3) review diagnostic accuracy of these screening measures in relation to detecting CI versus an accepted objective reference standard; and (4) evaluate methodology of these studies to determine risk of bias.

Aims 1 and 2: Overview of Cognitive Screening Measures in Adults with SUDs

The current review, to the authors' knowledge, is the first systematic overview of cognitive screening tools utilised in adult SUD populations as compared to a reference standard (e.g., neuropsychological assessment). Overall, 14 studies were identified which broadly aligned with definitions of cognitive screening as proposed by Roebuck-Spencer and colleagues (2017). In summary, more than three-quarters of studies made reference to specific SUD diagnostic criteria. The majority included individuals with poly-substance use, while the remaining specifically examined alcohol use. Alcohol was the most common substance used by participants, with a period of abstinence also frequently endorsed, and nearly half of studies were conducted in an inpatient setting. However, a number of studies failed to report key study details including study setting, specific substance of misuse, and abstinence status. Sample sizes were

| First author (year) | Comparisons | Summary of main findings* |
|-------------------------------|---|---|
| Berry et al. (2021) | Correlations between BEAT and executive tasks (reference battery) | BEAT total score significantly associated with Stroop Interference Trial ($r = 0.43$), FPT Unique Designs ($r = 0.47$), and Alpha Score ($r = 0.52$). |
| Braatveit et al. (2018) | Pearson correlations of HASI total score as compared to WAIS-IV FSIQ | HASI total score was significantly associated with the WAIS–IV FSIQ ($r = 0.70$). Data were used in support of convergent validity. |
| Bruijnen et al. (2019) | Pearson correlations of MoCA domain scores compared to domains scores on neuropsychological assessment | At <i>baseline</i> , correlations between MoCA and NPA domains were significant for executive functioning ($r = 0.24$), abstract reasoning ($r = 0.30$), and memory ($r = 0.42$). Correlations not reported for other domains. At <i>follow-up</i> , all MoCA domain scores were significantly correlated with the corresponding NPA domain: executive functioning: ($r = 0.33$); visuospatial abilities ($r = 0.24$); attention ($r = 0.40$); abstract reasoning ($r = 0.54$); memory ($r = 0.46$); and orientation ($r = 0.23$). |
| Cunha et al. (2010) | Spearman correlations of FAB total and domain scores with executive function tasks on NPA | Significant negative correlation between FAB score total and time to complete interference task on SCWT ($r_r = -0.41$). Significant negative correlation between FAB mental flexibility subtest and number of perseverative errors on WCST ($r_s = 0.44$). Significant positive correlation between FAB mental flexibility and performance on digits forward ($r_s = 0.57$). Significant positive correlation between performance on FAB motor programming subtest and failures to maintain set on WCST ($r_s = 0.41$). |
| Ewert et al. (2018) | Zero-order Pearson's correlations of total MoCA score with NPA tasks | Significant negative correlation between MoCA score and completion time ($r = -0.46$) and errors on the TMT ($r = -0.32$), as well as completion time on the Word-Colour part of the Stroop task ($r = -0.37$) and related interference measure ($r = -0.32$). MoCA score correlated positively with CVLT immediate ($r = 0.55$), long delay free ($r = 0.56$), long delay cued ($r = 0.45$), DSF ($r = 0.27$), and DSS ($r = 0.29$), ROCF (F(3,53) = 2.76, $p = 0.05$; multiple R = 0.02) and Letter ($r = 0.53$) and Animal Fluency $r = 0.37$). |
| Ridley et al. (2018) | Correlations (Pearson or Spearman) of MoCA and ACE–R domains with NPA total scores and domains across whole sample (clinical and control groups combined) | RBANS total score strongly correlated with MoCA ($r = 0.75$) and the ACE–R ($r = 0.81$), moderately correlated with the MMSE ($r = 0.38$). MoCA memory domain was signifi- cantly associated with NPA domains of delayed memory ($r = 0.61$), executive (r = 0.51), language ($r = 0.45$), and immediate memory ($r = 0.37$). MoCA visuospatial domain was significantly associated NPA domains of visuospatial ($r = 0.59$), executive ($r = 0.46$), and immediate memory ($r = 0.37$). MoCA language domain significantly associated with NPA domains of executive ($r = 0.55$) and attention ($r = 0.53$). MoCA executive domain significantly associated with NPA of delayed memory ($r = 0.49$), executive ($r = 0.47$), attention ($r = 0.37$), immediate memory ($r = 0.42$), and visuospatial ($r = 0.35$). MoCA attention/orientation was not significantly associated with any neuro- psychological domain. ACE–R memory domain significantly associated with NPA domains of delayed memory ($r = 0.67$), executive ($r = 0.50$), and immediate memory ($r = 0.48$). ACE–R visuospatial domain significantly associated with NPA executive |

Table 5. Summary of studies examining correlations between cognitive screening tools and comprehensive cognitive assessment

(Continued)

| Table 5. | (Continued) |
|----------|-------------|
|----------|-------------|

| First author (year) | Comparisons | Summary of main findings [*] | | |
|-------------------------------|---|--|--|--|
| | | (r = 0.64), visuospatial $(r = 0.59)$, delayed memory $(r = 0.48)$, immediate memory $(r = 0.42)$, and attention $(r = 0.35)$. ACE–R language significantly associated with immediate memory $(r = 0.56)$, executive $(r = 0.41)$, and visuospatial $(r = 0.37)$. ACE–R attention/orientation significantly associated with NPA attention $(r = 0.39)$ and executive $(r = 0.37)$ ACE–R fluency domain significantly associated with NPA language $(r = 0.73)$, delayed memory $(r = 0.68)$, immediate memory $(r = 0.61)$, executive $(r = 0.60)$, visuospatial $(r = 0.44)$, and attention $(r = 0.39)$. | | |
| Rojo-Mota et al. (2017) | Spearman's correlations between the LOTCA scale scores and scale scores obtained with the NPA | LOTCA orientation subscale significantly associated with the BCSE ($r_s = 0.26$), ROCF Copy ($r_s = 0.33$), Digit Span ($r_s = 0.33$), Symbol Span ($r_s = 0.28$), and Logical Memory I ($r_s = 0.27$). LOTCA Visuo-perceptual subscale significantly associated with ACLS–5 ($r_s = 0.34$), BCSE ($r_s = 0.55$), ROCF Copy ($r_s = 0.47$) and Immediate Recall ($r_s = 0.44$), WAIS–IV Coding ($r_s = .51$), Symbol Search ($r_s = .46$, PSI ($r_s = 0.54$), Digit Span r_s = 0.68), WMS–IV Symbol Span ($r_s = 0.50$), Logical Memory II ($r_s = 0.27$), FDT Reading ($r_s = 0.41$), FDT Counting ($r_s = 0.25$), and FDT Choosing ($r_s = 0.35$). LOTCA Psychomotor abilities significantly correlated with ACLS–5 ($r_s = 0.53$), BCSE (r_s = 0.54), ROCF Copy ($r_s = 0.55$) and Immediate Recall ($r_s = 0.56$), WAIS–IV Coding ($r_s = 0.59$), Symbol Search ($r_s = 0.64$), PSI ($r_s = 0.68$), Digit Span ($r_s = 0.59$), WMS–IV Symbol Span ($r_s = 0.52$), Logical Memory II ($r_s = 0.37$), FDT Reading ($r_s = 0.33$), Counting ($r_s = 0.31$), and Switching ($r_s = 0.25$). LOTCA thinking operations signifi- cantly associated with ACLS–5 ($r_s = 0.48$), BCSE ($r_s = 0.51$), ROCF Copy ($r_s = 0.43$) and Immediate Recall ($r_s = 0.41$), WAIS–IV Coding ($r_s = .59$), Symbol Search (r_s = 0.50), PSI ($r_s = 0.60$), Digit Span ($r_s = 0.51$), WMS–IV Symbol Span ($r_s = 0.54$), Logical Memory I ($r_s = 0.30$), and Logical Memory II ($r_s = 0.34$). LOTCA execution time significantly associated with ACLS–5 ($r_s = -0.44$), BCSE ($r_s = -0.58$), ROCF Copy ($r_s = -0.32$) and Immediate Recall ($r_s = -0.31$), WAIS–IV Symbol Search ($r_s = -0.37$), Digit Span ($r_s = -0.41$), WMS–IV Symbol Span ($r_s = -0.31$), Logical Memory I ($r_s = 0.30$), and Logical Memory II ($r_s = -0.31$), LOTCA execution time significantly associated with ACLS–5 ($r_s = -0.44$), BCSE ($r_s = -0.58$), ROCF Copy ($r_s = -0.32$) and Immediate Recall ($r_s = -0.31$), WAIS–IV Symbol Search ($r_s = -0.37$), Digit Span ($r_s = -0.41$), WMS–IV Symbol Span ($r_s = -0.31$), Logical Memor | | |
| To et al. (2015) | Pearson correlations between the HASI total and subscores, and WAIS-III FSIQ and subtest scores | Significant correlation between the total HASI score and WAIS–III FSIQ score ($r = 0.69$). Significant correlations between the correlations between the HASI subtests and the WAIS–III subtests were all significant: background information ($r = 0.58$), spelling ($r = 0.50$), puzzle ($r = 0.46$), and clock drawing ($r = 0.45$). The HASI also correlated significantly with the verbal subscale of the WAIS–III ($r = 0.70$) and the performance subscale of the WAIS–III ($r = 0.63$). | | |

*Included if at least significant at the .05 level or higher.

ACE–R, Addenbrooke's Cognitive Examination–Revised; ACLS–5, Allen Cognitive Level Screen–Fifth Edition; BCSE, Brief Cognitive Status Examination; BEAT, Brief Executive Function Assessment Tool; DSF, Digit Span Forward; DSS, Digit Span Sequencing; FAB, Frontal Assessment Battery; FDT, Five-Digit Test; FPT, Five-Point Test; FSIQ, Full Scale Intelligence Quotient; HASI, Hayes Ability Screening Index; LOTCA, Loewenstein Occupational Therapy Cognitive Assessment; MMSE, Mini-Mental State Examination; MoCA, Montreal Cognitive Assessment; NAB–SM, Neuropsychological Assessment Battery–Screening Module; NPA, neuropsychological assessment; PSI, Processing Speed Index; RBANS, Repeatable Battery for the Assessment of Neuropsychological Status; ROCF, Rey-Osterrieth Complex Figure; TMT, Trail Making Test; SCWT, Stroop and Colour Word Test; WAIS–III, Wechsler Adult Intelligence Scale–Third Edition; WCST, Wisconsin Card Sorting Test; WMS–IV, Wechsler Memory Scale–Fourth Edition.

| | Domain 1: Patient selection | | Domain 2: Index test(s) | | Domain 3: Reference standard | | Domain 4: Flow and timing |
|-------------------------------|-----------------------------|---------------|-------------------------|---------------|---------------------------------|---------------|---------------------------|
| First author, year | Risk of bias | Applicability | Risk of bias | Applicability | Risk of bias | Applicability | Risk of bias |
| Berry et al., 2021 | LR | LR | ? | LR | ? | LR | LR |
| Braatveit et al., 2018 | HR | LR | LR | LR | ? | HR | HR |
| Bruijnen et al., 2019 | ? | LR | LR | LR | ? | LR | ? |
| Copersino et al., 2009 | HR | LR | HR | LR | HR | LR | LR |
| Cunha et al., 2010 | HR | LR | ? | LR | ? | LR | LR |
| Ewert et al., 2018 | HR | LR | LR | LR | LR | LR | LR |
| Fals-Stewart, 1997 | HR | LR | ? | LR | ? | LR | LR |
| Grohman and Fals-Stewart 2004 | HR | LR | ? | LR | ? | LR | LR |
| Pelletier et al., 2018 | HR | LR | LR | LR | LR | LR | ? |
| Ridley et al., 2018 | HR | LR | LR | LR | HR | LR | LR |
| Ritz et al., 2015 | HR | LR | HR | LR | ? | LR | HR |
| Rojo-Mota et al., 2017 | HR | LR | LR | LR | LR | LR | HR |
| To et al., 2015 | HR | LR | HR | LR | HR | HR | LR |
| Wester et al., 2013 | HR | LR | LR | LR | LR | LR | HR |

Table 6. QUADAS-2 ratings for studies included in the review

QUADAS-2, Quality Assessment of Diagnostic Accuracy Studies - Version 2; HR, high risk; LR, low risk; ? = unclear risk.

generally small, with the exception of Berry and colleagues (2021), who commendably captured a large sample.

Across the 14 included studies, 10 different cognitive screening tools were evaluated. The majority of studies examined only one measure, and all were paper and pen tasks. The range of cognitive domains and tasks assessed by the screening tool and length of administration varied widely across studies. Most studies sought to validate existing cognitive screening tools in this new population, primarily the MoCA, with two seeking to validate new measures for specific use in SUD populations (BEAT and BEARNI). In over a quarter of studies, a particular syndrome of interest was evaluated within a SUD sample, with all other studies seeking to identify CI more generally.

Aim 3: Classification Accuracy

Out of the 14 retrieved studies, 12 reported on classification accuracy data. All reported sensitivity and specificity that varied widely (range: 36%–100%, 2%–100%, respectively). A further two used correlational data to assess associations between screening tools and reference measures. Of the 10 studies that provided overall AUC data, the six that related to the MoCA indicated an *acceptable* to *outstanding* ability to detect CI versus respective reference standards. By contrast, the MMSE, another well-known cognitive screening measure, yielded much lower accuracy that ranged from *no discrimination* to *acceptable* across two studies.

Although the HASI exhibited *excellent* discriminant ability (Braatveit et al., 2018; To et al., 2015), this measure was designed to screen specifically for intellectual disability and the reference standards in both instances were restricted to an intellectual disability diagnosis or intellectual cognitive

domains. Therefore, classification accuracy of the HASI for detecting CI more broadly in individuals with SUDs remains unknown.

Berry and colleagues (2021) reported the most extensive psychometric data on the BEAT, a measure specifically developed for use in those with SUDs. The BEAT demonstrated excellent discriminative ability and good sensitivity and specificity at the recommended cut-score of ≤ 24 . However, low prevalence of CI in this group (7.8%) contributed to a low positive predictive power (Abu-Akel, Bousman, Skafidas, & Pantelis, 2018). The authors suggest this was due to using a conservative reference standard (i.e., impairment across three tests); however, it is also possible that this relates to sample characteristics (residential participants only), as well as choice of the reference battery (executive tasks only). Given that the prevalence rate in this study is much lower than that in other treatment groups and the SUD literature more broadly (e.g., Copersino et al., 2009), further evaluation of the psychometric properties of the BEAT in broader SUD populations is necessary.

Another cognitive screening tool specifically designed for SUDs, the BEARNI, exhibited good discriminative abilities in the original study but a follow-up analysis demonstrated that despite good sensitivity, the specificity of the tool was poor (Pelletier et al., 2018). To this end and given superior MoCA data, Pelletier and colleagues (2018) suggest that the MoCA may be a more suitable choice for routine use.

Less than half of studies provided predictive value data for their screening tools. These data take into account the prevalence of CI in the population of interest. Including this information is crucial as in the event of a mismatch between the prevalence of CI in the screening sample and in individuals attending SUD treatment, classification data may misrepresent a tool's screening ability (Slater & Young, 2013). Of the studies that provided predictive data, the rate of CI across samples varied widely (8%–86%), which is unsurprising given the broad range of study settings and samples. Clinicians must be mindful of the expected prevalence of CI in their particular cohort, as screening tools examined in groups with different prevalence rates may differentially capture CI in their cohort (Molinaro, 2015).

Two studies relied on correlational data to examine screening measures in individuals with SUD. Medium to large correlations were reported between index and reference measures on the FAB (Cunha et al., 2010) and LOTCA (Rojo-Mota et al., 2017). However, not all domains corresponded, and no further classification data were provided to support these associations, resulting in only indirect inferences regarding criterion validity. As such, further evaluation of these screening measures is required. There were more consistent medium to large correlations between MoCA scores (total and domain) and reference measures, providing initial evidence for construct validity. The ACE–R, BEAT, and HASI also had promising correlational data between index and reference measures; however, these findings need to be replicated.

Aim 4: Risk of Bias

This review represents the first time that the methodological quality of studies evaluating cognitive screening tools in SUDs have been systematically reviewed. The studies included in this review were assessed overall to have high or unclear risk of bias according to the QUADAS-2, suggesting significant methodological limitations or insufficient reporting. Patient selection represented the most problematic domain with high or unclear risk of bias occurring in 85.71% of studies. This was primarily due to the use of convenience sampling methods and overly stringent exclusion criteria. This included exclusion of individuals who had a history of psychiatric illness, had head injuries, or were on opioid treatment programmes. As individuals with SUDs present clinically with a range of comorbid and co-occurring conditions (see Gooden et al., 2020), excluding participants based on these factors has the potential to significantly limit translation of study findings to applied clinical settings.

An additional source of high or unclear risk of bias related to whether index measures were interpreted without the knowledge of the reference standard (i.e., blind ratings), or vice versa, an issue in almost three-quarters of studies. Administration of both measures by the same individual is problematic, as it can inadvertently lead to bias in administration and scoring of measures. This is a clear area in need of improvement in future studies. There were also issues with flow and timing, with nearly half of studies classified as having high or unclear risk of bias. This was primarily due to inappropriate or unclear time intervals between the administration of the index test and reference standard. In a SUD population, administration of both the index and reference standard at a close time interval is critical given the potential of ongoing cognitive recovery with continued abstinence (Stavro, Pelletier, & Potvin, 2013). Other issues with flow and timing included not all participants receiving the reference standard, or not accounting for missing participant data, creating an unknown amount of potential bias. It is possible that participants who are more cognitively impaired may drop out more readily due to increased complexity of tasks. It is essential that all aspects of participant flow are tracked, including those invited, those that accept, and those that complete the study, in order to establish whether the population sample reflects the population at large.

The selection of the reference standard was also a source of discrepancy across studies. Given the aim of the study was to include *all* studies examining CI in general amongst individuals with SUD, three studies focused on specific sub-populations (intellectual disability, Korsakoff syndrome; Braatveit et al., 2018; Wester et al., 2013; To et al., 2015). For this reason, the reference standard used for classification was specific and focused, raising concerns on the applicability of identifying CI more broadly in individuals with SUDs.

Another concern regarding reference standards was inconsistency in impairment ratings. Some studies operationalised impairment as ranging from -1 to -2 standard deviations below normative data (the equivalent to a difference of 15th percentile versus 2nd percentile). Other studies required impairment on a certain number of tasks, by domain, or an average overall score. What constitutes psychometric impairment is a source of debate in the neuropsychological community (Schoenberg & Rum, 2017), and this inconsistency is reflected in the studies included. Interestingly, Godefroy et al. (2014) analysed various methods of classifying impairment and suggested that a cut-off of the 5th percentile is most appropriate. A consensus on how to characterise impairment is needed if analyses of screening measures are to be accurately compared.

Two studies (Berry et al., 2021; Wester et al., 2013) presented a unique methodological challenge that was not captured via the QUADAS-2. In both studies the reference standard focused on a limited area of cognitive functioning (executive functioning and memory, respectively), while the index measures evaluated cognitive function more broadly. In both instances, despite the apparent differences in screening and reference tasks, screening measures were able to detect individuals impaired on the reference standard at an acceptable level (AUC \geq 75). This suggests that the BEAT and MoCA were able to adequately detect executive functioning and memory impairments, respectively, despite assessing a broader range of skills. It remains unclear how accurately these screening measures detect impairment in other domains. Use of a comprehensive reference standard that includes a range of cognitive domains would allow evaluation of both how well these screens pick up on a range of impairments other than the domain of interest and establish further construct validity.

Overall, findings suggest that clear adherence to STARD criteria was generally poor in studies evaluating the utility of

cognitive screening tools in adults with SUDs (Bossuyt et al., 2015; Bossuyt et al., 2003; Cohen et al., 2016). The current results reflect the methodological limitations of the studies in this area to date; not one study was classified as low risk across all domains. This primarily relates to insufficient reporting of study procedure and data as in many cases domains were rated unclear, resulting in risk being unable to be determined. This is a novel and important finding. Not only does it provide direction for further research, but it highlights caution that should be taken in interpretation of the findings to date, given that studies with risk of bias are likely to over- or under-estimate diagnostic accuracy.

Limitations

The current study intentionally set broad limits on search strategy and inclusion criteria to capture as many studies as possible. A key limitation of applying the QUADAS–2 to evaluate risk of bias in captured studies includes dealing with studies that omit key methodological details versus those which report on these more comprehensively and which may be disproportionately penalised as a result. For example, a number of studies did not report how many individuals were approached versus declined to participate, and the reasons for refusal. This could have potentially introduced bias in the recruitment and exclusion of participants in some studies; however, with omission of this information, it cannot be evaluated.

Another limitation is that a comprehensive grey literature search was not incorporated in the search strategy. This was partially mitigated by searching the Psychology Preprint Archive, as well as through inclusion of a publication that was under review and subsequently published (Berry et al., 2021). A more comprehensive grey literature search would have enriched the evidence base for this review by including findings from a broader range of sources, as well as reduce potential publication bias (Paez, 2017).

Research Recommendations

While in a research context a homogeneous sample may be useful for interpreting study findings, this does not reflect the rich diversity and heterogeneity of individuals presenting for SUD treatment in health settings. To maximise ecological validity and translation of study findings to applied clinical practice, future studies seeking to validate cognitive screening tools in adults with SUDs should seek to adhere to STARD criteria (Bossuyt et al., 2015; Bossuyt et al., 2003), while also maximising ecological validity. This includes keeping exclusion criteria to a minimum, such as instances where reduced capacity to provide fully informed consent is evident, or there are acute medical, psychiatric, or other risks that warrant priority treatment. This ensures the study sample is more representative of individuals presenting for the treatment of SUD and captures the potential wide range of co-occurring conditions that may impact

cognition in a real-world setting. Studies should also clearly report on sample characteristics (including study setting, substances of misuse, abstinence status) to assist the clinician in understanding whether the cognitive screening tool may be useful for a particular client group. Additionally, including comprehensive classification accuracy statistics supports a greater understanding of the utility of the particular cognitive screening measure. Future studies should: use blinding to ensure objectivity of results; not integrate control with clinical data (as this can conflate results); and provide a clear description of study flow; accounting for any missing data. Future systematic reviews should also adhere to the recently updated PRISMA guidelines (Page et al., 2020), which were not available at the time of the current review. Given the wide number of cognitive screening tools available and critiqued in this review, some of which show good promise in use in a drug and alcohol clinical setting (e.g., the MoCA), comprehensively validating existing tools should be an initial first step.

Clinical Recommendations

CI is frequent in individuals with substance use and, given the negative impact on treatment outcomes, should be screened for on treatment entry and periodically over treatment engagement (Bates, Buckman, & Nguyen, 2013). A cognitive screen used in this context must also be able to capture a range of cognitive disorders, given that the heterogeneity of contributing factors (e.g., ABI, learning or intellectual disability) forms part of a typical client's presentation. Of the tools reviewed, the MoCA was the most studied and demonstrated the most consistently adequate diagnostic classification accuracy in this population. Practically, the MoCA is brief to administer, evaluates a broad range of cognitive domains, and is used widely across healthcare settings, which may aid in effective communication of health information. It also has alternate versions available, which facilitate re-testing over short periods of time (e.g., at the beginning and end of a detoxification stay, and then entry into outpatient clinics). While the MMSE is another well-known cognitive screening measure, it failed to demonstrate adequate classification accuracy in SUD populations and therefore cannot be recommended for clinical use in this setting. Other measures (e.g., ACE-R and BEAT) have initial promise for use in this area but require further empirical validation before they can be recommended for wide clinical use.

Regardless of the specific screening tool, a key consideration is the risk of over-reliance on a cut-off score alone to determine presence or absence of CI. A cognitive screen is not diagnostic and can only flag clients who may be in need of further assessment, and understanding the context in which the screen was administered is essential in an applied clinical setting. For example, capturing whether a client who performed below the cut-off score on a particular cognitive screen was acutely intoxicated, distracted, or in an uncharacteristically high state of distress is not an uncommon dilemma, and are all relevant details when deciding on clinical recommendations. Incorporating this information can reduce excessive use of resources but is also heavily reliant on clinical acumen. Given the complexities of clinical presentation and subtlety in interpreting cognitive screening results in adults with SUDs, it is recommended that a clinician skilled in the cognitive field provides oversight of cognitive screening processes within this setting, with clinical neuropsychologists ideally situated to perform such a role. Additionally, the development of clear clinical pathways within a service that provide further assessment and comprehensive support for clients identified as at risk of CI is also essential.

Conclusions

The current review provides the first systematic overview of cognitive screening tools utilised in adult SUD populations as compared to an accepted standard criterion of cognitive functioning (e.g., neuropsychological assessment). A strength of this review was the broad inclusion criteria; that is, that CI from a range of aetiologies and client populations was included. The MoCA was the most frequently reported cognitive screening tool, and it exhibited promising classification accuracy data across studies, suggesting the potential for use within this population. However, based on our search strategy and inclusion criteria, studies examining cognitive screening in adults with SUDs were significantly limited in number and quality. Other tools indicate initial promising classification data (ACE-R and BEAT); however, these also require further validation. Overall, given the very small number of studies, diverse sample characteristics, wide range of index measures and reference standards used, as well as high risk of bias identified amongst studies, further evaluation with stronger and more transparent methodological design and reporting is required. Without a clearly established tool to screen for CI, there is the potential that CI may not be systematically evaluated within SUD services, and this may negatively impact individual treatment outcomes. Crucially, future studies need to focus on adhering to the STARD criteria (Bossuyt et al., 2015; Bossuyt et al., 2003), while also seeking to maximise ecological validity. Positively, there remains promise in the measures utilised to date such as the MoCA, and it is pleasing to see the development of measures specifically designed for use in the SUD population (e.g., Berry et al., 2021; Ritz et al., 2015). Future studies should continue to build on the learnings of existing studies to ensure evidence-based measures are available to direct treatment and care of individuals who have comorbid substance use and cognitive needs.

SUPPLEMENTARY MATERIAL

To view supplementary material for this article, please visit https://doi.org/10.1017/S135561772100103X

ACKNOWLEDGEMENTS

The authors thank Hunter New England and South East Sydney Local Health District Drug and Alcohol Services for their support in producing this review.

AUTHOR CONTRIBUTIONS

JK and NR contributed equally to conception, design, interpretation, and writing of the manuscript. KK was involved in analysis, interpretation, and writing of the manuscript. AS designed and conducted the systematic search strategy. SB and KA contributed to design and revisions of the manuscript. All authors contributed to final manuscript review.

FINANCIAL SUPPORT

KA is supported by a Career Development Fellowship from the National Health and Medical Research Council (NHMRC; APP1141207). KK, NR, SB, AS, and JK have no financial support to disclose.

CONFLICTS OF INTEREST

The authors have nothing to disclose.

ETHICAL STANDARDS

This study did not involve human subject participation.

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