Disentangling the causal inter-relationship between negative life events and depressive symptoms in women: a longitudinal twin study

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Background. Negative life events are strongly associated with the development of depression. However, the etiologic relationship between life events and depression is complex. Evidence suggests that life events can cause depression, and depression increases the risk for life events. Additionally, third factors influencing both phenotypes may be involved. In this work we sought to disentangle these relationships using a genetically informative longitudinal design.

Method. Adult female twins (n = 536, including 281 twin pairs) were followed up for measurements of negative life event exposure and depressive symptoms. Four follow-ups were completed, each approximately 3 months apart. Model fitting was carried out using the Mx program.

Results. The best-fitting model included causal paths from life events to depressive symptoms for genetic and shared environmental risk factors, whereas paths from depressive symptoms to life events were apparent for shared environmental factors. Shared latent influence on both phenotypes was found for individual-specific effects.

Conclusions. Life events and depressive symptoms have complex inter-relationships that differ across sources of variance. The results of the model, if replicated, indicate that reducing life event exposure would reduce depressive symptoms and that lowering depressive symptoms would decrease the occurrence of negative life events.

Received 17 August 2010; Revised 8 December 2011; Accepted 15 December 2011; First published online 25 January 2012

Key words: Depressive symptoms, longitudinal studies, statistical modelling, stressful events, twins.

Introduction

There is strong accumulated evidence that negative life events play a role in the development of major depression (Kendler *et al.* 1999, 2001*a*; Rijsdijk *et al.* 2001; Paykel, 2003; Hammen, 2005). Interpersonal 'loss' factors, such as bereavement or separation, and other loss experiences such as loss of self-esteem, loss of employment or respected status in the community, or loss of cherished ideas and humiliation (Kendler *et al.* 2003*b*), are reported to be potent elicitors of depressive reactions (Brown *et al.* 1995; Farmer & McGuffin, 2003). However, the relationship between

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negative life events and depression is more complex and dynamic than can be accounted for solely by a causal effect of negative life events on depression. Several studies have shown that exposure to negative life events itself is partly under genetic control (Kendler *et al.* 1993*a*; Kendler & Baker, 2007; Vinkhuyzen *et al.* 2010).

Hammen (1991) introduced the concept of stress generation to describe the finding that people with a history of depression were more likely to expose themselves to negative life events than people without a history of depression (Kendler & Karkowski-Shuman, 1997). The fact that depressed individuals contribute to their experience of stress suggests a causal path from depression to life event exposure. A recent prospective study (n = 826) examined the temporal associations between initial chronic stress, neuroticism and follow-up depression severity in a

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sample of out-patients (Brown & Rosellini, 2011). Evidence was found for both causation (association between initial chronic stress and follow-up depression severity) and the stress generation hypothesis (association between initial depressive severity and follow-up levels of chronic stress). Two other studies, using latent modeling with cross-lagged paths, also found reciprocal associations between major life events and depressive symptoms, supporting the hypotheses of both the causation model and the stress generation model. However, in the latter study reciprocal associations were found only in girls (Ge et al. 1994; Cole et al. 2006). Furthermore, another recent study reported only modest support for 1-year lagged paths from depressive symptoms to major events, and little evidence for paths of the opposite direction (Pettit et al. 2011). Other studies have tried to examine directionality by separating the effects of 'dependent' from 'independent' life events. Dependent life events refer to events that the individual him/herself could have contributed to, such as interpersonal conflicts; independent events refer to those that were not under the control of the individual, such as death of a spouse or a child. Kercher et al. (2009) showed, using path analysis, that not only did depressive symptoms predict later dependent life events but also dependent life events mediated the effects of neuroticism on later depressive symptoms. Other studies (Kendler et al. 1999; Silberg et al. 2001) showed that independent life events predicted future onset of major depression. Thus, previous studies have tried to disentangle the direction of effects and found evidence for bidirectional associations.

However, instead of reciprocal causation, shared causal influence on both life events and depression may also explain the reported phenotypic correlation between negative life events and depression. Potential third factors that are related to both phenotypes are, for example, neuroticism (Kendler et al. 1993b, 2003a; Van Os & Jones, 1999) or socio-economic status (Brady & Matthews, 2002; Wang et al. 2010). Neuroticism has been shown phenotypically and genetically to be related to risk for major depression and depressive symptoms (Kendler et al. 1993b; Van Os & Jones, 1999). In addition, neuroticism is associated with negative life event exposure (Kendler et al. 2003a). As life events are subject to a degree of genetic control, shared genes that influence both life events and depression could be involved. The presence of shared causal influence can be examined using genetically sensitive designs, such as twins studies (see online supplementary material for an explanation of basic behavioral genetic principles and the co-twin control method). The fact that within monozygotic (MZ) twins pairs, matched for both genotype and family

environment, life event exposure increased the risk for onset of major depression suggests causal influence of life events on major depression. As this effect was smaller within MZ pairs than dizygotic (DZ) pairs or within the entire population, it was concluded that part of the association (about 1/3) was non-causal and explained by genetic factors that influence both depression liability and exposure to life events. However, in another study that also used the co-twin control method (Middeldorp et al. 2008), it was found that the genes that influenced anxious depression did not overlap with genes influencing life event exposure. One study (Thapar et al. 1998) used structural equation modeling to carry out bivariate genetic analyses of twin data on life events and depressive symptoms. The model supported the presence of shared genetic influences on both phenotypes. However, in this study, the bivariate model was not compared to other models of causal paths from life events to depressive symptoms or the other way around, which may or may not have fitted the data better.

Thus, studies have examined both directionality of effect and potential shared genetic influence, some (Kendler et al. 1999; Silberg et al. 2001; Middeldorp et al. 2008) within the same sample; however, there is disagreement on the role of shared genetic influence. The extent to which shared genes, instead of causal influence of depressive symptoms on new life events, explain the phenotypic association has clinical relevance because stress generation can be hypothesized as a contributing mechanism to the recurrence of depressive episodes (Hammen, 2005). Furthermore, multiple and contrasting pathways of causation may be operating simultaneously depending on the nature of the risk factor. However, no study has yet examined how several sources of variance (additive genetic effects, shared and individual-specific environmental effects) differentially impact on the association between life events and depression.

Therefore, the current longitudinal twin study examined the dynamic within and cross-time associations between negative life events and depressive symptoms using structural equation modeling. To test for directionality of effects and shared sources of influence several models were compared: (i) models with causal paths from life events to depressive symptoms, (ii) models with causal paths from depressive symptoms to life events, and (iii) models without causal paths but with factors for depressive symptoms and life events that were allowed to correlate (which would suggest shared influences). To our knowledge, this is the first longitudinal study using structural equation modeling to disentangle the nature of the dynamic associations between negative life events and depressive symptoms over time.

Sample

Subjects (n=621) were taking part in an ongoing, longitudinal, general population twin study on geneenvironment interaction in affective disorders, which has been described in detail elsewhere, and showed a very high degree of compliance with research procedures (Jacobs et al. 2005). The sample consisted of twins (n=575) and siblings of twins (n=46). Following exclusion of those individuals who were non-twin sisters and those with missing data on zygosity, 536 subjects (who were part of 281 twin pairs) had valid life events measurements at the first followup of the study. Given evidence for qualitative differences in the type of environmental stressors that are associated with depression in men and women (Kendler et al. 2001b, 2006) and potential gender differences in the temporal associations between stressors and depressive symptoms (Ge et al. 1994), a female-only sample was chosen to improve homogeneity. The study was approved by the standing ethics committee and subjects provided written informed consent. Zygosity was determined through sequential analysis based on sex, fetal membranes, blood groups and DNA fingerprints (Derom et al. 2006). In 81 pairs, determination of zygosity was based on self and mother's report of standard questions about physical similarity and the degree to which the twins are confused (Spitz et al. 1996; Peeters et al. 1998; Christiansen et al. 2003) and, if necessary, on examination of DNA fingerprints.

Study design

Subjects were assessed five times at approximately 3to 4-monthly intervals. The average number of days between T0 and T1 was 132, between T1 and T2 n =91, between T2 and T3 n =116 and between T3 and T4 n =91. At T0 assessments were performed at the home of the individuals. For the collection of follow-up data, questionnaires were sent to the participants.

Measurements

An inventory of recent life events was made based on the event list of the Interview for Recent Life Events (Paykel, 1997). Participants reported on the occurrence of 61 events in the past 6 months (at baseline) and since the last measurement occasion (at follow-up) and rated their impact on a five-point scale (from 1 = very pleasant to 5 = very unpleasant). These recent life events were in the domain of 10 categories: work; education; finance; health; bereavement; migration; courtship, marriage and cohabitation; legal, family and social relationships, all representing dateable occurrences involving changes in the external social environment. Events rated as unpleasant (i.e. a score of 4=unpleasant or 5=very unpleasant) were included in the analysis, and a variable was constructed representing the number of such unpleasant events that had occurred since the last measurement occasion. In the analyses, a negative life event (LE) score was used and coded as follows: 0 LE=0, 1 LE=1, 2 LE=2, 3 LE=3, 4 LE=4, $\geq 5 \text{ LE}=5$, resulting in six categories of life event exposure. As the first (baseline) measurement of negative life events represented oc-

measurement of negative life events represented occurrences in the past 6 months whereas the follow-up measurements all represented occurrences since the past measurement occasion (approximately 3-month intervals), only the four follow-up measurements were used for the analyses to ensure that all four measurements used the same phenotype.

For the measurements of depressive symptoms, a validated self-report measure was used. At baseline and at each of the four follow-ups, subjects filled in the 90-item Symptom Checklist (SCL-90; Derogatis et al. 1973). The dimension of depressive symptomatology consists of 16 items such as 'feeling low in energy or slowed down', 'feeling no interest in things' or 'experiencing feelings of worthlessness'. Subjects were instructed to rate the degree of discomfort associated with each depressive symptom during the past week on a five-point scale ranging from 'not at all' to 'extremely'. A continuous weighted depression score (sum of scores of the depression items divided by number of items filled in) was calculated at each measurement occasion. Table 1 shows the number of subjects at each time point, the time interval between time points and the average life event and SCL-90 scores at each time point. The mean age of the sample at T1 was 28 years (range=18-46 years). For information on attrition, see supplementary online material. As the measurements of life events and depressive symptoms were analyzed within one model, only the four follow-up measurements of depressive symptoms were used and data were transformed into six categories of symptoms, each with equal numbers of observations (see online material for a further explanation of this choice).

Analyses

Model fitting was performed using the Mx program (Neale *et al.* 2003). Several different plausible models that may explain the observed phenotypic association between depressive symptoms and life events were modeled and compared. The models were chosen to reflect the different possibilities of how negative life events and depressive symptoms may be associated

T1 536 (MZ: 329) 1.34 (1.46; 0-17) 476 (MZ: 2 99) 1.50 (0.57; 1-4.77) 0.91 - T2 535 (MZ: 327) 1.24 (1.43; 0-9) 473 (MZ: 293) 1.48 (0.55; 1-4.31) 0.90 91 T3 532 (MZ: 324) 1.00 (1.40; 0-21) ^c 448 (MZ: 281) 1.47 (0.56; 1-4.78) 0.91 116 T4 535 (MZ: 328) 0.88 (1.37; 0-13) ^c 437 (MZ: 272) 1.46 (0.60; 1-4.46) 0.92 91		mean (s.D.; range) ⁻	Cronbach's α	Time interval (days)
T2 535 (MZ: 327) 1.24 (1.43; 0-9) 473 (MZ: 293) 1.48 (0.55; 1-4.31) 0.90 91 T3 532 (MZ: 324) 1.00 (1.40; 0-21) ^c 448 (MZ: 281) 1.47 (0.56; 1-4.78) 0.91 116 T4 535 (MZ: 328) 0.88 (1.37; 0-13) ^c 437 (MZ: 272) 1.46 (0.60; 1-4.46) 0.92 91	536 (MZ: 329) 1.34 (1.46; 0–17) 476 (MZ: 2 99)	1.50 (0.57; 1–4.77)	0.91	I
T3 532 (MZ: 324) $1.00(1.40; 0-21)^{\circ}$ 448 (MZ: 281) $1.47 (0.56; 1-4.78)$ 0.91 116 T4 535 (MZ: 328) 0.88 (1.37; 0-13)^{\circ} 437 (MZ: 272) $1.46 (0.60; 1-4.46)$ 0.92 91	535 (MZ: 327) 1.24 (1.43; 0–9) 473 (MZ: 293)	1.48 (0.55; 1–4.31)	0.90	91
T4 535 (MZ: 328) 0.88 (1.37; 0–13) ^c 437 (MZ: 272) 1.46 (0.60; 1–4.46) 0.92 91	532 (MZ: 324) 1.00 (1.40; 0–21) ^c 448 (MZ: 281)	1.47 (0.56; 1-4.78)	0.91	116
	535 (MZ: 328) 0.88 (1.37; 0–13) ^c 437 (MZ: 272)	$1.46\ (0.60\ ;\ 1-4.46)$	0.92	91

Significant differences (p < 0.05) between life event score at T1 and life event score at T3 and T4.

with one another (whether they might be causally affecting each other (bi- or unidirectionally) or whether a common latent construct is affecting both phenotypes simultaneously. Similar types of models have been tested in a previous study on the relationship between peer deviance and conduct disorder (Kendler *et al.* 2008).

The models are illustrated in Fig. 1 for the genetic paths, but apply equally to shared and individualspecific environmental paths. The first one assumes one latent factor for each phenotype that influences either exposure to life events or experience of depressive symptoms at all time points. This model is called the 'causal factor model' (Fig. 1a). It also includes causal paths between life events and depressive symptoms. Three different causal factor models were tested: (1) a model with bidirectional cross-time causal paths from life events to depressive symptoms (from LE₁ to DS₂; from LE₂ to DS₃; from LE₃ to DS₄) and vice versa (from DS_1 to LE_2 ; from DS_2 to LE_3 ; from DS_3 to LE_4 ; (2) a model including only unidirectional paths from life events to depressive symptoms; and (3) a model including only unidirectional paths from depressive symptoms to life events. The latter two models are nested within the first bidirectional model.

The second model (Fig. 1*b*), which is called the 'simple causal model', assumes that separate independent latent factors influence life events and depressive symptoms at each of the four measurements occasions instead of having one common factor for each phenotype. Three versions of this model were also tested, one with bidirectional paths between the phenotypes and two with either one of the unidirectional paths.

The third model (Fig. 1*c*) is called the 'correlated factor model'. This model postulates that phenotypic correlations between life events and depression arise from a correlation between the latent factors that influence life events and depression.

Initially, all models include cross-time withinphenotype paths (hereafter 'simplex paths': for example, paths from T1 to T2, T2 to T3 and T3 to T4 within both phenotypes). See online material for further details on model characteristics.

Nested models were compared by evaluating decline in fit using χ^2 (df) tests. For the evaluation of nonnested models the Bayesian Information Criterion (BIC; Schwartz, 1978) was used. The model with the lowest BIC is considered as the model with the best balance between explanatory power and parsimony. Another frequently used criterion is Akaike's Information Criterion (AIC). The optimum fit as considered by the AIC is typically shifted more towards high explanatory power and less to parsimony as compared to the BIC. The BIC is the criterion used in

able 1. Descriptive information of measurements of negative life events and depressive symptoms



Fig. 1. (*a*) The causal factor model; (*b*) the simple causal model; and (*c*) the correlated factor model. * This model was also tested with the lower arrows pointing in the opposite direction (from depressive symptoms towards life events: DS t1 to LE t2, DS t2 to LE t3 and DS t3 to LE t4) and with the lower arrows in both directions simultaneously.

this study because it performs well with such complex models (Markon & Krueger, 2004). For completeness, however, both the AIC and the BIC values are shown in the description of the results.

First, we fitted a fully saturated model with separate means, variances and covariances, as a baseline for model comparisons. Second, we then tested assumptions of the twin modeling such as equality of means and variances by twin order and zygosity. Third, a triple Cholesky decomposition model for genetic factors (A), shared environmental (C) and individualspecific factors (E) was fitted. This is a saturated model of the observed genetic and environmental variances and covariances (see online material for a further explanation of the Cholesky decomposition and an overview of fit indices for all models tested). We simplified genetic factors first, then shared environmental and then individual-specific factors (see Table 2). The same model testing procedure was followed for each variance component. First, for both the causal factor and simple causal models, the bidirectional models were compared to the unidirectional models to see whether fit would significantly deteriorate after removing one of the directional paths. The best of all causal factor models and the best of all simple causal models were retained. Second, the non-nested bestfitting causal factor, simple causal model and correlated factor model were compared using the BIC. Finally, we tested whether simplex paths could be dropped from the best-fitting model resulting from the

(a) Variance component A	Model								
		Description	-2LL	df	-2LL change (df)	Result ^a			
			i- and unidirec	tional ca	usal factor models				
Causal factor: nested model evaluation	Ι	LE⇔DS	10934.698	3874	_	_			
	II	LE→DS	10935.432	3875	0.73 (1)	← can be dropped			
	III	LE←DS	10949.188	3875	14.49 (1)	\rightarrow cannot be dropped			
		Variance component A: bi- and unidirectional simple causal models							
Simple causal: nested model evaluation	IV	LE⇔DS	10936.614	3874	_	_			
•	V	LE→DS	10937.018	3875	0.40 (1)	← can be dropped			
	VI	LE←DS	10937.753	3875	1.14 (1)	\rightarrow can be dropped			
		Variance component A: comparison non-nested best models							
Comparison of non-nested best models using BIC		Description	-2LL	df	AIC	BIC			
	II	CF: LE→DS	10935.432	3875	3185.432	-5463.479			
	V	$SC: LE \rightarrow DS$	10937.018	3875	3187.018	-5462.686			
	VI	SC: LE←DS	10937.753	3875	3187.753	-5462.318			
	VII	Corr. factor	10939.443	3875	3189.443	-5461.473			
		Variance component A: c	omparison sim	plex patl	ns of best-fitting mode	1			
Simplex paths: nested model evaluation of best-fitting model		Description	-2LL	df	-2LL change (df)	Result			
	П		10935.432	3875	_	_			
	IIc	Drop of simplex paths	10935.515	3877	0.08 (2)	Simplex paths can be dropped			
		Description	-2LL	df	AIC	BIC			
Best model for A	IIc	CF: LE→DS Drop of simplex paths	10935.515	3877	3181.515	-5469.080			

(b) Variance component C using best model for A (IIc)	Model						
		Description	-2LL	df	-2LL change (df)	Result ^a	
		Variance component C: bi- and unidirectional causal factor models					
Causal factor: nested model evaluation	VIII	LE⇔DS	10954.459	3901	-	_	
	VIIII	LE→DS	10963.657	3902	9.20 (1)	\leftarrow cannot be dropped	
	Х	LE←DS	10954.477	3902	0.02 (1)	\rightarrow can be dropped	
		Variance component C: h	oi- and unidired	tional si	mple causal models		
Simple causal: nested model evaluation	XI	LE⇔DS	10955.115	3901	-	_	
I	XII	LE→DS	10959.770	3902	4.66 (1)	\leftarrow cannot be dropped	
	XIII	LE←DS	10962.909	3902	7.79 (1)	\rightarrow cannot be dropped	
		Variance component C: c	comparison non	-nested	best models		
Comparison of non-nested best models using BIC criterion		Description	-2LL	df	AIC	BIC	
	x	CF: LE←DS	10954.477	3902	3150.477	-5530.122	
	XI	$SC: LE \leftrightarrow DS$	10955.115	3901	3153.115	-5526.982	
	XIV	Corr. factor	10964.009	3902	3160.009	-5525.356	
		Variance component C: c	comparison sim	plex pat	hs of best-fitting mode	1	
Simplex paths: nested model evaluation of best-fitting model		Description	-2LL	df	-2LL change (df)	Result ^a	
	х	Best model CF: LE←DS	10954.477	3902	_	_	
	Xc	Drop of simplex paths	10957.498	3904	3.02 (2)	Simplex paths can be dropped	
		Description	-2LL	df	AIC	BIC	
Best model for C	Xc	CF: LE←DS Drop of simplex paths	10957.498	3904	3149.498	-5534.253	
(c) Variance component E using best model for A and C (Xc)	Model						
		Description	-2LL	df	-2LL change (df)	Result ^a	
		Variance component E: b	oi- and unidirec	tional ca	usal factor model		
Causal factor: nested model evaluation	XV	$LE \leftrightarrow DS$	11040.907	3920	-	_	
	XVI	$LE \rightarrow DS$	11053.902	3921	12.995 (1)	\leftarrow cannot be dropped	
	XVII	LE←DS	11043.102	3921	2.195 (1)	\rightarrow can be dropped	

(c) Variance component E using best model for A and C (Xc)	Model					
		Variance component E: bi- and unidirectional simple causal models				
Simple causal: nested model evaluation	XVIII	$LE \leftrightarrow DS$	11091.971	3928	-	_
	XIX	LE→DS	11099.040	3929	7.069 (1)	\rightarrow cannot be dropped
	XX	LE←DS	11101.807	3929	9.836 (1)	\rightarrow cannot be dropped
		Variance component E: c	omparison non	-nested b	pest models	
		Description	-2LL	df	AIC	BIC
Comparison of non-nested best models using BIC	XVII	$CF: LE \leftrightarrow DS$	11043.102	3921	3201.102	-5539.408
	XVIII	$SC: LE \leftrightarrow DS$	11091.971	3928	3235.971	-5534.720
	XXI	Corr. factor	10998.747	3921	3156.747	-5561.585
		Variance component E: comparison simplex paths of best-fitting model				
		Description	-2LL	df	-2LL and df change	Result*
Simplex paths: nested model evaluation of best-fitting model	XXI	Best model: corr. factor	10998.747	3921	-	-
	XXIc	E Simplex paths	11007.896	3923	9.149 (2)	Simplex paths cannot be dropped
		Description	-2LL	df	AIC	BIC
Best model for E	XXI	Corr. factor	10998.747	3921	3156.747	-5561.585

A, Additive genetic effects; C, shared environmental effects; E, individual-specific effects; -2LL, -2 log likelihood; df, degrees of freedom; AIC, Akaike Information Criterion; BIC, Bayesian Information Criterion; LE, life events; DS, depressive symptoms; Corr. factor, correlated factor model; CF, causal factor model; SC, simple causal model.

First, nested models were evaluated by the decline in fit (-2LL) in relation to df using χ^2 tests. In a second step, the best models resulting from these evaluations were compared to other non-nested models and evaluated using the BIC. In a third step, the best of these models was selected and we tested whether simplex paths could be dropped, again by evaluating the decline in fit using χ^2 tests. The best-fitting A model was used for further testing of the C model and, similarly, the best resulting C model was used for further testing of the E model (see online material for additional information). Note that we may repeat the statistics of the models for different comparisons; however, the model number reflects this.

^a Critical χ^2 values are 3.84 (df=1) and 5.99 (df=2). Above the critical values there is a significant deterioration of fit.

Best-fitting models are marked in bold.

Life events t1	Life events t2	Life events t3	Life events t4
0.37	0.27	0.24	0.18
0.22	0.30	0.24	0.20
0.26	0.27	0.33	0.28
0.24	0.23	0.30	0.32
	Life events t1 0.37 0.22 0.26 0.24	Life events t1 Life events t2 0.37 0.27 0.22 0.30 0.26 0.27 0.24 0.23	Life events t1Life events t2Life events t30.37 0.270.24 0.220.30 0.24 0.260.270.330.240.230.30

Table 3. Within- and cross-time cross-phenotype correlations

Numbers in bold represent correlations between depressive symptoms and life events later in time. Numbers in italics represents correlations between life events and depressive symptoms later in time. Diagonals represent within-time cross-phenotype correlations.

latter evaluation. The final model was then used in further testing of other (C and E) variance components. Because directional effects were not hypothesized to be different across time points, the simplex paths in addition to the causal paths were *a priori* set to the same value.

As temporal changes in variance in longitudinal studies with repeated measurements are informative about the underlying developmental process (Eaves *et al.* 1986), path coefficients of the model on the first occasion only were standardized so that the phenotypic variance is unity. Variances at subsequent occasions were expressed relative to their initial values. Therefore, the path coefficients can exceed unity, particularly when variances are increasing over time.

Results

Phenotypic correlations

Table 3 shows the within- and across-time phenotypic correlations between negative life events and depressive symptoms. Correlations below the diagonal depict the prediction of depressive symptoms by previously experienced life events. Those above the diagonal are informative for the prediction of life event exposure by prior depressive symptoms. These two sets of correlations are broadly similar in magnitude, suggesting that causal effects are probably operating in both directions. However, the hypothesis that a third factor influences both phenotypes is also consistent with this pattern of correlations.

Model fitting

First, fully saturated models were fitted on the observed variables. The model in which means and variances were equated across groups (twin 1, twin 2, MZ and DZ twins) did not fit significantly worse than the fully saturated model as evaluated by the BIC. Second, a Cholesky model with equal means and variances across groups but with different thresholds per time point was tested and compared to the same model but with equal thresholds across time. The latter fitted best and was used for further model testing (for information on fit statistics of these models see online material). Table 2 summarizes the results from the correlational and causal models tested (see online material for further information on the model testing procedure).

For the genetic factor (A), the causal factor model with paths going from negative life events to depressive symptoms (model II) provided the best fit. In addition, simplex paths could be dropped from this model without deterioration in fit. It should be noted that the nested model evaluation of the simple causal models could not distinguish the two unidirectional models. Both unidirectional models were preferred over the bidirectional model. The differences in BIC between model II on the one hand and models V and VI on the other were small (Table 2*a*). For the shared environmental factor (C), model X was the best fit. This model also had a causal factor structure, with one latent factor per phenotype influencing observations at all time points. In contrast to the previous model, this one included unidirectional causal paths from depressive symptoms to life events. Also here, simplex paths could be dropped without fit deterioration (Table 2*b*). Finally, the individual-specific factor (E) was simplified. The best-fitting model for the E factor was the correlated factor model (model XXI). Simplex paths could not be dropped from this model without significant deterioration in fit (Table 2*c*).

Best-fit model

The overall best-fit model (Fig. 2) had the following four key features. First, the genetic risk factors for negative life events and depressive symptoms could be best understood as two single common factors, so







Fig. 2. Resulting best-fit model of the additive genetic (A), shared environmental (B) and individual-specific (E) influences on negative life event exposure and depressive symptoms. (*a*) The causal factor model with paths from negative life events to depressive symptoms (see Table 2*a*). (*b*) The causal factor model with paths from depressive symptoms to negative life events (see Table 2*b*). (*c*) The correlated factor model with simplex paths and additional observation-specific effects (see Table 2*c*).

that the same genetic factors influenced the phenotypes at all four time points (Fig. 2*a*). Second, the causal paths in the genetic portion of the model went from negative life events to depression. Third, the shared environmental factors showed a similar structure to that seen for genetic factors, but with causal

Table 4. Variance components for life events and depressive symptoms at all time points

	Percentage explained variance for A, C and E									
	Life e	vents			Depressive symptoms					
	T1	T2	T3	T4	T1	T2	T3	T4		
A C E	0.17 0.03 0.80	0.10 0.34 0.56	0.24 0.15 0.60	0.42 0.12 0.46	0.40 0 0.60	0.28 0.05 0.68	0.25 0.10 0.66	0.27 0.21 0.52		

A, Additive genetic effects; C, shared environmental effects; E, individual-specific effects.

paths going in the opposite direction, that is from depressive symptoms to life events (Fig. 2*b*). Fourth, by contrast, the individual-specific environmental influences on life events and depressive symptoms could best be modeled as two correlated latent factors (Fig. 2*c*). In addition, forward transmission was present in the E model for both negative life events and depressive symptoms from T1 to T2, T2 to T3 and T3 to T4. That is, levels of life event exposure at one time point had a direct impact on levels of life event exposure at the next time point. Furthermore, depressive symptoms at one time period directly impacted on depressive symptoms at the next time point.

Table 4 shows the estimates for a^2 , c^2 and e^2 for negative life events and depressive symptoms at all time points, obtained from the best-fit model (model XXI).

Discussion

Findings

This study sought to clarify the causal relationship between exposure to negative life events and depressive symptoms using a longitudinal genetically informative design. The most striking features of the best-fitting model are the following. Life events and depressive symptoms had a complex interrelationship that differed depending on the source of variance considered. The best-fitting model was a combination of model specifications with directional paths across time between the two phenotypes and with a correlated factor structure, implying shared influences on both phenotypes. Thus, both causal paths and shared influences explained the phenotypic correlations between negative life events and depressive symptoms. Furthermore, causal paths between the two phenotypes went in both directions, depending on the source of variance. Genetic factors impacted on exposure to life events, which in turn influenced the risk for depressive symptoms. Thus, although the exposure to life events itself was influenced by genetic factors, the life events were causal to the development of depressive symptoms. This model thus shows two different paths by which genes may influence depressive symptoms. First, there is the direct path of additive genetic influences on depressive symptoms. These are probably genes that affect people's vulnerability to depression by acting on biological, cognitive or psychological processes, for example by influencing people's affective processing or genes associated with increased stress responses to negative situations (Wichers et al. 2007, 2009). These genes thus act 'inside the skin'. Second, there is an indirect path from genes to depressive symptoms via exposure to negative life events. These genes act on depression by creating an environment (outside the skin) that exposes the individual to negative life events. Genes that influence the ability to decide on the important choices in life (e.g. choosing one's marital partner, study, job) may impact on exposure to life events. Moreover, genetic influences on having low emotional intelligence or a difficult personality (e.g. high levels of neuroticism) may set people up for the loss of relationships, friendships or jobs. The indirect path is a typical example of active geneenvironment correlation (Plomin et al. 1977). According to the current model, however, it can be calculated from the path coefficients and the total variance that only around 1-2% of the variance at each time point was explained by indirect (outside the skin) paths and 98-99% by the direct (inside the skin) pathway (exact numbers of standardized total effects of the direct and indirect genetic paths available upon request). However, this does not mean that the indirect pathway is non-existent. This model examined effects across time. It is possible that the effects of this path may be larger when examining the effects of life events on depressive symptoms within time, examining the effects of life events of the past 3 months, instead of the life events as reported one time point before.

For shared environmental influences the coefficients of the latent factor on depressive symptoms were modest; however, the strength of the causal path from depressive symptoms to life events was fairly large (see Fig. 2). An example of such shared environmental influences is parental divorce or shared adverse upbringing leading to adult depressive symptoms. The expression of these depressive symptoms may then further increase the risk for life events, such as having a divorce or having relational problems themselves. Thus, also here there is a direct path to the experience of negative life events and an indirect path via the experience of depressive symptoms. The indirect paths (from DS T2 to LE T3 and from DS T3 to LE T4) contribute meaningful percentages of the total effect of the shared environment, 31% and 57% respectively. This finding is consistent with Hammen's hypothesis of stress generation and the idea that the relationship between life events and depressive symptoms is bidirectional (Hammen, 1991, 2005). The fact that experience of depressive symptoms themselves selects an environment of increased stress exposure would imply that lowering depressive symptoms also positively impacts on the environment people create around themselves (Hammen, 1991). This finding also emphasizes the need to resolve residual symptoms following a depressive episode to prevent recurrence (Kennedy & Paykel, 2004).

For individual-specific effects the correlated factor structure fitted best. Thus, phenotypic correlations between life events and depression arise partly from a correlation between the E latent factors that influence life events and depression. Shared individual-specific effects are both affecting risk for life events and depressive symptoms. These could involve experiences such as a physical illness, simultaneously leading to the loss of a job and to feeling down. Another example is being bullied in childhood, leading both to an altered way of coping with daily life situations, resulting in an increased level of life events, and to a more active stress system, resulting in mood symptoms.

Causal paths versus shared causal influence

The finding that genes impact on depressive symptoms through their effect on exposure to negative life events is in agreement with previous studies. Kendler & Karkowski-Shuman (1997) concluded that genetic risk factors for major depressive disorder increase the probability of experiencing stressful life events. That study, however, mentioned the possibility that not only causal effects from life events to depression but also shared causal influence on both phenotypes might have explained the findings. Neuroticism seems a likely candidate because effects of neuroticism on both phenotypes have been shown (Kendler et al. 1993b, 2003a) and neuroticism is partially heritable (Viken et al. 1994; Jang et al. 1996). However, a study using path analyses (Kercher et al. 2009) showed that the best-fitting model did not include any direct paths from neuroticism to depression. Instead it included paths from neuroticism to dependent negative life events and negative thoughts, which in turn, had causal paths to depression. These findings are consistent with those of the current study, in which the best-fitting model showed genetic effects on negative life events, which in turn had causal paths to depressive symptoms.

The current model thus suggests that shared genetic and environmental influences on negative life events and depressive symptoms are expressed at the phenotypic level (at the level of observed negative life events that cause depressive symptoms) and not at the latent level (the level of the latent genetic factor that has direct paths to both negative life events and depressive symptoms). The fact that the current design is able to distinguish models favoring shared influences expressed at the phenotypic level from those favoring shared genetic influence at the latent level is not a trivial or purely theoretical advantage. These two interpretations have substantially different implications for the prevention of depressive symptoms or negative life events. If shared genes indeed impact on depressive symptoms at the phenotypic level, through the generation of negative life events, it follows that decreasing life event exposure should decrease the risk for depressive symptoms, whereas this would not be the case when shared genes exert their effects at the latent level. Likewise, it follows that decreasing the level of depressive symptoms should decrease the risk for negative life events. The current results revealed a fairly large effect of the indirect path to life events, but small effects of the indirect path to depressive symptoms. Because of the clinical relevance of the outcome, there is an urgent need for replication and further examination of these effects.

A drawback of the current model, however, is that it leaves little room for the effect of independent negative life events on depressive symptoms. Previous work has clearly shown that fateful negative events that are uncontrollable by people themselves impact on the risk for depression (Kendler et al. 1999, 2000). The model, therefore, must be interpreted with caution. Moreover, explanations for the lack of effects of independent life events were investigated further. When life events were split into dependent and independent life events, regression analyses showed that more items in the list of life events were dependent (marital discord, fights and arguments with friends or family, loss of job, etc.) than independent (death of partner or close relatives, serious illness, partner or close relatives with serious illness, etc.). In addition, the dependent life events were far more frequent than the independent ones, and the dependent life events had a much stronger depressogenic influence. Therefore, it may be that dependent life events dominated the impact of independent life events in the process of model fit comparisons. Separate analyses for independent and dependent life events might have given a different picture for the associations between independent life events and depressive symptoms, with stronger causal paths of purely environmentally influenced life events on depressive symptoms.

However, because independent life events were rare, such an analysis would be likely to suffer from a lack of power. Dependent life events are usually more dominant in younger life. The mean age of the twins was 28 years, which is relatively young, and which might explain the high frequency of dependent compared to independent life events in this sample.

Methodological issues

Models II, V and VI, in the specification of the A variance component, showed only very small differences in BIC values (\triangle BIC <2). Therefore, caution is warranted regarding the resulting model. The study may have lacked the power to differentiate well between these models. Although not small, the current sample (n = 536) is smaller than the sample size of a previous study (n = 1492) that used similar model-fitting analyses (Kendler et al. 2008). The models, however, in the specification of the C and E variance component could be differentiated with sufficient confidence (Raftery, 1995). In addition, the measurements rely on retrospective self-report of negative life events and depressive symptoms. Finally, this was a female sample only. Therefore, the results may not be generalizable to men.

To summarize, the current study suggests that both reciprocal causation, using cross-time intervals of approximately 3 months, and shared latent influences explained the inter-relationship between negative life events and depressive symptoms. The results of our study should be interpreted with caution as these questions need to be addressed by further studies and replicated before they deserve wide acceptance.

Note

Supplementary material accompanies this paper on the Journal's website (http://journals.cambridge.org/psm).

Acknowledgements

This research was supported by the Dutch Organization for Scientific Research; the Fund for Scientific Research, Flanders and Twins, a non-profit association for scientific research in multiple births (Belgium) (to the East Flanders Prospective Survey); and the Dutch Medical Council (VENI grant no. 916.76.147) (to Dr M. Wichers). We thank all twins for their cooperation.

Declaration of Interest

None.

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