

## MOMENTS OF $k$ -HOP COUNTS IN THE RANDOM-CONNECTION MODEL

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### Abstract

We derive moment identities for the stochastic integrals of multiparameter processes in a random-connection model based on a point process admitting a Papangelou intensity. The identities are written using sums over partitions, and they reduce to sums over non-flat partition diagrams if the multiparameter processes vanish on diagonals. As an application, we obtain general identities for the moments of  $k$ -hop counts in the random-connection model, which simplify the derivations available in the literature.

*Keywords:* Point process; moments; random-connection model; random graph;  $k$ -hop

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### 1. Introduction

The random-connection model, see, e.g., [12, Chapter 6], is a classical model in continuum percolation. It consists of a random graph built on the vertices of a point process on  $\mathbb{R}^d$  by adding edges between two distinct vertices  $x$  and  $y$  with probability  $H(\|x - y\|)$ . In the case of the Rayleigh fading  $H_\beta(\|x - y\|) = e^{-\beta\|x - y\|^2}$  with  $x, y \in \mathbb{R}^2$ , the mean value of the number  $N_k^{x,y}$  of  $k$ -hop paths connecting  $x \in \mathbb{R}^d$  to  $y \in \mathbb{R}^d$  has been computed in [9], together with the variance of 3-hop counts. However, this argument does not extend to  $k \geq 3$  as the proof of the variance identity for 3-hop counts in [9] relies on the known Poisson distribution of the 2-hop count. As shown in [9], the knowledge of moments can provide accurate numerical estimates of the probability  $P(N_k^{x,y} > 0)$  of at least one  $k$ -hop path by expressing it as a series of factorial moments, and the need for a general theory of such moments was pointed out therein.

On the other hand, moment identities for Poisson stochastic integrals with random integrands have been obtained in [18] based on moment identities for Skorohod's integral on the Poisson space; see [16, 17], and also [19] for a review. These moment identities have been extended to point processes with Papangelou intensities in [5], and to multiparameter processes in [2]. Factorial moments have also been computed in [4] for point processes with Papangelou intensities.

In this paper we derive closed-form expressions for the moments of the number of  $k$ -hop paths in the random-connection model. In Proposition 4 the moment of order  $n$  of the  $k$ -hop count is given as a sum over non-flat partitions of  $\{1, \dots, nk\}$  in a general random-connection model based on a point process admitting a Papangelou intensity. Those results are then specialized to the case of Poisson point processes, with an expression for the variance of the  $k$ -hop count given in Corollary 2 using a sum over integer sequences. Finally, in the case of

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Rayleigh fading we show that some results of [9], such as the computation of variance for 3-hop counts, can be recovered via a shorter argument; see Corollary 4.

We proceed as follows. After presenting some background notation on point processes and Campbell measures, see [8], in Section 2 we review the derivation of moment identities for stochastic integrals using sums over partitions. In the multiparameter case we rewrite those identities for processes vanishing on diagonals, based on non-flat partition diagrams. In Section 3 we apply those results to the computation of the moments of  $k$ -hop counts in the random-connection model, and we specialize such computations to the case of Poisson point processes in Section 4. Section 5 is devoted to explicit computations in the case of Rayleigh fading, which result in simpler derivations than the current literature on moments in the random-connection model.

### 1.1. Notation on point processes

Let  $X$  be a Polish space with Borel  $\sigma$ -algebra  $\mathcal{B}(X)$ , equipped with a  $\sigma$ -finite non-atomic measure  $\lambda(dx)$ . We let

$$\Omega^X = \{\omega = \{x_i\}_{i \in I} \subset X : \#(A \cap \omega) < \infty \text{ for all compact } A \in \mathcal{B}(X)\}$$

denote the space of locally finite configurations on  $X$  whose elements  $\omega \in \Omega^X$  are identified with the Radon point measures  $\omega = \sum_{x \in \omega} \epsilon_x$ , where  $\epsilon_x$  denotes the Dirac measure at  $x \in X$ . A point process is a probability measure  $P$  on  $\Omega^X$  equipped with the  $\sigma$ -algebra  $\mathcal{F}$  generated by the topology of vague convergence.

Point processes can be characterized by their Campbell measure  $C$  defined on  $\mathcal{B}(X) \otimes \mathcal{F}$  by

$$C(A \times B) := \mathbb{E} \left[ \int_A \mathbf{1}_B(\omega \setminus \{x\}) \omega(dx) \right], \quad A \in \mathcal{B}(X), \quad B \in \mathcal{F},$$

which satisfies the Georgii–Nguyen–Zessin [14] identity

$$\mathbb{E} \left[ \int_X u(x; \omega) \omega(dx) \right] = \mathbb{E} \left[ \int_{\Omega^X} \int_X u(x; \omega \cup x) C(dx, d\omega) \right] \tag{1}$$

for all measurable processes  $u : X \times \Omega^X \rightarrow \mathbb{R}$  such that both sides of (1) make sense.

In the following we deal with point processes whose Campbell measure  $C(dx, d\omega)$  is absolutely continuous with respect to  $\lambda \otimes P$ , i.e.

$$C(dx, d\omega) = c(x; \omega) \lambda(dx) P(d\omega),$$

where the density  $c(x; \omega)$  is called the Papangelou density. We will also use the random measure  $\hat{\lambda}^n(d\mathbf{x}_n)$  defined on  $X^n$  by

$$\hat{\lambda}^n(d\mathbf{x}_n) = \hat{c}(\mathbf{x}_n; \omega) \lambda(dx_1) \cdots \lambda(dx_n),$$

where  $\hat{c}(\mathbf{x}_n; \omega)$  is the compound Campbell density  $\hat{c} : \Omega_0^X \times \Omega^X \rightarrow \mathbb{R}_+$  defined inductively on the set  $\Omega_0^X$  of finite configurations in  $\Omega^X$  by

$$\hat{c}(\{x_1, \dots, x_n, y\}; \omega) := c(y; \omega) \hat{c}(\{x_1, \dots, x_n\}; \omega \cup \{y\}), \quad n \geq 0; \tag{2}$$

see Relation (1) in [5]. In particular, the Poisson point process with intensity  $\lambda(dx)$  is a point process with Campbell measure  $C = \lambda \otimes P$  and  $c(x; \omega) = 1$ , and in this case the identity (1) becomes the Slivnyak–Mecke formula [20, 11]. Determinantal point processes are examples of point processes with Papangelou intensities, see, e.g., [6, Theorem 2.6], and they can be used for modeling wireless networks with repulsion; see, e.g., [7, 10, 13].

**2. Moment identities**

The moment of order  $n \geq 1$  of a Poisson random variable  $Z_\alpha$  with parameter  $\alpha > 0$  is given by

$$E[Z_\alpha^n] = \sum_{k=0}^n \alpha^k S(n, k), \quad n \in \mathbb{N}, \tag{3}$$

where the Stirling number of the second kind  $S(n, k)$  is the number of ways to partition a set of  $n$  objects into  $k$  non-empty subsets; see, e.g., [3, Proposition 3.1]. Regarding Poisson stochastic integrals of deterministic integrands, in [1] the moment formula

$$E\left[\left(\int_X f(x)\omega(dx)\right)^n\right] = n! \sum_{\substack{r_1+2r_2+\dots+nr_n=n \\ r_1,\dots,r_n \geq 0}} \prod_{k=1}^n \left(\frac{1}{(k!)^{r_k} r_k!} \left(\int_X f^k(x)\lambda(dx)\right)^{r_k}\right) \tag{4}$$

has been proved for deterministic functions  $f \in \bigcap_{p \geq 1} L^p(X, \lambda)$ .

The identity (4) has been rewritten in the language of sums over partitions, and extended to Poisson stochastic integrals of random integrands in [18, Proposition 3.1], and further extended to point processes admitting a Panpangelou intensity in [5, Theorem 3.1]; see also [4]. In the following, given  $\mathfrak{z}_n = (z_1, \dots, z_n) \in X^n$ , we will use the shorthand notation  $\epsilon_{\mathfrak{z}_n}^+$  for the operator

$$(\epsilon_{\mathfrak{z}_n}^+ F)(\omega) = F(\omega \cup \{z_1, \dots, z_n\}), \quad \omega \in \Omega,$$

where  $F$  is any random variable on  $\Omega^X$ . Given  $\rho = \{\rho_1, \dots, \rho_k\} \in \Pi[n]$  a partition of  $\{1, \dots, n\}$  of size  $|\rho| = k$ , we let  $|\rho_i|$  denote the cardinality of each block  $\rho_i$ ,  $i = 1, \dots, k$ .

**Proposition 1.** *Let  $u : X \times \Omega^X \rightarrow \mathbb{R}$  be a (measurable) process. For all  $n \geq 1$  we have*

$$E\left[\left(\int_X u(x; \omega)\omega(dx)\right)^n\right] = \sum_{\rho \in \Pi[n]} E\left[\int_{X^{|\rho|}} \epsilon_{\mathfrak{z}_n}^+ \prod_{l=1}^{|\rho|} u^{|\rho_l|}(z_l) \hat{\lambda}^{|\rho_l|}(d\mathfrak{z}_{|\rho_l|})\right],$$

where the sum runs over all partitions  $\rho$  of  $\{1, \dots, n\}$  with cardinality  $|\rho|$ .

Proposition 1 has also been extended, together with joint moment identities, to multiparameter processes  $(u_{z_1, \dots, z_r})_{(z_1, \dots, z_r) \in X^r}$ ; see [2, Theorem 3.1]. For this, let  $\Pi[n \times r]$  denote the set of all partitions of the set

$$\Delta_{n \times r} := \{1, \dots, n\} \times \{1, \dots, r\} = \{(k, l) : k = 1, \dots, n, l = 1, \dots, r\},$$

identified to  $\{1, \dots, nr\}$ , and let  $\pi := (\pi_1, \dots, \pi_m) \in \Pi[n \times r]$  denote the partition made of the  $n$  blocks  $\pi_k := \{(k, 1), \dots, (k, r)\}$  of size  $r$ , for  $k = 1, \dots, n$ . Given  $\rho = \{\rho_1, \dots, \rho_m\}$  a partition of  $\Delta_{n \times r}$ , we let  $\zeta^\rho : \Delta_{n \times r} \rightarrow \{1, \dots, m\}$  denote the mapping defined as

$$\zeta^\rho(k, l) = p \text{ if and only if } (k, l) \in \rho_p, \quad k = 1, \dots, n, l = 1, \dots, r, p = 1, \dots, m. \tag{5}$$

In other words,  $\zeta^\rho(k, l)$  denotes the index  $p$  of the block  $\rho_p \subset \Delta_{n \times r}$  to which  $(k, l)$  belongs.

Next, we restate Theorem 3.1 of [2] by noting that, in the same way as in Proposition 1, it extends to point processes admitting a Papangelou intensity using the arguments of [4, 5]. When  $(u(z_1, \dots, z_k; \omega))_{z_1, \dots, z_k \in X}$  is a multiparameter process, we will write

$$\epsilon_{\mathfrak{z}_k}^+ u(z_1, \dots, z_k; \omega) := u(z_1, \dots, z_k; \omega \cup \{z_1, \dots, z_k\}), \quad \mathfrak{z}_k = (z_1, \dots, z_k) \in X^n,$$

and in this case we may drop the variable  $\omega \in \Omega^X$  by writing  $\epsilon_{\mathfrak{z}_k}^+ u(z_1, \dots, z_k; \omega)$  instead of  $\epsilon_{\mathfrak{z}_k}^+ u(z_1, \dots, z_k; \omega)$ .

**Proposition 2.** *Let  $u : X^r \times \Omega^X \rightarrow \mathbb{R}$  be a (measurable)  $r$ -process. We have*

$$\mathbb{E} \left[ \left( \int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r) \right)^n \right] = \sum_{\rho \in \Pi[n \times r]} \mathbb{E} \left[ \int_{X^{|\rho|}} \epsilon_{\mathfrak{z}|\rho}^+ \prod_{k=1}^n u(z_{\pi_k}^\rho) \hat{\lambda}^{|\rho|}(d\mathfrak{z}|\rho) \right] \quad (6)$$

with  $z_{\pi_k}^\rho := (z_{\zeta^\rho(k,1)}, \dots, z_{\zeta^\rho(k,r)})$ ,  $k = 1, \dots, n$ .

*Proof.* The main change in the proof argument of [2] is to rewrite the proof of Lemma 2.1 therein by applying (2) recursively, as in the proof of [5, Theorem 3.1], while the main combinatorial argument remains identical.  $\square$

When  $n = 1$ , Proposition 2 yields a multivariate version of the Georgii–Nguyen–Zessin identity (1), i.e.

$$\begin{aligned} \mathbb{E} \left[ \int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r) \right] \\ = \sum_{\rho \in \Pi[1 \times r]} \mathbb{E} \left[ \int_{X^{|\rho|}} \epsilon_{\mathfrak{z}|\rho}^+ u(z_{\zeta^\rho(1,1)}, \dots, z_{\zeta^\rho(1,r)}; \omega) \hat{\lambda}^{|\rho|}(d\mathfrak{z}|\rho) \right]. \end{aligned}$$

### 2.1. Non-flat partitions

In the following we write  $\nu \leq \sigma$  when a partition  $\nu \in \Pi[n \times r]$  is finer than another partition  $\sigma \in \Pi[n \times r]$ , i.e. when every block of  $\nu$  is contained in a block of  $\sigma$ , and we let  $\hat{0} := \{\{1, 1\}, \dots, \{n, r\}\}$  denote the partition of  $\Delta_{n \times r}$  made of singletons. We write  $\rho \wedge \nu = \hat{0}$  when  $\mu = \hat{0}$  is the only partition  $\mu \in \Pi[n \times r]$  such that  $\mu \leq \nu$  and  $\mu \leq \rho$ , i.e.  $|\nu_k \cap \rho_l| \leq 1$ ,  $k = 1, \dots, n, l = 1, \dots, |\rho|$ . In this case we say that the partition diagram  $\Gamma(\nu, \rho)$  of  $\nu$  and  $\rho$  is *non-flat*; see [15, Chapter 4].

Here, a partition  $\rho \in \Pi[n \times r]$  is said to be *non-flat* if the partition diagram  $\Gamma(\pi, \rho)$  of  $\rho$  and the partition  $\pi$  is *non-flat*, where  $\pi := (\pi_1, \dots, \pi_n) \in \Pi[n \times r]$  with  $\pi_k := \{(k, 1), \dots, (k, r)\}$ ,  $k = 1, \dots, n$ . Figure 1 shows an example of a non-flat partition with  $n = 5, r = 4$ , and

$$\begin{aligned} \Delta &= \{(1, 2), (2, 1), (2, 2), (3, 3), (4, 2)\}, \\ \circ &= \{(1, 1), (3, 1), (4, 4), (5, 3)\}, \\ \square &= \{(1, 3), (2, 4), (3, 3), (4, 1), (5, 4)\}, \\ \diamond &= \{(1, 4), (2, 2)\}, \\ \times &= \{(2, 3), (3, 4), (4, 2), (5, 1)\} \\ \pi_k &= \{(k, 1), (k, 2), (k, 3), (k, 4), (k, 5)\}, \quad k = 1, 2, 3, 4, 5. \end{aligned}$$

### 2.2. Processes vanishing on diagonals

The next consequence of Proposition 2 shows that when  $u(z_1, \dots, z_r; \omega)$  vanishes on the diagonals in  $X^r$ , the moments of

$$\int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)$$

reduce to sums over non-flat partition diagrams.

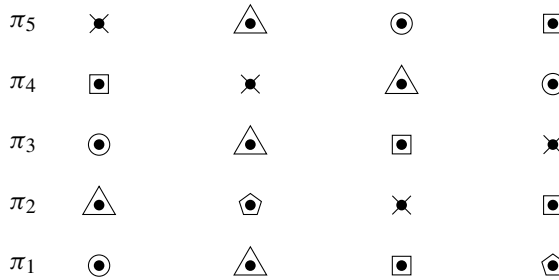


FIGURE 1: Example of a non-flat partition.

**Proposition 3.** Assume that  $u(z_1, \dots, z_r; \omega) = 0$  whenever  $z_i = z_j, 1 \leq i \neq j \leq r, \omega \in \Omega^X$ . Then we have

$$\mathbb{E}\left[\left(\int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)\right)^n\right] = \sum_{\substack{\rho \in \Pi[n \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon_{\mathfrak{z}|\rho}^+ \prod_{k=1}^n u(z_{\pi_k}^\rho) \hat{\lambda}^{|\rho|}(d\mathfrak{z}|\rho)\right].$$

*Proof.* Assume that  $u(z_1, \dots, z_r; \omega)$  vanishes on diagonals, and let  $\rho \in \Pi[n]$ . Then, for any  $z_1, \dots, z_r \in X$  we have

$$\prod_{k=1}^n u(z_{\pi_k}^\rho) = \prod_{k=1}^n u(z_{\zeta^\rho(k,1)}, \dots, z_{\zeta^\rho(k,r)}) = 0$$

whenever  $p := \zeta^\rho(k, a) = \zeta^\rho(k, b)$  for some  $k \in \{1, \dots, n\}$  and  $a \neq b \in \{1, \dots, r\}$ . According to (5) this implies that  $(k, a) \in \rho_p$  and  $(k, b) \in \rho_p$ ; therefore  $\rho$  is not a non-flat partition, and it should be excluded from the sum over  $\Pi[n]$ .  $\square$

When  $n = 1$ , the first moment in Proposition 3 yields the Georgii–Nguyen–Zessin identity

$$\begin{aligned} \mathbb{E}\left[\int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)\right] &= \sum_{\substack{\rho \in \Pi[1 \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon_{\mathfrak{z}|\rho}^+ u(z_{\pi_1}^\rho) \hat{\lambda}^{|\rho|}(d\mathfrak{z}|\rho)\right] \\ &= \mathbb{E}\left[\int_{X^r} \epsilon_{\mathfrak{z}^r}^+ u(z_1, \dots, z_r; \omega) \hat{\lambda}^r(d\mathfrak{z}^r)\right]; \end{aligned} \tag{7}$$

see [9, Lemma IV.1] and [2, Lemma 2.1] for different versions based on the Poisson point process. In the case of second moments, we find that

$$\begin{aligned} \mathbb{E}\left[\left(\int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)\right)^2\right] &= \sum_{\substack{\rho \in \Pi[2 \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon_{\mathfrak{z}|\rho}^+ u(z_{\pi_1}^\rho) u(z_{\pi_2}^\rho) \hat{\lambda}^{|\rho|}(d\mathfrak{z}|\rho)\right], \end{aligned}$$

and since the non-flat partitions in  $\Pi[2 \times r]$  are made of pairs and singletons, this identity can be rewritten as the following consequence of Proposition 3, in which for simplicity of notation we write  $\pi_1 = \{1, \dots, r\}$  and  $\pi_2 = \{r + 1, \dots, 2r\}$ .

**Corollary 1.** Assume that  $u(z_1, \dots, z_r; \omega) = 0$  whenever  $z_i = z_j$ ,  $1 \leq i \neq j \leq r$ ,  $\omega \in \Omega^X$ . Then the second moment of the integral of  $k$ -processes is given by

$$\begin{aligned} \mathbb{E} \left[ \left( \int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r) \right)^2 \right] &= \sum_{A \subset \pi_1} \frac{1}{(r - |A|)!} \\ &\times \sum_{\gamma: \pi_2 \rightarrow A \cup \{r+1, \dots, 2r - |A|\}} \mathbb{E} \left[ \int_{X^{2r - |A|}} \epsilon_{\mathfrak{z}^{2r - |A|}}^+ u(z_{\pi_1}) u(z_{\gamma(r+1)}, \dots, z_{\gamma(2r)}) \hat{\lambda}^{2r - |A|} (d_{\mathfrak{z}^{2r - |A|}}) \right], \end{aligned}$$

where the above sum is over all bijections  $\gamma: \pi_2 \rightarrow A \cup \{r + 1, \dots, 2r - |A|\}$ .

*Proof.* We express the partitions  $\rho \in \Pi[n \times r]$  with non-flat diagrams  $\Gamma(\pi, \rho)$  in Proposition 4 as the collections of pairs and singletons,

$$\rho = \{i, \gamma(i)\}_{i \in A} \cup \{i\}_{i \in \pi_1, i \notin A} \cup \{i\}_{i \in \pi_2, i \notin \gamma(A)},$$

for all subsets  $A \subset \pi_1 = \{1, \dots, r\}$  and bijections  $\gamma: \pi_2 \rightarrow A \cup \{r + 1, \dots, 2r - |A|\}$ . □

In the case of 2-processes, Corollary 1 shows that

$$\begin{aligned} &\mathbb{E} \left[ \left( \int_{X^2} u(z_1, z_2; \omega) \omega(dz_1) \omega(dz_2) \right)^2 \right] \\ &= \sum_{\substack{\rho \in \Pi[n \times 2] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E} \left[ \int_{X^{|\rho|}} \epsilon_{\mathfrak{z}^{|\rho|}}^+ \prod_{k=1}^n u(z_{\zeta^\rho(k,1)}, z_{\zeta^\rho(k,2)}) \hat{\lambda}^{|\rho|} (d_{\mathfrak{z}^{|\rho|}}) \right] \\ &= \sum_{\substack{A \subset \pi_1 \\ \gamma: \{3,4\} \rightarrow A \cup \{3, \dots, 4 - |A|\}}} \frac{1}{(r - |A|)!} \mathbb{E} \left[ \int_{X^{4 - |A|}} \epsilon_{\mathfrak{z}^{4 - |A|}}^+ u(z_1, z_2) u(z_{\gamma(3)}, z_{\gamma(4)}) \hat{\lambda}^{4 - |A|} (d_{\mathfrak{z}^{4 - |A|}}) \right] \\ &= \mathbb{E} \left[ \int_{X^4} \epsilon_{\mathfrak{z}^4}^+ (u(z_1, z_2) u(z_3, z_4)) \hat{\lambda}^4 (d_{\mathfrak{z}^4}) \right] \\ &\quad + \mathbb{E} \left[ \int_{X^3} \epsilon_{\mathfrak{z}^3}^+ (u(z_1, z_2) u(z_1, z_3)) \hat{\lambda}^3 (d_{\mathfrak{z}^3}) \right] + \mathbb{E} \left[ \int_{X^3} \epsilon_{\mathfrak{z}^3}^+ (u(z_2, z_1) u(z_3, z_1)) \hat{\lambda}^3 (d_{\mathfrak{z}^3}) \right] \\ &\quad + \mathbb{E} \left[ \int_{X^3} \epsilon_{\mathfrak{z}^3}^+ (u(z_1, z_2) u(z_2, z_3)) \hat{\lambda}^3 (d_{\mathfrak{z}^3}) \right] + \mathbb{E} \left[ \int_{X^3} \epsilon_{\mathfrak{z}^3}^+ (u(z_2, z_1) u(z_3, z_2)) \hat{\lambda}^3 (d_{\mathfrak{z}^3}) \right] \\ &\quad + \mathbb{E} \left[ \int_{X^2} \epsilon_{\mathfrak{z}^2}^+ (u(z_1, z_2) u(z_1, z_2)) \hat{\lambda}^2 (d_{\mathfrak{z}^2}) \right] + \mathbb{E} \left[ \int_{X^2} \epsilon_{\mathfrak{z}^2}^+ (u(z_1, z_2) u(z_2, z_1)) \hat{\lambda}^2 (d_{\mathfrak{z}^2}) \right]. \end{aligned}$$

Similarly, in the case of 3-processes we find

$$\begin{aligned} &\mathbb{E} \left[ \left( \int_{X^3} u(z_1, z_2, z_3; \omega) \omega(dz_1) \omega(dz_2) \omega(dz_3) \right)^2 \right] \\ &= \sum_{\substack{A \subset \{1,2,3\} \\ \gamma: \{4,5,6\} \rightarrow A \cup \{4, \dots, 6 - |A|\}}} \frac{1}{(3 - |A|)!} \mathbb{E} \left[ \int_{X^5} \epsilon_{\mathfrak{z}^5}^+ u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^5 (d_{\mathfrak{z}^5}) \right] \end{aligned}$$

$$\begin{aligned}
 &= \mathbb{E} \left[ \int_{X^6} \epsilon_{36}^+ u(z_1, z_2, z_3) u(z_4, z_5, z_6) \hat{\lambda}^6(d_{36}) \right] \\
 &+ \frac{1}{2} \sum_{\gamma: \{4,5,6\} \rightarrow \{1,5,6\}} \mathbb{E} \left[ \int_{X^5} \epsilon_{35}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^5(d_{35}) \right] \\
 &+ \frac{1}{2} \sum_{\gamma: \{4,5,6\} \rightarrow \{2,5,6\}} \mathbb{E} \left[ \int_{X^5} \epsilon_{35}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^5(d_{35}) \right] \\
 &+ \frac{1}{2} \sum_{\gamma: \{4,5,6\} \rightarrow \{3,5,6\}} \mathbb{E} \left[ \int_{X^5} \epsilon_{35}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^5(d_{35}) \right] \\
 &+ \sum_{\gamma: \{4,5,6\} \rightarrow \{1,2,6\}} \mathbb{E} \left[ \int_{X^4} \epsilon_{34}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^4(d_{34}) \right] \\
 &+ \sum_{\gamma: \{4,5,6\} \rightarrow \{1,3,6\}} \mathbb{E} \left[ \int_{X^4} \epsilon_{34}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^4(d_{34}) \right] \\
 &+ \sum_{\gamma: \{4,5,6\} \rightarrow \{2,3,6\}} \mathbb{E} \left[ \int_{X^4} \epsilon_{34}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^4(d_{34}) \right] \\
 &+ \sum_{\gamma: \{4,5,6\} \rightarrow \{1,2,3\}} \mathbb{E} \left[ \int_{X^3} \epsilon_{33}^+ u(z_1, z_2, z_3) u(z_\gamma(4), z_\gamma(5), z_\gamma(6)) \hat{\lambda}^3(d_{33}) \right].
 \end{aligned}$$

### 3. Random-connection model

Two point process vertices  $x \neq y$  are independently connected in the random-connection graph with the probability  $H(x, y)$  given  $\omega \in \Omega^X$ , where  $H : X \times X \rightarrow [0, 1]$ . In particular, the 1-hop count  $\mathbf{1}_{\{x \leftrightarrow y\}}$  is a Bernoulli random variable with parameter  $H(x, y)$ , and we have the relation

$$\mathbb{E} \left[ \epsilon_{\mathfrak{z}_r}^+ \prod_{i=0}^r \mathbf{1}_{\{z_i \leftrightarrow z_{i+1}\}}(\omega) \mid \omega \right] = \prod_{i=0}^r H(z_i, z_{i+1})$$

for any subset  $\{z_0, \dots, z_{r+1}\}$  of distinct elements of  $X$ , where  $\mathfrak{z}_r = \{z_1, \dots, z_r\}$  and  $x \leftrightarrow y$  means that  $x \in X$  is connected to  $y \in X$ .

Given  $x, y \in X$ , the number of  $(r + 1)$ -hop sequences  $z_1, \dots, z_r \in \omega$  of vertices connecting  $x$  to  $y$  in the random graph is given by the multiparameter stochastic integral

$$N_{r+1}^{x,y} = \int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)$$

of the  $\{0, 1\}$ -valued  $r$ -process

$$u(z_1, \dots, z_r; \omega) := \mathbf{1}_{\{z_i \neq z_j, 1 \leq i < j \leq r\}} \mathbf{1}_{\{z_1, \dots, z_r \in \omega\}} \prod_{i=0}^r \mathbf{1}_{\{z_i \leftrightarrow z_{i+1}\}}(\omega), \quad z_1, \dots, z_r \in X, \quad (8)$$

which vanishes on the diagonals in  $X^r$ , with  $z_0 := x$  and  $z_{r+1} := y$ . In addition, for any distinct  $z_1, \dots, z_r \in X$  and  $u(z_1, \dots, z_r; \omega)$  given by (8) we have

$$\mathbb{E} \left[ \epsilon_{\mathfrak{z}_r}^+ u(z_1, \dots, z_r; \omega) \mid \omega \right] = \mathbb{E} \left[ \epsilon_{\mathfrak{z}_r}^+ \prod_{i=0}^r \mathbf{1}_{\{z_i \leftrightarrow z_{i+1}\}}(\omega) \mid \omega \right] = \prod_{i=0}^r H(z_i, z_{i+1}), \quad (9)$$

and therefore the first-order moment of the  $(r + 1)$ -hop count between  $x \in X$  and  $y \in X$  is given as

$$\mathbb{E}\left[\int_{X^r} u(z_1, \dots, z_r; \omega)\omega(dz_1) \cdots \omega(dz_r)\right] = \mathbb{E}\left[\int_{X^r} \prod_{i=0}^r H(z_i, z_{i+1})\hat{\lambda}^r(d\mathfrak{z}_r)\right] \tag{10}$$

(see also [9, Theorem II.1]) as a consequence of the Georgii–Nguyen–Zessin identity (7).

In the next proposition we compute the moments of all orders of  $r$ -hop counts as sums over non-flat partition diagrams. The role of the powers  $1/n_{l,i}^\rho$  in (11) is to ensure that all powers of  $H(x, y)$  in (11) are equal to one, since all powers of  $\mathbf{1}_{\{z \leftrightarrow z'\}}$  in (12) below are equal to  $\mathbf{1}_{\{z \leftrightarrow z'\}}$ .

**Proposition 4.** *The moment of order  $n$  of the  $(r + 1)$ -hop count between  $x \in X$  and  $y \in X$  is given by*

$$\mathbb{E}[(N_{r+1}^{x,y})^n] = \sum_{\substack{\rho \in \Pi[n \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \prod_{l=1}^n \prod_{i=0}^r H^{1/n_{l,i}^\rho}(z_{\zeta^\rho(l,i)}, z_{\zeta^\rho(l,i+1)})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right], \tag{11}$$

where  $z_0 = x, z_{r+1} = y, \zeta^\rho(l, 0) = 0, \zeta^\rho(l, r + 1) = r + 1$ , and

$$n_{l,i}^\rho = \#\{(p, j) \in \{1, \dots, n\} \times \{0, \dots, r\} : \{\zeta^\rho(l, i), \zeta^\rho(l, i + 1)\} = \{\zeta^\rho(p, j), \zeta^\rho(p, j + 1)\}\},$$

$$1 \leq l \leq n, 0 \leq i \leq r.$$

*Proof.* Since  $u(z_1, \dots, z_r; \omega)$  vanishes whenever  $z_i = z_j$  for some  $1 \leq i < j \leq r$ , by Proposition 3 we have

$$\begin{aligned} & \mathbb{E}\left[\left(\int_{X^r} u(z_1, \dots, z_r; \omega)\omega(dz_1) \cdots \omega(dz_r)\right)^n\right] \\ &= \sum_{\substack{\rho \in \Pi[n \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \prod_{l=1}^n \prod_{i=0}^r \mathbf{1}_{\{z_{\zeta^\rho(l,i)} \leftrightarrow z_{\zeta^\rho(l,i+1)}\}}\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right] \\ &= \sum_{\substack{\rho \in \Pi[n \times r] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \prod_{l=1}^n \prod_{i=0}^r H^{1/n_{l,i}^\rho}(z_{\zeta^\rho(l,i)}, z_{\zeta^\rho(l,i+1)})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right], \end{aligned} \tag{12}$$

where we applied (9). □

As in Corollary 1, we have the following consequence of Proposition 4, which is obtained by expressing the partitions  $\rho \in \Pi[n \times r]$  with non-flat diagrams  $\Gamma(\pi, \sigma)$  as a collection of pairs and singletons.

**Corollary 2.** *The second moment of the  $(r + 1)$ -hop count between  $x \in X$  and  $y \in X$  is given by*

$$\begin{aligned} \mathbb{E}[(N_{r+1}^{x,y})^2] &= \sum_{\substack{A \subset \pi_1 \\ \gamma: \{1, \dots, r\} \rightarrow A \cup \{r+1, \dots, 2r-|A\}}} \frac{1}{(r - |A|)!} \\ &\times \mathbb{E}\left[\int_{X^{2r-|A|}} \prod_{i=0}^r H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \prod_{j=0}^r H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)})\hat{\lambda}^{2r-|A|}(d\mathfrak{z}_{2r-|A|})\right], \end{aligned}$$



where the above sum is over all bijections  $\gamma : \{1, \dots, r\} \rightarrow A \cup \{r + 1, \dots, 2r - |A|\}$  with  $\gamma(0) := 0, \gamma(r + 1) := r + 1, z_0 := x,$  and  $z_{r+1} := y,$  and

$$n_{1,i}^\gamma = \#\{j \in \{0, \dots, r\} : \{i, i + 1\} = \{\gamma(j), \gamma(j + 1)\}\},$$

$$n_{2,j}^\gamma = \#\{i \in \{0, \dots, r\} : (i, i + 1) = (\gamma(j), \gamma(j + 1))\},$$

for  $0 \leq i \leq r.$

**3.1. Variance of 3-hop counts**

When  $n = 2$  and  $r = 2,$  Corollary 2 allows us to express the variance of the 3-hop count between  $x \in X$  and  $y \in X$  as follows:

$$\begin{aligned} &\text{Var}[N_3^{x,y}] \\ &= \sum_{\substack{\emptyset \neq A \subset \{1,2\} \\ \gamma: \{1,2\} \rightarrow A \cup \{3,4-|A\}}} \frac{1}{(2 - |A|)!} \\ &\quad \times \mathbb{E} \left[ \int_{X^{4-|A|}} \prod_{i=0}^2 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \prod_{j=0}^2 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{4-|A|} (d\mathfrak{z}_{4-|A|}) \right] \\ &= \sum_{\gamma: \{1,2\} \rightarrow \{1,4\}} \mathbb{E} \left[ \int_{X^3} \prod_{i=0}^2 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \prod_{j=0}^2 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^3 (dz_1, dz_2, dz_4) \right] \\ &\quad + \sum_{\gamma: \{1,2\} \rightarrow \{2,4\}} \mathbb{E} \left[ \int_{X^3} \prod_{i=0}^2 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \prod_{j=0}^2 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^3 (dz_1, dz_2, dz_4) \right] \\ &\quad + \sum_{\gamma: \{1,2\} \rightarrow \{1,2\}} \mathbb{E} \left[ \int_{X^2} \prod_{i=0}^2 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \prod_{j=0}^2 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^2 (dz_1, dz_2) \right]. \end{aligned}$$

**3.2. Variance of 4-hop counts**

When  $r = 3$  and  $n = 2,$  Corollary 2 yields

$$\begin{aligned} \text{Var}[N_4^{x,y}] &= \sum_{\substack{\emptyset \neq A \subset \pi_1 \\ \gamma: \{1,\dots,3\} \rightarrow A \cup \{4,\dots,6-|A\}}} \frac{1}{(3 - |A|)!} \mathbb{E} \left[ \int_{X^{6-|A|}} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\ &\quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{6-|A|} (d\mathfrak{z}_{6-|A|}) \right] \\ &= \frac{1}{2} \sum_{\gamma: \{1,\dots,3\} \rightarrow \{1,5,6\}} \mathbb{E} \left[ \int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\ &\quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5 (dz_1, dz_2, dz_3, dz_5, dz_6) \right] \end{aligned}$$

$$\begin{aligned}
 & + \frac{1}{2} \sum_{\gamma:\{1,\dots,3\}\rightarrow\{2,5,6\}} \mathbb{E} \left[ \int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5(dz_1, dz_2, dz_3, dz_5, dz_6) \right] \\
 & + \frac{1}{2} \sum_{\gamma:\{1,\dots,3\}\rightarrow\{3,5,6\}} \mathbb{E} \left[ \int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5(dz_1, dz_2, dz_3, dz_5, dz_6) \right] \\
 & + \sum_{\gamma:\{1,\dots,3\}\rightarrow\{1,2,6\}} \mathbb{E} \left[ \int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\
 & + \sum_{\gamma:\{1,\dots,3\}\rightarrow\{1,3,6\}} \mathbb{E} \left[ \int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\
 & + \sum_{\gamma:\{1,\dots,3\}\rightarrow\{2,3,6\}} \mathbb{E} \left[ \int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\
 & + \sum_{\gamma:\{1,\dots,3\}\rightarrow\{1,\dots,3\}} \mathbb{E} \left[ \int_{X^3} \prod_{i=0}^3 H^{1/n_{1,i}^\gamma}(z_i, z_{i+1}) \right. \\
 & \quad \left. \times \prod_{j=0}^3 H^{1/n_{2,j}^\gamma}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^3(dz_1, dz_2, dz_3) \right].
 \end{aligned}$$

**4. Poisson case**

In this section and the next one we will work in the Poisson random-connection model, using a Poisson point process on  $X = \mathbb{R}^d$  with intensity  $\lambda(dx)$  on  $\mathbb{R}^d$ . We let

$$H^{(n)}(x_0, x_n) := \int_{\mathbb{R}^d} \cdots \int_{\mathbb{R}^d} \prod_{i=0}^{n-1} H(x_i, x_{i+1}) \lambda(dx_1) \cdots \lambda(dx_{n-1}), \quad x_0, x_n \in \mathbb{R}^d, \quad n \geq 1.$$

(13)

The 2-hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$  is given by the first-order stochastic integral

$$\int_{\mathbb{R}^d} u(z; \omega)\omega(dz) = \int_{\mathbb{R}^d} \mathbf{1}_{\{x \leftrightarrow z_1\}} \mathbf{1}_{\{z_1 \leftrightarrow y\}}(\omega)\omega(dz_1) = \int_{\mathbb{R}^d} \mathbf{1}_{\{x \leftrightarrow z_1\}} \mathbf{1}_{\{z_1 \leftrightarrow y\}}\omega(dz_1),$$

and its moment of order  $n$  is

$$\begin{aligned} \mathbb{E}\left[\left(\int_{\mathbb{R}^d} u(z_1; \omega)\omega(dz_1)\right)^n\right] &= \mathbb{E}\left[\left(\int_{\mathbb{R}^d} \mathbf{1}_{\{x \leftrightarrow z_1\}} \mathbf{1}_{\{z_1 \leftrightarrow y\}}\omega(dz_1)\right)^n\right] \\ &= \sum_{\rho \in \Pi[n \times 1]} \int_{X^{|\rho|}} \prod_{l=1}^{|\rho|} (H(x, z_l)H(z_l, y))\lambda^{|\rho|}(dz_1, \dots, dz_{|\rho|}) \\ &= \sum_{k=1}^n S(n, k) \left(\int_{\mathbb{R}^d} H(x, z)H(z, y)\lambda(dz)\right)^k \\ &= \sum_{k=1}^n S(n, k)(H^{(2)}(x, y))^k; \end{aligned}$$

therefore, from (3), the 2-hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$  is a Poisson random variable with mean

$$H^{(2)}(x, y) = \int_{\mathbb{R}^d} H(x, z)H(z, y)\lambda(dz).$$

By (10), the first-order moment of the  $r$ -hop count is given by

$$H^{(r)}(x, y) = \int_{X^{r-1}} \prod_{i=0}^{r-1} H(z_i, z_{i+1})\lambda^{r-1}(dz_1, \dots, dz_{r-1}).$$

**Corollary 3.** *The variance of the  $r$ -hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$  is given by*

$$\begin{aligned} \text{Var}[N_r^{x,y}] &= \sum_{p=1}^{r-1} \sum_{\substack{1 \leq k_1 < \dots < k_p < r \\ 1 \leq l_1 < \dots < l_p < r}} \sum_{\sigma \in \Sigma[p]} \int_{X^p} \prod_{0 \leq i \leq p} H^{(k_{i+1}-k_i)}(z_i, z_{i+1}) \\ &\quad \times \prod_{\substack{0 \leq j \leq p \\ l_{\sigma(j+1)}-l_{\sigma(j)}+k_{j+1}-k_j > 2 \\ \text{or } \{j, j+1\} \neq \{\sigma(j), \sigma(j+1)\}}} H^{(l_{\sigma(j+1)}-l_{\sigma(j)})}(z_{\sigma(j)}, z_{\sigma(j+1)})\lambda^p(d_{3p}), \end{aligned}$$

with  $k_0 = l_0 = 0$ ,  $k_{p+1} = l_{p+1} = r$ ,  $\sigma(0) = 0$ , and  $\sigma(r) = r$ , where the above sum is over all permutations  $\sigma \in \Sigma[p]$  of  $\{1, \dots, p\}$ .

*Proof.* We rewrite the result of Corollary 2 by denoting the set  $A \subset \pi_1$  as  $A = \{k_1, \dots, k_p\}$ , for  $1 \leq k_1 < \dots < k_p \leq r - 1$ , and we identify  $\gamma(A) \subset A \cup \{r + 1, \dots, 2r - |A|\}$  with  $\{l_1, \dots, l_p\}$ , which requires a sum over the permutations of  $\{1, \dots, p\}$  since  $1 \leq l_1 < \dots < l_p \leq r - 1$ , where  $1 \leq p \leq r - 1$ . In addition, the multiple integrals over contiguous index sets in  $A^c$  are evaluated using (13). □

### 4.1. Variance of 3-hop counts

When  $n = 2$  and  $r = 2$  Corollary 3 allows us to compute the variance of the 3-hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$ , as follows:

$$\begin{aligned} \text{Var}[N_3^{x,y}] &= 2 \int_{\mathbb{R}^d} H(x, z_1)H^{(2)}(z_1, y)H^{(2)}(z_1, y)\lambda(dz_1) \\ &\quad + 2 \int_{\mathbb{R}^d} H(x, z_1)H^{(2)}(x, z_1)H^{(2)}(z_1, y)H(z_1, y)\lambda(dz_1) \\ &\quad + \int_{X^2} H(x, z_1)H(z_1, z_2)H(z_2, y)H(x, z_2)H(z_1, y)\lambda^2(dz_1, dz_2) + H^{(3)}(x, y). \end{aligned} \tag{14}$$

By Corollary 3 the variance of 4-hop counts can be similarly computed explicitly as a sum of 33 terms.

### 5. Rayleigh fading

In this section we consider a Poisson point process on  $X = \mathbb{R}^d$  with flat intensity  $\lambda(dx) = \lambda dx$  on  $\mathbb{R}^d$ ,  $\lambda > 0$ , and a Rayleigh fading function of the form

$$H_\beta(x, y) := e^{-\beta\|x-y\|^2}, \quad x, y \in \mathbb{R}^d, \beta > 0.$$

Lemmas 1 and 2 can be used to evaluate the integrals appearing in Corollary 3 and in the variance (14) of the 3-hop counts.

**Lemma 1.** For all  $n \geq 1$ ,  $y_1, \dots, y_n \in \mathbb{R}^d$ , and  $\beta_1, \dots, \beta_n > 0$  we have

$$\begin{aligned} &\int_{\mathbb{R}^d} \prod_{i=1}^n H_{\beta_i}(x, y_i) dx \\ &= \left( \frac{\pi}{\beta_1 + \dots + \beta_n} \right)^{d/2} \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left( y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i} \right). \end{aligned}$$

*Proof.* We start by showing that for all  $n \geq 1$  we have

$$\begin{aligned} &\prod_{i=1}^n H_{\beta_i}(x, y_i) \\ &= H_{\beta_1 + \dots + \beta_n} \left( x, \frac{\beta_1 y_1 + \dots + \beta_n y_n}{\beta_1 + \dots + \beta_n} \right) \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left( y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i} \right). \end{aligned} \tag{15}$$

Clearly, this relation holds for  $n = 1$ . In addition, at the rank  $n = 2$  we have

$$\begin{aligned} H_{\beta_1}(x, y_1)H_{\beta_2}(x, y_2) &= e^{-\beta_1\|y_1-x\|^2} e^{-\beta_2\|x-y_2\|^2} \\ &= \exp\{-\beta_1\|y_1\|^2 - \beta_2\|y_2\|^2 + 2\langle\beta_1 y_1 + \beta_2 y_2, x\rangle - (\beta_1 + \beta_2)\|x\|^2\} \\ &= \exp\{-\beta_1\|y_1\|^2 - \beta_2\|y_2\|^2 - (\beta_1 + \beta_2)\|x - (\beta_1 y_1 + \beta_2 y_2)/(\beta_1 + \beta_2)\|^2 \\ &\quad + \|\beta_1 y_1 + \beta_2 y_2\|^2/(\beta_1 + \beta_2)\} \\ &= \exp\{-(\beta_1 + \beta_2)\|x - (\beta_1 y_1 + \beta_2 y_2)/(\beta_1 + \beta_2)\|^2 - \beta_1 \beta_2 \|y_1 - y_2\|^2/(\beta_1 + \beta_2)\} \\ &= H_{\beta_1 + \beta_2} \left( x, \frac{\beta_1 y_1 + \beta_2 y_2}{\beta_1 + \beta_2} \right) H_{\frac{\beta_1 \beta_2}{\beta_1 + \beta_2}}(y_1, y_2). \end{aligned}$$

Next, assuming that (15) holds at the rank  $n \geq 1$ , we have

$$\begin{aligned} \prod_{i=1}^{n+1} H_{\beta_i}(x, y_i) &= H_{\beta_{n+1}}(x, y_{n+1})H_{\beta_1+\dots+\beta_n}\left(x, \frac{\beta_1y_1 + \dots + \beta_ny_n}{\beta_1 + \dots + \beta_n}\right) \\ &\quad \times \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1+\dots+\beta_i)}{\beta_1+\dots+\beta_{i+1}}}\left(y_{i+1}, \frac{\beta_1y_1 + \dots + \beta_iy_i}{\beta_1 + \dots + \beta_i}\right) \\ &= H_{\beta_1+\dots+\beta_{n+1}}\left(x, \frac{\beta_1y_1 + \dots + \beta_{n+1}y_{n+1}}{\beta_1 + \dots + \beta_n}\right) \\ &\quad \times \prod_{i=1}^n H_{\frac{\beta_{i+1}(\beta_1+\dots+\beta_i)}{\beta_1+\dots+\beta_{i+1}}}\left(y_{i+1}, \frac{\beta_1y_1 + \dots + \beta_iy_i}{\beta_1 + \dots + \beta_i}\right). \end{aligned}$$

As a consequence, we find that

$$\begin{aligned} \int_{\mathbb{R}^d} \prod_{i=1}^n H_{\beta_i}(x, y_i) dx &= \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1+\dots+\beta_i)}{\beta_1+\dots+\beta_{i+1}}}\left(y_{i+1}, \frac{\beta_1y_1 + \dots + \beta_iy_i}{\beta_1 + \dots + \beta_i}\right) \\ &\quad \times \int_{\mathbb{R}^d} H_{\beta_1+\dots+\beta_n}\left(x, \frac{\beta_1y_1 + \dots + \beta_ny_n}{\beta_1 + \dots + \beta_n}\right) dx \\ &= \left(\frac{\pi}{\beta_1 + \dots + \beta_n}\right)^{d/2} \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1+\dots+\beta_i)}{\beta_1+\dots+\beta_{i+1}}}\left(y_{i+1}, \frac{\beta_1y_1 + \dots + \beta_iy_i}{\beta_1 + \dots + \beta_i}\right). \quad \square \end{aligned}$$

In particular, applying Lemma 1 for  $n = 2$  yields

$$\begin{aligned} \int_{\mathbb{R}^d} H_{\beta_1}(y_1, x)H_{\beta_2}(x, y_2) dx &= \left(\frac{\pi}{\beta_1 + \beta_2}\right)^{d/2} H_{\frac{\beta_1\beta_2}{\beta_1+\beta_2}}(y_1, y_2) \\ &= \left(\frac{\pi}{\beta_1 + \beta_2}\right)^{d/2} e^{-\beta_1\beta_2\|y_1-y_2\|^2/(\beta_1+\beta_2)}, \quad y_1, y_2 \in \mathbb{R}^d, \quad (16) \end{aligned}$$

and the 2-hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$  is a Poisson random variable with mean

$$\begin{aligned} H_{\beta}^{(2)}(x, y) &= \lambda \int_{\mathbb{R}^d} H_{\beta}(x, z)H_{\beta}(z, y) dz \\ &= \lambda \left(\frac{\pi}{2\beta}\right)^{d/2} H_{\beta/2}(x, y) \\ &= \lambda \left(\frac{\pi}{2\beta}\right)^{d/2} e^{-\|x-y\|^2/2}. \end{aligned}$$

By an induction argument similar to that of Lemma 1, we obtain the following lemma.

**Lemma 2.** For all  $n \geq 1$ ,  $x_0, \dots, x_n \in \mathbb{R}^d$ , and  $\beta_1, \dots, \beta_n > 0$  we have

$$\begin{aligned} \int_{\mathbb{R}^d} \dots \int_{\mathbb{R}^d} \prod_{i=1}^n H_{\beta_i}(x_{i-1}, x_i) dx_1 \dots dx_{n-1} \\ = \left(\frac{\pi^{n-1}}{\sum_{i=1}^n \beta_1 \dots \beta_{i-1}\beta_{i+1} \dots \beta_n}\right)^{d/2} H_{\frac{\beta_1 \dots \beta_n}{\sum_{i=1}^n \beta_1 \dots \beta_{i-1}\beta_{i+1} \dots \beta_n}}(x_0, y_n). \end{aligned}$$

*Proof.* Clearly the relation holds at the rank  $n = 1$ . Assuming that it holds at the rank  $n \geq 1$  and using (16), we have

$$\begin{aligned} & \int_{\mathbb{R}^d} \cdots \int_{\mathbb{R}^d} \prod_{i=1}^{n+1} H_{\beta_i}(x_{i-1}, x_i) \, dx_1 \cdots dx_n \\ &= \int_{\mathbb{R}^d} H_{\beta_{n+1}}(x_n, x_{n+1}) \int_{\mathbb{R}^d} \cdots \int_{\mathbb{R}^d} \prod_{i=1}^n H_{\beta_i}(x_{i-1}, x_i) \, dx_1 \cdots dx_n \\ &= \left( \frac{\pi^{n-1}}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n} \right)^{d/2} \\ & \quad \times \int_{\mathbb{R}^d} H_{\frac{\beta_1 \cdots \beta_n}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n}}(x_0, x_n) H_{\beta_{n+1}}(x_n, x_{n+1}) \, dx_n \\ &= \left( \frac{\pi^{n-1}}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n} \right)^{d/2} \\ & \quad \times \left( \frac{\pi}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n + \beta_{n+1}} \right)^{d/2} H_{\frac{\beta_1 \cdots \beta_{n+1}}{\sum_{i=1}^{n+1} \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_{n+1}}}(x_0, x_{n+1}). \quad \square \end{aligned}$$

In particular, the first-order moment of the  $r$ -hop count between  $x_0 \in \mathbb{R}^d$  and  $x_r \in \mathbb{R}^d$  is given by

$$\begin{aligned} H_{\beta}^{(r)}(x_0, x_r) &= \int_{\mathbb{R}^d} \cdots \int_{\mathbb{R}^d} \prod_{i=0}^{r-1} H_{\beta}(x_i, x_{i+1}) \lambda(dx_1) \cdots \lambda(dx_{r-1}) \\ &= \lambda^{r-1} \left( \frac{\pi^{r-1}}{r\beta^{r-1}} \right)^{d/2} H_{\beta/r}(x, y) \\ &= \lambda^{r-1} \left( \frac{\pi^{r-1}}{r\beta^{r-1}} \right)^{d/2} e^{-\beta\|x-y\|^2/r}, \quad x, y \in \mathbb{R}^d. \end{aligned} \tag{17}$$

### 5.1. Variance of 3-hop counts

Corollary 3 and Lemma 2 allow us to recover Theorem II.3 of [9] for the variance of 3-hop counts by a shorter argument, while extending it from the plane  $X = \mathbb{R}^2$  to  $X = \mathbb{R}^d$ .

**Corollary 4.** *The variance of the 3-hop count between  $x \in \mathbb{R}^d$  and  $y \in \mathbb{R}^d$  is given by*

$$\begin{aligned} \text{Var}[N_3^{x,y}] &= 2\lambda^3 \left( \frac{\pi^3}{8\beta^3} \right)^{d/2} e^{-\beta\|x-y\|^2/2} + \lambda^2 \left( \frac{\pi^2}{3\beta^2} \right)^{d/2} e^{-\beta\|x-y\|^2/3} \\ & \quad + 2\lambda^3 \left( \frac{\pi^3}{12\beta^3} \right)^{d/2} e^{-3\beta\|x-y\|^2/4} + \lambda^2 \left( \frac{\pi^2}{8\beta^2} \right)^{d/2} e^{-\beta\|x-y\|^2}. \end{aligned}$$

*Proof.* By (17) and Lemma 2 we have

$$\begin{aligned} & \int_{\mathbb{R}^d} H_{\beta}(x, z_1) H_{\beta}^{(2)}(z_1, y) H_{\beta}^{(2)}(z_1, y) \lambda(dz_1) \\ &= \lambda^2 \left( \frac{\pi^2}{4\beta^2} \right)^{d/2} \int_{\mathbb{R}^d} H_{\beta}(x, z_1) H_{\beta/2}^2(z_1, y) \lambda(dz_1) \end{aligned}$$

$$\begin{aligned}
&= \lambda^3 \left( \frac{\pi^2}{4\beta^2} \right)^{d/2} \int_{\mathbb{R}^d} H_\beta(x, z_1) H_\beta(z_1, y) \lambda(dz_1) = \lambda^3 \left( \frac{\pi^3}{8\beta^3} \right)^{d/2} H_{\beta/2}(x, y); \\
&\int_{\mathbb{R}^d} H_\beta(x, z_1) H_\beta^{(2)}(x, z_1) H_\beta^{(2)}(z_1, y) H_\beta(z_1, y) \lambda(dz_1) \\
&= \lambda^2 \left( \frac{\pi^2}{4\beta^2} \right)^{d/2} \int_{\mathbb{R}^d} H_{3\beta/2}(z_1, y) H_{3\beta/2}(x, z_1) \lambda(dz_1) = \lambda^3 \left( \frac{\pi^3}{12\beta^3} \right)^{d/2} H_{3\beta/4}(x, y); \\
&\int_{X^2} H_\beta(x, z_1) H_\beta(z_1, z_2) H_\beta(z_2, y) H_\beta(x, z_2) H_\beta(z_1, y) \lambda^2(dz_1, dz_2) \\
&= \lambda \left( \frac{\pi}{3\beta} \right)^{d/2} H_\beta(x, y) \int_{\mathbb{R}^d} H_{2\beta/3}(z_2, (x+y)/2) H_{2\beta}(z_2, (x+y)/2) \lambda(dz_2) \\
&= \lambda^2 \left( \frac{\pi^2}{8\beta^2} \right)^{d/2} H_\beta(x, y);
\end{aligned}$$

and we conclude by (14). □

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