

# Parental mental illness and fatal birth defects in a national birth cohort

R. T. Webb<sup>1,2\*</sup>, A. R. Pickles<sup>2</sup>, S. A. King-Hele<sup>1</sup>, L. Appleby<sup>1</sup>, P. B. Mortensen<sup>3</sup> and K. M. Abel<sup>1</sup>

<sup>1</sup> Centre for Women's Mental Health Research, The University of Manchester, Manchester, UK

<sup>2</sup> Biostatistics/Health Methodology Research Group, The University of Manchester, Manchester, UK

<sup>3</sup> National Centre for Register-based Research, University of Aarhus, Denmark

**Background.** Few large studies describe links between maternal mental illness and risk of major birth defect in offspring. Evidence is sparser still for how effects vary between maternal diagnoses and no previous study has assessed risk with paternal illnesses.

**Method.** A population-based birth cohort was created by linking Danish national registers. We identified all singleton live births during 1973–1998 ( $n=1.45$  m), all parental psychiatric admissions from 1969 onwards, and all fatal birth defects until 1 January 1999. Linkage and case ascertainment were almost complete. Relative risks were estimated using Poisson regression.

**Results.** Risk of fatal birth defect was elevated in relation to history of any maternal admission and also with affective disorders specifically, although the strongest effect found was with maternal schizophrenia. The rate was more than doubled in this group compared to the general population [relative risk (RR) 2.34, 95% confidence interval (CI) 1.45–3.77], which also represented a significant excess risk compared with all other admitted maternal disorders ( $p=0.018$ ). Risk of death from causes other than birth defect was no higher with schizophrenia than with other maternal conditions. There was no elevation in risk of fatal birth defect if the father was admitted with schizophrenia or any other psychiatric diagnosis.

**Conclusions.** There are many possible explanations for a higher risk of fatal birth defect with maternal schizophrenia and affective disorder. These include genetic effects directly linked with maternal illness, lifestyle factors (diet, smoking, alcohol and drugs), poor antenatal care, psychotropic medication toxicity, and gene–environment interactions. Further research is needed to elucidate the causal mechanisms.

Received 8 May 2007; Revised 16 October 2007; Accepted 30 October 2007; First published online 13 December 2007

**Key words:** Abnormalities, epidemiology, maternal exposure, mood disorder, schizophrenia.

## Introduction

The link between minor physical anomalies and adult onset of schizophrenia has been widely described (McNeil *et al.* 2000; Weinberg *et al.* 2007). Such anomalies occur more frequently in a range of developmental deficits as well as in schizophrenia, which is indicative of a neurodevelopmental aetiology (Lloyd *et al.* 2003). However, there have been few large epidemiological investigations of major birth defects in offspring of parents with mental illness. For example, a national Danish study (Bennedsen *et al.* 2001) reported a 70% higher risk of any type of registered birth defect with maternal schizophrenia. Studies have also

attempted to identify teratogenic effects of *in utero* exposure to psychotropic medication (e.g. Kallen & Tanberg, 1983; McKenna *et al.* 2005). However, in general these studies have been based on relatively small samples with possible selection and reporting biases. Strong population-based evidence is needed to describe risk of major birth defect in offspring of mentally ill parents, across a range of parental diagnostic categories and between maternal *versus* paternal psychopathology. In particular, evidence for effects linked with paternal disorders is lacking.

Our recent critical review found no published studies of fatal birth defect risk (Webb *et al.* 2005). For this empirical study we aimed to estimate relative risks associated with history of parental psychiatric admission. Investigation of fatal birth defect was selected as an objective and clinically important outcome measure with complete case ascertainment through infancy and childhood. We were primarily

\* Address for correspondence: Dr R. T. Webb, Centre for Women's Mental Health Research/Health Methodology Research Group, Williamson Building, The University of Manchester, Oxford Road, Manchester M13 9PL, UK.  
(Email: roger.webb@manchester.ac.uk)

interested in estimating risks linked with parental schizophrenia and related disorders, with affective disorders and alcohol/drug-related illnesses also assessed as comparison groups of parents with severe mental illness.

## Method

### *Description of the study cohort, exposures and outcomes*

The birth cohort is described in greater detail in our previous paper (Webb *et al.* 2006), which reported all-cause offspring mortality risks from the same cohort. It consisted of all singleton live births in Denmark during 1973–1998 to Danish-born mothers ( $n=1.45$  million). As with our previous reports from this cohort (Webb *et al.* 2006, 2007; King-Hele *et al.* 2007) exposure was defined by date of first maternal or paternal psychiatric admission. All admissions nationally from 1969 onwards were identified through the Psychiatric Central Register (Munk-Jørgensen & Mortensen, 1997), according to the following diagnostic categories:

- (1) All psychiatric diagnoses: ICD-8: 290–315; ICD-10: F00–F99.
- (2) Schizophrenia and related disorders (schizophrenia, schizophrenia-like or schizo-affective): ICD-8: 295, 296.8, 297, 298.39, 301.83; ICD-10: F20–F29.
- (3) Affective disorders (including bipolar disorder): ICD-8: 296.09, 296.19, 296.29, 296.39, 296.99, 298.09, 298.19, 300.49, 301.19; ICD-10: F30–F39.
- (4) Alcohol/drug-related disorders: ICD-8: 291, 294.30, 294.38, 303; ICD-10: F10, F11, F16, F18, F19.

The diagnostic category ‘schizophrenia and related disorders’ is referred to as ‘schizophrenia’ for the remainder of the paper, with the broader diagnostic group selected to maximize statistical power and precision. Although diagnoses recorded in the Psychiatric Central Register are routinely checked against those entered in the medical notes (Munk-Jørgensen & Mortensen, 1997), they are nonetheless made heterogeneously in the course of routine clinical practice and without reference to standardized research criteria (Strudsholm *et al.* 2005). For example, a registered diagnosis with schizophrenia has a high positive predictive value (Munk-Jørgensen, 1995), but the consistency of an affective disorder diagnosis over time is less certain (Kessing, 1998).

The terms ‘birth defect’, ‘congenital malformation’, ‘congenital anomaly’ and ‘congenital abnormality’ are used interchangeably in the literature; in this paper we refer consistently to ‘fatal birth defects’. All types of fatal birth defect were ascertained by linkage

to the national Causes of Death Register (Juel & Helweg-Larsen, 1999) using ICD-8 codes 740–759 (‘congenital anomalies’) for years 1973–1993 (WHO, 1967) and ICD-10 Q00–Q99 (‘congenital malformations, deformations and chromosomal abnormalities’) for 1994 onwards (WHO, 1992). Risks of death from all other causes were also estimated to assess the specificity with particular parental diagnostic categories of observed effects on fatal birth defect risk. Predominant among these other causes were: immaturity, anoxia/brain injury and other perinatal conditions; cancer; sudden infant death syndrome; and unnatural causes (mostly accidents). Classification was made according to the underlying cause of death, with coding being over 99% complete up to 31 December 1998. Since 1976, all Danish death certificates have to be completed by a physician but, as with any national mortality registration system, validity relies on the quality of physicians’ notification and the accuracy of subsequent classification and coding performed at the National Board of Health (Juel & Helweg-Larsen, 1999). No external validation of the cause of death coding could be performed because of the historical nature of the data. Inconsistency in classification between the eighth and tenth ICD revisions was minimized by use of a broad range of codes to delineate both exposures and outcomes.

We could find no standardized age range for measuring risk of fatal birth defect reported in the paediatric literature. However, we opted to use 0–4 years in line with the age stratification applied in our previous studies of this cohort (Webb *et al.* 2006, 2007; King-Hele *et al.* 2007), and also because 96% (3867/4042) of all fatal birth defects in childhood (before 16th birthday) occurred at ages 0–4 years. Stillbirths due to birth defect were excluded from our analyses because paternal identity could not be established for these cases, and one of our primary aims was to compare risks of fatal birth defect with maternal *versus* paternal mental illness.

Maternal identity was registered for every birth in the cohort, whereas paternal identity was missing for 1.2% of all live births, 15% of fatal birth defects at 0–4 years and 18% of deaths due to any other causes in this age range. This missing paternal identity among offspring deaths is due to an artefact in the Danish Civil Registration System; the younger the age of death the higher the percentage of missing paternities (and so 27% of all early neonatal deaths lacked a registered father). Following birth, a child’s mother is automatically registered, whereas parents must themselves register the father’s identity at the mother’s local parish church some time later. For married people, the child’s father is always assumed to be the mother’s husband, so rates of unregistered fatherhood

are especially high in offspring of unmarried mothers. In cases of perinatal death, paternal registration only occurs if requested by the mother, or if this information is required for legal purposes. The rate of unregistered paternity has fallen since 1990. It should also be noted that the registered father may not necessarily have been the biological one as this information is not recorded in the Civil Registration System (Pedersen *et al.* 2006).

### Statistical analyses

These were conducted using Stata software version 9 (StataCorp, College Station, TX, USA). Relative risks were estimated by log-linear Poisson regression models without overdispersion (Gardner *et al.* 1995). The models were adjusted for calendar year of death (in 5-year bands) and finely stratified offspring age, and the interaction between these terms. The reference group for calculating relative risk was offspring whose mothers or fathers had not been admitted for psychiatric treatment at any time before or during the observation period. A pooled estimate from birth to early childhood (0–4 years) was made to achieve adequate precision in comparing effects between parental diagnostic categories. Effects specific to a particular parental diagnostic category (e.g. schizophrenia *versus* all other admitted maternal conditions) were assessed by two-tailed Wald *z* tests for linear comparison of the regression coefficients. To eliminate the possibility of reverse causality bias, offspring subjects who died were only classified as being exposed if first parental admission occurred prior to their death.

## Results

### Descriptive analyses

Exposure prevalence by parental diagnosis, numbers of fatal birth defects *versus* other causes of death and frequency by type of birth defect, across the whole birth cohort, are shown in Table 1. With parental disorder measured according to parental admission before the child's fifth birthday or death, the prevalence of exposure was low (more than 2% for any maternal or paternal admission and less than 1% for specific diagnostic groups). Fifty per cent of all cases were defects of the circulatory system, 89% of which were cardiac defects; the next most common type was nervous system defects (15% of the total). Distributions by type of fatal birth defect and according to parental illness exposure status are presented in Table 2. The pattern by fatal birth defect type was generally similar across the categories of exposed and unexposed offspring subjects.

**Table 1.** Descriptive data for the whole national birth cohort: exposure prevalence, number of deaths and type of fatal birth defect

Exposure prevalence, person-years (%)	
Maternal disorders <sup>a</sup>	
All psychiatric admissions	155 215 (2.4)
Schizophrenia (and related disorder)	15 932 (0.2)
Affective disorder	36 001 (0.6)
Alcohol/drug-related disorders	28 192 (0.4)
Paternal disorders <sup>a</sup>	
All psychiatric admissions	152 085 (2.3)
Schizophrenia (and related disorder)	14 519 (0.2)
Affective disorder	26 636 (0.4)
Alcohol/drug-related disorders	64 867 (1.0)
Deaths <sup>b</sup> , <i>n</i> (%)	
Fatal birth defects	3867 (31.6)
All other deaths	8371 (68.4)
Type of birth defect <sup>b</sup> , <i>n</i> (%)	
1. Nervous system (ICD-8 740–743; ICD-10 Q00–07)	571 (15.2)
2. Eye, ear, face and neck (ICD-8 744–745; ICD-10 Q10–18)	2 (0.0)
3. Circulatory system (ICD 746–747; ICD-10 Q20–28)	1936 (50.1)
4. Respiratory system (ICD-8 748; ICD-10 Q30–34)	105 (2.7)
5. Digestive system (ICD-8 749–751; ICD-10 Q35–45)	262 (6.8)
6. Genital organs (ICD-8 752; ICD-10 Q50–56)	2 (0.0)
7. Urinary system (ICD-8 753; ICD-10 Q60–64)	234 (6.1)
8. Musculoskeletal system (ICD-8 754–756; ICD-10 Q65–79)	279 (7.2)
9. Other type (ICD-8 757–759; ICD-10 Q80–99)	476 (12.3)
Total	3867 (100)

<sup>a</sup> Exposure prevalence estimated using the total person-years denominator for the 1.45 million live births, i.e. 6.48 million person-years at 0–4 years (Note: registered paternal identity was missing for 55592.9 person-years, 0.9% of the total denominator).

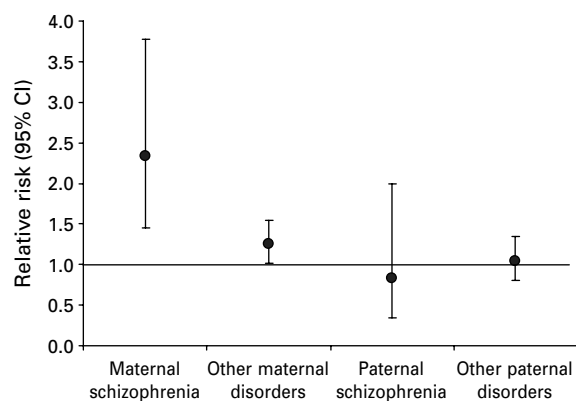
<sup>b</sup> Numerators relate to the whole study population irrespective of exposure status.

### Relative risk of fatal birth defect with maternal admission

These results are shown in Table 3. Risk was elevated with any maternal admission history but the increase in risk was modest (35%). Risk was significantly elevated by around 50% with affective disorders but there was no evidence of higher risk with alcohol/drug-related disorders. The highest risk was associated with schizophrenia [*n* = 17, relative risk (RR) 2.34, 95% confidence interval (CI) 1.45–3.77] whereas the

**Table 2.** Frequency counts for type of fatal birth defect (0–4 years) and by parental psychiatric admission status

Fatal birth defect by type (ICD-8 and ICD-10 code)	Parent not admitted	All disorders	Schizophrenia (and related)	Affective disorder	Alcohol/drug related
<b>Maternal admission status</b>					
1. Nervous system	559	12	1	4	5
2. Eye, ear, face and neck	2	0	0	0	0
3. Circulatory system	1881	55	11	13	7
4. Respiratory system	104	1	1	0	0
5. Digestive system	256	6	0	2	0
6. Genital organs	2	0	0	0	0
7. Urinary system	228	6	0	1	0
8. Musculoskeletal system	272	7	1	2	3
9. Other birth defect	461	15	3	4	1
Total	3765	102	17	26	16
<b>Paternal admission status</b>					
1. Nervous system	461	11	1	2	3
2. Eye, ear, face and neck	2	0	0	0	0
3. Circulatory system	1658	35	2	8	13
4. Respiratory system	81	2	1	0	1
5. Digestive system	201	7	0	1	1
6. Genital organs	2	0	0	0	0
7. Urinary system	178	2	0	1	0
8. Musculoskeletal system	216	1	0	0	0
9. Other birth defect	412	8	1	2	4
Total	3211	66	5	14	22

**Fig. 1.** Relative risk of fatal birth defect at 0–4 years: maternal versus paternal admission with schizophrenia and all other admitted parental disorders.

increase in risk with all maternal diagnoses other than schizophrenia was modest ( $n=85$ , RR 1.25, 95% CI 1.01–1.55); linear comparison between effects for these two diagnostic categories indicated an excess risk specific to schizophrenia ( $z=2.4$ ,  $p=0.018$ ). We assessed whether this key result was specific to our selection of 0–4 years as the follow-up age range. In a sensitivity analysis we estimated relative risks with maternal schizophrenia and with all other admitted maternal disorders using different age ranges: infancy (first year); 0–2 years; 0–4 years; 0–15 years. These

estimates are shown in Table 4. The excess risk with maternal schizophrenia was not specific to 0–4 years; the estimates were consistent across the various age ranges (although in infancy the excess risk was non-significant because of the smaller numbers of observed events). A higher risk with schizophrenia was also seen in both earlier (1973–1989:  $n=12$ , RR 2.69, 95% CI 1.52–4.74) and later (1990–1998:  $n=5$ , RR 1.78, 95% CI 0.74–4.29) study periods, with homogeneity in these effect sizes confirmed by an interaction test ( $\chi^2=0.6$ ,  $df=1$ ,  $p=0.44$ ).

Relative risks of offspring death due to any cause other than birth defect associated with history of maternal admission are also shown in Table 3. For this outcome the effect linked with schizophrenia was statistically comparable to that with all other maternal diagnoses ( $z=-0.6$ ,  $p=0.58$ ). The highest risk of death by any other cause was seen with maternal alcohol/drug-related disorders, which constituted a significant excess risk versus all other admitted maternal illnesses ( $z=6.1$ ,  $p<0.001$ ).

#### **Relative risk of fatal birth defect with paternal admission**

These results are also presented in Table 3. There was no evidence of higher fatal birth defect risk if the father was admitted with schizophrenia or any other

**Table 3.** Relative risk of fatal birth defect and death from any other cause at 0–4 years, by parental diagnostic categories<sup>a,b</sup>

Exposure category	Cause of death	
	Birth defect	Any other cause
Mother not admitted (reference, RR 1.0)		
Rate per 1000 person-years	0.6	1.3
Number of deaths	3765	8070
All maternal psychiatric admissions		
Rate per 1000 person-years	0.7	1.9
Number of deaths	102	301
RR (95% CI)	1.35 (1.11–1.65)	1.85 (1.64–2.07)
Maternal schizophrenia (and related disorder)		
Rate per 1000 person-years	1.1	1.6
Number of deaths	17	26
RR (95% CI)	2.34 (1.45–3.77) <sup>c</sup>	1.66 (1.13–2.44)
Maternal affective disorder		
Rate per 1000 person-years	0.7	2.1
Number of deaths	26	76
RR (95% CI)	1.54 (1.05–2.27)	2.05 (1.64–2.57)
Maternal alcohol/drug-related disorder		
Rate per 1000 person-years	0.6	3.4
Number of deaths	16	97
RR (95% CI)	1.16 (0.71–1.90)	3.24 (2.65–3.96) <sup>d</sup>
Father not admitted (reference, RR 1.0)		
Rate per 1000 person-years	0.5	1.0
Number of deaths	3211	6637
All paternal psychiatric admissions		
Rate per 1000 person-years	0.4	1.3
Number of deaths	66	203
RR (95% CI)	1.02 (0.80–1.30)	1.50 (1.30–1.72)
Paternal schizophrenia (and related disorder)		
Rate per 1000 person-years	0.3	1.2
Number of deaths	5	17
RR (95% CI)	0.83 (0.34–1.99)	1.35 (0.84–2.17)
Paternal affective disorder		
Rate per 1000 person-years	0.5	1.2
Number of deaths	14	33
RR (95% CI)	1.24 (0.74–2.10)	1.38 (0.98–1.94)
Paternal alcohol/drug-related disorder		
Rate per 1000 person-years	0.3	1.2
Number of deaths	22	80
RR (95% CI)	0.83 (0.54–1.26)	1.42 (1.14–1.77)

RR, Relative risk; CI, confidence interval.

<sup>a</sup> All models adjusted for offspring age by time period interaction.

<sup>b</sup> Paternity was unregistered for 15% of fatal birth defects and 18% of deaths from any other cause.

<sup>c</sup> Excess risk of fatal birth defect (*versus* all other maternal diagnoses):  $z = 2.4$ ,  $p = 0.02$ .

<sup>d</sup> Excess risk of death from any other cause (*versus* all other maternal diagnoses):  $z = 6.1$ ,  $p < 0.001$ .

psychiatric disorder. The relative risks with paternal schizophrenia and alcohol/drug-related disorders were below unity, whereas for each maternal

diagnostic category the effects were consistently in a positive direction (whether they reached statistical significance or not). Fig. 1 highlights the marked

**Table 4.** Sensitivity analysis indicating consistency of effect sizes for maternal schizophrenia versus other admitted disorders by different offspring age ranges

Exposure group by different age ranges	<i>n</i>	RR (95% CI)
Infancy (first year)		
No admission	3292	1.00 (–)
All other disorders	77	1.33 (1.06–1.66)
Schizophrenia (and related disorder)	12	1.95 (1.10–3.43)
<i>z</i> test for excess risk		<i>z</i> = 1.2, <i>p</i> = 0.22
0–2 years		
No admission	3664	1.00 (–)
All other disorders	84	1.28 (1.03–1.59)
Schizophrenia (and related disorder)	16	2.29 (1.40–3.74)
<i>z</i> test for excess risk		<i>z</i> = 2.1, <i>p</i> = 0.03
0–4 years		
No admission	3765	1.00 (–)
All other disorders	85	1.25 (1.01–1.55)
Schizophrenia (and related disorder)	17	2.34 (1.45–3.77)
<i>z</i> test for excess risk		<i>z</i> = 2.4, <i>p</i> = 0.02
0–15 years		
No admission	3933	1.00 (–)
All other disorders	91	1.23 (1.00–1.51)
Schizophrenia (and related disorder)	18	2.34 (1.41–3.56)
<i>z</i> test for excess risk		<i>z</i> = 2.3, <i>p</i> = 0.02

RR, Relative risk; CI, confidence interval.

difference in risk of fatal birth defect between maternal versus paternal schizophrenia. The event data were too sparse to enable assessment of effects by number of affected parents as there was just one fatal birth defect if both parents were admitted for schizophrenia. There was evidence of modestly raised risk of offspring death by any other cause with paternal admission history, although these effects were weaker than if the mother had been admitted.

#### *Specific types of fatal birth defect: cardiac and neural tube defects*

We compared risks for cardiac defect (ICD-8 746; ICD-10 Q20–24) with all other types of fatal birth defect. With maternal affective disorder the effect size was the same for cardiac (*n* = 12, RR 1.56, 95% CI 0.88–2.75) as for non-cardiac defects (*n* = 14, RR 1.54, 95% CI 0.91–2.61). The relative risk of fatal cardiac defect with maternal schizophrenia (*n* = 11, RR 3.43, 95% CI 1.90–6.21) was higher than the relative risk of all other types of fatal defect (*n* = 6, RR 1.48, 95% CI 0.66–3.29), although joint analysis of competing risks (Pierce & Preston, 1993) indicated no statistically significant difference between these effect sizes ( $\chi^2 = 2.6$ , 1 df, *p* = 0.11). Nonetheless, compared to the 11 cases of fatal cardiac defect with maternal schizophrenia, there were no cases of this type with paternal schizophrenia (Fisher's exact test: *p* = 0.026).

Fatal neural tube defects were identified according to the three most common specific types: anencephaly (ICD-8 740, ICD-10 Q00); spina bifida (ICD-8 741, ICD-10 Q05); and encephalocele (ICD-8 743.0, ICD-10 Q01). These events were extremely sparse; across the whole 26-year study period only four cases were seen among offspring of mothers admitted with any diagnosis, and three cases among those with admitted fathers. The only parental diagnostic groups to contain any such events were maternal affective disorder (two cases) and maternal alcohol/drug-related disorder (two cases).

#### Discussion

To our knowledge, this is the first population-based investigation to describe the risks of fatal birth defect associated with maternal and paternal psychiatric illness. Among a range of maternal conditions resulting in admission prior to offspring death, the highest risk was linked to schizophrenia and related disorders. For this diagnostic category there was an excess risk compared to all other groups of admitted mothers. Our results indicate that the excess risk with maternal schizophrenia could be due to cardiac defects. That no fatal cardiac defects were found with paternal schizophrenia strengthens this notion, but this finding requires replication in other large cohorts. The excess



risk with maternal schizophrenia was confined to birth defect inasmuch as the risk of death from any other cause was similar to that seen with other maternal psychiatric diagnoses.

An elevated risk of fatal birth defect was also seen with maternal affective disorder. The likely external validity of this observed association is, however, limited by the fact that most women in this diagnostic range would not have been admitted during the study period. Thresholds for admission with affective disorder many also vary greatly between countries. Maternal admission with alcohol/drug-related disorders conferred a large excess risk of death by any cause other than birth defect; further detailed analyses of this cohort has shown that this excess risk was largely explained by cases of sudden infant death syndrome in the first year of life (King-Hele *et al.* 2007). We found no evidence of higher risk of fatal birth defect if the father was admitted with schizophrenia or any other diagnostic category.

Our study was conducted using a large national birth cohort. The Danish Civil Registration System enables record linkage between all mothers (but not all fathers) and their offspring, and complete ascertainment of all cases of fatal birth defect, and all other causes of death, nationally (Pedersen *et al.* 2006). We could also appropriately censor the data to account for loss to follow-up due to emigration. Pooling the estimates across ages 0–4 years increased statistical power and precision.

However, the study had several important weaknesses. First, 15% of fathers were unregistered among cases of fatal birth defect because of imperfections in the registration system. Thus, our relative risk estimates with maternal admission are more robust than those with paternal admission. The degree of any resulting bias is difficult to assess, although unknown fathers with admission histories are most likely to have attenuated the effect estimates. Furthermore, we could not compare effects of maternal *versus* paternal admission on risk of stillborn fatal birth defect, again because of missing paternity data; however, in line with our results from the live-born cohort, higher risks for stillborn birth defect with maternal schizophrenia and affective disorder were indicated. Second, the rarity of fatal birth defect meant that we lacked sufficient event data and statistical power to examine effects linked with specific parental diagnoses, such as tightly defined schizophrenia or bipolar disorder. Small numbers also precluded assessment of exposure by timing of admission across diagnostic categories. Third, parental diagnoses recorded in the register were made during routine clinical practice rather than by standardized research tools. Fourth, as with any national mortality registration system, the classi-

fication of cause of death was not entirely accurate (Juel & Helweg-Larsen, 1999).

From our data we can only speculate about the likely causal mechanisms involved. We might consider possible genetic factors linked directly with maternal disorder. Higher rates of minor physical anomaly have been consistently reported in people with schizophrenia and their close relatives (Ismail *et al.* 1998). Although the topic has not been widely studied, genetic mechanisms linking parental schizophrenia and fatal birth defect in offspring cannot be excluded. These may be complex, involving gene–environment interaction or pleiotropic effects, whereby a range of pathological outcomes (other than schizophrenia itself) could be linked with maternal disorder (Moldin, 1994; Hallmayer, 2000). We observed a markedly elevated risk with maternal schizophrenia (RR 2.34), but found no evidence of higher risk with paternal schizophrenia (RR 0.83), which points towards a greater environmental than genetic aetiological contribution. It could be that failure to find a link between paternal schizophrenia and risk of fatal birth defect occurred because of bias by unregistered fatherhood. However, a trend towards higher risk of death from any other cause was seen with paternal schizophrenia, which perhaps suggests otherwise.

Environmental determinants with higher prevalence in mentally ill women could include health-care factors such as lower antenatal attendance and poorer standards of care (Kelly *et al.* 1999), and lifestyle factors such as diet, smoking and substance misuse. These factors may apply to women diagnosed with schizophrenia or affective disorder. Elevated risk of birth defect has been linked with both malnutrition (Lam & Torfs, 2006) and obesity (Scialli, 2006). For example, a body mass index of 30 or greater confers a doubling in risk of neural tube defects. Honein *et al.* (2001) reported modest effect sizes linking maternal smoking with non-fatal defects such as gastroschisis, oral clefts and clubfoot, although Tuthill *et al.* (1999) observed no higher risk of fatal birth defect with this exposure. We were unable to assess the effects of antenatal smoking as this was not routinely recorded in the registers.

Particularly high risks have been shown with heavy alcohol consumption in pregnancy. Habbick *et al.* (1997) reported an elevated mortality risk in cases of foetal alcohol syndrome, with nine out of 12 observed deaths being due to birth defects. We found no association between maternal admission for alcohol/drug-related disorders and higher risk of fatal birth defect, although actual levels of substance misuse in pregnancy were not recorded. Furthermore, rates of diagnosed substance misuse disorders among Danish psychiatric in-patients are likely to be underestimated

by around a half (Hansen *et al.* 2000). Without detailed measurements of exposure, we cannot rule out the possibility of causal links with maternal substance misuse during pregnancy. Furthermore, specific anomalies such as neural tube defect that have been studied in relation to parental substance misuse (Shaw *et al.* 1996) may be too rare for investigation, even with the large sample size that was available to us.

Foetal exposure to prescribed psychotropic drugs may also be important but ethical constraints have precluded the conduct of randomized evaluations of psychotropic medication in pregnancy (Howard *et al.* 2004), and there are few large population-based observational studies that have measured these exposures directly. Munk *et al.* (2005) conducted a large cohort study across four Danish counties (a combined population of 1.4 million people) by linking prescription registry data to the national Medical Birth Register. With just four exposed cases, they reported a suggestion of elevated risk of major birth defect with foetal exposure to antipsychotic medication [odds ratio (OR) 4.1, 95% CI 0.5–34.4]. However, this estimate was imprecise and far from being statistically significant. Observational data from sources other than population-based registers may be prone to selection bias, reporting bias and low statistical power. For example, McKenna *et al.* (2005) reported finding no apparent link between *in utero* exposure to higher potency atypical antipsychotic drugs and risk of major birth defect. However, their sample of 151 exposed pregnant women (and 151 unexposed controls) was based on maternal self-reported medication usage; they observed only one exposed and two unexposed cases. To have adequate power to detect a twofold higher risk, a sample size of several thousand subjects in the exposure group would be required.

Despite the lack of direct evidence, the possibility that psychotropic medication may be linked to fatal teratogenic risk should be of great interest to clinicians. Use of atypical antipsychotics, which are less likely to cause hyperprolactinaemia-related side-effects such as reduced fertility, is increasing (McKenna *et al.* 2004). This implies that more women with schizophrenia and other psychotic illnesses will become pregnant. Our finding that relative risks of fatal birth defect were homogeneous between earlier (1973–1989) and later (1990–1998) time periods, in relation to maternal schizophrenia, is difficult to interpret without greater understanding of the causal mechanisms. Although risk of fatal birth defect in our study was lower with maternal affective disorders than with schizophrenia, it was still 50% higher than the general population rate. There is some evidence indicating links between *in utero* exposure to selective serotonin reuptake inhibitors (SSRI)

antidepressants and risk of major birth defect (e.g. cardiac malformation with paroxetine use; Bar-Oz *et al.* 2007) and other perinatal complications (e.g. persistent pulmonary hypertension of the newborn with SSRI use after the 20th gestational week; Chambers *et al.* 2006). These effects may require further investigation using large population-based samples.

Our study has highlighted important links between maternal schizophrenia and affective disorder and higher risk of fatal birth defect. As we failed to observe elevated risk with paternal admission for these disorders, our findings may indicate a greater environmental than genetic aetiology. This result requires replication in national registry studies from other countries, ideally without missing registered fatherhood data, although such a large and comprehensive national cohort with complete paternity data may not currently exist in any country. Future work should focus on generating high-quality observational cohorts, including detailed prospective measurement of medication usage during pregnancy, to enable precise determination of the causal mechanisms involved.

#### Declaration of Interest

None.

#### Acknowledgements

We are grateful to Carsten B. Pedersen, Thomas M. Laursen and Heine Gøtzsche, National Centre for Register-based Research (Aarhus, Denmark), for extracting and linking the data and providing statistical advice. The study was funded by the Wellcome Trust, London, UK (grant ref. 073935) and by the Stanley Medical Research Institute, Chevy Chase, MD, USA.

#### References

- Bar-Oz B, Einarson T, Einarson A, Boskovic R, O'Brien L, Malm H, Bérard A, Gideon K (2007). Paroxetine and congenital malformations: meta-analysis and consideration of potential confounding factors. *Clinical Therapeutics* **29**, 918–926.
- Bennedsen BE, Mortensen PB, Olesen AV, Henriksen TB (2001). Congenital malformations, stillbirths, and infant deaths among children of women with schizophrenia. *Archives of General Psychiatry* **58**, 674–679.
- Chambers CD, Hernandez-Diaz S, Van Marter LJ, Werler MM, Louik C, Jones KL, Mitchell AA (2006). Selective serotonin-reuptake inhibitors and risk of persistent pulmonary hypertension of the newborn. *New England Journal of Medicine* **354**, 579–587.
- Gardner W, Mulvey EP, Shaw EC (1995). Regression analyses of counts and rates: Poisson, overdispersed Poisson, and negative binomial models. *Psychological Bulletin* **118**, 392–404.



- Habbick BF, Nanson JL, Snyder RE, Casey RE** (1997). Mortality in foetal alcohol syndrome. *Canadian Journal of Public Health* **88**, 181–183.
- Hallmayer J** (2000). The epidemiology of the genetic liability for schizophrenia. *Australian and New Zealand Journal of Psychiatry* **34** (Suppl.), S47–S55.
- Hansen SS, Munk-Jørgensen P, Gulbæk B, Solgård T, Lauszus KS, Albrechtsen N, Borg L, Egander A, Faurholdt K, Gilberg A, Godsen NP, Lorenzen J, Richelsen B, Weischer K, Bertelsen A** (2000). Psychoactive substance use among psychiatric in-patients. *Acta Psychiatrica Scandinavica* **102**, 432–438.
- Honein MA, Paulozzi LJ, Watkins ML** (2001). Maternal smoking and birth defects: validity of birth certificate data for effect estimation. *Public Health Reports* **116**, 327–335.
- Howard L, Webb R, Abel K** (2004). Safety of antipsychotic drugs for pregnant and breastfeeding women with non-affective psychosis. *British Medical Journal* **329**, 933–934.
- Ismail B, Cantor-Graae E, McNeil TF** (1998). Minor physical anomalies in schizophrenic patients and their siblings. *American Journal of Psychiatry* **155**, 1695–1702.
- Juel K, Helweg-Larsen K** (1999). The Danish registers of causes of death. *Danish Medical Bulletin* **46**, 354–357.
- Kallen B, Tanberg A** (1983). Lithium and pregnancy. A cohort study on manic-depressive women. *Acta Psychiatrica Scandinavica* **68**, 134–139.
- Kelly RH, Danielson BH, Golding JM, Anders TF, Gilbert WM, Zatzick DF** (1999). Adequacy of prenatal care among women with psychiatric diagnoses giving birth in California in 1994 and 1995. *Psychiatric Services* **50**, 1584–1590.
- Kessing LV** (1998). A comparison of ICD-8 and ICD-10 diagnoses of affective disorder – a case register study from Denmark. *European Psychiatry* **13**, 342–345.
- King-Hele SA, Abel KM, Webb RT, Mortensen PB, Appleby L, Pickles AR** (2007). Risk of sudden infant death syndrome with parental mental illness. *Archives of General Psychiatry* **64**, 1323–1330.
- Lam PK, Torfs CP** (2006). Interaction between maternal smoking and malnutrition in infant risk of gastroschisis. *Birth Defects Research* **76**, 182–186.
- Lloyd T, Doody G, Brewin J, Park B, Jones P** (2003). Minor physical anomalies in schizophrenia: is age a confounding factor? *Schizophrenia Research* **61**, 67–73.
- McKenna K, Einarson A, Levinson A, Koren G** (2004). Significant changes in antipsychotic drug use during pregnancy. *Veterinary and Human Toxicology* **46**, 44–46.
- McKenna K, Koren G, Tetelbaum M, Wilton L, Shakir S, Diav-Citrin O, Levinson A, Zipursky RB, Einarson A** (2005). Pregnancy outcome of women using atypical antipsychotic drugs: a prospective comparative study. *Journal of Clinical Psychiatry* **66**, 444–449.
- McNeil TF, Cantor-Graae E, Ismail B** (2000). Obstetric complications and congenital malformation in schizophrenia. *Brain Research Reviews* **31**, 166–178.
- Moldin SO** (1994). Indicators of liability to schizophrenia: perspectives from genetic epidemiology. *Schizophrenia Bulletin* **20**, 169–184.
- Munk EM, Norgaard B, Gislum M, Mortensen PB, Sorensen HT** (2005). Use of antipsychotic drugs during pregnancy and the risk of adverse birth outcomes: a population-based cohort study. *Schizophrenia Bulletin* **31**, 233.
- Munk-Jørgensen P** (1995). Decreasing first admission rates for schizophrenia in Denmark 1970–1991 [in Danish]. Thesis, Department of Psychiatric Demography, University of Copenhagen.
- Munk-Jørgensen P, Mortensen PB** (1997). The Danish Psychiatric Central Register. *Danish Medical Bulletin* **44**, 82–84.
- Pedersen CB, Gøtzsche H, Moller JO, Mortensen PB** (2006). The Danish Civil Registration System. A cohort of eight million persons. *Danish Medical Bulletin* **53**, 441–449.
- Pierce DA, Preston DL** (1993). Joint analysis of site-specific cancer risks for the atomic bomb survivors. *Radiation Research* **134**, 134–142.
- Scialli AR** (2006). Teratology public affairs committee position paper: maternal obesity and pregnancy. *Birth Defects Research* **76**, 73–77.
- Shaw GM, Velie EM, Morland KB** (1996). Parental recreational drug use and risk for neural tube defects. *American Journal of Epidemiology* **144**, 1155–1160.
- Strudsholm U, Johannessen L, Foldager L, Munk-Jørgensen P** (2005). Increased risk for pulmonary embolism in patients with bipolar disorder. *Bipolar Disorders* **7**, 77–81.
- Tuthill DP, Stewart JH, Coles EC, Andrews J, Carlidge PH** (1999). Maternal cigarette smoking and pregnancy outcome. *Paediatric and Perinatal Epidemiology* **13**, 245–253.
- Webb R, Abel K, Pickles A, Appleby L** (2005). Mortality in offspring of parents with psychotic disorders: a critical review and meta-analysis. *American Journal of Psychiatry* **162**, 1045–1056.
- Webb RT, Abel KM, Pickles AR, Appleby L, King-Hele SA, Mortensen PB** (2006). Mortality risk among offspring of psychiatric inpatients: a population-based follow-up to early adulthood. *American Journal of Psychiatry* **163**, 2170–2177.
- Webb RT, Pickles AR, Appleby L, Mortensen PB, Abel KM** (2007). Death by unnatural causes during childhood and early adulthood in offspring of psychiatric inpatients. *Archives of General Psychiatry* **64**, 345–352.
- Weinberg SM, Jenkins EA, Marazita ML, Maher BS** (2007). Minor physical anomalies in schizophrenia: a meta-analysis. *Schizophrenia Research* **89**, 72–85.
- WHO** (1967). *Manual of the International Classification of Diseases (ICD-8)*. World Health Organization: Geneva.
- WHO** (1992). *The ICD-10 Classification of Mental and Behavioural Disorders*. World Health Organization: Geneva.