The Everyday Compensation (EComp) Questionnaire: Construct Validity and Associations with Diagnosis and Longitudinal Change in Cognition and Everyday Function in Older Adults

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Abstract

Objective: The Everyday Compensation scale (EComp) is an informant-rated questionnaire designed to measure cognitively based compensatory strategies that support both everyday memory and executive function in the context of completing instrumental activities of daily living (IADLs). Although previous findings provided early support for the usefulness of the initial version of EComp, the current paper further describes the development, refinement, and validation of EComp as a new assessment tool of compensation for IADLs. Method: Confirmatory factor analysis (CFA) was used to examine its factor structure. Convergent and predictive validity was evaluated by examining the relationship between EComp and markers of disease, including diagnosis, cognitive change, and trajectories of functional abilities. Results: CFA supported a general compensation factor after accounting for variance attributable to IADL domain-specific engagement. The clinical groups differed in compensatory strategy use, with those with dementia using significantly fewer compensatory strategies as compared to individuals with normal cognition or mild cognitive impairment. Greater levels of compensation were related to better cognitive functions (memory and executive function) and functional abilities, as well as slower rates of cognitive and functional decline over time. Importantly, higher levels of compensation were associated with less functional difficulties and subsequently slower rate of functional decline independent of the level of cognitive impairment. Conclusions: Engagement in compensatory strategies among older adults has important implications for prolonging functional independence, even in those with declining cognitive functioning. Results suggest that the revised EComp is likely to be useful in measuring cognitively based compensation in older adults.

Keywords: Activities of daily living, Dementia, Mild cognitive impairment, Executive function, Factor analysis, Statistical, Surveys and questionnaires

INTRODUCTION

The older adult population in the United States is growing rapidly, and the prevalence of disorders of aging, including Alzheimer's disease (AD) and related disorders, is also increasing (Alzheimer's Association, 2018). A hallmark of AD is gradual and progressive loss of independence in the ability to perform various instrumental activities of daily living (IADLs) such as driving, cooking, and managing appointments. Additionally, subtle changes in functional abilities are evident among individuals with prodromal AD, including mild cognitive impairment (MCI; Marshall et al., 2001; Schmitter-Edgecombe, McAlister, & Weakley, 2012; Schmitter-Edgecombe & Parsey, 2014) and "pre-MCI" (Lau, Parikh, Harvey, Huang, & Farias, 2015; McAlister & Schmitter-Edgecombe, 2016; Tomaszewski Farias et al., 2013; Tomaszewski Farias et al., 2017). Understanding the factors that contribute to the rate of decline in functional abilities is important, particularly to inform the development of interventions to support and prolong independence.

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Cognitive impairment clearly plays a major role in the development of functional impairment and disability (Freedman, Martin, & Schoeni, 2002). In particular, greater deficits in executive functioning and episodic memory have been most strongly and consistently associated with worse everyday functioning among older adults (Pereira, Yassuda, Oliveira, & Forlenza, 2008; Tomaszewski Farias et al., 2009). However, cognitive test performance typically only accounts for about a third or less of the variance in functional abilities (see McAlister, Schmitter-Edgecombe, & Lamb, 2016; Royall et al., 2007), and some individuals with substantial cognitive impairment maintain a good level of independence in IADLs, while others with less cognitive impairment are more functionally disabled. Thus, it is clear that there are other non-cognitive factors that contribute to the maintenance of functional abilities. The degree to which one utilizes compensatory strategies in their everyday life is likely to promote a higher level of functioning.

Although there is a rich rehabilitation literature describing various compensation strategies that help individuals accommodate for cognitive and functional deficits (Ciceron et al. 2005), a lack of reliable tools to measure such strategies has hindered the study of compensation. Two memory-based compensation measures have developed with older adults in mind: the Multifactorial Memory Questionnaire, Memory Strategies Scale (MMQ-MSS; Troyer & Rich, 2002) and the Memory Compensation Questionnaire (MCQ; de Frias & Dixon, 2005). Although helpful, MMQ-MSS and MCQ have a number of limitations. First, both were designed as self-report measures. Informant ratings, however, can be valuable particularly when an individual is starting to have cognitive difficulties that reduce ones' insight or awareness. MMQ-MSS and MCQ include strategies that rely on the recruitment of others for assistance, which may not represent a compensation strategy as much as a loss of independence. Further, they focus solely on memory compensation rather than a broad array of compensation strategies (including those that support various executive functions). Finally, MMQ-MSS and MCQ do not examine compensation explicitly within the context of completing IADLs (e.g., transportation, cooking).

With these limitations in mind, we previously reported on a new assessment tool designed to measure a broad range of compensatory strategies used within the context of competing IADLs (the Everyday Compensation scale (EComp); Tomaszewski Farias et al., 2018). Using this initial version of EComp, we showed that greater strategy use was associated with a higher level of functional independence, and this effect was independent of the degree of cognitive impairment present. That is, regardless of the cognitive status, the more individuals utilized compensation strategies, the better their level of everyday function. These initial findings provide early support for the usefulness of EComp.

The primary purpose of the current paper was to describe the further refinement and validation of EComp. The original/ pilot version of EComp (Tomaszewski Farias et al., 2018) was further refined in order to reduce the number of questions

thereby optimizing (a) the efficiency (reduce the time commitment on the part of the responder) and (b) the psychometric properties of the test (deleting items with poor variance or content validity). Confirmatory factor analysis (CFA) was used to examine the factor structure of EComp and whether compensation is best represented as a unidimensional or multidimensional construct. Based on previous work (de Frias & Dixon, 2005), we anticipated finding evidence of a multidimensional structure including compensationspecific categories (e.g., external strategies, taking more time to complete tasks). We also expected that older adults would vary in terms of their routine engagement in IADLs due to physical limitations, division of labor (e.g., minimal involvement in finances), etc. Thus, we also examined the variance attributable to individual IADL domains. We next examined the relationship between compensatory strategy use on EComp and diagnosis (normal cognition, MCI, dementia) and how compensation relates to change in cognition and functional abilities over time (convergent and predictive validity). We hypothesized that greater compensation would be present among individuals with normal cognition or MCI as compared to dementia, and that greater compensation would be associated with a slower rate of decline in cognitive and functional abilities.

METHOD

Instrument Development

The goals of developing EComp were two-fold. First, we sought to comprehensively assess cognitive and behaviorally based compensation strategies that support both everyday memory and everyday executive abilities (i.e., strategies that rely on enhanced organization, planning, prioritization, self-monitoring, and use of routines). A second goal was to measure compensation strategies used explicitly in the context of completing common IADLs critical to the maintenance of functional independence. As previously described (Tomaszewski Farias et al., 2018), an initial pool of possible items was developed first by surveying existing measures and reviewing relevant literature. Additionally, experts (authors S.T.F., J.G., A.W., M.S.W., T.G. and an independent rater who was not an author) in cognitive aging and rehabilitation generated items thought to support memory and executive functioning and that represent compensation-specific categories identified in the literature: external strategies (e.g., calendars), internal strategies (e.g., mnemonics), environmental strategies (e.g., visual cues), organizational strategies (e.g., additional structure), planning strategies (e.g., approaching a problem ahead of time), routinization strategies (e.g., habit formation), simplification strategies (e.g., reducing task complexity), reducing distractions (e.g., avoiding interfering information), and increasing time/effort (e.g., allocating more time to tasks). A total of 70 items were initially developed and covered six IADL-specific domains: appointments, shopping, cooking, finances, transportation, and medication. Each item measured how often the individual utilized a given strategy in their daily life using a four-point scale: 0 = never, 1 = rarely, 2 = sometimes, 3 = frequently, 4 = always. A response choice indicating that the rater "could not say/strategy was not applicable" was also included.

Further refinement of the initial 70 items included multiple steps. First, experts independently rated items based on importance/utility to the maintenance of IADL independence (content validity) on a three-point scale (1 = low, 2 = moderate,3 = high utility). They also categorized what type of compensation they believed each item reflected (e.g., external, internal, environmental, organizational, etc.). Descriptive statistics on these ratings were compiled and the experts collectively reviewed and discussed items with poor ratings. Items were deleted if they were deemed to have low utility or had poor agreement in terms of type of compensation to ensure content validity of the scale. Items were further deleted or revised if they showed poor response variability; specifically, items were deleted if there as a high frequency (>25%) of "could not say/strategy was not applicable" responses. Disagreements were settled through consensus. These processes retained 41 items and this version of EComp was used in the current validation study. Of those that were removed, nine were removed due to poor response variability, six were removed because they were internal strategies that were likely to not be reliably rated by informants, nine were removed because they could not be identified as a specific compensation type (e.g., avoid left-hand turns, shop at same store), and five were removed secondary to low utility (i.e., cook same meals each week, shop frequently). A table of all deleted items can be found in online supplementary materials. In order to have the items well distributed across the types of compensation, based on theoretical grounds and expert consensus, the items were collapsed into four compensation-specific categories for subsequent factor analysis: External/Environmental, Organizational/Planning, Routine/Simplification/Reducing Distraction, and Increased Time/Effort.

Participants

Participants were part of a longitudinal research cohort at the University of California, Davis Alzheimer's Disease Center (UCD ADC) previously described (e.g., Early et al., 2013; Hinton et al., 2010; Mungas et al., 2010). All participants were living within the community (not in a nursing home or other institution). Exclusion criteria included unstable major medical illness, severe/debilitating psychiatric disorder, or another existing neurologic condition (e.g., idiopathic Parkinson's disease). Participants received annual clinical evaluations that included neurological exam, neuroimaging, lab work, and neuropsychological testing. Annual diagnosis was categorized as normal cognition, MCI, or dementia according to standard clinical criteria (Morris et al., 2006). Neuropsychological and everyday functional tests used to make clinical diagnoses were separate from the neuropsychological tests used as variables in the current study.

All participants provided written informed consent, and involvement was approved by institutional review boards at the University of California, Davis.

Procedure

EComp was administered to an informant (e.g., spouse, adult child) during participants' annual visit. Cognitive and functional measures described below are routinely administered as part of each annual visit.

Instruments

Cognitive function

Neuropsychological functioning was assessed using the Spanish and English Neuropsychological Assessment Scales battery (SENAS; Mungas, Reed, Crane, Haan, & González., 2004; Mungas, Reed, Haan, & González, 2005). This study used two composite indices from SENAS: episodic memory and executive function. The Episodic Memory Index is derived from a multi-trial word list-learning test. The Executive Function Index is constructed from component tasks of verbal fluency and working memory. These measures do not have appreciable floor or ceiling effects for participants in this sample and have linear measurement properties across a broad ability range (Mungas et al., 2004).

Everyday function

Everyday function was assessed using two instruments to capture the full range and breadth of functional abilities. The Everyday Cognition (ECog) scale (Tomaszewski Farias et al., 2008) is a 39-item informant-based questionnaire designed specifically to be sensitive to mild functional limitations that predate the loss of independence and has been shown to be relevant to functional changes associated with MCI (Lau et al., 2015; Rog et al., 2014; Tomaszewski Farias et al., 2017). It has been shown to have good content, convergent and discriminant, and external validity (Tomaszewski Farias et al., 2008). On each item of the ECog scale, informants were asked to assess the participant's current level of everyday functioning in comparison to their own baseline. Each item on ECog is rated on a four-point scale (1 = better or no change, to 4 = consistently muchworse). The total score was calculated by summing all completed items and dividing by the number of items completed to get an average (this allows for missing items, but at least half needed to be completed to calculate a score), with higher scores reflecting worse everyday cognition.

Everyday function was also assessed using the Instrumental Activities of Daily Living Scale (Lawton & Brody, 1969), which is a widely used informant-based measure used to rate participants' abilities across eight IADLs (i.e., telephone use, shopping, cooking, housework, laundry, transportation, medication management, and finances). Each item was rated as follows: 0 = can complete the task independently, 1 = the task requires some assistance, 2 = the task must now be completed by someone else. Level of IADL independence was measured via a summary total score (e.g., the average rating across the items that were rated; if fewer than half of the items were rated, the total score was missing). Higher scores reflect greater dependence on caregivers.

Analyses

CFA was used to test whether the EComp factor structure was best represented as a unidimensional or multidimensional model. Items varied along two major dimensions, IADLspecific and compensation-specific domains. We tested three a priori models to explain covariance among items. We first used a single general factor to account for intercorrelations among all items. This model essentially tested for a strong general compensation dimension that did not vary by IADL-specific or compensation-specific domains. The second model was a bifactor model (Holzinger & Swineford, 1937; Jennrich & Bentler, 2011) that had a general compensation type factor but added six IADL-specific factors (i.e., Appointments, Cooking, Medications, Transportation, Shopping, Finances) to account for residual intercorrelation due to within-IADL domain similarities of items. The general factor was uncorrelated with IADL-specific factors in this model, and IADL factors represent specific methods factors. In the third model, we examined whether adding four compensation-specific factors (External/Environmental, Organizational/Planning, Routine/Simplification/Reducing Distraction, and Increased time) to the IADL-specific factors improved model fit over the bifactor model with a single general factor and six IADL-specific factors. In sum, competing models were as follows: (1) a single general compensation factor model; (2) seven-factor model including a single general compensation factor and six IADL-specific factors; (3) 10-factor model including four compensationspecific factors and six IADL-specific factors.

The Mplus application (Muthen & Muthen, 2007) was used, and EComp items were modeled as categorical indicators of latent factors. This approach assumes that there is a latent continuous variable underlying each categorical variable, with the categories defined by threshold or cut-off values related to the underlying continuous variable. Latent continuous variables are assumed to have a multivariate normal distribution. Thresholds are estimated along factor loadings, which, like traditional loadings for continuous variables, relate the continuous variable underlying the categorical variables to latent factors. There is no single accepted criterion index to judge model fit, so we report several goodness-of-fit indices identified by Hu and Bentler (1998) for continuous indicators and by Yu (2002) for categorical indicators. These indices included the comparative fit index (CFI; Bentler, 1990), the root mean square error of approximation (RMSEA; Cudeck & Browne, 1983), and the Tucker-Lewis

index (TLI; Tucker & Lewis, 1973). CFI and TLI measure the fit of the model relative to the null model. CFI incorporates a correction for model complexity, and TLI takes the degrees of freedom into account. RMSEA takes model parsimony into account, which is important because goodness-of-fit values can sometimes be artificially inflated as the number of parameters in the model is increased. Guidelines for the interpretation of these indices are similar for analyses involving continuous and categorical indicators. TLI and CFI range from 0 (poor fit) to 1 (perfect fit); values \geq .95 are indicative of a good model fit. RMSEA values .06–.08 are considered to reflect adequate fit, and values >.08 indicate good fit. A mean and variance-adjusted weighted least squares estimator (WLSMV; Muthen, 1993) was used for all analyses.

Models were estimated in the following order. Loadings of all items on a single, common, primary dimension were freely estimated with the variance of the latent dimension constrained to unity. Then, IADL-specific factors were added to test the relative ability of a secondary factor structure to account for residual intercorrelation of items. IADL-specific factors were constrained to be uncorrelated with the primary general factor and to have variances of 1.0, but intercorrelations of secondary, domain-specific factors were freely estimated. Lastly, to test the hypotheses of specific overarching types of compensation, a 10-bifactor model (four compensatory strategies by six IADL-specific) was fit and compared to the seven-bifactor (one general compensation by six IADLspecific) model. Competing secondary factor models were not nested, but relative fit was evaluated using fit indices and residual correlations. Factor scores were generated from the model that was chosen as providing the best fit, and these factor scores were then used as variables in subsequent analyses; individuals missing more than half of the individual EComp items were excluded from these analyses (n = 33).

Simple bivariate correlation coefficients were used to characterize the strength of association of EComp with demographic variables (age, education, gender). Linear regression was used to assess differences in compensation between diagnostic groups after adjusting for age and education. To test the hypotheses about how compensation use was associated with the level and change in episodic memory or executive function, mixed effects models were used in which time was centered at the time of EComp assessment. Annual research visits and associated data points prior to that assessment were used (with negative time) to get a better estimate of slope. These models included person-specific random intercepts and slopes to account for between-person variability in level and change not accounted for by compensation use or demographics (age, education, gender). Similar models were used to assess associations between compensation and average IADL score. Model assumptions for both linear regression and mixed effects models were checked and met by the data. To evaluate the association between compensation and total ECog, we rescaled ECog scores to fall between 0 (no impairment reported) and 1 (worst impairment possible) (see Tomaszewski Farias et al., 2013 for more details). This ECog impairment index represents the proportion of

	Cognitively normal $(n = 179)$	MCI $(n = 66)$	Dementia $(n = 60)$		
Age	78.0 (7.0)	80.8 (6.4)	82.1 (7.6)		
Education	14.9 (3.5)	15.3 (3.3)	14.8 (3.5)		
Female, n (%)	117 (65.4)	40 (60.6)	27 (45.0)		
Ethnicity, n (%)					
African American	36 (20.1)	20 (30.3)	8 (13.3)		
Caucasian	93 (52.0)	38 (57.6)	42 (70.0)		
Hispanic	39 (21.8)	5 (7.6)	6 (10.0)		
Other	11 (6.1)	3 (4.5)	4 (6.7)		
Neuropsychological composites					
Episodic memory	0.6 (0.8)	-0.6(0.8)	-1.5(0.8)		
Executive function	0.2 (0.6)	-0.3(0.6)	-1.0(0.7)		
Everyday functional limitations					
ECog total score	1.4 (0.5)	1.8 (0.7)	2.9 (0.9)		
IADL total score	0.1 (0.4)	0.4 (0.4)	1.3 (0.6)		
EComp total compensation	2.0 (0.7)	2.2 (0.7)	1.2 (1.2)		

Table 1. Participant characteristics

Note: Means and standard deviations presented unless otherwise stated. Neuropsychological scores are expressed as *z*-scores with a mean of 0 and standard deviation of 1. Episodic memory is available for 170 cognitively normal, 59 MCI, and 34 dementia. Executive function is available for 176 cognitively normal, 65 MCI, and 41 dementia. ECog = Everyday Cognition (Tomaszewski Farias et al., 2008); higher ECog represents worse everyday cognition. EComp = Everyday Compensation (Tomaszewski Farias et al., 2018). The EComp score is the average score across all items, and a higher score represents more compensation.

Table 2. Relative fit of indices for the models tested

Model	Chi-squared (df)	CFI	TLI	RMSEA	RMSEA 90% CI	
One-factor model (one general compensation)	2463.23 (779)*	0.87	0.87	0.08	0.079–0.086	
Seven-factor model (one general compensation and six IADL-specific)	1303.41 (723)*	0.96	0.95	0.05	0.046-0.055	
10-factor model (six IADL-specific and four compensation-specific)	1284.46 (717)*	0.96	0.95	0.05	0.046-0.054	

*Values significant at p < .05.

CFI = comparative fit index, df = degree of freedom, RMSEA = root mean square error of approximation, TLI = Tucker-Lewis Index. TLI and CFI range from 0 (poor fit) to 1 (perfect fit); values \geq .95 are indicative of a good model fit. RMSEA values .06–.08 are considered to reflect adequate fit; values >.08 indicate good fit.

impairment reported and was analyzed using generalized linear mixed models with a logistic link. As for the linear mixed effects models, these models centered time at the time of EComp assessment and included person-specific random effects. Positive coefficients for predictors indicate an increase in the log "odds," which implies an increase in impairment proportion (and overall worse functioning). All analyses were conducted using SAS software, version 9.4 and a *p*-value <.05 was considered statistically significant.

RESULTS

Sample Characteristics

Participants were 305 older adults with an average age of 79.4 (standard deviation (SD) = 7.2), and had an average of 15.0 years (SD = 3.4) of education; although average education was fairly high, it ranged from 0 to 20 years within the sample with 27.9% having less than a high school education. Table 1

presents baseline demographic information by diagnostic group (cognitively normal, MCI, dementia) as well as average EComp, neuropsychological, and ECog scores. The majority of informants were either a spouse (48.0%) or adult child (32.8%). Informants spent 79.5 (SD = 68.4) hours per week on average with each participant.

Factor Structure

Table 2 presents the fit indices for the three models evaluated. All three models yielded a significant chi-square statistic, indicating a less-than-perfect fit by this metric. However, the chi-square statistic has a tendency to overstate significance in large samples (N > 200) suggestive of a poor model fit (Bollen, 1989). For this reason, the model fit was primarily evaluated using fit indices and residual correlations. The one-factor model (one general compensation factor) did not fit well. The seven-factor model that



Fig. 1. Diagram showing significant loadings and factor correlations. App = appointments; Trans = transportation; specific questions for each factor loading are presented in Table 3.

included a general compensation factor and six IADLspecific factors and the 10-factor model that included four compensatory-specific and six IADL-specific factors both fit well. The 7- and 10-factor models were near identical in TLI, CFI, and RMSEA fit indices; so consequently, the seven-factor model (one compensation and six IADL-specific) was selected as the best model for parsimony. Additional review of the 10-factor model showed very high intercorrelations among the proposed compensatory-specific strategies (r's > .90), suggesting these factors were not very distinct from one another, further supporting the seven-factor model.

Figure 1 presents a diagrammatic representation of the final model. Table 3 shows standardized factor loadings for the final model that include the general factor plus the six IADL-specific factors (seven-factor model). Standardized loadings of individual items on the general compensation factor ranged from 0.30 to 0.84 (average = 0.63), and therefore accounted for substantial variance in all items (9–71%). The majority of items showed standardized loadings on IADL factors that exceeded .30, a generally accepted threshold for a salient loading (McDonald, 1999). The subsequent analyses (below) focus on the general compensation factor (which is not specific to any given IADL domain) and how it relates to demographic factors, cognition, and everyday function.

Demographic Correlates

The EComp general factor score was not associated with age (r = -0.1, p = .9) or education (r = -0.02, p = .7), but females tended to use more compensation strategies than males. We further split the sample according to whether or not the participants completed high school and found no difference in the EComp general factor score between these groups (<12 years: mean = 0.002, SD = 0.96; ≥ 12 years: mean = -0.05, SD = 0.96; d = 0.05; p = 0.7).

Compensation Effects on Diagnosis, Cognition, and Everyday Functioning

The EComp general factor score differed by diagnostic group (p < .001). Although compensation did not significantly differ between MCI (mean = 0.2, SD = 0.8) and cognitively normal groups (mean = 0.2, SD = 0.7; d = 0; p = .9), individuals with dementia (mean = -0.9, SD = 1.3) used compensation strategies less than both the MCI (d = 0.27; p < .001) and cognitively normal (d = 0.28; p < .001) groups.

A subset of participants had longitudinal cognitive and EComp data (225 with episodic memory and EComp, and 230 with executive function and EComp) over an average of 5.8 years (SD = 3.8). In models adjusted for age, education, and gender, the general compensation factor score was significantly associated with both baseline and change in episodic memory and executive function. Specifically, higher general compensation factor scores (more compensation use) were associated with stronger episodic memory and executive function ($\beta = 0.37$, SE = 0.06, p < .001; and $\beta = 0.23$, SE = 0.04, p < .001, respectively) and less decline over time ($\beta = 0.017$, SE = 0.007, p = .01; and $\beta = 0.023$, SE = 0.004, p < .001, respectively).

We further considered the association between compensation and functional change, measured by ECog and IADL measures. We had longitudinal ECog on 234 participants over an average of 6.4 years (SD = 3.7) and longitudinal IADL on 196 participants over an average of 5.8 years (SD = 3.1). Using the ECog and IADL score, a higher general compensation was associated with less functional impairment ($\beta = -0.87$, SE = 0.08, p < .001; and $\beta = -0.39$, SE = 0.03, p < .001, respectively) and slower functional decline ($\beta = -0.04$, SE = 0.004, p < .001, respectively). When also including cognition as a predictor of level and change in ECog and IADL score in the model, EComp remained associated with concurrent level of ECog and IADL score (less functional impairment) (p < .001), and change over time Table 3. Factor loadings for the seven-factor model

Description	G Comp	App	Cook	Fin	Med	Shop	Trans
Keep a calendar (paper or electronic) in a central location or on one's	0.70	0.50					
person at most times							
Put things needed for an appointment in a visible location to remember to bring them	0.80	0.51					
Prepare ahead by gathering items that need to be brought to appointment	0.76	0.54					
Check calendar or day-planner (paper or electronic) on a routine/regular	0.72	0.38					
Keen a very routine or regular schedule as a way to remember appointments	0.68	0.15					
(e.g. every Tuesday go to yoga)	0.00	0.15					
Set alarm or request a reminder about appointments	0.51	0.13					
Audio-record or take notes during appointments to remember the information	0.51	0.09					
Write appointment on a calendar (paper or electronic calendar)	0.77	0.49					
Set a timer when items are cooking to not forget about them	0.59		0.43				
Keep the kitchen very organized so items can be found easily	0.76		0.10				
Set out ingredients/supplies ahead of time when cooking	0.68		0.28				
Stay in the kitchen to not forget about something on the stove	0.59		0.53				
Limit cooking to the microwave	0.30		0.39				
Make only simple, familiar meals	0.55		0.43				
Refer to written recipes more often so items or steps are not forgotten	0.56		0.49				
Use written or electronic reminders to help remember to pay bills	0.54			0.57			
Place bills in a specific location until they get paid	0.78			-0.08			
Limit distractions when working on finances (e.g., clear work space,	0.82			0.39			
go to a quict room) Have a routine to new hills on a certain day of the week or month	0.66			0.00			
Litilize online and/or automatic navment methods as a way to remember	0.00			0.00			
to pay bills	0.47			0.45			
Pay bills immediately when they arrive so bills are not forgotten	0.71			0.32			
Allow more time or double-check work to avoid errors when managing finances	0.84			0.43			
Use written notes, an alarm on smartphone, or other reminders to take medications	0.50				0.39		
Use a strategy to help know which pills have been taken (e.g., turn bottle upside down pill box is empty)	0.47				0.81		
Prepare pill box at the beginning of each week to organize and monitor medications	0.66				0.22		
Keep medications in a visible location (e.g. on the kitchen counter)	0.42				0 50		
Limit distractions when organizing or taking medications (e.g., turn off music/TV, clear a place to work)	0.67				0.50		
Keep a running shopping list to track what items need to be purchased	0.63					0.58	
Use a shopping list (written or in phone)	0.66					0.58	
Check off items on a shopping list as they are found in the store	0.65					0.60	
Plan shopping trip stops prior to leaving the house	0.79					0.14	
Keep shopping list in a specific location	0.70					0.61	
Take more time or double-check shopping list before checking out to make sure all items have been obtained	0.71					0.51	
Use a GPS device in the car or on phone to help plan out or remember	0.50						-0.14
a foure Drive mostly during non-most hours to quoid traffic	0.55						0.55
Limit distractions while driving (e.g. radio and/or talking atc.)	0.55						0.55
Keen car keys in a specific location so as not to lose them	0.59						0.51
Park in the same general location in a parking lot to remember where	0.00						0.51
the car is parked	0.04						0.50
Drive mostly on familiar roads and/or to familiar destinations	0.62						0.53
Leave earlier to allow for more time	0.63						0.47
Drive more slowly	0.49						0.59

Note: Factor loadings statistically significant at p < .05 presented in bold. G Comp = general compensation factor; App = appointments; Cook = cooking; Fin = finances; Med = medications; Shop = shopping; Trans = transportation.

for IADL score ($\beta = -0.02$, SE = 0.004, p < .001) but not ECog (p > .05). That is, greater general compensation was associated with slower subsequent decline in IADL impairment independent of the level of cognitive impairment.

DISCUSSION

Functional independence is a major priority for older adults (Zissimopoulos, Crimmins, & Clair, 2015), and understanding factors contributing to better everyday function is important. Compensation is one possible factor that may enhance functional independence and buffer effects of cognitive decline on everyday function. However, research examining the efficacy of compensation is limited by a lack of assessment tools to measure compensation broadly and explicitly in the context of IADLs. To address this limitation, EComp was previously piloted by our group (Tomaszewski Farias et al., 2018). Preliminary results indicate that greater compensation is associated with better functional ability, independent of the degree of cognitive impairment (Tomaszewski Farias et al., 2018). Using a refined EComp (described here), we sought to rigorously examine its psychometric properties and utility.

Previous theoretical models of compensation and some limited empirical work suggest that people utilize different types of compensation. As such, we first examined whether compensation was best described as a unitary or multidimensional construct. Modeling EComp as a single, unidimensional construct did not produce good model fit. However, we suspect that this may be due to variability in the degree to which individuals engage in the six IADL domains covered by EComp (e.g., spouse manages finances by tradition). Therefore, we next tested a bifactor model to delineate both a general compensation factor as well as IADL-specific factors. This model fit the data well, suggesting that after accounting for the degree of variability in IADL engagement, compensation was well represented by a general compensation factor. We then compared this model with another bifactor model in which compensation was modeled as four, a priori-defined types of compensation (use of external aids, organizational and planning strategies, simplification strategies, and allowing more time and effort), while also accounting for variance attributable to IADL engagement. This model showed no improvement in fit over the previous, simpler model. Furthermore, there was considerable multicollinearity among the compensation factors (r's > .90), suggesting that the strategies were not very distinct. Based on these findings, we concluded that a general compensation factor, while accounting for variance attributable to IADL engagement, was the most parsimonious approach.

Our identification of primarily a general compensation factor differs somewhat from that of de Frias and Dixon (2005) who reported support for multiple types of compensation (external, internal, reliance, time, and effort) using MCQ. Differences across studies may be due to the content covered by each instrument. EComp was organized according to IADL domains, whereas MCQ was not. Additionally, strategies measured by EComp were meant to broadly cover memory- and executively-based compensation strategies, while MCQ measured strategies specific to memory support. Finally, EComp was used as an informant-based measure, and MCQ relied on self-report.

We next examined the association between demographic factors and compensation. The EComp general factor score was not associated with age or education. Similar results have been found with MMQ-MSS (Troyer & Rich, 2002). Women in our sample had a tendency to use more compensation strategies than men. Dixon et al. (2001) described an interaction effect with gender and age on compensation use. Specifically, they found that older men tended to use help from others as a form of compensation more than women did. We excluded help as a form of compensation from our questionnaire, considering it more related to a loss of independence than a form of compensation.

Next, we examined the association between cognitive function and compensation. Diagnosis (normal cognition, MCI, or dementia) was associated with compensation to some degree. Specifically, those diagnosed with dementia compensated less than individuals with either normal cognition or MCI, supporting the construct validity of EComp. However, compensation use did not differ across the latter two groups. These findings are in line with research examining real-world compensation behavior (Aronov et al., 2015). Further, individuals with MCI have been observed to use the same number of external (but fewer internal) strategies as healthy controls on neuropsychological tests of retrospective and prospective memory (Hutchens et al., 2012). However, our finding differs somewhat from other informant report measures that include reliance on others as a form of compensation (Schmitter-Edgecombe, Parsey & Lamb, 2014). Therefore, like any construct, how studies define or examine compensation may result in reported differences.

Next we examined the association between general compensation and episodic memory and executive function performance. We found greater compensation was related to better memory and executive function when measured concurrently. Such findings are not surprising since greater cognitive resources should enable better and more consistent use of compensation strategies. The additional finding that a greater compensation was associated with a subsequent slower rate of decline in cognition (both memory and executive function) is interesting. One possible interpretation is that greater use of compensation has a protective effect against cognitive decline. An alternate explanation is that better baseline cognitive function is associated with both greater compensation use and less cognitive decline, and that the link between cognitive decline and compensation is mediated by cognitive status. Intervention studies examining whether compensation training can positively impact subsequent cognitive trajectories will be critical for answering this difficult causality question.

Finally, and arguably most important, we examined the degree to which greater compensation is associated with

better concurrent functional abilities and rate of functional change over time. Using two separate measures of everyday function (ECog which measures everyday cognitive abilities, and a more traditional IADL measure), greater compensation was associated with a higher level of functional abilities and was associated with a slower rate of functional decline. Additionally, even when accounting for the degree of memory and executive impairment, greater compensation was associated with fewer functional difficulties (on both the ECog and IADL measure) and slower rate of IADL decline. These latter findings are particularly notable because they suggest that even in individuals who have some degree of cognitive impairment, greater compensation may prolong functional independence. This finding is important as prolonging independence likely translates into better quality of life and less economic and caregiver burden.

Because compensation has an effect on functional abilities independent of cognition, it has considerable potential as a target of intervention. In fact, there is emerging evidence that teaching various compensation strategies may help mitigate functional decline (Chudoba, Sawaqdeh, Dahmen, Brown, & Schmitter-Edgecombe, under review; Greenaway, Duncan, & Smith, 2013; Schmitter-Edgecombe & Dyck, 2014; Troyer, Murphy, Anderson, Moscovitch, & Craik, 2008). EComp is well poised to measure the malleability of compensation through directed intervention/rehabilitation, and we have found it to be sensitive to change in relation to a compensatory training intervention (Denny, Barbra, & Tomaszewski-Farias, 2017; Denny & Tomaszewski Farias, 2017).

As with any study, there are limitations. First, measurement of compensation was based on knowledgeable informant report, which is subject to a range of biases including variables such as caregiver burden (Dassel & Schmitt, 2008; Richardson, Nadler, & Malloy, 1995; Tomaszewski Farias et al., 2006). Further, outside observers may not be aware of or recognize all of the strategies that an individual is engaging in while completing daily activities. This could also vary by the type of relationship between informant and participant. In this study, informants included spouses, adult children, and friends. Preliminary analysis (not shown) suggests that accounting for relationship type did not alter results, but more work in this area is needed. Further, it will be important in the future to investigate self-reported compensation use and how it compares to informant report. Future work should also compare informant and self-reported compensation to more objective and real-world measures of compensation. Recent preliminary data suggest that subjective and objective measures of compensation are related (Chudoba et al., 2018). Our analyses may also have been limited by the measure of executive function used. Executive functions encompass a number of disparate abilities, many of which were not captured by our composite measure (e.g., problem solving). Finally, since greater compensation use is associated with slower rate of functional and cognitive decline, examination of whether it forestalls diagnostic progression (e.g., conversion from MCI to dementia) will be important.

In summary, the degree to which older adults utilize compensatory strategies likely has important implications for functional independence over time, even in the face of declining cognition. EComp is a new tool that is likely to prove useful in measuring, and hence better understanding, how the use of compensation strategies impacts various outcomes. The presented factor analysis is an extension of our previous findings using a simple summary score (Tomaszewski et al., 2018). Given that a general factor emerged in the current analysis, it is reasonable to use the simple summary score in future analyses if engagement in all IADLs measures is confirmed. Overall, EComp may serve as a useful outcome measure in interventions that aim to teach or otherwise improve compensation among various clinical groups.

CONFLICT OF INTEREST

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