

Prospective relationships of adolescents' screen-based sedentary behaviour with depressive symptoms: the Millennium Cohort Study

Original Article

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Abstract

Background. Frequent use of screen-based devices could be a modifiable risk factor for adolescent depression, but findings have been inconsistent and mostly from cross-sectional studies. We examined prospective associations of video gaming, social media, and internet use with depressive symptoms in adolescents.

Methods. A total of 11 341 adolescents from the Millennium Cohort Study, a representative, UK population-based. The main outcome was depressive symptoms from a Moods and Feelings Questionnaire (age 14). Exposures were frequency of video game, social media, and internet use (age 11). Physical activity (effect modifier) was measured by self-report.

Results. The fully adjusted models indicated that boys playing video games most days, at least once a week, and at least once a month at age 11 had lower depression scores at age 14 by 24.2% (IRR = 0.77, 95% CI 0.66–0.91), 25.1% (IRR = 0.75, 95% CI 0.62–0.90), and 31.2% (IRR = 0.69, 95% CI 0.57–0.83), compared with playing less than once a month/never. In girls, compared with less than once a month/never, using social media most days at age 11 was associated with 13% higher depression scores at age 14 (IRR = 1.13, 95% CI 1.05–1.22). We found some evidence of associations between using the internet most days and depressive symptoms compared with less than once a month/never in boys (IRR = 0.86, 95% CI 0.75–1.00). More frequent video game use was consistently associated with fewer depressive symptoms in boys with low physical activity, but not in those with high physical activity.

Conclusions. Different types of screen-time may have contrasting associations with depressive symptoms during adolescence. Initiatives to address adolescents' screen-time may require targeted approaches.

Introduction

Depression is the leading cause of global disability (World Health Organisation, 2017). It has long-term effects on daily functioning and can increase risk of multiple serious physical health problems (Batelaan, Seldenrijk, Bot, van Balkom, & Penninx, 2016; Machado et al., 2018; Walker, McGee, & Druss, 2015). The onset of depression tends to first occur during adolescence (Patton et al., 2014; Thapar, Collishaw, Pine, & Thapar, 2012), with an estimated prevalence during this period of 11–14% (Merikangas et al., 2010; Mojtabai, Olfson, & Han, 2016). Depressive symptoms during adolescence are associated with an increased risk of depression, other mental health disorders, and behavioural problems in later life (Bertha & Balázs, 2013; McLeod, Horwood, & Fergusson, 2016). Identifying modifiable risk factors for depressive symptoms during adolescence is an essential step towards reducing the future incidence and burden of depression.

Lower levels of physical activity and higher volumes of sedentary behaviour have consistently been associated with an increased risk of depression in prospective population-based studies of adults (Huang et al., 2020; Schuch et al., 2018; Teychenne, Ball, & Salmon, 2010; Zhai, Zhang, & Zhang, 2015). Sedentary behaviour is any waking activity in a sitting, lying, or reclining position with low energy expenditure (≤ 1.5 metabolic equivalents) (Tremblay et al., 2017). Time spent in sedentary behaviour is high in young people and increases throughout adolescence (Steene-Johannessen et al., 2020; van Ekris et al., 2020). The majority of sedentary behaviour during adolescence is due to screen time, such as television watching (Tremblay et al., 2011). High sedentary behaviour could influence depressive symptoms through several pathways, such as limiting neuroplasticity in the hippocampal brain region,

increasing oxidative stress, or reducing social interactions and support (Kandola, Ashdown-Franks, Hendrikse, Sabiston, & Stubbs, 2019). However, there have been few prospective studies of these associations in adolescence, and findings have generally been inconsistent.

Screen-based devices are embedded in modern life and have many important practical and cultural applications, but there may be risks associated with excessive use. A 2016 meta-analysis of 12 cross-sectional and four longitudinal studies suggested that high screen time-based sedentary behaviours are associated with higher odds of depression in adolescents (Liu, Wu, & Yao, 2016). These findings align with some systematic reviews that suggest high screen time is associated with increased risk of depressive symptoms in adolescents (Hoare, Milton, Foster, & Allender, 2016), but other reviews found no associations (Suchert, Hanewinkel, & Isensee, 2015). Most of those studies were cross-sectional and unable to adjust for reverse causality. A recent prospective study of device-measured activity found that an additional hour of total sedentary behaviour per day between the ages of 12 and 16 was associated with 8–12% increase in depressive symptoms by age 18 (Kandola, Lewis, Osborn, Stubbs, & Hayes, 2020).

The available evidence suggests that high volumes of sedentary behaviour and screen time could increase the risk of depressive symptoms in adolescents. However, previous studies use total sedentary behaviour or screen time as their exposure or have focused on a particular behaviour, such as television-watching (Hoare *et al.*, 2016; Liu *et al.*, 2016; Suchert *et al.*, 2015). The factors contributing to relationships of screen time with mental health in adolescents are complex (Orben & Przybylski, 2019; Przybylski & Weinstein, 2017), and the type of screen time may affect mental health differently.

For example, in video gaming there are social, cooperative, and engaging elements that are absent from other screen time activities, such as general computer use. Screen time modalities with social elements could have mental health benefits that mitigate some of the potential risks of high sedentary behaviour. There is evidence in adults that mentally-passive sedentary behaviours, such as television-watching, are associated with a higher risk of depression than mentally-active sedentary behaviours, such as working at a computer (Hallgren *et al.*, 2019, 2018; Hallgren, Dunstan, & Owen, 2020; Huang *et al.*, 2020). More stimulating forms of screen time could potentially mitigate some of the possible brain and mental health risks of high sedentary behaviours (Hallgren *et al.*, 2020). Different types of screen use, and their differential effects on mental health indicators could account for some of the inconsistencies in previous studies with self-report measures of sedentary behaviour in adolescents (Suchert *et al.*, 2015).

In the 2016 meta-analysis of screen time-based sedentary behaviour, subgroup analyses indicated that increased computer use was modestly but significantly associated with higher depression risk in adolescents (Liu *et al.*, 2016). A UK-based prospective cohort study found that computer use at age 16 was associated with a small increase in the risk of anxiety symptoms at age 18 (Khouja *et al.*, 2019). However, there were no associations for television watching or texting. Recent trial data from adolescents in Canada showed that social media, computer, and television use at age 12 were all associated prospectively with a higher risk of depressive symptoms (Boers, Afzali, Newton, & Conrod, 2019). The same study found no association between increased video gaming and depressive symptoms. A recent systematic review of

12 cross-sectional and one longitudinal study found that high social media use was associated with depression and anxiety symptoms (Keles, McCrae, & Grealish, 2020).

Whilst evidence is emerging to suggest that there are varying associations between different types of screen time and depressive symptoms, findings are inconsistent and primarily based on cross-sectional data (Hoare *et al.*, 2016; Liu *et al.*, 2016; Suchert *et al.*, 2015). A previous meta-analysis identified gender as an effect modifier of associations between screen time and depressive symptoms, with an association only present in boys (Liu *et al.*, 2016). Another meta-analysis that included adults also found the association between screen time and depressive symptoms was not present in males (Wang, Li, & Fan, 2019). Depressive symptoms occur at a higher rate in women, a trend that begins in mid-adolescence and may reflect divergent internal and external influences (Bone, Lewis, & Lewis, 2020). Screen time may differentially influence the risk of depressive symptoms depending on gender, but prospective studies of associations between screen time and depressive symptoms rarely examine gender as a moderator (Boers *et al.*, 2019; Khouja *et al.*, 2019).

Furthermore, structured physical activity can reduce depressive symptoms in adolescents (Bailey, Hetrick, Rosenbaum, Purcell, & Parker, 2018) and high physical activity volumes are associated with a lower risk of depression in the general population (Schuch *et al.*, 2018). Regular physical activity could mitigate some of the mental health risks associated with high sedentary behaviour or screen time, as it does with physical health risks (Ekelund *et al.*, 2020).

We conducted a prospective study with data from a large population-based cohort of adolescents to examine associations of particular forms of screen time with depressive symptoms. We aimed to: (1) assess associations of frequency of video game, social media, and internet use at age 11 with depressive symptoms at age 14; (2) determine the extent to which associations between screen time and depressive symptoms may differ by gender; and, (3) examine whether the physical activity may moderate any associations between screen time and depressive symptoms. We expected that more frequent social media use at age 11, but not video game or computer use would be associated with increased depressive symptoms at age 14. This is based on the video game and computer use being mentally-active behaviours (Hallgren *et al.*, 2020, 2019, 2018; Huang *et al.*, 2020) and systematic review evidence of a positive association between social media use and depressive symptoms (Keles *et al.*, 2020). We expect that there may be gender differences in these associations based on prior systematic review evidence (Liu *et al.*, 2016; Wang *et al.*, 2019), and no association between screen time and depressive symptoms in those with high physical activity given its capacity to reduce depressive symptoms in adolescents (Bailey *et al.*, 2018).

Methods

Participants

We used data from the Millennium Cohort Study (MCS), a representative sample of 18,552 families and 18,818 children born in the UK between September 2000 and January 2002, described in full elsewhere (Connelly & Platt, 2014). Those from socially deprived areas and ethnic minority groups were oversampled to increase representation. The ongoing study currently includes six waves of data collection covering a range of demographic,

psychosocial, environmental, and biological factors. Our study focuses on adolescent behaviour and includes data from sweeps 5 (January 2012–February 2013) with 13 469 participants aged 11 (71.5% of the original sample) and sweep 6 (January 2015–March 2016) with 11 872 aged 14 (63.1%). We defined our sample as all with a completed outcome measure ($n = 11\,341$).

The National Health Service Research Ethics Committee provided ethical approval for MCS. We obtained all MCS data from the UK Data Archive.

Exposure(s)

Our exposure was the self-reported frequency of three different types of screen use at age 11: video games, social media, and leisure-time internet use. Participants were asked: *How often do you [play games on a computer or games console/use the internet (not for school)/visit a social networking website on the internet]?* The possible categorical responses included: *most days, at least one a week, at least once a month, less often than once a month, or never*. The question does not specify a time period. Due to low numbers, we combined *less often than once a month* with *never* to create a 4-point Likert scale.

Outcome

Depressive symptoms were measured using a short Moods and Feelings Questionnaire (sMFQ) at age 14. The sMFQ is a self-report measure of DSM-IV depressive symptoms over the past 2 weeks (Sharp, Goodyer, & Croudace, 2006). It includes 13 questions, with responses including not true (0 score), somewhat true (1 score) to true (2 score) with scores ranging from 0 to 26. Higher scores indicate more severe symptoms. It is validated for assessing depressive symptoms in adolescents in population-based research (Sharp et al., 2006). We used sMFQ scores as a continuous outcome measure to maximise statistical power.

Confounding and moderating variables

We determined all possible confounding variables *a priori*. We mapped causal assumptions between screen-time, depressive symptoms, and all confounding variables using the Directed Acyclic Graph (DAG) in Figure 1 of the Supplementary Materials (page 1) and adjusted models accordingly. Possible confounding variables included: gender, socioeconomic position (household income), baseline emotional symptoms [emotional symptoms subscale from the Strengths and Difficulties Questionnaire (SDQ)], self-reported maternal history of a depression or anxiety diagnosis, the self-reported experience of bullying, self-reported physical activity (frequency of playing sports or active games inside or outside on the same 4-point Likert scale as exposure variables), and standardised body mass index (BMI). The direction of causality between BMI and sedentary behaviour is unclear in young people (Biddle, García Bengoechea, & Wiesner, 2017). We chose to adjust for BMI as a confounding variable due to the substantial genetic influences on adiposity (Rohde et al., 2019) that potentially suggests BMI could cause high screen time in young people. We did not adjust for physical activity as a confounding variable due to evidence that sedentary behaviour and physical activity are unlikely to displace one another in young people (Pearson, Braithwaite, Biddle, van Sluijs, & Atkin, 2014).

Analyses

Main analysis

The main analysis examined how the frequency of screen time use for each activity at age 11 was associated with depressive symptoms at age 14 (aim 1). The outcome distribution had a high positive skew (see Figure 2, page 2 of the Supplementary Materials) and was over-dispersed (variance > mean). To account for this, we used negative binomial regression models. These models are commonly used for count data, but as the sMFQ scores are discrete, independent, and have no negative values, models using count distributions are still applicable (Green, 2020; Kandola et al., 2020). The outcome for these models is interpretable as a percentage change in sMFQ scores.

We entered each categorical exposure variable (video gaming, social media, or internet use) into separate models with the same continuous outcome (depressive symptoms). We ran each model with an interaction term for gender (aim 2) and stratified models accordingly. We present models fully-adjusted for all confounding variables in the main text and crude models in the online Supplementary Materials.

***Secondary and sensitivity analysis

The secondary analysis investigated the extent to which associations between each screen time type and depressive symptoms varied by physical activity (aim 3). We dichotomized the physical activity variable to create a 'high activity' group from the *most days* responses and a 'low activity' group from combining all other responses: *at least once a week, at least once a month, less often than once a month, or never*. We then reran the adjusted models from the main analysis with the physical activity variable as a multiplicative interaction term. Where interaction terms were significant, we presented the models stratified by physical activity.

We also conducted sensitivity analyses that included using psychosocial adjustment (total SDQ score) instead of emotional symptoms as an alternative method of adjusting for baseline mental health. We also reran fully-adjusted models for the video game exposure with a larger reference group (16% of participants) by combining the *less often than once a month* or *never* and *at least once a month* categories. This was due to large differences in the video games reference group (6% of participants) and some of the comparison groups (54%, 30%, and 10%) in the main analysis. We also used multiple imputations by the chained equation to examine how missing data could have affected our main findings through selection bias. We reran the main analysis in a full cohort with imputed missing data.

All analyses were conducted in Stata (version 13) and weighted according to sampling design.

Results

Participants

There were 11 341 participants in the total pool of participants and the fully adjusted models included 7701 (68%) participants with complete data. The mean sMFQ score at follow-up was 6.04 (SD = 5.22). Table 1 contains the baseline characteristics of participants included in this study according to gender.

Main analysis

The interaction terms for gender were significant for all exposures ($p < 0.05$), and stratified, fully-adjusted models are presented in Table 2. We provide crude models in the online Supplementary

Table 1. Baseline participant characteristics by gender

Variable	Overall <i>n</i> (%)	Male <i>n</i> (%)	Female <i>n</i> (%)
Gender			
Male	4877 (48.85)		
Female	5106 (51.15)		
Ethnicity			
White	8482 (84.98)	4162 (85.36)	4320 (84.62)
Indian, Pakistani, or Bangladeshi	951 (9.53)	450 (9.22)	501 (9.81)
Black or Black British	297 (2.98)	147 (3.01)	150 (2.94)
Other or mixed	251 (2.51)	117 (2.4)	134 (2.62)
Maternal history of depression or anxiety			
Yes	2521 (27.08)	1187 (26.11)	1334 (28.01)
No	6787 (72.92)	3359 (73.89)	3428 (71.99)
Experience of bullying			
Certainly true	350 (4.09)	178 (4.30)	172 (3.88)
Somewhat true	1644 (19.19)	818 (19.78)	826 (18.64)
Not true	6573 (76.72)	3139 (75.91)	3434 (77.48)
Household income			
Lowest quintile	1963 (18.77)	835 (19.72)	1007 (18.45)
Second	2021 (19.32)	936 (19.19)	971 (19.02)
Third	2094 (20.02)	1013 (20.77)	990 (19.39)
Fourth	2190 (20.94)	1026 (21.04)	1078 (21.11)
Highest quintile	2179 (20.84)	1065 (21.84)	1053 (20.62)
BMI			
Mean (s.d.)	19.16 (3.60)	18.93 (3.5)	19.37 (3.65)
Physical activity			
Most days	5602 (59.48)	3061 (66.98)	2541 (52.41)
At least once a week	2678 (28.43)	1037 (22.69)	1641 (33.85)
At least once a month	593 (6.30)	224 (4.90)	369 (7.61)
Less often than once a month or never	545 (5.79)	248 (5.43)	297 (6.12)
SDQ			
Emotional symptoms mean score (s.d.)	1.71 (2.00)	1.58 (1.92)	1.77 (2.00)
Total problems mean score	7.06 (5.74)	7.4 (5.80)	6.48 (5.38)
Video gaming			
Most days	5074 (53.68)	3079 (67.08)	1995 (41.02)
At least once a week	2861 (30.27)	1152 (25.10)	1709 (35.14)
At least once a month	944 (9.99)	219 (4.77)	725 (14.91)
Less than once a month/never	574 (6.07)	140 (3.05)	434 (8.92)
Social media			
Most days	1734 (18.41)	729 (15.95)	1005 (20.73)
At least once a week	1211 (12.86)	553 (12.10)	658 (13.57)
At least once a month	545 (5.79)	280 (6.13)	265 (5.47)
Less than once a month/never	5929 (62.95)	3008 (65.82)	2921 (60.24)
Internet use			
Most days	5237 (55.44)	2576 (56.21)	2661 (54.71)

(Continued)

Table 1. (Continued.)

Variable	Overall <i>n</i> (%)	Male <i>n</i> (%)	Female <i>n</i> (%)
At least once a week	3027 (32.04)	1454 (31.73)	1573 (32.34)
At least once a month	720 (7.62)	341 (7.44)	379 (7.79)
Less than once a month/never	463 (4.90)	212 (4.63)	251 (5.16)

BMI, body mass index; SDQ, strengths and difficulties questionnaire.

Table 2. Associations between screen-time activity and depressive symptoms stratified by gender

Exposure	Variable	Depression scores (sMFQ)					
		Male (<i>n</i> = 3710)			Female (<i>n</i> = 3991)		
		IRR	95% CI	<i>p</i>	IRR	95% CI	<i>p</i>
Video gaming	Most days	0.768	0.645–0.913	0.003	1.001	0.909–1.120	0.816
	At least once a week	0.749	0.624–0.898	0.002	0.928	0.833–1.033	0.171
	At least once a month	0.688	0.569–0.833	<0.001	0.923	0.820–1.040	0.189
	Less than once a month/never	Reference					
Social media	Most days	1.071	0.976–1.174	0.148	1.130	1.050–1.217	0.001
	At least once a week	1.024	0.903–1.162	0.705	0.964	0.883–1.052	0.416
	At least once a month	1.012	0.869–1.178	0.877	0.966	0.853–1.094	0.588
	Less than once a month/never	Reference					
Internet	Most days	0.864	0.746–1.000	0.051	1.100	0.955–1.267	0.181
	At least once a week	0.872	0.751–1.014	0.076	0.973	0.843–1.122	0.703
	At least once a month	0.969	0.807–1.166	0.744	0.944	0.803–1.11	0.486
	Less than once a month/never	Reference					

IRR, incident rate ratios; 95% CI = 95% confidence intervals.

All models are adjusted for BMI, bullying, emotional symptoms at baseline, socioeconomic position, maternal depression or anxiety diagnoses, and physical activity.

Materials (Table 1, page 3). Compared with less than once a month/never, playing video games most days, at least once a week, and at least once a month at age 11 were associated with 24.2% (IRR = 0.77, 95% CI 0.65–0.91), 25.1% (IRR = 0.75, 95% CI 0.62–0.89), and 31.2% (IRR = 0.69, 95% CI 0.57–0.83) lower depression scores in boys at age 14, respectively. There were no clear associations between more frequent *v.* less frequent video gaming and depression scores in girls. Using social media most days at age 11 was associated with 13% (IRR = 1.13, 95% CI 1.05–1.22) higher depression scores at age 14 compared with less than once a month/never in girls. There were no clear associations between other frequency of use categories and depression scores in girls or any associations in boys. There was some indication of associations between internet use most days (IRR = 0.86, 95% CI 0.75–1.00) and at least once a week (IRR = 0.87 95% CI 0.75–1.01) and depression scores compared with less than once a month/never in boys. There were no associations between more frequent *v.* less frequent internet use and depression scores in girls.

Secondary and sensitivity analysis

In the secondary analysis, there was no evidence of an interaction with physical activity for social media or internet use frequency and depressive symptoms ($p > 0.05$). There was evidence of an interaction with physical activity for video gaming frequency and depressive symptoms in boys only ($p = 0.024$).

In fully adjusted models for boys with low physical activity ($n = 1226$), using video games for most days was associated with 32.2% (IRR = 0.68; 95% CI = 0.54–0.86; $p < 0.001$), at least once a week with 35.2% (IRR = 0.65; 95% CI = 0.50–0.83; $p < 0.001$), and at least once a month with 38.7% (IRR = 0.61; 95% CI = 0.46–0.82; $p < 0.001$) lower depression scores than less than once a month/never. In boys with high physical activity ($n = 2484$), there were some associations between using video games at least once a month and depressive symptoms (IRR = 0.75; 95% CI = 0.58–0.98; $p = 0.034$) compared with less than once a month/never, but not with more frequent video game use.

These results were consistent in a full cohort with imputed missing data (see Table 2, page 4 of the online Supplementary Materials). The results of the sensitivity analysis were similar when using total SDQ to adjust for baseline mental health (see Table 3 of the Supplementary Materials). The associations between video gaming and depressive symptoms were attenuated when using the larger combined reference group (see Table 4, page 6 of the Supplementary Materials).

Discussion

Main findings

This prospective study examined associations of three types of screen time in girls and boys at age 11 with depressive symptoms

at age 14. We found that using video games most days, at least once a week, and at least once a month were associated with 24.2–31.2% lower depressive symptom scores compared to less than once a month/never in boys, but not in girls. There was some evidence that physical activity moderated this association as the associations were consistent in boys with low physical activity, but not in those with high physical activity. Using social media most days was associated with 13% higher depressive symptom scores than less than once a month/never in girls. The relationship between internet use and depressive symptoms was unclear in our results.

Few studies have examined associations between the frequency of video gaming and depressive symptoms in adolescents. A previous meta-analysis of mostly cross-sectional data provided some indications that more frequent video gaming was associated with a lower risk of depression (OR = 0.89, 95% CI 0.74–1.06) (Liu et al., 2016). A recent longitudinal study found no associations between video gaming and depressive symptoms (Boers et al., 2019), but this study did not examine gender as a potential effect modifier.

Our results also suggest the novel finding that more frequent video gaming is associated with lower depression symptom scores in boys who are less physically active, but not in those who were physically active. Adolescents who spend less time playing sports and active games may derive more enjoyment and social interaction from playing video games more frequently. We also found some associations between increased social media use and depressive symptoms in girls, which aligns with prior, mostly cross-sectional studies (Boers et al., 2019; Keles et al., 2020). This finding may again be influenced by social factors. For example, studies in adults suggest that women are more likely than men to report using social media for maintaining social ties and gather social information (Krasnova, Veltri, Eling, & Buxmann, 2017). Frequent social media use is associated with greater feelings of social isolation than less frequent use (Primack et al., 2017). Adolescent girls with frequent social media use may experience increased social isolation, which can increase the risk of depressive symptoms (Santini et al., 2020). Some studies have indicated that associations between social media use and poorer mental health are stronger in female adolescents than boys (Blomfield Neira & Barber, 2014), but other studies have not found this (Keles et al., 2020).

Strengths and limitations

Our findings are based on data from a large, representative cohort of adolescents with a 3-year follow up. The use of an sMFQ is another strength as it allows the assessment of clinical and sub-clinical symptoms in participants that may not be present to mental health services. The prospective study design and adjustment for baseline symptoms lower the risk of reverse causation. We used DAGs determined *a priori* to inform each analysis, which improves our capacity to estimate causal effects (Hernan & Robins, 2020).

A limitation of our study includes the high attrition, which could have introduced selection bias. However, the results remained consistent in a full sample with imputed missing data. This suggests that selection bias within our sample is unlikely to have increased due to the attrition, but selection bias is still possible in the wider Millennium Cohort sample. Another limitation is the lack of data on the duration of screen-time use, which could moderate the association between frequency of use and

depressive symptoms. For example, there could be a difference in the risk of depression symptoms between participants who played video games most days for several hours *v.* those who played for just 1 hour. As no timeframe is specified in the question, participants' reported use could refer to different periods. Screen time use in young adolescents may also have changed since they were measured in 2012 and 2013 in the Millennium Cohort Study.

There were also large differences in the size of some comparison groups, which could cause unstable estimates when comparing groups. One sensitivity analysis indicated that associations between video gaming and depressive symptoms were attenuated in boys when using a larger reference group from combining the two least frequent use groups. However, it is not possible to determine whether this is due to the inclusion of boys who play video games semi-regularly, *i.e.*, more than once a month. A larger sample with more evenly distributed groups will be necessary to determine the extent to which our findings are affected by a random error in the reference groups.

There could also have been a measurement error with the physical activity data. We used self-reported physical activity data that are prone to biases, such as attention and recall bias (Prince et al., 2008). Another possible source of measurement error includes using the SDQ emotional symptom subscale to assess baseline depressive symptoms. While the outcome measure (sMFQ) directly assesses depressive symptoms, the SDQ subscale captures the broader concept of depression. It may miss specific depressive symptoms and allow for potential confounding from baseline depression. However, as depression is relatively uncommon before puberty, measuring the broader concept of depression could be sufficient.

Implications and future directions

Sedentary behaviour is high in young people and increases during adolescence (Steene-Johannessen et al., 2020; van Ekris et al., 2020) with the growing use of screen-based devices (Tremblay et al., 2011), which may contribute to a higher subsequent risk for depression (Kandola et al., 2020). More-passive compared to more mentally-active sedentary behaviours can have varying relationships with the risk of depression in adults, with mentally active sedentary behaviours in some cases being protective (Hallgren et al., 2020, 2019, 2018; Huang et al., 2020). Our findings suggest that there may be such relationships in adolescents. Approaches that aim to broadly reduce sedentary behaviour or screen-time in young people can overlook these complexities and may not maximise the potential impact on mental health risks.

Our findings suggest that a more targeted approach to screen time may be necessary for the context of risk of depression in adolescents. For example, targeting high social media use could produce a greater effect on reducing depression risk than video gaming, particularly in girls. Our results suggest that interventions may benefit from a gender-specific approach and considering related factors that improve adolescent mental health, such as physical activity (Bailey et al., 2018). Adolescents may interact differently with screen-based devices depending on their gender and warrants further research to determine whether different recommendations would be helpful.

The relationships between screen-time and mental health are complex, and their nuances warrant more careful consideration. Inconsistent findings in previous studies could be due to not

examining different types of screen-time in relation to depression risk in adolescents (Hoare et al., 2016; Liu et al., 2016; Suchert et al., 2015). More evidence is needed on how different types of screen-time may affect the risk of depression in young people. Each type of screen-time provides broadly different experiences that are likely to have a divergent effect on mental health.

For example, video games can involve complex, immersive experiences with detailed and interactive storylines. Many games involve problem-solving, co-operation, and offer a platform for socialization. The use of video games as a social platform could be particularly important for adolescents who participate in fewer sports and active games. Several studies have found that commercial video gaming is associated with improvements in performance on attention, problem-solving, and memory tasks (Choi et al., 2020) and structural changes in brain plasticity, such as growth in hippocampal and prefrontal areas (Kühn et al., 2014b; Kühn, Gleich, Lorenz, Lindenberger, & Gallinat, 2014a). These elements of video gaming may translate into mental health benefits in some young people with mild to moderate use. Infrequent video game use in this study may also reflect environmental factors that could also contribute to the risk of depression, such as financial difficulties or highly restrictive parenting.

However, excessive video game use may nevertheless be harmful to mental health in young people. Similarly, excessive social media use could be detrimental, particularly if it increases perceptions of social isolation (Primack et al., 2017). Contextual factors of social media use may also be relevant to adolescents' risk of depressive symptoms. For example, using social media for social comparisons could affect self-esteem, leading to depressive symptoms (Robinson et al., 2019).

Conclusions

In this prospective cohort study, we found that more-frequent video gaming at age 11 was associated with a lower risk of depressive symptoms at age 14 for boys but not girls. More frequent social media use at 11 was associated with a higher risk of depressive symptoms in adolescent girls but not boys. Approaches aimed at reducing sedentary behaviour or screen-time should consider the differential associations between activity type and depressive symptoms. More research is necessary to understand how different types of screen-time affect the risk of depression in young people.

Data availability. Details for accessing the data used in this study are available from the UK Data Service.

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Ethical standards. The authors assert that all procedures contributing to this work comply with the ethical standards of the relevant national and institutional committees on human experimentation and with the Helsinki Declaration of 1975, as revised in 2008.

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