

Does Geography Matter? The Health Service Use and Unmet Health Care Needs of Older Canadians*

James Ted McDonald and Heather Conde
University of New Brunswick

RÉSUMÉ

Le coût croissant de soins de santé et le changement des profils démographiques ont entraîné le déplacement et la redistribution du financement et des services entre les zones rurales et urbaines. La plupart des analyses économétriques de l'utilisation de services de santé au Canada incluent des contrôles larges selon la province et l'état rural/urbain; mais relativement peu du travail économétrique a porté sur la variation géographique dans l'utilisation de services de santé. À l'aide de l'Enquête sur la santé dans les collectivités canadiennes (ESCC 2.1), nous avons examiné les déterminants de diverses mesures d'utilisation des services de santé par les Canadiens âgés de 55 ou plus d'une gamme de zones urbaines et rurales de résidence. Notre analyse de régression a montré que les anciens résidents dans les zones rurales font moins visites chez un omnipraticien, chez un spécialiste et chez un dentiste par rapport aux résidents urbains. Tout étant égal, il n'existe aucune différence significative parmi nuits passées à l'hôpital ou dans les besoins de soins de santé non satisfaits. Cependant, après contrôle pour les caractéristiques démographiques, le statut socioéconomique, l'assurance santé privée et l'état de santé, ces différences sont importantes.

ABSTRACT

The rising cost of health care and changing demographic profiles have resulted in the relocation and redistribution of funding and services between rural and urban areas. Most econometric analyses of Canada's health service use include broad controls by province and rural/urban status, but relatively little econometric work has focused on geographical variation in health service use. Using the Canadian Community Health Survey 2.1, we examined determinants of various measures of health services use by Canadians aged 55 or older across a range of urban and rural areas of residence. Our regression analysis showed that older residents in rural areas made fewer visits to a general practitioner, to a specialist, and to a dentist relative to urban residents. All else being equal, there are no significant differences in hospital nights or in unmet healthcare needs. These differences are significant after controlling for demographic characteristics, socioeconomic status, private health insurance, and health status.

* The authors thank Brian Ferguson, Logan McLeod, Carolyn Rosenthal, and Byron Spencer for helpful comments. Conde acknowledges financial support from a Social Sciences and Humanities Research Council master's fellowship. McDonald acknowledges financial support from SSHRC through the Social and Economic Dimensions of an Aging Population (SEDAP) program at McMaster University. Analysis of confidential Canadian data was conducted at the Canadian Research Institute for Social Policy – University of New Brunswick (CRISP-UNB) Research Data Centre in Fredericton.

Manuscript received: / manuscrit reçu : 25/08/08

Manuscript accepted: / manuscrit accepté : 30/06/09

Mots clés : utilisation de services de santé, différences rurales/urbaines, les personnes âgées, besoins de soins de santé non satisfaits

Keywords: health service use, rural/urban differences, older and aged individuals, unmet health care needs

Correspondence concerning this article should be addressed to / La correspondance concernant cet article doit être adressées à:

James Ted McDonald
Department of Economics
University of New Brunswick
PO Box 4400
Fredericton, New Brunswick
Canada E3B 5A3
(tedmcdon@unb.ca)

Introduction

According to the Canada Health Act, all Canadians are entitled to equitable access to health services, regardless of where in Canada they live. Although the organization and delivery of hospital and medical services are the responsibility of the individual provinces and territories, access to these services must be universal, comprehensive, accessible, portable, and publicly administered in order for provinces and territories to receive federal funding to cover a portion of the associated costs of providing health services (Jennissen, 1992). Nevertheless, differences both in the provision of health services and in health outcomes are well documented across provinces and regions of Canada, and in particular between rural and urban areas. The Romanow (2001) report entitled *Building our Values: The Future of Health Care in Canada* identified access to health care in rural areas and remote communities as a major problem due to both distance and retention of health workers. Similar conclusions were drawn in the Kirby report that noted access issues were the most serious problems for residents of rural and remote areas, and also that the health of rural residents was worse than that of their urban counterparts.¹ While the percentage of the population living in rural areas fell from 29.2 per cent in 1991 to 22.2 per cent in 1996, the percentage of physicians practicing in rural areas fell from 14.9 per cent to 9.8 per cent over the same period. Further, the ratio of physicians per 1,000 residents in rural areas has been forecast to fall from 0.79 in 1999 to 0.53 in 2021 (Laurent, 2002).

Ongoing rationalization of health care provision by provincial governments, including the closure of hospital beds, emergency wards, and the replacement of hospitals with community care centres in less populated areas, has been well documented in the media, giving at least the impression that people in rural areas are experiencing longer waiting or travel times, lower levels of technology, and more uneven resource distribution than in other areas. As Cloutier-Fisher and Joseph (2000) noted, government funding reductions and downsizing may have contributed to the devolution of responsibility for health care to local communities and individuals, making the provision of health care to vulnerable populations such as those in rural areas more challenging.²

In this context, there has been continued interest in the health outcomes and health services use of older individuals since as a group they are the most frequent users of health services (Martin-Matthews, 2002; Rosenberg & James, 2000). As Canada's population ages, there will be increasing pressure on the public health care system because of the rate at which older individuals use publicly funded health services (Lassey, Lassey, & Jinks, 1997, quoted in Martin-Matthews). The health service use of older individuals is particu-

larly important for rural areas, because demographic and socio-economic changes have meant that older individuals are increasingly over-represented in rural and small town areas (Jennissen, 1992). However, most of the recent research on the use of health services by older Canadians in rural and urban areas has been limited to specific provinces (Fakhoury & Roos, 1996, and Peterson, Shapiro, & Roos, 2005, for Manitoba; Cloutier-Fisher & Joseph, 2000, for Ontario; Liu, Hader, Broussart, White, & Lewis, 2001, for Saskatchewan; Allan & Cloutier-Fisher, 2006, for British Columbia). The research that has been of national scope has not focused specifically on the health care use or unmet health care needs of older individuals in rural areas.³

This article examines whether the use of basic health services and the incidence of unmet health care needs experienced by Canadians aged 55 years or older vary across urban and rural areas of Canada, and analyzes possible reasons for any observed differences. The underlying motivation for the article is to provide additional evidence on whether residents of rural areas are relatively more likely to face barriers in obtaining health care than residents of more urban areas. In the work we report on here, we controlled for a range of demographic, socio-economic and health status characteristics that may differ between rural and urban residents and thus might account for observed disparities in health service use.⁴ Since disparities in health service use among otherwise comparable residents of rural and urban areas may still not necessarily indicate access barriers, we found it useful to analyze a range of different health services that vary in terms of who makes the decision to obtain care. Thus, we considered the determinants both of health services that are typically obtained at the discretion of the patient and those that are typically obtained following a joint decision of doctor and patient. We also analyzed the prevalence of unmet health care needs. Such unmet needs, particularly those arising from services not being available in the time required, provide arguably a more direct measure of barriers in access to appropriate health care (Chen, Hou, Sanmartin, Houle, Tremblay, & Berthelot, 2002; Nelson & Park, 2006).

Conceptual Framework

The Andersen framework (Andersen, 1968; Andersen & Newman, 1973) is commonly used in studies of health service use and has been both refined and criticized by various authors in later research (e.g., Andersen, 1995; Wolinsky, 1994; Wolinsky & Johnson, 1991). The basic model identifies three types of factors likely to be important determinants of an individual's demand for health services: (1) predisposing factors such as age and gender; (2) needs factors such as health

status; and (3) enabling factors such as income. Wolinsky extended the notion of enabling factors to how medical care is organized and included, in the basic model, variables reflecting community resources generally and specific measures such as physician density. In the same vein, health insurance should also be included as an enabling factor. Similarly, Andersen (1995) argued that the main focus of the model for policy should be on enabling factors, which are most mutable. He expanded the model to include measures of health behaviours as enabling factors affecting health service use, but also recognized the dynamic interaction of health behaviours, health service use, and health outcomes. Geography can be considered an enabling factor since living in a rural area would imply access issues arising from longer distances that must be travelled to obtain certain health services and possibly longer waiting times.

Dependent Variables – Measuring Health Service Use

Certain measures of health services are basic services that the average Canadian should have regardless of the state of his/her physical health. For example, it is recommended that all Canadians, particularly those individuals over 20 years of age, see a doctor once a year for a health check-up, and most dentists recommend at least an annual check-up for good oral health (Peckins, 2006). Other types of health service use reflect particular medical needs in which the state of one's physical health is likely to be an important determinant. For example, a visit to a specialist or a night's stay in the hospital is likely to arise in response to a particular medical condition. By considering how a range of health services differs in use between rural and urban areas, insights into the nature of possible barriers to access can be gained. Additional insights can also be gained by considering use of health services that are not typically covered by provincial or territorial health insurance systems, such as dental care or visits to alternative health care providers.

We considered a range of measures of basic health service use as our dependent variables. These included binary indicators for whether an individual had a family doctor, whether the individual had visited a general practitioner (GP), a medical specialist (such as a surgeon, allergist, gynaecologist, or psychiatrist) or a dentist in the previous 12 months, whether the individual had spent at least one night in hospital, had received home care, or had received alternative health care. Our set of dependent variables also included measures of the frequency during the previous 12 months of GP visits, specialist visits, dental visits, and nights in hospital conditional on at least some use

of that service during the year. As well, we analyzed an indicator of whether the individual experienced unmet health care needs in the previous 12 months.

Independent Variables

Predisposing Factors

Following the expanded Andersen model, we included variables for age, gender, marital status, and immigrant status as predisposing factors. We controlled for age using a set of indicator variables for five-year intervals (age 55–59, age 60–64, etc.) as well as indicator variables for immigrant status and for the different categories of marital status (married, widowed, separated/divorced, never married). Unfortunately, small sample sizes precluded our use of controls for ethnicity for many measures of health service use since our focus is on older individuals in rural areas.

Need Factors

Need factors (e.g., health status) are likely to be the most important and immediate determinants of health service use, and self-perceived health is widely used in the literature as a proxy for health status (Newbold, Eyles, & Birch, 1995). We included indicator variables for the different categories of self-reported general health specified by the respondent (excellent, very good, good, fair, or poor) and indicator variables for each of nine different (self-reported) chronic conditions: (1) cancer, (2) Alzheimer's disease, (3) high blood pressure, (4) asthma, (5) stroke, (6) heart disease, (7) diabetes, (8) arthritis, and (9) glaucoma.

Enabling Factors

We included controls for the highest level of education attained (less than secondary school graduation, secondary school graduation, some post-secondary, or university degree or more), and indicator variables for five levels of household income adequacy.⁵ It is noteworthy that causality between health (and also health service use) and income can work in both directions (Buckley, Denton, Robb, & Spencer, 2004; Fuchs, 2004). As pointed out by Case, Fertig, and Paxson (2005), being in poor health, even at an early age, can determine one's level of socio-economic status in the future.

We expanded the set of enabling factors to include four sets of variables identified in the literature as potentially important. The first set of variables reflected types of private health insurance coverage that are available in Canada to cover the cost of those services that are not insured under Medicare, such as dentist visits, pharmaceuticals, eye exams and glasses, and

private hospital rooms. We included in the statistical analysis indicators for health insurance covering each of these aspects of health service use.⁶ The second set of variables was for health behaviours related to smoking, including whether the person was currently a daily smoker or was ever a daily smoker. The third set of variables measured differences in the potential availability of physicians across health regions, and included both the number of general practitioners (GPs) and the number of specialists per 1000 residents.

The fourth set of variables reflected region of residence, the main focus of the analysis. Region of residence is also an enabling factor since it can reflect ease of access to health services among other factors. We included a set of indicators for the degree of urbanization of the individual's community. Using Statistics Canada (2006) definitions, we identified four types of areas based on population and proximity to a Census metropolitan area (CMA) or Census agglomeration (CA). For ease of expression, we hereafter refer to the group of CMAs and CAs simply as CMAs. The four categories follow: (1) urban CMA: areas with a population of at least 10,000 and delineated within a Census metropolitan area or Census agglomeration; (2) rural fringe of CMA: areas within a CMA but with a population of less than 10,000; (3) urban not CMA: areas with a population of at least 1,000 but less than 10,000 and with no fewer than 400 people per square kilometre; and (4) rural not CMA: areas outside of CMAs but not otherwise classified as urban. (See Allan & Cloutier-Fisher, 2006, for further discussion of these categories.)⁷ We also included a set of indicator variables for province of residence that will capture provincial-wide differences in the provision of health care for older Canadians since provinces differ markedly in their management of health care systems and in terms of the extent of publicly provided health insurance for various health services, prescriptions, and other services.

Data and Methods of Analysis

Data Sources

The data used in our study are from the Statistics Canada Master file of Cycle 2.1 of the Canadian Community Health Survey (CCHS) from 2002–2003. The CCHS focuses on Canadians aged 12 and older who live in private dwellings in all of the provinces and territories and does not sample those living on Indian Reserves, Crown Land or private institutions, or members of the Armed Forces. We restricted our attention only to residents of Canada's provinces who are aged 55 or older. Given the possible influence of outliers in the data arising from the very old, in results not reported here we repeated the analysis after restricting the sample of Canadians to those aged 55–79. There

was very little impact on the results that we report here. Our final sample for estimation consisted of 39,974 observations across the 10 Canadian provinces. Note that all data in the CCHS, including data on both health service use and on health outcomes such as chronic conditions, are based on self-reports by the respondents. We also obtained data by health region on the number of active registered GPs and number of active registered specialists per 1,000 residents (both full-time and part-time) for the year 2002 from the Canadian Institute of Health Information. We included these data as rough measures of the average supply of physicians in an individual's particular health region of residence.⁸

Empirical Approach

For the statistical analysis, estimation of the determinants of binary dependent variables was by multivariate Logistic regression. For measures of health service use where we are also interested in the frequency of use, we employed the "two-part" approach in which frequency is divided into whether there was use at all and the frequency of use conditional on some use. Ordinary Least Squares (OLS) estimation on the number of visits or days conditional on positive use constituted the second part. This method is commonly used in the literature (see, for example, Van Houtven & Norton, 2004; Escarce, Shea & Chen, 1997; Hurd & McGarry, 1997). Results of the Logistic estimation are reported as odds ratios relative to the specified base case. We transformed the frequency measures with a log transformation in order to reduce the possible influence of large numbers of visits or nights in hospital. For frequency measures conditional on use, OLS coefficient estimates were reported and since the dependent variables are in log form, these OLS estimates can be interpreted approximately as the proportional change in the frequency of use of the service relative to the specified base case. In all regressions, results were obtained using population weights, and standard errors were calculated after allowing for clustering of observations by health region.

Results

Descriptive Statistics

We first illustrate overall differences in health service by residents of different types of rural and urban areas. Figures 1 and 2 show the proportion of older Canadians who used particular health services in the year prior to the survey date, as well as the proportion of older Canadians reporting unmet health care needs. Figure 3 shows the average frequency of use of particular health services, conditional on at least some use. Although the proportion of older Canadians reporting

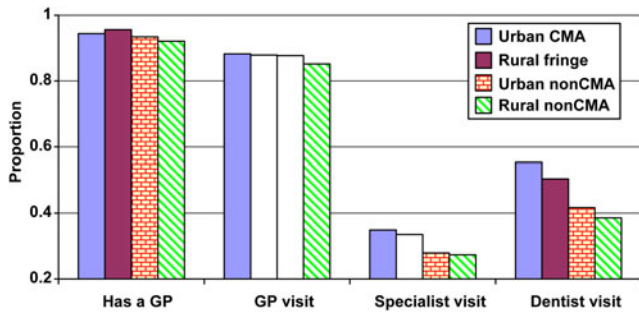


Figure 1: Proportion of older Canadians using health services in the past 12 months. A clear bar indicates not significantly different from Urban CMA at the 5 per cent level of significance. Results are for the adult non-institutional population aged 55 or older.

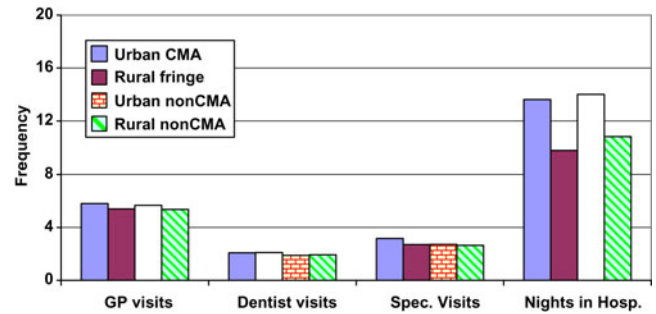


Figure 3: Frequency of use of health services by older Canadians in the past 12 months (conditional on some use). A clear bar indicates not significantly different from Urban CMA at the 5 per cent level of significance. Results are for the adult non-institutional population aged 55 or older.

unmet health care needs was not statistically different across rural and urban areas, residents of rural non-CMA areas were less likely to have a GP, to have visited a GP, a specialist, or a dentist, have had fewer visits with a GP, a dentist, or a specialist, and have spent fewer nights in hospital.

Table 1 shows a comparison of measures of predisposing, enabling, and need factors from the (weighted) sample of people aged 55 or older across the rural/urban categories. Residents of rural fringe areas of CMAs and rural areas outside of CMAs were actually a little younger on average than urban residents, and were also more likely to be married. CMA urban core areas had the highest proportion of residents born outside of Canada while rural non-CMA areas had the lowest. Perhaps not surprisingly, the biggest difference in socio-economic status across regions was between CMA areas (including urban core and rural

fringe) and non-CMA areas (including urban non-CMA and rural non-CMA), rather than between urban and rural areas. More than 20 per cent of CMA residents (both urban core and rural fringe) were in the highest income adequacy quintile for Canadian households, and more than 40 per cent of these residents also had university degrees. Comparable figures for regions outside of CMAs were less than 15 and 35 per cent respectively. Related to this, a greater percentage of CMA residents had health insurance to cover drugs, dental care, eye care, and hospital care than non-CMA residents although the gap is smaller for drug coverage than for the other forms of insurance.

Self-reported health appeared to be marginally better in the urban core and rural fringe of CMAs than in areas outside of CMAs. In contrast, the prevalence of certain chronic conditions (asthma, hypertension, diabetes, cancer, and heart disease) was somewhat higher in the urban core areas of CMAs than in other areas, and for none of the nine chronic conditions we considered was the prevalence in rural areas outside of CMAs higher than in the more populous urban areas. The prevalence of current daily smoking was lower in urban CMA areas than in other areas.

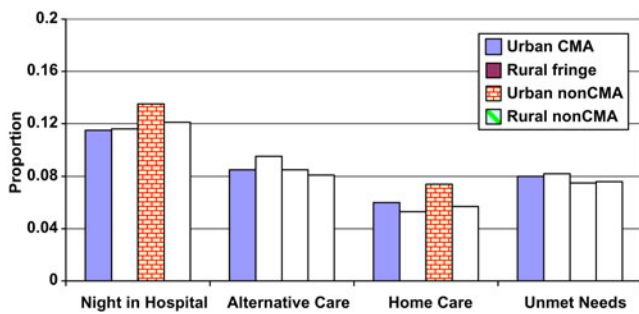


Figure 2: Proportion of older Canadians using health services in the past 12 months. A clear bar indicates not significantly different from Urban CMA at the 5 per cent level of significance. Results are for the adult non-institutional population aged 55 or older.

Regression Results

Table 2 presents selected results from a multivariate Logistic regression of each of the binary indicators of health service use expressed as a function of our full list of enabling, need, and predisposing factors. The first part of Table 2 reports results for the measures of rural/urban status, our main variables of interest. In column 1 of Table 2, it can be seen that other things being equal, residents of rural areas outside of CMAs were significantly less likely to have a GP (OR: 0.732;

Table 1: Descriptive statistics on predisposing, enabling, and need factors associated with the health service use of Canadians aged 55 or older by urban/rural classification of residence (percentages)

Explanatory Variables	Urban Core in CMA	Rural Fringe in CMA	Urban outside CMA	Rural outside CMA
Predisposing Factors				
Age (years)	67.2	65.5	67.7	66.6
Male	45.8	48.7	45.8	51.5
Married	67.0	79.2	69.4	74.6
Widowed/Separated/Divorced	27.7	17.7	26.8	20.7
Never Married	5.0	3.1	3.5	4.5
Immigrant	31.5	16.7	10.4	8.8
Enabling Factors				
Household Income Adequacy				
Lowest quintile	2.3	1.4	2.9	2.9
2nd quintile	6.7	5.8	9.5	8.7
Middle quintile	20.7	22.4	25.3	28.3
4th quintile	28.7	30.5	28.6	27.7
Highest quintile	22.1	21.7	14.9	13.3
Income not reported	19.6	18.2	19.0	19.2
Education				
Less than high school	33.0	34.3	45.4	47.8
High school graduate	17.2	16.0	12.9	11.8
Post-secondary	4.9	4.8	4.8	4.2
University	41.7	41.2	34.2	33.3
Education not reported	3.2	3.7	2.8	2.9
Private Insurance				
Pharmaceutical	79.0	77.0	76.9	73.3
Dental	43.2	42.4	33.5	31.0
Eye care	43.6	44.7	36.8	35.9
Hospital care	51.7	51.7	46.1	42.8
Need Factors				
Self-Reported Health (SRH)				
Excellent	14.9	15.7	12.7	12.5
Very good	27.8	25.3	28.2	28.0
Good	34.9	34.2	34.1	35.6
Fair	16.4	16.2	19.3	18.1
Poor	5.8	4.5	5.5	5.6
Chronic Conditions				
Asthma	9.1	7.8	7.8	8.0
Arthritis	43.4	41.2	44.8	41.6
Hypertension	40.7	34.3	37.9	36.9
Glaucoma	4.4	4.0	5.2	4.0
Diabetes	14.7	11.0	12.2	12.1
Cancer	6.1	4.9	4.4	4.2
Alzheimer's	1.4	0.9	1.3	1.1
Stroke	3.0	2.5	3.1	3.0
Heart disease	17.5	14.0	15.4	15.4
Smoking				
Current smoker	13.9	15.9	14.9	15.4
Former smoker	55.3	56.5	54.4	56.5
Never smoked	30.5	26.9	30.0	27.3
Composition of sample by age range				
Percentage of population aged 55 or older	69.1	6.7	7.2	14.6
Percentage of population aged 12 or older	71.9	6.6	6.8	12.4
Percentage of population aged 55–80	68.8	6.1	7.1	14.8

Notes. Data are drawn from the CCHS 2.1. The CCHS data exclude individuals who are resident in institutions, and on reserve. This table also excludes residents of Canada's territories.

For the percentage of the smoking population aged 55 or older, 12 or older, and between the ages of 55 and 80, the CCHS data differentiate further between urban core CMA and urban fringe CMA. Individuals classified as living in the urban fringe of a CMA are not included in Table 1 owing to small sample sizes, so that the percentages in the last three rows add to less than 1. Percentages for people resident in "urban fringe" are 2.5, 2.4, and 2.5 respectively.

Table 2: Multivariate Logistic regression results of the determinants of health service use and unmet health care needs in the past 12 months by Canadians aged 55 years or older (odds-ratios relative to urban core of a CMA)

Explanatory Variables	Has a GP		Has Visited a GP		Has Visited a Specialist		Has Visited a Dentist		Has Unmet Health Care Needs	
	OR	p value	OR	p value	OR	p value	OR	p value	OR	p value
Enabling Factors: Geography										
Urban/Rural (urban CMA = 1)										
Rural fringe of CMA	1.166	0.279	0.962	0.646	0.946	0.349	0.844	0.016	1.059	0.619
Urban non-CMA	0.806	0.043	0.955	0.602	0.803	0.001	0.894	0.077	0.996	0.968
Rural non-CMA	0.732	0.000	0.818	0.000	0.740	0.000	0.817	0.020	0.959	0.610
Province (Ontario = 1)										
Newfoundland	0.485	0.000	1.669	0.003	0.975	0.829	0.352	0.000	1.074	0.679
Nova Scotia	0.971	0.884	1.490	0.000	1.032	0.637	0.694	0.000	1.289	0.049
New Brunswick	0.525	0.000	1.135	0.171	1.139	0.179	0.542	0.000	1.188	0.240
Quebec	0.353	0.000	0.806	0.004	1.217	0.001	0.536	0.000	1.425	0.002
Manitoba	0.397	0.000	0.951	0.469	0.916	0.072	0.716	0.000	1.335	0.049
Saskatchewan	0.485	0.000	1.669	0.003	0.975	0.829	0.352	0.000	1.074	0.679
Alberta	0.545	0.000	0.999	0.984	0.712	0.000	0.625	0.000	1.023	0.846
BC	0.726	0.013	1.236	0.088	0.894	0.109	0.985	0.872	1.437	0.001
Socio-economic status										
Income Adequacy (highest = 1)										
Lowest Income quintile	0.970	0.868	0.713	0.029	0.705	0.001	0.339	0.000	1.128	0.465
2nd lowest quintile	0.654	0.021	0.591	0.000	0.674	0.000	0.358	0.000	0.922	0.535
Middle quintile	0.861	0.222	0.691	0.001	0.719	0.000	0.425	0.000	0.825	0.050
2nd highest quintile	0.979	0.832	0.823	0.089	0.867	0.015	0.652	0.000	0.950	0.597
Education Obtained (< HS = 1)										
High school graduate	1.090	0.392	1.334	0.000	1.285	0.001	1.620	0.000	0.981	0.839
Other post-secondary	0.932	0.612	1.406	0.006	1.460	0.000	1.785	0.000	1.196	0.154
University degree	1.020	0.820	1.303	0.000	1.534	0.000	2.347	0.000	1.497	0.000
Private Health Insurance (no = 1)										
Prescription drugs	1.301	0.004	1.341	0.000	1.342	0.000	1.032	0.513	0.865	0.098
Dental	1.187	0.084	1.052	0.424	0.906	0.018	1.872	0.000	0.884	0.167
Eye care	1.034	0.724	1.137	0.027	1.108	0.045	1.005	0.912	1.194	0.041
Private hospital room	1.109	0.270	0.983	0.840	1.041	0.436	1.113	0.017	0.931	0.304
Physicians in Health Region										
GPs per 1,000 pop.	1.005	0.043	1.000	0.816	0.998	0.182	0.996	0.011	0.995	0.041
Specialist MDs per 1,000 pop.	0.998	0.060	1.001	0.463	1.002	0.000	1.003	0.000	1.001	0.280
Need Factors										
SRH (excellent = 1)										
Very good	1.406	0.000	1.530	0.000	1.296	0.000	0.933	0.121	1.175	0.215
Good	1.580	0.000	1.939	0.000	1.722	0.000	0.821	0.000	1.966	0.000
Fair	1.651	0.000	2.437	0.000	2.400	0.000	0.738	0.000	3.273	0.000
Poor	1.549	0.050	2.192	0.000	3.399	0.000	0.606	0.000	6.369	0.000
Chronic Condition (no = 1)										
Asthma	1.666	0.000	1.237	0.252	1.185	0.008	0.872	0.038	1.202	0.032
Arthritis	1.440	0.000	1.716	0.000	1.379	0.000	1.067	0.049	1.436	0.000
Hypertension	2.282	0.000	2.074	0.000	1.090	0.030	1.010	0.729	0.895	0.083
Glaucoma	1.231	0.245	1.154	0.400	1.009	0.906	0.974	0.638	1.106	0.481
Diabetes	1.974	0.000	1.398	0.000	1.083	0.168	0.794	0.000	0.811	0.011
Cancer	1.081	0.701	2.277	0.000	3.894	0.000	1.033	0.663	1.248	0.049
Alzheimer's	1.245	0.671	1.676	0.055	1.140	0.607	1.088	0.490	0.852	0.583
Stroke	1.604	0.268	1.481	0.117	1.024	0.818	1.159	0.145	1.168	0.314
Heart disease	1.638	0.000	2.426	0.000	2.065	0.000	0.938	0.191	1.250	0.002
Smoking (never smoked = 1)										
Current smoker	0.590	0.000	0.642	0.000	0.846	0.001	0.432	0.000	1.263	0.027
Former smoker	1.001	0.990	1.080	0.197	1.199	0.001	0.933	0.039	1.066	0.375

Notes. Regression equations also include controls for predisposing conditions (age, gender, marital status, immigrant status). These results are not reported but are available on request from the authors.

Residents of the Territories and PEI are omitted from the estimating sample for data confidentiality reasons.

Bold font denotes statistical significance at the five per cent level.

The Wald Chi-squared test of the overall significance of the regression has a p value of 0.000 in each case.

Source of Data: Canadian Community Health Survey 2.1. Data on numbers of GPs and Specialists per 1,000 residents are from the Canadian Institute of Health Information and are included at the level of the health region.

p value: 0.000). This was also true of urban residents outside of CMAs (OR: 0.806; *p* value: 0.043), but it was not true of residents in the rural fringe of a CMA (OR: 1.166; *p* value: 0.279). In columns 2 to 4, comparable results show that rural non-CMA residents were significantly less likely than urban CMA residents to have visited a GP (OR: 0.818; *p* value: 0.000), to have visited a specialist (OR: 0.740; *p* value: 0.000), and to have visited a dentist (OR: 0.817; *p* value: 0.020). Thus, differences in demographic, socio-economic, and health status characteristics between rural and urban residents did not account for the differences in the likelihood of contact with doctors, specialists, and dentists. In contrast, column 5, shows that despite lower levels of health service use, there were no significant differences in the occurrence of unmet health care needs between rural and urban residents. The estimated odds ratios for the urban and rural categories were all very close to 1.0, with large *p* values. There were also no significant differences across rural and urban areas in the prevalence of having spent at least one night in hospital, in receiving home care, or in utilizing alternative methods of health care, and so these results were not reported.

It is also clear from Table 2 that all other things being equal, health service use varied significantly across provinces but not uniformly. Residents of all other provinces were less likely to have a family GP than residents of Ontario. Relative to Ontario, residents of Quebec were less likely to have visited a GP while residents of Newfoundland, Nova Scotia, and Saskatchewan were more likely. In contrast, Quebec residents were actually more likely to have visited a specialist, as were Alberta residents. Ontario residents were also more likely to have visited a dentist than residents of the other provinces. In terms of unmet health care needs, residents of Nova Scotia, Quebec, Manitoba, and BC were more likely to have such needs than Ontario residents.

Among the other enabling explanatory variables included in the regressions, we found that a higher number of GPs per 1,000 residents was associated with a greater likelihood of a person having a GP, and a higher number of specialists per 1,000 residents was associated with a greater likelihood of visiting a specialist in the past year. As well, a higher number of GPs per 1,000 residents was found to be associated with a lower prevalence of having unmet health care needs. Controls for both income adequacy and education level were highly significant and indicated clearly that the prevalence of having visited a GP, a specialist, or a dentist at least once during the previous year was higher for people with more education and people in households with a higher income adequacy. Odds ratios for education levels of high school graduate or higher were all significantly greater than 1.0, while

odds ratios for income adequacy levels below the top category were almost all significantly less than 1.0. The patterns were particularly pronounced for having visited a dentist, and it should be noted that these results were based on regressions that also controlled for various forms of medical insurance, including dental insurance.

Not unexpectedly, variables reflecting need factors were also significant determinants of health service use. Individuals with lower levels of self-reported health were more likely to have visited a GP and a specialist, and were more likely to have unmet health care needs. They were also less likely to have visited with a dentist. Similarly, many of the nine chronic conditions we considered were positively correlated with a person's visit to a GP or a specialist as well as to that person's having unmet health care needs, in particular heart disease, cancer, and arthritis. Controls for smoking had no significant effect on health service use after controlling for chronic conditions and the other factors.

In Table 3, we present unstandardized OLS regression results for the frequency of health service use, as measured by the number of GP visits, specialist visits, and dentist visits in the past year, conditional on at least one visit to each. Consistent with the results from Table 2, the frequency of visits to each medical professional was significantly lower for residents of rural non-CMA regions than for urban CMA residents: all other things being equal, rural residents had 6.3 per cent fewer visits to a GP⁹ (coefficient: -0.061; *p* value: 0.002), 8.0 per cent fewer visits to a specialist (coefficient: -0.077; *p* value: 0.006), and 5.3 per cent fewer visits to a dentist (coefficient: -0.052; *p* value: 0.003). Frequency of visits was also lower for residents of rural fringe CMA areas and for urban areas outside of CMAs, but the difference was not always significant at the five per cent level. Though not reported in Table 3, the results for the frequency of nights in hospital showed no significant variation across rural or urban regions.

Results for the province variables again gave a somewhat mixed picture, although generally it appeared that the frequency of visits to health care professionals was lower in other provinces than in Ontario. Further, older residents of Quebec had fewer visits to a GP, fewer visits to a specialist, and fewer visits to a dentist than residents of any other province, other factors being equal. The concentration of GPs per 1,000 residents in the health region was positively associated with the number of GP visits by individuals, while the concentration of specialists per 1,000 residents was positively associated with the number of specialist visits (though GPs per 1,000 was negatively associated with this measure).

Table 3: Multivariate OLS regression results of the determinants of the number of health service visits in the past 12 months, conditional on at least one visit, by Canadians aged 55 years or older

Explanatory Variables	Number of GP Visits		Number of Specialist Visits		Number of Dentist Visits	
	Coefficient	p value	Coefficient	p value	Coefficient	p value
Enabling Factors Geography						
Urban/Rural (urban CMA = 0)						
Rural fringe of CMA	-0.060	0.012	-0.106	0.008	-0.017	0.558
Urban non-CMA	-0.034	0.111	-0.058	0.069	-0.050	0.026
Rural non-CMA	-0.061	0.002	-0.077	0.006	-0.052	0.003
Province (Ontario = 0)						
Newfoundland	0.109	0.013	0.035	0.281	-0.084	0.181
Nova Scotia	0.029	0.447	-0.028	0.435	-0.042	0.049
New Brunswick	-0.050	0.027	-0.113	0.038	-0.073	0.048
Quebec	-0.305	0.000	-0.125	0.000	-0.144	0.000
Manitoba	-0.128	0.000	-0.034	0.074	-0.035	0.184
Saskatchewan	-0.020	0.368	-0.035	0.201	-0.128	0.000
Alberta	-0.035	0.047	0.059	0.003	-0.133	0.000
BC	0.007	0.758	0.030	0.255	-0.023	0.385
Socio-economic Status						
Income Adequacy (highest = 0)						
Lowest income quintile	0.025	0.624	0.118	0.194	-0.032	0.587
2nd lowest quintile	0.012	0.672	-0.038	0.342	-0.149	0.000
Middle quintile	0.043	0.027	0.069	0.020	-0.081	0.000
2nd highest quintile	0.014	0.425	0.032	0.099	-0.067	0.000
Education Obtained (< HS = 0)						
High school graduate	-0.026	0.236	0.037	0.237	0.030	0.164
Other post-secondary	0.034	0.188	0.087	0.087	0.105	0.000
University degree	0.017	0.263	0.074	0.000	0.074	0.000
Private Health Insurance (no = 0)						
Prescription drugs	0.083	0.000	0.046	0.079	0.021	0.233
Dental	-0.059	0.001	-0.015	0.542	0.062	0.000
Eye care	0.066	0.001	0.053	0.082	-0.027	0.021
Private hospital room	-0.008	0.644	-0.015	0.647	-0.006	0.666
Physicians in Health Region						
GPs per 1,000 pop.	0.000	0.664	-0.002	0.043	0.000	0.318
Specialist MDs per 1,000 pop.	0.001	0.000	0.001	0.000	0.000	0.640
Need Factors						
SRH (excellent = 0)						
Very good	0.173	0.000	0.033	0.269	0.017	0.264
Good	0.382	0.000	0.129	0.001	0.051	0.001
Fair	0.620	0.000	0.296	0.000	0.038	0.043
Poor	0.984	0.000	0.529	0.000	0.071	0.037
Chronic Condition (no = 0)						
Asthma	0.181	0.000	0.025	0.271	0.029	0.235
Arthritis	0.160	0.000	0.040	0.022	0.031	0.001
Hypertension	0.206	0.000	-0.023	0.200	0.005	0.610
Glaucoma	0.030	0.199	0.023	0.552	-0.011	0.571
Diabetes	0.202	0.000	0.086	0.005	0.020	0.128
Cancer	0.381	0.000	0.445	0.000	0.015	0.532
Alzheimer's	0.027	0.768	-0.081	0.456	-0.099	0.204
Stroke	0.129	0.000	-0.026	0.522	-0.052	0.289
Heart disease	0.226	0.000	0.030	0.102	-0.005	0.784
Smoking (never smoked = 0)						
Current smoker	-0.010	0.584	0.045	0.177	0.110	0.000
Former smoker	0.034	0.003	0.046	0.017	0.049	0.002

Notes. Regression equations also include controls for predisposing conditions (age, gender, marital status, immigrant status). These results are not reported but are available on request from the authors.

Residents of the Territories and PEI are omitted from the estimating sample for data confidentiality reasons.

Bold font denotes statistical significance at the five per cent level.

The F-test of the overall significance of each OLS regression has a p value of 0.000

Source of Data: Canadian Community Health Survey 2.1. Data on numbers of GPs and Specialists per 1,000 residents are from the Canadian Institute of Health Information and are included at the level of the health region.

The relationship between the frequency of medical visits and socio-economic status was less clear, although there was some indication that higher socio-economic status as measured by higher income and more education was associated with more visits. Need factors reflecting self-reported health status and chronic conditions were found to be strongly associated with the frequency of physician consultations, and in fact the number of GP visits and number of specialist visits both increased monotonically with decreasing levels of self-reported health.

It was not surprising that health services use was found to vary significantly across provinces since the provision of health services and the timing of health care rationalization can vary widely. For example, the provision of basic health care in Quebec is much more likely to be through community health centres than in other provinces (see Richard, Gauvin, Ducharme, Gosselin, Sapinski & Trudel, 2005). It is therefore quite possible that rural-urban differences may also vary by province of residence. To investigate this possibility, we next split the sample by province and estimated the same specification as before only run separately for each sub-sample of individuals. For this analysis, we defined five Canadian provincial groups because of confidentiality restrictions: the Atlantic Provinces, Quebec, Ontario, the Prairies (including Alberta), and British Columbia. For brevity, we only present results for the indicators of rural or urban residence although the regressions included all of the same explanatory variables we described earlier except for the province indicators. For regions that include multiple provinces, we included indicator variables for each province that is part of that provincial group.

Table 4 reports results for the same binary measures of health care use as in Table 2. From column one of Table 4, we see that older residents of rural areas outside of CMAs were significantly less likely to have a GP than comparable residents of urban CMAs for each provincial group except Quebec (though the odds ratio for rural Ontario was only significant at the 10% level [OR: 0.759; p value: 0.056]). Residents of smaller urban centers outside of CMAs in Atlantic Canada were also significantly less likely to have a GP (OR: 0.531; p value: 0.022). Results for whether a rural non-CMA resident had visited a GP in the past year show odds ratios less than 1.0 for all provincial groups, but it was only for Ontario that the difference was significant (OR: 0.760; p value: 0.019). Consulting with a specialist was significantly less likely for older residents of rural non-CMA areas in Atlantic Canada, Quebec, and Ontario, and it was also less likely for residents of urban areas outside of CMAs in Ontario, the Prairies, and British Columbia. The likelihood of visiting a dentist was lower for rural non-CMA residents in Ontario and the Prairies,

and for urban non-CMA residents in the Prairies and British Columbia. Overall, Table 4 shows that after controlling for other factors, health service use in rural areas was almost always lower than in urban CMA areas, and was never significantly higher than in urban CMA areas. In Table 4 as in Table 2, rural residents outside of CMAs, however, were no less likely than residents of urban CMAs to have unmet health care needs in any provincial group, despite differences in the prevalence of a visit to a GP or specialist. The odds ratio for rural residents of Ontario was significantly lower than 1.0 only at the 10 per cent level (OR: 0.771; p value: 0.052). There was also no significant difference between rural and urban areas in any provincial group for a night in hospital, for alternative care, and for home care, and so again these results were not reported.

Table 5 gives selected results for the same frequency measures as in Table 3, although again we report only those regression results for the rural and urban indicators. Table 5 shows that differences in GP visits between rural non-CMA regions and urban CMA regions were inconsistent across provincial groups, with significant differences only for rural residents of Atlantic Canada (13.7% fewer GP visits: coefficient: -0.128 ; p value: 0.044) and Ontario (9.1% fewer GP visits: coefficient: -0.087 ; p value: 0.013). Rural residents of Atlantic Canada also had 22.4 per cent fewer visits with a specialist than urban CMA residents (coefficient: -0.202 ; p value: 0.008), while rural residents of the Prairies had 14.8 per cent fewer visits with a specialist than urban CMA residents (coefficient: -0.138 ; p value: 0.036). There was no significant difference in the frequency of specialist visits between rural non-CMA and urban non-CMA regions for Quebec, Ontario, and British Columbia. Finally, although rural non-CMA residents overall had a lower frequency of dentist visits based on Table 2, it was only rural non-CMA residents of Ontario for whom this discrepancy was significant.

Discussion

The results clearly indicate that a number of important measures of health service use are lower among older Canadians living in rural areas outside of CMAs than among those living in the urban core of CMAs: rural residents are less likely to have a GP; to have visited a GP, a specialist or a dentist in the past year; and have significantly fewer visits with a GP, a specialist, or a dentist for those people who have had at least one visit during the year. These discrepancies between rural and urban residents are significant after controlling for a wide range of predisposing, need, and enabling factors, including the concentration of physicians and specialists at the level of the health region and the health of the individual. When the sample is disaggregated by province and the same equations estimated

Table 4: Selected multivariate Logistic regression results of the effect of rural/urban status on health service use in the past 12 months by Canadians aged 55 years or older; separate regressions by province of residence (odds-ratios relative to urban core of a CMA)

Explanatory Variables	Has a GP		Has Visited a GP		Has Visited a Specialist		Has Visited a Dentist		Has Unmet Health Care Needs	
	OR	p value	OR	p value	OR	OR	OR	p value	OR	p value
CMA Urban Core = 1										
Atlantic Provinces										
Rural fringe of CMA	1.337	0.470	1.260	0.462	0.807	0.244	0.836	0.343	0.963	0.893
Urban non-CMA	0.531	0.022	1.008	0.976	0.872	0.281	0.902	0.608	0.915	0.690
Rural non-CMA	0.492	0.002	0.854	0.496	0.783	0.050	0.814	0.306	0.870	0.471
Quebec										
Rural fringe of CMA	1.112	0.702	0.727	0.030	0.958	0.729	0.809	0.187	1.672	0.034
Urban non-CMA	1.248	0.286	1.204	0.292	0.855	0.341	0.844	0.199	1.386	0.117
Rural non-CMA	0.952	0.754	0.862	0.095	0.652	0.002	0.907	0.614	1.079	0.672
Ontario										
Rural fringe of CMA	1.140	0.508	1.110	0.327	0.952	0.596	0.890	0.165	1.051	0.765
Urban non-CMA	0.718	0.083	0.756	0.027	0.800	0.047	0.994	0.957	0.948	0.728
Rural non-CMA	0.759	0.056	0.760	0.019	0.737	0.005	0.782	0.009	0.771	0.052
Prairies										
Rural fringe of CMA	0.924	0.849	0.905	0.660	0.736	0.004	0.551	0.009	0.386	0.028
Urban non-CMA	0.726	0.139	0.942	0.777	0.745	0.003	0.758	0.002	0.687	0.084
Rural non-CMA	0.569	0.029	0.860	0.528	0.867	0.170	0.669	0.002	1.183	0.404
British Columbia										
Rural fringe of CMA	1.289	0.444	0.964	0.897	1.234	0.146	0.832	0.422	0.976	0.941
Urban non-CMA	0.958	0.896	1.243	0.422	0.697	0.015	0.721	0.040	0.796	0.350
Rural non-CMA	0.619	0.047	0.819	0.234	0.859	0.472	0.771	0.232	1.030	0.881

Notes. Equations are estimated separately for each of the provinces or group of provinces identified in the table. Each province-specific regression also includes controls for predisposing conditions (age, gender, marital status, immigrant status), predisposing factors (education level, household income quintile, type of private insurance, if any, concentration of GPs and specialist MDs per 1,000 residents), and need factors (self-assessed health, chronic conditions, smoking status, age started smoking). These results are not reported but are available on request from the authors.

The reference individual is a married male aged 60–64 with less than high school education, a household income in the highest quintile of income adequacy, no private health insurance, and living in the urban core of a Census metropolitan area in Ontario.

Residents of the Territories and PEI are omitted from the estimating sample for data confidentiality reasons.

Bold font denotes statistical significance at the five percent level.

The Wald Chi-squared test of the overall significance of the regression has a *p* value of 0.000 in each case.

Source of Data: Canadian Community Health Survey 2.1. Data on numbers of GPs and Specialists per 1,000 residents are from the Canadian Institute of Health Information and are included at the level of the health region.

separately for each provincial group, health services use in rural areas is almost always estimated to be less than in urban CMA areas (though not always significantly so) and is certainly not significantly greater than in urban CMA areas for any measure of health services in any provincial group. In contrast, there do not appear to be any differences between older residents of rural and urban areas in terms of unmet health care needs after controlling for other factors (a result consistent with the research of Wilson & Rosenberg, 2004). As well, the prevalence and frequency of nights in hospital are not any lower for older residents of rural areas compared to those in urban areas. (Allan & Cloutier-Fisher [2006] actually found higher rates of hospital stays for older residents of rural BC based on BC administrative data.)

That there are no discrepancies between rural and urban areas in terms of hospital nights or self-reported unmet health care needs is an encouraging result and suggests that barriers to access of health services may not be any more pronounced in rural areas, all other things being equal. However, possible reasons for having unmet health care needs are varied, and it may well be that unmet needs specifically related to access barriers do differ between rural and urban areas. For example, waiting times might be longer in rural areas or travel to visit specialists in urban areas might be more difficult. To investigate this possibility, we recast this variable to include only unmet health care needs arising from waiting times that were too long, services that were not available in the locality, the required distance to be traveled was too far, or services that

Table 5: Selected multivariate OLS regression results of the effect of rural/urban status on of the number of health service visits in the past 12 months by Canadians aged 55 years or older; separate regressions by province of residence (OLS estimates relative to urban core of a CMA)

Explanatory Variables	Number of Visits to a GP		Number of Visits to a Specialist		Number of Visits to a Dentist	
	Coefficient	p value	Coefficient	p value	Coefficient	p value
Atlantic Provinces						
Rural fringe of CMA	-0.118	0.028	-0.137	0.047	-0.006	0.945
Urban non-CMA	-0.124	0.022	-0.145	0.106	-0.032	0.475
Rural non-CMA	-0.128	0.044	-0.202	0.008	-0.076	0.186
Quebec						
Rural fringe of CMA	0.003	0.963	-0.031	0.561	-0.040	0.659
Urban non-CMA	0.017	0.677	-0.072	0.096	-0.028	0.585
Rural non-CMA	-0.030	0.227	-0.029	0.558	0.005	0.796
Ontario						
Rural fringe of CMA	-0.107	0.002	-0.140	0.023	-0.039	0.343
Urban non-CMA	-0.009	0.804	-0.017	0.756	-0.078	0.035
Rural non-CMA	-0.087	0.013	-0.052	0.229	-0.123	0.000
Prairies and Alberta						
Rural fringe of CMA	0.009	0.742	-0.150	0.274	0.096	0.242
Urban non-CMA	-0.045	0.242	-0.103	0.259	0.006	0.894
Rural non-CMA	0.006	0.877	-0.138	0.036	0.024	0.507
British Columbia						
Rural fringe of CMA	0.008	0.875	-0.051	0.699	0.013	0.824
Urban non-CMA	-0.064	0.455	0.129	0.342	-0.061	0.336
Rural non-CMA	-0.074	0.360	-0.017	0.759	-0.006	0.902

Notes. Equations are estimated separately for each of the provinces or group of provinces identified in the table. Each province-specific regression also includes controls for predisposing conditions (age, gender, marital status, immigrant status), predisposing factors (education level, household income quintile, type of private insurance, if any, concentration of GPs and specialist MDs per 1,000 residents), and need factors (self-assessed health, chronic conditions, smoking status, age started smoking). These results are not reported but are available on request from the authors.

The reference individual is a married male aged 60–64 with less than high school education, a household income in the highest quintile of income adequacy, no private health insurance, and living in the urban core of a Census metropolitan area in Ontario.

Residents of the Territories and PEI are omitted from the estimating sample for data confidentiality reasons.

Bold font denotes statistical significance at the five percent level.

The Wald Chi-squared test of the overall significance of each regression has a *p* value of 0.000 in each case.

Source of Data: Canadian Community Health Survey 2.1. Data on numbers of GPs and Specialists per 1,000 residents are from the Canadian Institute of Health Information and are included at the level of the health region.

could not be obtained in the time required.¹⁰ Notably, results were qualitatively the same as what we have reported. Specifically, there was no significant difference between rural and urban areas overall or for each provincial group with the sole exception of British Columbia where unmet needs were significantly more likely in rural non-CMA areas than urban CMA areas at the 5 per cent level.

Given that there are no widespread differences in the prevalence of unmet health care needs, it is of course possible that differences in health service use are not indicative of under-utilization relative to what is medically appropriate and may instead simply reflect differences in the interaction between patients and physicians in rural and urban areas. For example, if doctor

visits in rural areas are longer, more thorough, or address multiple medical complaints, then fewer visits might be required in rural areas for the same effective level of health care. However, the lower likelihood that a rural resident will visit a GP or a dentist *at all* during the year is of serious concern. Good health practices should involve annual check-ups with a GP and a dentist even if there are no apparent health problems, and this is particularly the case for older individuals where regular blood pressure and cholesterol and cancer screening tests, as well as other measures that take place in a doctor's office or require a doctor's referral, are strongly recommended. The true prevalence of unmet health care needs may be understated if individuals do not obtain basic health care and so may not be aware of conditions requiring treatment, such as high blood

pressure or diabetes. Similarly, the lower likelihood of visiting a specialist among rural residents after controlling for health status might also suggest access barriers since specialist visits are at the instigation of the family physician in the Canadian health system. These results for access to specialists are consistent with anecdotal evidence of the under-supply of specialist physicians in rural areas. Additional research on how health service use patterns react to specific health sector restructuring (such as Cloutier-Fisher & Joseph, 2000, and Liu et al., 2001) would be informative in this regard.

A number of caveats should be emphasized when considering these results. First, the use of health services among older Canadians at one point in time may give only a partial look at the extent of barriers in access even after controlling for physical health status. One reason is that barriers in access to health services at earlier ages can contribute to worse health later in life, leading to greater need for and reliance on health services by older individuals. Although some of our measures of health service use – such as visiting a doctor and a dentist at least once during the year – are recommended for all adults, more specific measures of preventive health service use among younger people – such as cancer screening – would provide a useful complement to the results we have reported here.

A more general caveat is that all information on health service use and health outcomes is self-reported in the CCHS. Allan and Cloutier-Fisher (2006) instead used administrative data on health care based on provincial Medicare records that do not suffer from recall bias or reporting errors. However, they were constrained by a relatively limited set of control variables so that the rural/urban indicators in their analysis may in fact reflect differences in predisposing, enabling, or need factors between rural and urban residents.

A third caveat is that despite the large sample size available in the CCHS, the sampling frame specifically excludes individuals living in institutions such as nursing homes. It is thus the case that the extent of health service use among older individuals will be under-stated, and if there are differences in the proportion of older individuals residing in institutions between rural and urban areas, this will affect our results. Related to this, if older individuals living in rural areas move to more urban areas to take advantage of what they believe will be more accessible health care, then again the discrepancies in health service use between rural and urban areas will be under-stated.

In this regard, we can get a preliminary sense of the potential problem by examining data of the migration decisions of older Canadians from the 2001 Census. Overall, 95 per cent of all people aged 65 or older and 92.1 per cent of all people aged 45 to 64 have the same

address as one year ago. Figures for remaining at the same address in the past five years are 81 and 71 per cent respectively. While migration is therefore not a common occurrence for older Canadians, it may still be the individuals in relatively poorer health who are more likely to move. Analysis of a longitudinal data set such as the National Population Health Survey could focus on the timing of a move *vis-à-vis* health status and would provide a useful complement to the work reported on in this article.

A final caveat is that, although more detailed than much previous work on the topic, our definition of rural areas outside of CMA/CAs, even differentiated by province, is still relatively broad. In this regard, Propper, Damiani, Leckie, and Dixon (2007) examined the links between distance traveled for in-patient treatment and socio-economic status using data from the UK, while McLean, Guthrie, and Sutton (2007) studied differences in the quality of primary medical care by remoteness from urban settlements.

Notwithstanding these caveats, if part of what underpins differences in basic health service use such as a visit to a GP relates to differences in perceptions of the need for health services, then policy makers may need to stress the importance of access to timely and preventive health care for people in rural areas through, for example, information campaigns about the importance of regular check-ups. One direction for future research that might inform this issue is the role of distance and remoteness in affecting basic preventive and diagnostic health service use. Finally, it should be noted that since the concentration of physicians in a health region is found to be a significant determinant of both health service use and unmet health care needs, the extent to which rural areas have lower concentrations of physicians than urban areas will only exacerbate the important differences in health service use that we have documented.

Notes

- 1 See Nagarajan (2004) for a summary of both reports as they relate to health care in rural and remote areas of Canada.
- 2 The effect of rationalization of service provision in rural areas depends on community and government responses to the changes and does not necessarily worsen the health status of residents of the affected community. For example, Liu, Hader, Broussart, White, & Lewis (2001) found that the closure of 52 small rural hospitals in Saskatchewan in 1993 did not adversely affect the health of local residents or their access to in-patient hospital services.
- 3 Recent work using nationally representative data includes Deri (2005), Wilson and Rosenberg (2004), and Newbold, Eyles, & Birch (1995). Since rural health service use is not the focus of this work, rural-urban differences in health service use are not explored beyond the inclusion of a single indicator variable for rural residence.

- 4 In a similar vein, a general conclusion of the literature on health outcomes and geography is that regional differences in health outcomes are small once differences in socio-economic factors and health-related behaviours are taken into account (Beaujot & Niu, 2005; Boyle & Willms, 1999; Buckley, Denton, Robb, & Spencer, 2006; Mitura & Bollman, 2003; Omariba & Rasugu, 2006; Rosenberg & Wilson, 2000; Shields & Tremblay, 2002; Tremblay, Ross, & Berthelot, 2002).
- 5 Household income adequacy is a Statistics Canada-derived variable that classifies each household into one of five categories (which they term quintiles) based on total household income and household size. (See http://www.statcan.gc.ca/imdb-bmdi/document/3226_D2_T9_V2-eng.pdf for a complete definition of the variable.) In our analysis, we include indicator variables for each income adequacy quintile.
- 6 Provinces often make different types of publicly funded health insurance available to seniors for services not covered by Medicare, but the availability of insurance is often means tested. For example, the province of New Brunswick covers most of the cost of pharmaceuticals for low-income individuals who are over the age of 65.
- 7 Dansky, Brannon, Shea, Vasey, and Dirani (1998) discussed the limitations of the dichotomous rural/urban classification that typically feature in US research. They utilized a five-category classification of counties based on population and adjacency to metropolitan areas that is broadly similar to what we used for this article. It is also notable that rural areas are typically defined much more narrowly in the US literature: a place is classified as rural if it is an unincorporated area or has fewer than 2,500 people.
- 8 Data by health region are available for 95 Canadian health regions but are suppressed for 28 other health regions due to confidentiality restrictions. Regression results are robust to how the missing data are addressed – omitting individuals residing in health regions with missing physician data yields estimation results comparable to estimating over the full range of health regions but including a binary variable that takes the value one if the data are missing and zero otherwise. Results reported here were based on the first approach.
- 9 The percentage change for a discrete variable is given by $[\exp(b)-1] \times 100$, where b is the estimated coefficient.
- 10 Other reasons cited include the cost of the service; the person was too busy; the person did not get around to it; language barriers; and personal/family reasons.

References

- Allan, D., & Cloutier-Fisher, D. (2006). Health service utilization among older adults in British Columbia: Making sense of geography. *Canadian Journal on Aging*, 25(2), 219–232.
- Andersen, R. (1968). *A behavioral model of families' use of health services*. Research Series No. 25. Chicago: Center for Health Administration Studies, University of Chicago.
- Andersen, R. (1995). Revisiting the behavioural model and access to medical care: Does it matter? *Journal of Health and Social Behaviour*, 36, 1–10.
- Andersen, R., & Newman, J. (1973). Societal and individual determinants of medical care utilization in the United States. *Milbank Quarterly*, 51(1), 95–124.
- Beaujot, R., & Niu, J. (2005). *Aggregate level community characteristics and health*. Ottawa: Population Studies Centre Discussion Paper 05-14. University of Ottawa.
- Boyle, M., & Willms, D. (1999). Place effects for areas defined by administrative boundaries. *American Journal of Epidemiology*, 149(6), 577–585.
- Buckley, N., Denton, F., Robb, C., & Spencer, B. (2004). The transition from good to poor health: An econometric study of the older population. *Journal of Health Economics*, 23(5), 1013–1034.
- Buckley, N., Denton, F., Robb, C., & Spencer, B. (2006). Socio-economic influences on the health of older Canadians. *Canadian Public Policy*, 32(1), 59–81.
- Case, A., Fertig, A., & Paxson, C. (2005). The lasting impact of childhood health and circumstance. *Journal of Health Economics*, 24(2), 365–389.
- Chen, J., Hou, F., Sanmartin, C., Houle, C., Tremblay, S., & Berthelot, J. (2002). Unmet health care needs. *Canadian Social Trends*, 11–008 18–22.
- Cloutier-Fisher, D., & Joseph, A. (2000). Long-term care restructuring in rural Ontario: Retrieving community service user and provider narratives. *Social Science and Medicine*, 50(7-8), 1037–1045.
- Dansky, K., Brannon, D., Shea, D., Vasey, J., & Dirani, R. (1998). Profiles of hospital, physician, and home health care service use by older persons in rural areas. *The Gerontologist*, 38(3), 320–330.
- Deri, C. (2005). Social networks and health service utilization. *Journal of Health Economics*, 24(6), 1076–1107.
- Escarce, J., Shea, J., & Chen, W. (1997). Segmentation of hospital markets: where do HMO enrollees get care? *Health Affairs*, 16(6), 181–192.
- Fakhoury, W., & Roos, L. (1996). Access to and use of physician resources by the rural and urban populations in Manitoba. *Canadian Journal of Public Health*, 87(4), 248–252.
- Fuchs, V. (2004). Reflections on the socio-economic correlates of health. *Journal of Health Economics*, 23(4), 653–661.
- Hurd, M., & McGarry, K. (1997). Medical insurance and the use of health care services by the elderly. *Journal of Health Economics*, 16(2), 129–154.
- Jennissen, T. (1992). *Health issues in rural Canada*. Ottawa: Political and Social Affairs Division BP-325E, Government of Canada.
- Lassey, M., Lassey, W., & Jinks, M. (1997). *Healthcare systems around the world: characteristics, issues, reforms*. Upper Saddle River, NJ: Prentice Hall.

- Laurent, S., (2002). *Rural Canada: Access to health care*. Ottawa: Parliamentary Research Branch Occasional Paper. Government of Canada. Retrieved from <http://dsp-psd.pwgsc.gc.ca>
- Liu, L., Hader, J., Broussart, B., White, R., & Lewis, S. (2001). Impact of rural hospital closures in Saskatchewan, Canada. *Social Science and Medicine*, 52(12), 1793–1804.
- Martin-Matthews, A. (2002). *Seniors' health. Sharing the learning: The health transition fund*. Ottawa: Synthesis series, Health Canada.
- McLean, G., Guthrie, B., & Sutton, M. (2007). Differences in the quality of primary medical care services by remoteness from urban settlements. *Quality and Safety in Health Care*, 16(6), 446–449.
- Mitura, C., & Bollman, R. (2003). The health of rural Canadians: A rural-urban comparison of health indicators (Catalogue No. 21-006-XIE). *Rural and Small Town Canada Analysis Bulletin*, 4(6). 1–23.
- Nagarajan, K., (2004). Rural and remote community health care in Canada: Beyond the Kirby Panel Report, the Romanow Report and the federal budget of 2003. *Canadian Journal of Rural Medicine*, 9(4), 245–251.
- Nelson, C., & Park, J. (2006). The nature and correlates of unmet health care needs in Ontario, Canada. *Social Science and Medicine*, 62(9), 2291–2300.
- Newbold, B., Eyles, J., & Birch, S. (1995). Equity in health care: Methodological contributions to the analysis of hospital utilization within Canada. *Social Science and Medicine*, 40(9), 1181–1192.
- Omariba, D., & Rasugu, W. (2006). *Neighbourhood characteristics, individual and household attributes and health perception among elderly Canadians*. Ottawa: Population Studies Centre Discussion Paper 06-01, University of Ottawa.
- Peckins, C. (2006). *Annual doctor visits: Many benefits*. Cambridge, MA: Mount Auburn Hospital Review, Winter Edition, Mount Auburn Hospital, 1–3.
- Peterson, S., Shapiro, E., & Roos, N. (2005). Regional variation in home care use in Manitoba. *Canadian Journal on Aging*, 24(Suppl. 1), 69–80.
- Propper, C., Damiani, M., Leckie, G., & Dixon, J. (2007). Impact of patients' socioeconomic status on the distance travelled for hospital admission in the English National Health Service. *Journal of Health Services Research & Policy*, 12, 153–159.
- Richard, L., Gauvin, L., Ducharme, F., Gosselin, C., Sapinski, J.P., & Trudel, M. (2005). Health promotion and disease prevention for older adults: Intervention themes and strategies used in Quebec local community health centres and seniors' day centres. *Canadian Journal of Public Health*, 96(6), 467–470.
- Romanow, R. (2001). *Building our values: The future of health care in Canada*. Ottawa: Commission on the Future of Health Care in Canada, Government of Canada.
- Rosenberg, M., & James, A. (2000). Medical services utilization patterns by seniors. *Canadian Journal on Aging*, 19(Suppl. 1), 125–142.
- Rosenberg, M., & Wilson, K. (2000). Gender, poverty and location: How much difference do they make in the geography of health inequalities. *Social Science and Medicine*, 51(2), 275–587.
- Shields, M., & Tremblay, S. (2002). The health of Canada's communities. *Supplement to Health Reports*, 13, 1–24.
- Statistics Canada. (2006). *Geographic units*. Ottawa: Statistics Canada. Retrieved August 20, 2006, from <http://www12.statcan.ca/english/census01/products/reference/dict/geo049.html>
- Tremblay, S., Ross, N., & Berthelot, J.-M. (2002). *Regional socio-economic context and health*. Ottawa: Supplement to Health Reports 13, Statistics Canada, Catalogue 82-003.
- Van Houtven, C., & Norton, E. (2004). Informal care and health care use of older adults. *Journal of Health Economics*, 23(6), 1159–1180.
- Wilson, K., & Rosenberg, M. (2004). Accessibility and the Canadian health care system: Squaring perceptions and realities. *Health Policy*, 67(2), 137–148.
- Wolinsky, F. (1994). Health services utilization among older adults: Conceptual, measurement and modeling issues in secondary analysis. *The Gerontologist*, 34(4), 470–475.
- Wolinsky, F., & Johnson, R. (1991). The use of health services by older adults. *Journal of Gerontology: Social Sciences*, 46(6), S345–S357.