

# Microsatellite variation of mussels (*Mytilus galloprovincialis*) in central and eastern Mediterranean: genetic panmixia in the Aegean and the Ionian Seas

IOANNIS A. GIANTSIS<sup>1</sup>, NADIA MUCCI<sup>2</sup>, ETTORE RANDI<sup>2</sup>, THEODORE J. ABATZOPOULOS<sup>3</sup>  
AND APOSTOLOS P. APOSTOLIDIS<sup>1</sup>

<sup>1</sup>Laboratory of Ichthyology and Fisheries, Department of Animal Production, Faculty of Agriculture, Aristotle University of Thessaloniki, Thessaloniki, 54124, Greece, <sup>2</sup>Laboratorio di Genetica, Istituto Superiore per la Protezione e la Ricerca Ambientale (ISPRA), Via Ca Fornacetta 9, I-40064 Ozzano dell'Emilia (Bologna), Italy, <sup>3</sup>Department of Genetics, Development and Molecular Biology, School of Biology, Aristotle University of Thessaloniki, Thessaloniki, Greece

*Genetic variation of mussels Mytilus galloprovincialis in central–eastern Mediterranean Sea is investigated in this study. A total of 550 individuals sampled from two cultured and 11 wild populations from Italy, Croatia, Greece and Turkey were genotyped at 10 microsatellite loci. Significant deviations from Hardy–Weinberg expectations were observed in more than 75% of the tests performed. All populations showed extensive heterozygote deficits, which remained at significant levels even after correction for null alleles, providing evidence that null alleles were only partly responsible for deviations from Hardy–Weinberg equilibrium in these molluscs. Moreover, null alleles seem to have limited influence on the population genetic differentiation. Similar levels of multi-locus heterozygosity and allelic richness were observed in all populations, cultured and wild, implying the sustainability of the exploited populations. Lack of isolation by distance and markedly low genetic differentiation between the nine Greek sampling sites (shoreline >1000 km) was revealed by Mantel tests,  $F_{ST}$  values, exact tests and analyses of molecular variance, indicating that mussels from these regions are either at or close to panmixia. Similarly, patterns of genetic homogeneity were also found between the two Italian samples, whereas the observed genetic differentiation of the populations from Turkey and Croatia probably reflects the specific topographic and oceanographic conditions of these regions.*

**Keywords:** genetic structure, heterozygote deficiency, microsatellites, Mediterranean, *Mytilus galloprovincialis*, panmixia

Submitted 6 December 2013; accepted 21 January 2014; first published online 11 March 2014

## INTRODUCTION

The Mediterranean mussel *Mytilus galloprovincialis* (Lamarck, 1819) is widespread along the Mediterranean coasts, from where it is supposed to originate (Riginos & Cunningham, 2005), while outside of these regions, it also occurs at the Atlantic coasts of Western Europe hybridizing with *M. edulis* (Hilbish *et al.*, 2002; Bierne *et al.*, 2003; Kijewski *et al.*, 2011). Moreover, it has been widely translocated around the world by human activity, mainly for aquaculture purposes (Westfall & Gardner, 2010; Smietanka *et al.*, 2013), as mussels are nowadays one of the most popular cultured species (Karayücel *et al.*, 2013). Italy is the predominant country in mussel culture within the Mediterranean Sea, employing more than 250 companies and reaching 123,000 tons annually (Parisi *et al.*, 2012). In Greece, the annual production in 2008 was 36,000 tons, with a continuous trend of expansion by licensing new farming sites (Theodorou

*et al.*, 2011). In other coastal countries of the eastern Mediterranean, although mussel culture has not reached yet such high levels—for example, Turkey and Croatia produce around 1000 tons  $\text{yr}^{-1}$  (Oraic & Zrncic, 2005; Karayücel *et al.*, 2010)—there is a very good potential for further development. Further, beyond their commercial interest, mussels are considered organisms of great ecological importance (Vidal *et al.*, 2009).

Despite the great economic and biological importance of the species, the genetic structure of Mediterranean mussel populations from the central–eastern Mediterranean Sea is essentially unknown. Yet, such information is crucial for the development of appropriate strategies for management of both cultured and wild populations. In fact, existing management models may sometimes be refined, taking into account new observations regarding the genetic composition of the species (Graves, 1998). The few genetic studies that have been published so far have not yielded a clear outcome about the genetic status of mussel populations from the above area. The scopes of those studies were either limited to comparisons of a few local populations such as those from the north Aegean Sea using mtDNA and allozyme analyses (see Karakousis & Skibinski, 1992; Kravva *et al.*, 2000),

**Corresponding author:**  
A.P. Apostolidis  
Email: apaposto@agro.auth.gr

from the Venice Lagoon using allozymes (Venier *et al.*, 2003) and from Croatian coasts using microsatellites (Stambuk *et al.*, 2013), or yield controversial findings. For instance, while a mtDNA based typing survey found genetic homogeneity among Aegean mussel populations (Ladoukakis *et al.*, 2002), a recent study based on RAPD markers (Giantsis *et al.*, 2012) revealed significant levels of genetic heterogeneity among them. However, neither of these two studies involved microsatellite markers, which are supposed to be more informative in resolving spatial genetic structures (Loughheed *et al.*, 2000). Notably, Kijewski *et al.* (2011), who used nuclear and mtDNA markers to investigate the genetic structure of *Mytilus* species on a large European scale, did not include in their analysis samples from the Aegean, Ionian and Adriatic Seas.

Microsatellites are considered to be one of the most useful molecular markers for addressing questions in population genetics (Hauffe & Sbordoni, 2009). Nonetheless, these markers have their own limitations, due to homoplastic alleles, allelic drop-out and null alleles (Van Oosterhout *et al.*, 2004; Pompanon *et al.*, 2005). A null allele can be defined as any allele that fails to be amplified during the polymerase chain reaction (PCR) by a given pair of primers (Dakin & Avise, 2004). Microsatellite null alleles have been detected in a variety of species, but they occur with unusual high frequencies in some groups, particularly in molluscs (Carlsson *et al.*, 2008; Lemer *et al.*, 2011). Highly frequent null alleles may be caused by frequent single nucleotide polymorphisms in flanking non-coding regions, as a consequence of the huge effective population sizes of marine bivalves (Lallias *et al.*, 2009). Marine bivalves are also characterized by high levels of heterozygote deficiency (Vidal *et al.*, 2009), which

has been predominantly attributed to the presence of null alleles (Launey *et al.*, 2002; Lemer *et al.*, 2011).

The present study was designed to clarify the genetic structure of wild and cultured populations of *M. galloprovincialis* sampled from the central-eastern Mediterranean Sea and genotyped using microsatellite markers. Furthermore, we attempted to evaluate the influence of null alleles on the genetic variability of mussels, as well as the impact of anthropogenic transplantations on the population differentiation and genetic make-up of these populations.

## MATERIALS AND METHODS

### Sample collection and DNA extraction

A total of 550 *M. galloprovincialis* individuals were collected by hand or diving, from 13 localities across the coasts of the central-eastern Mediterranean Sea, originating from Italy, Croatia, Greece and Turkey (Figure 1; Table 1). The selected sampling sites are representatives of the greatest part of *M. galloprovincialis* geographical distribution in the above area, including the Ligurian, Adriatic, Ionian, Aegean and Marmara Seas. Of particular importance in studying the genetic structure of Mediterranean mussels is the Gulf of Thermaikos, located at the northern part of the Aegean Sea. This is a semi-enclosed basin plentiful in organic material due to the discharge of approximately  $150 \text{ m}^3 \text{ s}^{-1}$  water by three rivers and with a maximum depth of 45 m, the west part of which comprises the most important shellfish cultivating area (Koukaras & Nikolaidis, 2004), offering excellent conditions for growth and preservation of mussel stocks

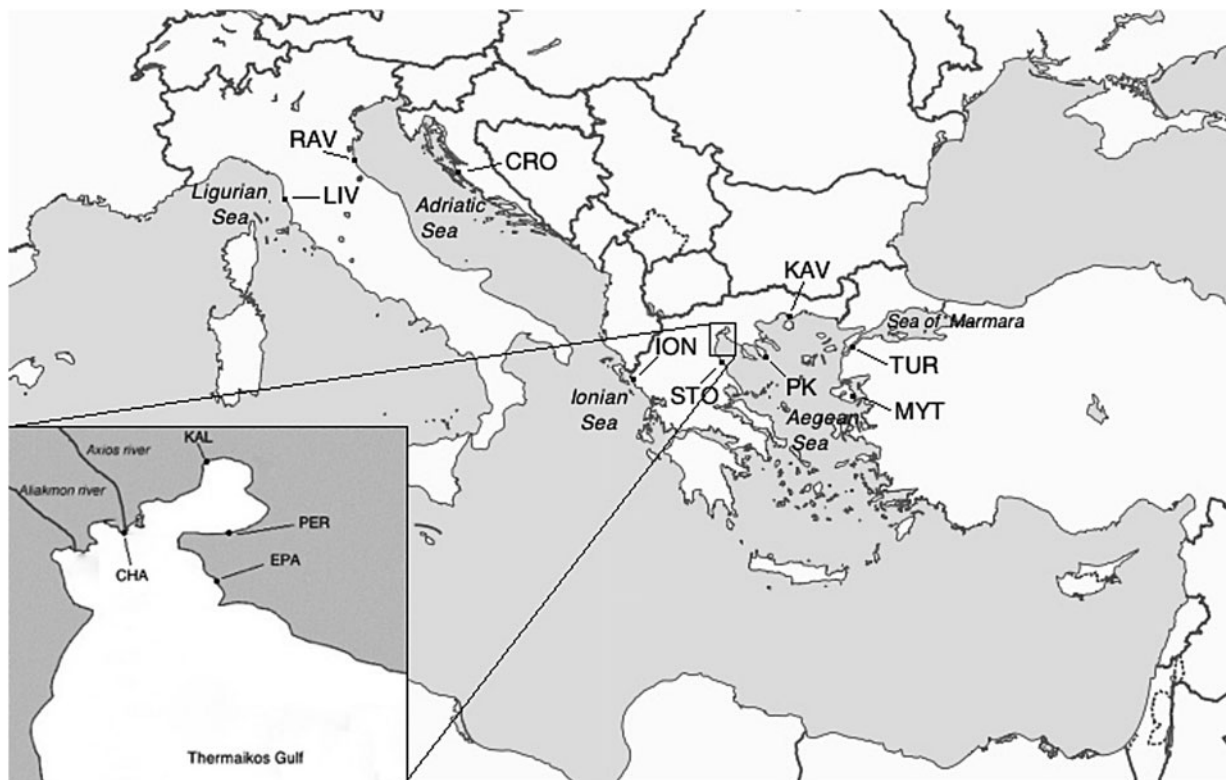


Fig. 1. Sampling locations of *Mytilus galloprovincialis* in the central-eastern Mediterranean: code explanations and additional information for each population are reported in Table 1.

**Table 1.** Sampling sites' additional information.

| No. | Population  | Sample size | Code | Status   | Basin                           | Country | Sampling time |
|-----|-------------|-------------|------|----------|---------------------------------|---------|---------------|
| 1   | Chalastra   | 45          | CHA  | Cultured | Northern Aegean—Thermaikos Gulf | Greece  | February 2009 |
| 2   | Kalohori    | 45          | KAL  | Wild     | Northern Aegean—Thermaikos Gulf | Greece  | February 2009 |
| 3   | Peraia      | 43          | PER  | Wild     | Northern Aegean—Thermaikos Gulf | Greece  | February 2009 |
| 4   | Epanomi     | 45          | EPA  | Cultured | Northern Aegean—Thermaikos Gulf | Greece  | February 2009 |
| 5   | Stomio      | 45          | STO  | Wild     | Western Aegean Sea              | Greece  | February 2009 |
| 6   | Porto Koufo | 45          | PK   | Wild     | Northern Aegean Sea             | Greece  | February 2009 |
| 7   | Kavala      | 45          | KAV  | Wild     | Northern Aegean Sea             | Greece  | March 2009    |
| 8   | Mytilene    | 45          | MYT  | Wild     | Eastern Aegean Sea              | Greece  | March 2009    |
| 9   | Canakkale   | 45          | TUR  | Wild     | Sea of Marmara                  | Turkey  | November 2008 |
| 10  | Igoumenitsa | 45          | ION  | Wild     | Ionian Sea                      | Greece  | February 2009 |
| 11  | Zadar       | 36          | CRO  | Wild     | Eastern Adriatic Sea            | Croatia | October 2010  |
| 12  | Ravenna     | 36          | RAV  | Wild     | Western Adriatic Sea            | Italy   | March 2011    |
| 13  | Livorno     | 30          | LIV  | Wild     | Ligurian Sea                    | Italy   | March 2011    |

(Arsenoudi *et al.*, 2003). Thus, four of the samples were located within this Gulf, with two of them representing cultured populations (Chalastra and Epanomi) that were collected with the permission of the owners of the mussel farms. The remaining nine samples were collected from wild stocks (Table 1), none of which was part of a national park, privately owned area or protected in any other way, thus no specific permission was required. Mussels were transported alive to the laboratory where a small piece of the mantle tissue was removed from each specimen and used for DNA extraction. The standard phenol extraction protocol of Hillis *et al.* (1996) was applied for isolation of total genomic DNA and the extracted DNA was quantified by Infinite<sup>®</sup> 200 PRO NanoQuant spectrophotometer (TECAN). All individuals were taxonomically identified to belong to *M. galloprovincialis*, using the nuclear DNA marker Me15/17 (Bierne *et al.*, 2003).

### Selection of microsatellite primers

In order to find the most appropriate loci for the genetic analysis of mussel populations, 22 polymorphic microsatellite primer pairs, all specific for the three recognized species of closely related *Mytilus*, that is, *M. galloprovincialis*, *M. edulis* and *M. trossulus* (Beaumont *et al.*, 2008), were chosen from the literature and tested on 20–30 randomly selected individuals. These were: Mgμ1, Mgμ2, Mgμ3, Mgμ4, Mgμ5, Mgμ6, (Presa *et al.*, 2002), MGE001, MGE002, MGE005, MGE006, MGE007, MGE008 (Yu & Li, 2007), Mg181, Mg220 (Varela *et al.*, 2007), Med362, Med367, Med737, Med740, Med744 (Lallias *et al.*, 2009), MT203 and MT282 (Gardeström *et al.*, 2008). Amplicons were initially checked in agarose gels and afterwards on an ABI-Prism 3130xL automatic sequencer (Life Technologies, Foster City, USA). Based on the clarity and reproducibility of the banding patterns, 10 of these primers (Mg181, MT203, MT282, MGE001, MGE005, MGE006, MGE008, Med367, Mgμ3 and Mgμ7) were selected for further screening.

### Tailed PCR method and PCR reaction conditions

Forward primers were synthesized with the addition of a short tail at the 5' end. This tail is complementary to a fluorescent labelled oligonucleotide sequence joining the primer during

the PCR reaction. Three tails labelled with different fluorescent labels were used to enable the analysis on the ABI-Prism 3130xL automatic sequencer. This procedure, described by Schuelke (2000), decreases the total cost of the method, as it avoids the need of a large number of primers carrying a fluorescent dye label. Each PCR mixture consisted of 25–50 ng of template DNA, one unit of *Taq* DNA polymerase, 1 μl 10 × *Taq* DNA polymerase buffer including 1.5 mM MgCl<sub>2</sub>, 2 μg BSA, 0.2 mM each of dNTP, 1 pmol unlabelled forward primer, 10 pmol unlabelled reverse primer and 10 pmol labelled tail primer in 10 μl of total reaction volume. Amplification reactions were carried out in an ABI 9700 thermocycler (Life Technologies) consisting of an initial denaturation for 2 min at 94°C, followed by: (a) 10 cycles of 40 s at 94°C, 90 s at a temperature depending on the primer with a touchdown procedure (decreasing 0.5 or 1°C per cycle) and 60 s at 72°C; (b) 35 cycles of 40 s at 94°C, 90 s at the lowest temperature value of the touchdown scale and 60 s at 72°C, for the amplification of the target sequence; (c) 10 cycles of 40 s at 94°C, 90 s at 50°C and 60 s at 72°C, for the hybridization of the tailed primers; and (d) 10 min at 72°C for a final extension. All reactions were performed in a 96-well plate using always at least one negative sample (no DNA in PCR) and all PCR products were analysed in the ABI-Prism 3130xL automatic sequencer.

### Data analysis

Alleles were scored using the software GENEMAPPER v.4.0 (Applied Biosystems) and named according to their nucleotide size. Input files for several statistical computer programs were created using CREATE v.2 (Coombs *et al.*, 2008). The software MICROCHECKER (Van Oosterhout *et al.*, 2004) was used to investigate the presence of PCR errors like stuttering, large allele dropouts and null alleles. As analysis with MICROCHECKER indicated strong evidence of null alleles at all loci, allele frequencies for each locus and population were estimated by the expectation maximization algorithm (Dempster *et al.*, 1977) implemented in the package FreeNA. This software also provides a correction method for the bias of null alleles on F<sub>ST</sub> estimation (Chapuis & Estoup, 2007). Since the reliability of this method depends on its capacity to integrate the missing data, only samples with seven or more successful loci amplifications were included in the genetic structure analysis.

The number of effective alleles (allelic richness) per locus in each population (standardized to a minimum of 26 genes per population) was estimated without frequency adjustments for null alleles using HP-RARE v.1.0 (Kalinowski, 2005). Deviations from Hardy–Weinberg equilibrium (HWE) were estimated following Hedrick's (2000) procedures implemented in GENALEX v.6.41 (Peakall & Smouse, 2006), whereas linkage disequilibrium and allele frequency differences among populations or between pairs of populations were estimated by Fisher exact tests using GENEPOP v.4.1.0 (Raymond & Rousset, 1995). Expected heterozygosity ( $H_e$ ), observed heterozygosity ( $H_o$ ) and inbreeding coefficient ( $F_{IS}$ ) were computed with and without correction for null alleles using GENETIX (Belkhir *et al.*, 2004). Corrected  $H_e$ ,  $H_o$  and  $F_{IS}$  for each population was computed taking account only loci with null allele frequencies lower than 0.2, as these loci are not considered as potentially problematic for the analysis (Dakin & Avise, 2004; Chapuis & Estoup, 2007; Lawson Handley *et al.*, 2007). Since significant heterozygote deficiency is frequent in invertebrates, and in our case-study it did not disappear even after the withdrawal of loci with high proportions of null alleles, a test for self-fertilization was carried out by the calculation of selfing rate for each population with the program RMES (David *et al.*, 2007). This computer program is not sensitive to null alleles and scoring errors.

An analysis of molecular variance (AMOVA) was carried out to partition genetic variation among and within groups using ARLEQUIN v.3.1 (Excoffier *et al.*, 2005) and their significance was assessed by 10,000 permutations. Population structure was also investigated by the calculation of Reynolds *et al.*, (1983) genetic distance with 1000 bootstrap permutations, harbouring null alleles after the including null alleles (INA) correction method of allele frequencies (Chapuis & Estoup, 2007). Values of genetic distance computed by this method were then compared with those that had not been corrected for null alleles, computed by GENDIST v.3.69 in the PHYLIP package (Felsenstein, 2005). In order to visualize these comparisons two UPGMA dendrograms were constructed by the software TREEVIEW (Page, 1996). Discriminant analysis of principal component (DAPC) was computed using the ADEGENET (Jombart, 2008) in R (<http://cran.r-project.org/web/packages/fields/index.html>). A total of 100 principal components and five discriminant functions were retained for the computations. The genetic relationships among populations were also visually evaluated through a principal component analysis (PCA) carried out on gene frequency data with the program PCAGEN 1.2.1 (Goudet, 2005).  $F_{ST}$  values among populations were computed before and after correction for null alleles using the excluding null alleles (ENA) method (Chapuis & Estoup, 2007) implemented in FREENA. Moreover,  $\Phi_{ST}$  values (analogous to  $F_{ST}$ ), were estimated using the software GENOTYPE (Meirmans & Van Tienderen, 2004). Finally, a Mantel test embedded in GENALEX v.6.41 (Peakall & Smouse, 2006) was conducted to examine the isolation by distance model (IBD), that is, to correlate geographical and genetic distances. Its significance was assessed by 999 random permutations.

As both the DAPC diagram and pairwise exact tests suggested genetic homogeneity among discrete samples, the software POWSIM v.4.1 (Ryman & Palm, 2006) was used to assess this presumption. This is a simulation-based method,

which measures the probability of incorrectly rejecting the hypothesis of no population structuring and the statistical power for detecting genetic differentiation with a  $\chi^2$  and a Fisher exact test. The tests were performed both for the global scenario (all samples) and for the Aegean Sea (8 populations) using adjusted (by FreeNA) allele frequencies. A thousand simulations were run applying combinations of  $N_e$  and  $t$  leading to  $F_{ST}$  values of 0.001, 0.0025, 0.005 and 0.01. Number of dememorizations, batches and iterations were set as default—that is, 1000, 100 and 1000 respectively.

## RESULTS

### Data quality

No evidence for large allelic dropout or scoring errors due to stuttering was found by MICROCHECKER. However, both MICROCHECKER and FreeNA revealed evidence for the presence of null alleles in all loci. Null allele frequencies for each locus and population varied between 0.00 and 0.34 (Table 2) as computed by the EM algorithm, so genetic structure was analysed with and without correction for null alleles. Given that null alleles could be attributed to point mutations of primers binding sites (Chapuis & Estoup, 2007), primers designed to amplify target sequences from closely related species could have potential disadvantages. Nevertheless, in our study despite the use of three non-specific primers (Med367, MT203 and MT282) the two loci that exhibited an average null allele frequency greater than 0.2 (Mg181 and Mgu7 with 0.20952 and 0.22123, respectively) were those designed specifically for *M. galloprovincialis*. Further, due to the influence of null alleles, through an increase of  $F_{ST}$  values on the genetic differentiation between population pairs of small size (<20 individuals, White *et al.*, 2011), all mussel populations utilized in this study consisted of at least 30 individuals (Table 1). Therefore, as indicated by comparisons of corrected and non-corrected  $F_{ST}$  and genetic distance, although null alleles have been detected in all loci, they did not seem to affect the results of interpopulation genetic differentiation.

### Genetic and genotyping diversity

With the exception of MGE006 locus, which was monomorphic, substantial levels of polymorphism were observed. Table 2 shows the basic parameters of genetic diversity estimated in each population. The overall number of alleles per locus ranged from 5 at Mg181, to 31 at Mgu7, with a mean value of 11.8. Although there were cases of fewer alleles for particular loci across populations, average allelic richness did not differ highly among the 13 geographical regions, with the mean value ranging from 4.33 in the Croatian population to 5.52 in that of Livorno. Further, approximately equal values of observed and expected heterozygosity were revealed among the 13 populations (Table 2), indicating a pattern of similar genetic diversity in all populations. Instead, a strong discrepancy between observed and expected heterozygosity was detected in all populations, albeit these values were reduced when recalculated using the correction for null alleles (Table 2). This heterozygote deficit was also reflected in  $F_{IS}$  values and was directly related to multiple deviations from HWE (in 91 out of 117 tests performed; Table 2).

**Table 2.** Genetic diversity parameters of thirteen *Mytilus galloprovincialis* populations from central–eastern Mediterranean estimated by 10 microsatellites. Number of alleles per locus (N), allelic richness ( $A_R$ ), expected ( $H_E$ ) and observed heterozygosity ( $H_O$ ), inbreeding coefficient ( $F_{IS}$ ),  $P$  values for tests of departure from HWE ( $P_{HWE}$ ) and the frequency of null alleles (Null) are shown for each population, for each locus and across all loci. Values in parentheses represent the corrected estimates of  $H_E$ ,  $H_O$  and  $F_{IS}$ .

| Pop. | Locus     | Locus   |         |         |         |         |         |         |         |         |         | Mean*           |
|------|-----------|---------|---------|---------|---------|---------|---------|---------|---------|---------|---------|-----------------|
|      |           | Mg181   | Med367  | MT203   | MT282   | MGE001  | MGE005  | MGE006  | MGE008  | Mgμ3    | Mgμ7    |                 |
| CHA  | N         | 5       | 9       | 9       | 9       | 4       | 9       | 1       | 3       | 3       | 15      | 6.7             |
|      | $A_R$     | 3.2851  | 6.5500  | 6.5612  | 6.4674  | 3.3104  | 6.5391  | 1.0000  | 2.9141  | 2.9625  | 11.8634 | 5.15            |
|      | $H_E$     | 0.5336  | 0.6669  | 0.6732  | 0.7684  | 0.5085  | 0.5911  | 0.0000  | 0.5414  | 0.3981  | 0.9079  | 0.5589 (0.4616) |
|      | $H_O$     | 0.1892  | 0.6500  | 0.2821  | 0.3421  | 0.3611  | 0.5385  | 0.0000  | 0.3077  | 0.1282  | 0.3871  | 0.3186 (0.3715) |
|      | $F_{IS}$  | 0.653   | 0.038   | 0.590   | 0.564   | 0.339   | 0.110   | 0.000   | 0.450   | 0.686   | 0.595   | 0.4025 (0.1915) |
|      | $P_{HWE}$ | 0.000   | 0.000   | 0.000   | 0.000   | 0.121   | 0.065   | –       | 0.000   | 0.000   | 0.000   |                 |
|      | Null      | 0.22705 | 0.05961 | 0.23855 | 0.24495 | 0.10054 | 0.00001 | 0.00100 | 0.16329 | 0.22182 | 0.27372 |                 |
| KAL  | N         | 3       | 9       | 9       | 8       | 4       | 8       | 1       | 4       | 3       | 13      | 6.2             |
|      | $A_R$     | 2.7095  | 6.3438  | 6.2735  | 5.9252  | 3.5561  | 5.5918  | 1.0000  | 3.2688  | 2.5078  | 10.3564 | 4.75            |
|      | $H_E$     | 0.4849  | 0.7659  | 0.7482  | 0.6925  | 0.5773  | 0.4925  | 0.0000  | 0.4744  | 0.2087  | 0.8786  | 0.5323 (0.5273) |
|      | $H_O$     | 0.3333  | 0.5263  | 0.4286  | 0.3750  | 0.2308  | 0.4750  | 0.0000  | 0.2500  | 0.1282  | 0.5185  | 0.3266 (0.3372) |
|      | $F_{IS}$  | 0.324   | 0.325   | 0.439   | 0.468   | 0.608   | 0.048   | 0.000   | 0.483   | 0.397   | 0.425   | 0.3817 (0.3642) |
|      | $P_{HWE}$ | 0.000   | 0.039   | 0.001   | 0.000   | 0.000   | 0.009   | –       | 0.002   | 0.000   | 0.000   |                 |
|      | Null      | 0.11701 | 0.13971 | 0.17412 | 0.19774 | 0.23144 | 0.00599 | 0.00100 | 0.16679 | 0.11116 | 0.19063 |                 |
| PER  | N         | 5       | 8       | 8       | 10      | 5       | 7       | 1       | 6       | 4       | 14      | 6.8             |
|      | $A_R$     | 3.5511  | 5.5703  | 6.0456  | 7.6882  | 3.6658  | 4.6808  | 1.0000  | 4.2871  | 3.3362  | 11.5738 | 5.14            |
|      | $H_E$     | 0.4298  | 0.6670  | 0.7091  | 0.7632  | 0.5139  | 0.4444  | 0.0000  | 0.6001  | 0.5667  | 0.9013  | 0.5595 (0.445)  |
|      | $H_O$     | 0.1429  | 0.4872  | 0.3077  | 0.3158  | 0.2895  | 0.4615  | 0.0000  | 0.3684  | 0.1667  | 0.5172  | 0.3057 (0.3213) |
|      | $F_{IS}$  | 0.705   | 0.290   | 0.558   | 0.611   | 0.447   | –0.020  | 0.000   | 0.406   | 0.724   | 0.440   | 0.416 (0.278)   |
|      | $P_{HWE}$ | 0.000   | 0.003   | 0.000   | 0.000   | 0.009   | 1.000   | –       | 0.000   | 0.000   | 0.000   |                 |
|      | Null      | 0.23557 | 0.10688 | 0.23099 | 0.26474 | 0.15428 | 0.00001 | 0.00100 | 0.16264 | 0.26725 | 0.20352 |                 |
| EPA  | N         | 4       | 6       | 9       | 7       | 3       | 8       | 1       | 5       | 4       | 15      | 6.2             |
|      | $A_R$     | 3.3999  | 5.7284  | 7.7292  | 6.2137  | 2.9748  | 5.9787  | 1.0000  | 3.8394  | 3.2760  | 12.0998 | 5.22            |
|      | $H_E$     | 0.5147  | 0.7433  | 0.8119  | 0.7397  | 0.5151  | 0.6241  | 0.0000  | 0.4797  | 0.3365  | 0.8861  | 0.5651 (0.4612) |
|      | $H_O$     | 0.1786  | 0.5455  | 0.5714  | 0.3125  | 0.3438  | 0.5000  | 0.0000  | 0.2647  | 0.2121  | 0.4444  | 0.3373 (0.3483) |
|      | $F_{IS}$  | 0.668   | 0.310   | 0.313   | 0.531   | 0.306   | 0.223   | 0.000   | 0.476   | 0.514   | 0.549   | 0.389 (0.113)   |
|      | $P_{HWE}$ | 0.000   | 0.007   | 0.005   | 0.000   | 0.151   | 0.010   | –       | 0.000   | 0.000   | 0.000   |                 |
|      | Null      | 0.23796 | 0.12125 | 0.13723 | 0.24741 | 0.11511 | 0.06345 | 0.00100 | 0.17690 | 0.13243 | 0.23551 |                 |
| STO  | N         | 4       | 8       | 9       | 6       | 4       | 7       | 1       | 4       | 3       | 16      | 6.2             |
|      | $A_R$     | 3.2914  | 6.3821  | 6.9491  | 5.6067  | 3.3113  | 5.0337  | 1.0000  | 3.1822  | 2.9020  | 12.9431 | 5.06            |
|      | $H_E$     | 0.5058  | 0.7353  | 0.7878  | 0.6704  | 0.3672  | 0.5102  | 0.0000  | 0.4673  | 0.5342  | 0.9045  | 0.5483 (0.4310) |
|      | $H_O$     | 0.2432  | 0.6667  | 0.4000  | 0.2647  | 0.2647  | 0.5405  | 0.0000  | 0.3429  | 0.2353  | 0.4783  | 0.3436 (0.3430) |
|      | $F_{IS}$  | 0.534   | 0.120   | 0.488   | 0.619   | 0.368   | –0.070  | 0.000   | 0.289   | 0.593   | 0.513   | 0.345 (0.2041)  |
|      | $P_{HWE}$ | 0.000   | 0.520   | 0.000   | 0.000   | 0.008   | 0.779   | –       | 0.417   | 0.000   | 0.000   |                 |
|      | Null      | 0.19280 | 0.06160 | 0.21272 | 0.24835 | 0.13089 | 0.00000 | 0.00100 | 0.09263 | 0.20112 | 0.23785 |                 |
| PK   | N         | 4       | 10      | 9       | 10      | 4       | 7       | 1       | 5       | 3       | 20      | 7.3             |
|      | $A_R$     | 3.2166  | 5.7328  | 7.1371  | 7.3500  | 3.2251  | 4.8816  | 1.0000  | 4.0377  | 2.9591  | 13.0690 | 5.26            |
|      | $H_E$     | 0.5208  | 0.6341  | 0.7084  | 0.7766  | 0.4037  | 0.4222  | 0.0000  | 0.5943  | 0.5409  | 0.8994  | 0.5501 (0.4923) |
|      | $H_O$     | 0.1714  | 0.4359  | 0.3000  | 0.4250  | 0.1842  | 0.3750  | 0.0000  | 0.3846  | 0.1622  | 0.6364  | 0.3075 (0.3360) |
|      | $F_{IS}$  | 0.679   | 0.324   | 0.585   | 0.463   | 0.553   | 0.124   | 0.000   | 0.364   | 0.707   | 0.307   | 0.411 (0.3175)  |
|      | $P_{HWE}$ | 0.000   | 0.409   | 0.000   | 0.000   | 0.000   | 0.466   | –       | 0.000   | 0.000   | 0.000   |                 |
|      | Null      | 0.24200 | 0.12387 | 0.23579 | 0.20008 | 0.18033 | 0.00988 | 0.00100 | 0.14459 | 0.25277 | 0.14198 |                 |
| KAV  | N         | 3       | 8       | 11      | 8       | 3       | 9       | 1       | 3       | 4       | 18      | 6.8             |
|      | $A_R$     | 2.8944  | 6.8046  | 8.5878  | 6.1560  | 2.9352  | 6.4149  | 1.0000  | 4.1739  | 2.8920  | 12.5339 | 5.44            |
|      | $H_E$     | 0.5328  | 0.7643  | 0.8222  | 0.7334  | 0.4829  | 0.5659  | 0.0000  | 0.5952  | 0.3704  | 0.9026  | 0.577 (0.5144)  |
|      | $H_O$     | 0.1290  | 0.6129  | 0.4667  | 0.3750  | 0.2963  | 0.5313  | 0.0000  | 0.3750  | 0.2581  | 0.4783  | 0.3522 (0.3629) |
|      | $F_{IS}$  | 0.765   | 0.214   | 0.446   | 0.501   | 0.402   | 0.077   | 0.000   | 0.384   | 0.318   | 0.487   | 0.359 (0.2945)  |
|      | $P_{HWE}$ | 0.000   | 0.688   | 0.000   | 0.000   | 0.054   | 0.839   | –       | 0.000   | 0.000   | 0.001   |                 |
|      | Null      | 0.27210 | 0.08062 | 0.18919 | 0.21137 | 0.11768 | 0.04321 | 0.00100 | 0.15142 | 0.12280 | 0.22007 |                 |
| MYT  | N         | 4       | 8       | 9       | 9       | 3       | 6       | 1       | 5       | 3       | 16      | 6.4             |
|      | $A_R$     | 3.2875  | 5.5347  | 6.8184  | 7.0381  | 2.9067  | 5.0193  | 1.0000  | 4.0439  | 2.6562  | 12.1544 | 5.05            |
|      | $H_E$     | 0.5042  | 0.6613  | 0.6177  | 0.7534  | 0.3751  | 0.5266  | 0.0000  | 0.5760  | 0.2378  | 0.8994  | 0.5151 (0.4852) |
|      | $H_O$     | 0.3793  | 0.6765  | 0.3125  | 0.5000  | 0.3714  | 0.4444  | 0.0000  | 0.3333  | 0.0000  | 0.4231  | 0.3441 (0.3864) |
|      | $F_{IS}$  | 0.264   | –0.008  | 0.506   | 0.350   | 0.024   | 0.170   | 0.000   | 0.433   | 1.00    | 0.544   | 0.328 (0.0988)  |
|      | $P_{HWE}$ | 0.013   | 1.000   | 0.000   | 0.000   | 0.961   | 0.019   | –       | 0.006   | 0.000   | 0.000   |                 |
|      | Null      | 0.09620 | 0.00001 | 0.20460 | 0.15187 | 0.00002 | 0.06552 | 0.00100 | 0.15849 | 0.23763 | 0.25326 |                 |

Continued

Table 2. Continued

| Pop. | Locus            |         |         |         |         |         |         |         |         |         |         | Mean*           |
|------|------------------|---------|---------|---------|---------|---------|---------|---------|---------|---------|---------|-----------------|
|      |                  | Mg181   | Med367  | MT203   | MT282   | MGE001  | MGE005  | MGE006  | MGE008  | Mgμ3    | Mgμ7    |                 |
| TUR  | N                | 5       | 12      | 9       | 11      | 4       | 7       | 1       | 4       | 4       | 14      | 7.1             |
|      | A <sub>R</sub>   | 4.3551  | 7.6326  | 6.3205  | 6.5892  | 3.6882  | 5.0292  | 1.0000  | 3.5436  | 3.5637  | 10.3312 | 5.21            |
|      | H <sub>E</sub>   | 0.5918  | 0.7216  | 0.6859  | 0.6125  | 0.3965  | 0.4572  | 0.0000  | 0.6147  | 0.6243  | 0.8522  | 0.5557 (0.4895) |
|      | H <sub>O</sub>   | 0.1714  | 0.6750  | 0.2105  | 0.5000  | 0.3684  | 0.4250  | 0.0000  | 0.3500  | 0.3421  | 0.2258  | 0.3268 (0.3801) |
|      | F <sub>IS</sub>  | 0.717   | -0.007  | 0.700   | 0.196   | 0.084   | 0.083   | 0.000   | 0.441   | 0.463   | 0.742   | 0.35 (0.2235)   |
|      | P <sub>HWE</sub> | 0.000   | 0.003   | 0.000   | 0.000   | 0.115   | 0.847   | -       | 0.004   | 0.000   | 0.000   |                 |
|      | Null             | 0.27385 | 0.04046 | 0.28200 | 0.07238 | 0.05138 | 0.01837 | 0.00100 | 0.16600 | 0.17263 | 0.33782 |                 |
| ION  | N                | 5       | 9       | 11      | 9       | 3       | 7       | 1       | 6       | 3       | 18      | 7.2             |
|      | A <sub>R</sub>   | 3.8838  | 6.0598  | 7.7246  | 6.9976  | 2.7328  | 5.4366  | 1.0000  | 4.9989  | 2.9506  | 13.1979 | 5.50            |
|      | H <sub>E</sub>   | 0.5606  | 0.6782  | 0.7361  | 0.7861  | 0.3924  | 0.5677  | 0.0000  | 0.6532  | 0.4204  | 0.9114  | 0.5706 (0.4096) |
|      | H <sub>O</sub>   | 0.2105  | 0.6923  | 0.3889  | 0.2857  | 0.2222  | 0.5641  | 0.0000  | 0.3333  | 0.0588  | 0.5152  | 0.3271 (0.3697) |
|      | F <sub>IS</sub>  | 0.633   | -0.008  | 0.483   | 0.645   | 0.402   | 0.019   | 0.000   | 0.499   | 0.864   | 0.464   | 0.4 (0.0614)    |
|      | P <sub>HWE</sub> | 0.000   | 1.000   | 0.000   | 0.000   | 0.018   | 0.604   | -       | 0.000   | 0.000   | 0.000   |                 |
|      | Null             | 0.22734 | 0.01105 | 0.20849 | 0.28013 | 0.12284 | 0.00000 | 0.00100 | 0.20173 | 0.27704 | 0.21117 |                 |
| CRO  | N                | 3       | 6       | 6       | 4       | 5       | 6       | 1       | 5       | 2       | 8       | 4.6             |
|      | A <sub>R</sub>   | 2.9286  | 5.1874  | 5.5862  | 4.0000  | 4.8666  | 5.2058  | 1.0000  | 4.8545  | 1.9887  | 7.6459  | 4.33            |
|      | H <sub>E</sub>   | 0.4259  | 0.6190  | 0.6178  | 0.6243  | 0.7222  | 0.6054  | 0.0000  | 0.5000  | 0.1938  | 0.7993  | 0.5108 (0.4873) |
|      | H <sub>O</sub>   | 0.4444  | 0.5714  | 0.3333  | 0.3846  | 0.2000  | 0.8182  | 0.0000  | 0.1429  | 0.1304  | 0.5294  | 0.3555 (0.3727) |
|      | F <sub>IS</sub>  | -0.015  | 0.101   | 0.487   | 0.417   | 0.739   | -0.331  | 0.000   | 0.732   | 0.347   | 0.364   | 0.2841 (0.2351) |
|      | P <sub>HWE</sub> | 0.861   | 0.000   | 0.084   | 0.018   | 0.000   | 0.839   | -       | 0.000   | 0.117   | 0.000   |                 |
|      | Null             | 0.02720 | 0.0000  | 0.00001 | 0.08135 | 0.22381 | 0.00000 | 0.00100 | 0.15330 | 0.11236 | 0.05884 |                 |
| RAV  | N                | 4       | 8       | 7       | 7       | 4       | 7       | 1       | 5       | 3       | 12      | 5.8             |
|      | A <sub>R</sub>   | 3.8118  | 6.8388  | 6.3561  | 7.0000  | 3.7551  | 5.6169  | 1.0000  | 4.8434  | 2.5416  | 11.1173 | 5.29            |
|      | H <sub>E</sub>   | 0.5844  | 0.6578  | 0.7511  | 0.7273  | 0.4949  | 0.3956  | 0.0000  | 0.6389  | 0.4867  | 0.8580  | 0.556 (0.3823)  |
|      | H <sub>O</sub>   | 0.0667  | 0.5333  | 0.2000  | 0.3636  | 0.3571  | 0.4000  | 0.0000  | 0.1667  | 0.7333  | 0.5385  | 0.2977 (0.3272) |
|      | F <sub>IS</sub>  | 0.849   | 0.340   | 0.782   | 0.508   | 0.342   | 0.119   | 0.000   | 0.655   | -0.281  | 0.522   | 0.384 (0.1441)  |
|      | P <sub>HWE</sub> | 0.000   | 0.000   | 0.000   | 0.019   | 0.209   | 0.000   | -       | 0.000   | 0.446   | 0.000   |                 |
|      | Null             | 0.32950 | 0.10185 | 0.32721 | 0.22175 | 0.12014 | 0.00000 | 0.00100 | 0.25964 | 0.00000 | 0.20257 |                 |
| LIV  | N                | 4       | 9       | 13      | 9       | 5       | 6       | 1       | 3       | 4       | 18      | 7.2             |
|      | A <sub>R</sub>   | 3.5494  | 6.4644  | 8.8312  | 7.9866  | 4.4524  | 4.3636  | 1.0000  | 2.9970  | 3.1445  | 12.3897 | 5.52            |
|      | H <sub>E</sub>   | 0.5661  | 0.6361  | 0.7996  | 0.6867  | 0.5227  | 0.5813  | 0.0000  | 0.5910  | 0.3277  | 0.8913  | 0.5603 (0.5077) |
|      | H <sub>O</sub>   | 0.1818  | 0.5217  | 0.5652  | 0.3889  | 0.5000  | 0.7000  | 0.0000  | 0.1667  | 0.2857  | 0.3043  | 0.3614 (0.423)  |
|      | F <sub>IS</sub>  | 0.661   | 0.247   | 0.357   | 0.457   | -0.001  | -0.178  | 0.000   | 0.755   | 0.180   | 0.711   | 0.319 (0.2071)  |
|      | P <sub>HWE</sub> | 0.000   | 0.000   | 0.000   | 0.001   | 0.034   | 0.989   | -       | 0.000   | 0.742   | 0.000   |                 |
|      | Null             | 0.24515 | 0.06222 | 0.11851 | 0.17213 | 0.05733 | 0.00000 | 0.00100 | 0.27277 | 0.04833 | 0.30911 |                 |

Overall expected ( $H_e$ ) and observed ( $H_o$ ) heterozygosity values at species level were 0.577 and 0.459 without correction, and 0.332 and 0.342 with null allele correction, respectively, while the corresponding  $F_{IS}$  values were 0.365 and 0.207, respectively. The selfing rate as estimated by RMES was 0.043 at species level, suggesting the rejection of self fertilization hypothesis in mussel populations. Finally, no significant disequilibrium between pairs of loci within each population was detected (468 tests) whereas tests between pairs of loci from all 13 populations (36 tests) disclosed significant disequilibria only between Med367 and MT282 loci.

**Genetic differentiation**

The UPGMA dendrogram, constructed from Reynolds' genetic distances after correction for null alleles, is shown in Figure 2, whereas the DAPC diagram is displayed in Figure 3. Notably, the UPGMA dendrogram constructed without applying the corrections or by using other distance methods, such as Nei's (1978) and Cavalli-Sforza's chord measure  $D_{CE}$  (Cavalli-Sforza & Edwards, 1967), as well as the PCA plot derived from PCAgen 1.2.1 (not shown)

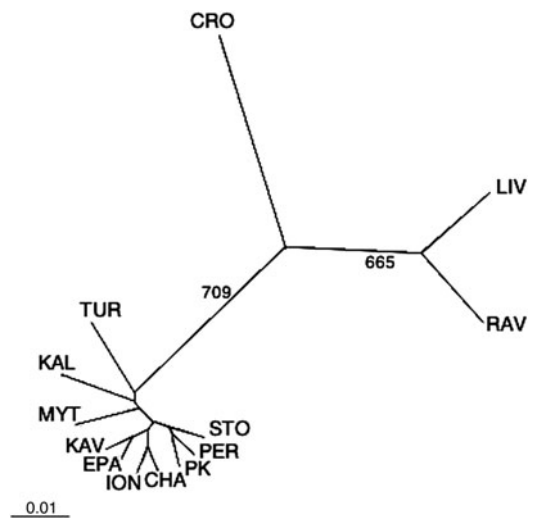
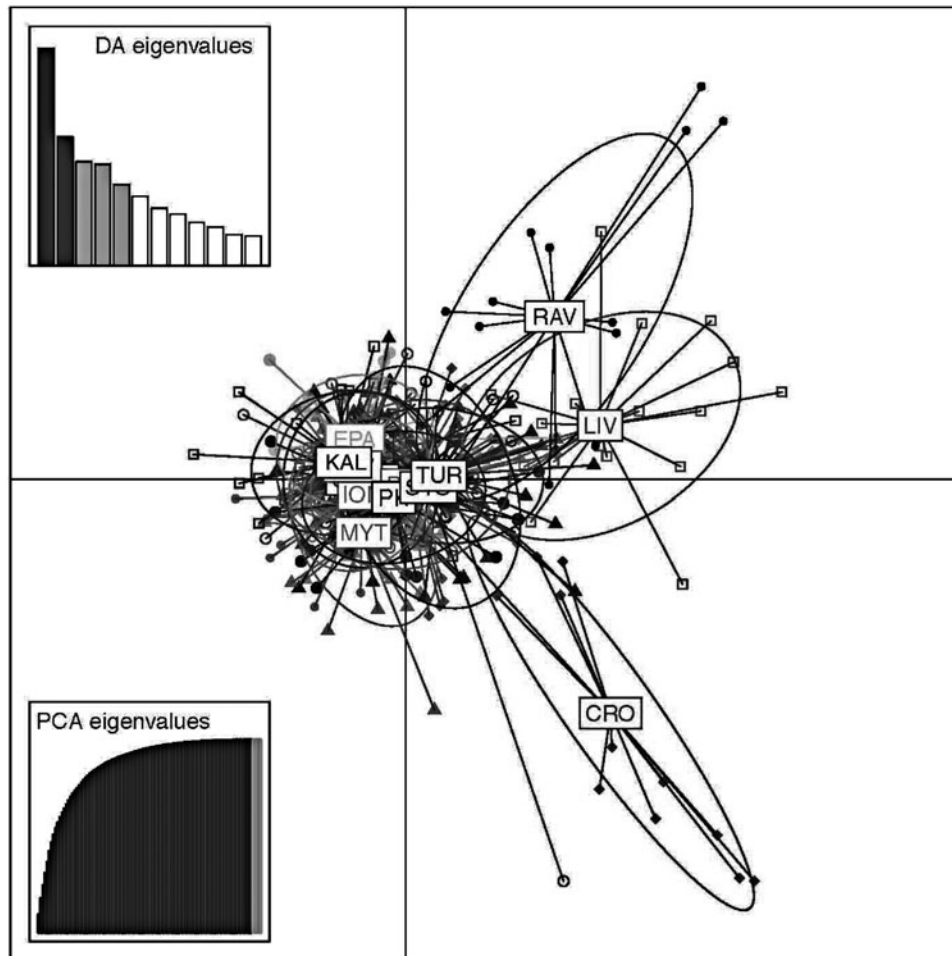


Fig. 2. UPGMA tree indicating Reynolds' genetic distances among the 13 *Mytilus galloprovincialis* populations after gene frequency adjustment for null alleles. Bootstrap values greater than 500 are shown on the branches. The constructed tree before the null allele correction was equal therefore is not shown.



**Fig. 3.** DAPC plot of mussel populations. Dots represent individuals, whereas coloured ellipses correspond to geographical populations. DA and PCA scatterplots in the left side of the graphs indicate the number of discriminant functions and PCs retained for the computations. CHA, Chalastra; KAL, Kalohori; PER, Peraia; EPA, Epanomi; STO, Stomio; PK, Porto Koufo; KAV, Kavala; MYT, Mytilene; TUR, Canakkale; ION, Igoumenitsa; CRO, Zadar; RAV, Marina di Ravenna; LIV, Livorno.

yielded very similar patterns, clustering the populations into three different groups. The first group comprised all Greek populations and the Turkish one; the second the two Italian populations, whereas the third group comprised only the population from Croatia. Fisher's exact tests and  $F_{ST}$  values among all populations revealed significant genetic differentiation regardless of whether the ENA correction method was used or not ( $F_{ST} = 0.024$  and  $F_{ST} = 0.022$ ,  $P < 10^{-5}$ , respectively). Significant differentiation, although smaller, was also revealed by  $\Phi$  statistics ( $\Phi_{ST} = 0.013$ ,  $P = 0.011$ ). In contrast, as it concerns the Aegean populations both the  $F_{ST}$  and  $\Phi_{ST}$  values were not significant i.e. the  $F_{ST}$  values (with and without correction) were 0.006 and 0.004 ( $P = 0.282$  and  $P = 0.356$ , respectively), while the  $\Phi_{ST}$  value was 0.001 ( $P = 0.30$ ). None of the pairwise comparisons for genetic heterogeneity among populations within Aegean Sea was significant (Table 3), after applying the Benjamini & Hochberg's false discovery rate procedure (Verhoeven *et al.*, 2005). In addition, no significant genetic differentiation was observed between the two Italian populations, as well as between the Ionian and the Aegean populations. On the other hand, differentiation of Italian, Croatian and Turkish populations versus Aegean populations was predominately highly significant (Table 3).

The AMOVA was used to partition the variation among and within group (population) components. At first, populations were pooled into three geographical groups: all Aegean populations, together with the one from the Sea of Marmara, were treated as one group; populations from the Adriatic Sea (Zadar and Ravenna) constituted the second group, whereas Livorno (Ligurian Sea) has been considered as a separated third group. The sample from the Ionian Sea (Igoumenitsa) was omitted from the AMOVA, as it was evidenced that nearby mussel transfers may have influenced it (see Discussion below). The results demonstrated that most variation exists within populations (97.16%) followed by that attributed to among groups (2.08%) and among populations within groups (0.76%) (Table 4). Despite the small amounts of genetic variation that are attributable to grouping of populations by and within basins, variation in each partition was significantly different from zero ( $P < 0.05$ ). Interestingly in a second AMOVA, when only Aegean populations were considered, the  $F_{ST}$  value fell to 0.004 and among populations variation was found to be non-significant ( $P = 0.36$ ) accounting only for 0.49% of the total variance (Table 4). These results indicate panmixia of mussels along Aegean coasts and are in consistency with the assignment analyses performed with the GENECLASS2 (Piry *et al.*, 2004)

**Table 3.** Pairwise exact tests. *P* values for each pairwise exact test for genetic heterogeneity across all loci (Fisher's method). Significant values after Benjamini and Hochberg correction ( $P \leq 0.023$ ) are shown in bold.

|     | CHA               | KAL               | PER               | EPA               | STO               | PK                | KAV               | MYT               | TUR               | ION               | CRO               | RAV   | LIV |
|-----|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------------------|-------|-----|
| CHA | –                 |                   |                   |                   |                   |                   |                   |                   |                   |                   |                   |       |     |
| KAL | 0.414             | –                 |                   |                   |                   |                   |                   |                   |                   |                   |                   |       |     |
| PER | 0.693             | 0.141             | –                 |                   |                   |                   |                   |                   |                   |                   |                   |       |     |
| EPA | 0.596             | 0.362             | 0.158             | –                 |                   |                   |                   |                   |                   |                   |                   |       |     |
| STO | 0.133             | 0.203             | 0.585             | 0.058             | –                 |                   |                   |                   |                   |                   |                   |       |     |
| PK  | 0.607             | 0.101             | 0.917             | 0.401             | 0.574             | –                 |                   |                   |                   |                   |                   |       |     |
| KAV | 0.745             | 0.640             | 0.207             | 0.825             | 0.099             | 0.073             | –                 |                   |                   |                   |                   |       |     |
| MYT | 0.712             | 0.085             | 0.241             | 0.403             | 0.151             | 0.579             | 0.195             | –                 |                   |                   |                   |       |     |
| TUR | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>0.002</b>      | <b>0.0002</b>     | 0.036             | <b>0.008</b>      | <b>0.006</b>      | <b>0.0004</b>     | –                 |                   |                   |       |     |
| ION | 0.653             | 0.202             | 0.065             | 0.348             | 0.109             | 0.488             | 0.944             | 0.302             | <b>0.003</b>      | –                 |                   |       |     |
| CRO | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>0.002</b>      | –                 |       |     |
| RAV | <b>0.0009</b>     | <b>&lt;0.0001</b> | 0.049             | <b>&lt;0.0001</b> | 0.048             | 0.044             | <b>&lt;0.0001</b> | <b>0.0006</b>     | <b>0.006</b>      | <b>0.002</b>      | <b>&lt;0.0001</b> | –     |     |
| LIV | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>0.0005</b>     | <b>&lt;0.0001</b> | <b>0.012</b>      | <b>0.0002</b>     | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>&lt;0.0001</b> | <b>0.001</b>      | 0.283 | –   |

**Table 4.** Analysis of molecular variance.

| Source of variation             | All populations <sup>1</sup> |                     |                         | Aegean populations <sup>2</sup> |                     |                         |
|---------------------------------|------------------------------|---------------------|-------------------------|---------------------------------|---------------------|-------------------------|
|                                 | df                           | Variance components | Percentage of variation | df                              | Variance components | Percentage of variation |
| Among groups                    | 2                            | 0.01820             | 2.08                    | –                               | –                   | –                       |
| Among populations within groups | 9                            | 0.00666             | 0.76                    | 7                               | 0.00597             | 0.49                    |
| Within populations              | 774                          | 0.85142             | 97.16                   | 592                             | 1.20778             | 99.51                   |

Two-level hierarchical AMOVA. <sup>1</sup>Considering all populations in three main groups: Aegean and Marmara Sea; Adriatic and Ligurian Sea. <sup>2</sup>Solely for the Aegean populations (one group).

where only 16.2% of the Aegean mussels were assigned correctly to the population from which they were sampled. In addition, the Mantel test revealed a significant correlation between geographical and genetic distances ( $r = 0.822$ ;  $P = 0.004$ ) which, however, does not hold for the eastern Mediterranean samples—the eight Aegean populations and the one from the sea of Marmara—where no such correlation was found ( $r = -0.045$ ;  $P = 0.723$ ).

## Power analysis

POWSIM analysis estimated that there was more than 96% statistical chance of detecting genetic differentiation when global  $F_{ST}$  (first case: including all samples) was 0.0025, and 100% when  $F_{ST}$  was 0.005 or higher. Concerning the case of Aegean populations, 78.6% of the  $\chi^2$  tests and 64.3% of the Fisher tests indicated significant genetic differentiation when  $F_{ST}$  was set to 0.005. Only when expected  $F_{ST}$  was increased to 0.01, were values of statistical power found to be 100% for the  $\chi^2$  test and 98.7% for the Fisher test. Thus, regarding mussels from Aegean Sea, the hypothesis of genetic homogeneity cannot be excluded.

## DISCUSSION

### Intra-population genetic diversity

High levels of genetic diversity are common within marine bivalves (Marin *et al.*, 2013 and references therein). On the other hand, several studies have demonstrated a reduction in genetic variability within cultured aquatic populations

over a relatively short period of time as a result of inbreeding (e.g. Sekino *et al.*, 2002; Lind *et al.*, 2009; De La Cruz *et al.*, 2010). In mussels, although no artificial fertilization of eggs is carried out (Gosling, 2003), loss of the mean multi-locus heterozygosity has been reported in suspension-cultured populations of *M. edulis*, in Iles-de-la-Madeleine (Quebec, Canada) (Myrand *et al.*, 2009 and references therein). The phenomenon occurred mostly in the periphery of the sleeves and was attributed to fall-offs of heterozygous individuals. Nevertheless, our results indicate that the two cultured populations of *M. galloprovincialis* exhibit similar level of genetic diversity (i.e. heterozygosity and allelic richness values) with the wild Mediterranean populations (Table 2), information that may be extremely useful in future management practices. A similar conclusion was drawn in a previous analysis of the present cultured populations with RAPD markers (Giantsis *et al.*, 2012) suggesting that the mussel culture systems employed in the eastern Mediterranean (as well as in the western Mediterranean, Diz & Presa, 2009) do not lead to bottlenecks of exploited populations.

### Deviation from HWE and heterozygote deficiency

Significant departures from HWE associated with heterozygote deficiencies were observed in the majority of the loci examined, in all populations (Table 2). Heterozygote deficiencies have been frequently observed in marine bivalves (e.g. Marin *et al.*, 2013) including those of *Mytilus* (Addison *et al.*, 2008; Diz & Presa, 2009). The cause of these observations is probably a combination of technical features and biologically based factors (i.e. null alleles, allele homoplasy,



inbreeding and self-fertilization, selection, population sub-structure and aneuploidy), but they have yet to be adequately explained (Wei *et al.*, 2013).

Self-fertilization during spawning constitutes a phenomenon closely related to inbreeding, which could explain heterozygote deficiencies, and it has been noted in several aquatic invertebrates, including clam shrimp (Chasnov, 2010), scallops (Martinez *et al.*, 2007) and clams (Kurihara, 2010). In mussels, hermaphroditism has been reported at very low prevalences, although it is likely to be increased when they are exposed to toxic chemicals (Ortiz-Zarragoitia & Cajaraville, 2010). Nevertheless, this seems to be the least possible scenario in *Mytilus*, where effective population sizes are large enough to prevent inbreeding (Hoarau *et al.*, 2005). Moreover, the selfing rates at species level, as revealed by the maximum likelihood estimation of  $S$  (RMES, David *et al.*, 2007), were not significantly different from zero ( $S = 0.043$ ) indicating that there was no evidence for self-fertilization in the *M. galloprovincialis* populations studied. The hypothesis that introgression may be responsible for the observed deviations from Hardy–Weinberg expectations also receives little support since there was no evidence of introgressive hybridization among the analysed *M. galloprovincialis* samples and other *Mytilus* species.

In theory, microsatellite loci are generally considered selectively neutral. Sometimes however, they do not act singularly but are linked to genes networks, and hence may be affected by selective pressure (Rhode *et al.*, 2013). In such cases selection may affect the genetic structure of polymorphism at specific loci causing heterogeneity among loci, with high levels of differentiation at some and low or none at others (Launey *et al.*, 2002). However, the similar  $F_{ST}$  estimates of the polymorphic loci (data not shown) coupled with simulation results in ARLEQUIN 3.5 where only locus  $Mg\mu 3$  appear to be a candidate for balancing selection ( $P < 10^{-5}$ ), suggest that the loci used were acting as selectively neutral markers and therefore selection is unlikely to be the main cause of the heterozygote deficiency observed in this study.

*Mytilus* species have a spawning period that may reach up to six months exhibiting excellent survival capability to various environmental conditions in this phase, whereas females are able to produce up to 40 million eggs individually (Gosling, 2003). Furthermore, they show a high dispersal capability during pelagic larvae stage. Consequently, even though an overall Wahlund effect (including all populations) could be probably ruled out as overall  $F_{ST}$  (0.024) is much lower than mean  $F_{is}$  (0.207), a cryptic subpopulation admixture generated by pooling populations from large marine areas might explain the HWE departures and heterozygote deficits in bivalves of the genus *Mytilus* (see Diz & Presa, 2008; Addison *et al.*, 2008).

Regarding null alleles, they are nearly always mentioned as a potential explanation for the occurrence of heterozygote deficits in microsatellite (e.g. Carlsson *et al.*, 2008; Polato *et al.*, 2010). However, as stated earlier, in most loci examined in this study, significant deficits remained after the correction for null alleles, suggesting that null alleles are only partially responsible for these observations.

The high incidences of aneuploidy phenomena in bivalves, mainly in oysters and veneroids (Teixeira de Sousa *et al.*, 2012) lead us to suggest that aneuploidy could also contribute to the observed heterozygote deficiencies of mussels although

further studies toward this direction are clearly needed to clarify the prevalence of this phenomenon in mussels.

Lastly, although stuttering and allelic dropout were statistically ruled out and a correction procedure for null alleles took place, there is still a noteworthy potential artefact at microsatellites: mispairing of repeated units during DNA replication due to high levels of polymorphism could potentially generate allelic size homoplasies (Culver *et al.*, 2001), responsible for erroneous overestimation of homozygotes and subsequently for the observed heterozygote deficiency.

## Spatial patterns and genetic structure of populations

Microsatellite analysis revealed that although most variation was attributed to within populations' variance (Table 4), significant genetic differentiation among all populations was present. However, this genetic divergence, as indicated by exact tests and pairwise  $F_{ST}$  estimates, was mostly caused by the strong genetic heterogeneity of the samples from the Adriatic and the Sea of Marmara (Table 3; Figures 2, 3), whereas regarding the region of Aegean Sea, neither global  $F_{ST}$  or  $\Phi_{ST}$ , nor pairwise tests revealed any genetic differentiation. Our study differs from previous population genetic studies of *M. galloprovincialis* that propose a restricted dispersal ability of mussels around north Greek coasts (Karakousis & Skibinski, 1992; Kravva *et al.*, 2000) and the existence of significant genetic structuring among Aegean populations (Giantsis *et al.*, 2012). However, these discrepancies could potentially be caused by several factors, such as the resolving power of the markers used to detect patterns of population differentiation. On the other hand, the results confirm and extend those of Ladoukakis *et al.* (2002) who, based on mtDNA analyses, found that *M. galloprovincialis* populations from the Aegean Sea form a homogeneous collection and are differentiated from those of the Black and Adriatic Seas. The complete lack of genetic differentiation among the Aegean populations suggests that mussels from the Aegean Sea are either at or close to panmixia. This was also evidenced by the second AMOVA, considering only Aegean populations, which revealed that less of 0.5% of the total variation detected was attributed among populations (Table 4).

*Mytilus galloprovincialis* is not the only marine species whose populations display genetic homogenization in the Aegean Sea. For instance, a mtDNA-RFLP analysis of European lobster (*Homarus gammarus*) populations found homogeneity among Aegean populations and suggested that the Aegean may act as a metapopulation for European lobster, in which these samples represent a dynamic system of sub-populations that are linked by gene flow, and with individual sub-populations becoming extinct and being recolonized (Triantafyllidis *et al.*, 2005). However the panmictic model that has been observed in *M. galloprovincialis* populations from the Aegean Sea and the Ionian Sea correspondingly, does not represent a kind of metapopulation but can be attributed mainly to the species' inherent ability to disperse over long distances through its planktonic larval phase. Indeed, widespread marine organisms with pelagic larval dispersal usually show little, if any, genetic differentiation, due to the absence of natural or artificial barriers that could impede their larval dispersal, resulting in low levels of interpopulation genetic variability. In particular, with regards to marine

mussels, several studies have demonstrated genetic homogenization or panmixia among populations within broad geographic areas such as on the Californian coasts and the Galician Rias (Addison *et al.*, 2008; Diz & Presa 2009, respectively). However, in addition to species high dispersal ability, anthropogenic interferences, that is, unrecorded transplantations of spat for aquaculture purposes and, to a lesser degree, accidental transfer with the ballast waters of ships, may have also played a role in shaping the patterns of genetic homogeneity among the Greek populations, as well as among the two Italian samples.

In the process of mussel culture, mussel seed (juveniles) is collected from traps or collector ropes and transferred to other areas for growing (Kijewski *et al.*, 2011). In particular, as it concerns *Mytilus*, transplantations seem to have influenced the genetic composition of several native mussel populations, such as those along northern Chilean (Toro *et al.*, 2004) and north-western European coasts (Kijewski *et al.*, 2011). Correspondingly, the past three decades when mussel culture has spread widely in Greece, spat from Chalastra (Thermaikos Gulf, Figure 1) has been transferred to most mussel farms around Greece, and these transplantations may have affected the genetic structure of wild (naturally occurring) local populations and hence contributed to the genetic homogeneity among Greek populations, overcoming, for example, the physical barriers that prevent mass gene flow between the Aegean and Ionian basins. This hypothesis would be further reinforced if there were observed populations from southern Greece, an area where the mussel culture has not yet developed, that were different from both Aegean and Ionian ones. However, since no such population was found, nor is there any certainty about the annual numbers of mussels moved around the country or the duration and frequency of these transfers, it is very difficult to support this scenario with confidence. Similarly, to explain the observed genetic homogeneity between the two distant Italian populations (shoreline >3000 km, Figure 1), further analyses including populations from several locations around the Italian coasts are needed.

On the other hand, there are cases where gene flow among marine organisms can be constrained by biological, physical and ecological factors (Borrero-Perez *et al.*, 2011; Sanna *et al.*, 2013). The Mediterranean biogeographical boundaries play a significant role in shaping the genetic structuring in marine species (Sanna *et al.*, 2013). Particularly the region of the Aegean Sea (eastern Mediterranean) is of special interest, as a number of studies have revealed the existence of a genetic break or limitation to larval dispersal between the Aegean populations and those in other basins (Ladoukakis *et al.*, 2002; Nikula & Väinölä, 2003; Borrero-Pérez *et al.*, 2011), and the existence of natural frontiers that may prevent the invasion of non-native species (Panucci-Papadopoulou *et al.*, 2012). Such an example may be the population from the Sea of Marmara (Canakkale) which appears significantly differentiated from most Aegean populations (Table 3). The outflow of low salinity waters from the Dardanelles into the Aegean (Olson *et al.*, 2007) in combination with the narrow width (1.2–6 km) probably forms a physical barrier preventing gene flow from the Aegean towards the Sea of Marmara. Likewise the topographic and oceanographic conditions prevailing on the Dalmatian coasts may form physical barriers inhibiting larval dispersal and thus promoting the significant differentiation of the

Croatian sample from all the other samples, including that of Ravenna situated at the opposite side of the Adriatic (Table 4; Figure 1). Indeed, the bay of Zadar is surrounded by a complex of longish islands, which stand in the way of direct contact of the bay with the open sea. The region comprises several distinct marine habitats such as karst marine lakes and submerged caves and pits (Radović *et al.*, 2009). Moreover the sea currents of the two Adriatic sides are different as they flow in a cyclonic circulation, with those of the Italian coasts flowing southwards far away from the Croatian coast (Mantziafou & Lascaratos, 2004).

In conclusion, the current research presents a first attempt to investigate the genetic structure of *Mytilus galloprovincialis* mussels throughout the coasts of the central–eastern Mediterranean Sea using microsatellite markers. The results indicate broad patterns of genetic homogeneity between Greek populations, as well as between the two Italian ones, and the differentiation of those from Turkey and Croatia. However, to elucidate the processes that led to the observed patterns of genetic structure and to ensure the sustainability of populations of the species, further research will be required, including key locations throughout the Mediterranean Sea and strategies to deal with human transplantations.

## ACKNOWLEDGEMENTS

The authors are grateful to Drs Gloria Isani and Emilio Carpenè for their kind hospitality in the Department of Veterinary Medical Science, University of Bologna. We are also indebted to Dr Natalija Topic Popovic for her valuable help in collection of the Croatian samples.

## FINANCIAL SUPPORT

This research received no specific grant from any funding agency, commercial or not-for-profit sectors.

## REFERENCES

- Addison J.A., Ort B.S., Mesa K.A. and Pogson G.H. (2008) Range-wide genetic homogeneity in the California sea mussel (*Mytilus californianus*): a comparison of allozymes, nuclear DNA markers and mitochondrial DNA sequences. *Molecular Ecology* 17, 4222–4232.
- Arsenoudi P., Scouras Z. and Chintiroglou C.C. (2003) First evaluation of *Mytilus galloprovincialis* LMK, natural populations in Thermaikos Gulf: Structure and distribution. *Fresenius Environmental Bulletin* 12, 1384–1393.
- Beaumont A.R., Hawkins M.P., Doig F.L., Davies I.M. and Snow M. (2008) Three species of *Mytilus* and their hybrids identified in a Scottish Loch: natives, relicts and invaders? *Journal of Experimental Marine Biology and Ecology* 367, 100–110.
- Belkhir K., Borsa P., Goudet J., Chikhi L. and Bonhomme F. (2004) *Genetix v. 4.05, Logiciel sous Windows™ pour la génétique des populations*. Montpellier: Laboratoire Génome et Population, Université Montpellier 2.
- Bierne N., Borsa P., Daguin C., Jollivet D., Viard F., Bonhomme F. and David P. (2003) Introgression patterns in the mosaic hybrid zone between *Mytilus edulis* and *M. galloprovincialis*. *Molecular Ecology* 12, 447–461.

- Borrero-Pérez G.H., González-Wangüemert M., Marcos C. and Pérez-Ruzafa A.** (2011) Phylogeography of the Atlanto-Mediterranean sea cucumber *Holothuria (Holothuria) mammata*: the combined effects of the historical processes and current oceanographic pattern. *Molecular Ecology* 20, 1964–1975.
- Carlsson J.** (2008) Effects of microsatellite null alleles on assignment testing. *Journal of Heredity* 99, 616–623.
- Cavalli-Sforza L.L. and Edwards A.W.F.** (1967) Phylogenetic analysis: models and estimation procedures. *American Journal of Human Genetics* 19, 233–257.
- Chapuis M.P. and Estoup A.** (2007) Microsatellite null alleles and estimation of population differentiation. *Molecular Biology and Evolution* 24, 621–631.
- Chasnov J.R.** (2010) The evolution from females to hermaphrodites results in a sexual conflict over mating in androdioecious nematode worms and clam shrimp. *Journal of Evolutionary Biology* 23, 539–556.
- Coombs J.A., Letcher B.H. and Nislow K.H.** (2008) CREATE: a software to create input files from diploid genotyping data for 52 genetic software programs. *Molecular Ecology Resources* 8, 578–580.
- Culver M., Menotti-Reymond M.A. and O'Brien S.J.** (2001) Patterns of size homoplasy at 10 microsatellite loci in Pumas (*Puma concolor*). *Molecular Biology and Evolution* 17, 489–458.
- Dakin E.E. and Avise J.C.** (2004) Microsatellite null alleles in parentage analysis. *Heredity* 93, 504–509.
- David P., Pujol B., Viard F., Castella V. and Goudet J.** (2007) Reliable selfing rate estimates from imperfect population genetic data. *Molecular Ecology* 16, 2474–2487.
- De La Cruz F.L., Rio-Portilla M.A.D. and Gallardo-Escarate C.** (2010) Genetic variability of cultured populations of red abalone in Chile: an approach based on heterologous microsatellites. *Journal of Shellfish Research* 29, 709–715.
- Dempster A.P., Laird N.M. and Rubin D.B.** (1977) Maximum likelihood from incomplete data via the EM algorithm. *Journal of the Royal Statistical Society: Series B* 39, 1–38.
- Diz A.P. and Presa P.** (2008) Regional patterns of microsatellite variation in *Mytilus galloprovincialis* from the Iberian Peninsula. *Marine Biology* 154, 277–286.
- Diz A.P. and Presa P.** (2009) The genetic diversity pattern of *Mytilus galloprovincialis* in Galician Rias (NW Iberian estuaries). *Aquaculture* 287, 278–285.
- Excoffier L., Laval G. and Schneider S.** (2005) Arlequin ver. 3.0: An integrated software package for population genetics data analysis. *Evolutionary Bioinformatics Online* 1, 47–50.
- Felsenstein J.** (2005) *PHYLIP (Phylogeny Inference Package) Version 3.69*. Seattle, WA: Department of Genome Sciences, University of Washington.
- Gardeström J., Pereyra R.T. and André C.** (2008) Characterization of six microsatellite loci in the Baltic blue mussel *Mytilus trossulus* and cross-species amplification in North Sea *Mytilus edulis*. *Conservation Genetics* 9, 1003–1005.
- Giantsis I.A., Kravva N. and Apostolidis A.P.** (2012) Genetic characterization and evaluation of anthropogenic impacts on genetic patterns in cultured and wild populations of mussels (*Mytilus galloprovincialis*) from Greece. *Genetics and Molecular Research* 11, 3814–3823.
- Gosling E.M.** (2003) *Bivalve molluscs. Biology, ecology and culture*. Oxford: Fishing News Books.
- Goudet J.** (2005) PCAGEN 1.2. Available at: <http://www2.unil.ch/popgen/softwares/pcagen.htm> (accessed 31 January 2014).
- Graves J.E.** (1998) Molecular insights into the population structures of cosmopolitan marine fishes. *Journal of Heredity* 89, 427–437.
- Hauffe H.C. and Sbordoni V.** (2009) The synergy between conservation biology and genetics. In Bertorelle G., Bruford M.W., Hauffe H.C., Rizzoli A. and Vernesi C. (eds) *Population genetics for animal conservation*. Cambridge: Cambridge University Press, pp. 3–9.
- Hedrick P.W.** (2000) *Genetics of populations*. 2nd edition. Boston, MA: Jones and Bartlett.
- Hilbish T.J., Carson E.W., Plante J.R., Weaver L.A. and Gilg M.R.** (2002) Distribution of *Mytilus edulis*, *M. galloprovincialis*, and their hybrids in open-coast populations of mussels in southwestern England. *Marine Biology* 140, 137–142.
- Hillis D.M., Moritz C. and Mable B.K.** (1996) *Molecular systematics*. 2nd edition. Sunderland, MA: Sinauer Associates Inc.
- Hoarau G., Boon E., Jongma D.N., Ferber S., Palsson J., Van Der Veer H.W., Rijnsdorp A.D., Stam W.T. and Olsen J.L.** (2005) Low effective population size and evidence for inbreeding in an overexploited flatfish, plaice (*Pleuronectes platessa* L.). *Proceedings of the Royal Society B: Biological Sciences* 272, 497–503.
- Jombart T.** (2008) ADEGENET: an R package for the multivariate analysis of genetic markers. *Bioinformatics* 24, 1403–1405.
- Kalinowski S.T.** (2005) HP-RARE 1.0: a computer program for performing rarefaction on measures of allelic richness. *Molecular Ecology Notes* 5, 187–189.
- Karakousis Y. and Skibinski D.O.F.** (1992) An analysis of allozyme, mitochondrial DNA and morphological variation in mussel (*Mytilus galloprovincialis*) populations from Greece. *Experientia* 48, 878–881.
- Karayucel S., Çelik M.Y., Karayücel I. and Erik G.** (2010) Growth and production of raft cultivated Mediterranean mussel (*Mytilus galloprovincialis* Lamarck, 1819) in Sinop, Black Sea. *Turkish Journal of Fisheries and Aquatic Sciences* 10, 09–17.
- Karayucel S., Çelik M.Y., Karayücel I., Ozturk R. and Eyuboglu B.** (2013) Effects of stocking density on survival, growth and biochemical composition of cultured mussels (*Mytilus galloprovincialis*, Lamarck 1819) from an offshore submerged longline system. *Aquaculture Research*. doi:10.1111/are.12291
- Kijewski T., Smietanka B., Zbawicka M., Gosling E., Hummel H. and Wenne R.** (2011) Distribution of *Mytilus* taxa in European coastal areas as inferred from molecular markers. *Journal of Sea Research* 65, 224–234.
- Koukaras K. and Nikolaidis G.** (2004) *Dinophysis* blooms in Greek coastal waters (Thermaikos Gulf, NW Aegean Sea). *Journal of Plankton Research* 4, 445–457.
- Kravva N., Staikou A. and Triantaphyllidis C.** (2000) Genetic composition and temporal genetic variation in *Mytilus galloprovincialis* populations of the Thermaikos gulf (Northern Aegean Sea). *Biologia* 55, 289–297.
- Kurihara T., Fuseya R., Katoh M. and Inoue K.** (2010) Possibility of self-fertilization during hatchery culturing of giant clam, *Tridacna crocea*. *Plankton and Benthos Research* 5, 11–16.
- Ladoukakis E.D., Saavedra C., Magoulas A. and Zouros E.** (2002) Mitochondrial DNA variation in a species with two mitochondrial genomes: the case of *Mytilus galloprovincialis* from the Atlantic, the Mediterranean and the Black Sea. *Molecular Ecology* 11, 755–769.
- Lallias D., Stockdale R., Boudry P., Lapègue S. and Beaumont A.R.** (2009) Characterization of ten microsatellite loci in the blue mussel *Mytilus edulis*. *Journal of Shellfish Research* 28, 547–551.
- Launey S., Ledu C., Bourdy P., Bonhomme F. and Naciri-Graven Y.** (2002) Geographic structure in the European Flat Oyster (*Ostrea*

- edulis* L.) as revealed by microsatellite polymorphism. *Journal of Heredity* 93, 331–338.
- Lawson Handley L.J., Byrne K., Santucci F., Townsend S., Taylor M., Bruford M.W. and Hewitt G.M.** (2007) Genetic structure of European sheep breeds. *Heredity* 99, 620–631.
- Lemer S., Rochel E. and Planes S.** (2011) Correction method for null alleles in species with variable microsatellite flanking regions, a case study of the black-lipped pearl oyster *Pinctada margaritifera*. *Journal of Heredity* 102, 243–246.
- Lind C.E., Evans B.S., Knauer J., Taylor J.U. and Jerry D.R.** (2009) Decreased diversity and a reduced effective population size in cultured silver-lipped oysters (*Pinctada maxima*). *Aquaculture* 286, 12–19.
- Lougheed S.C., Gibbs H.L., Prior K.A. and Weatherhead P.J.** (2000) A comparison of RAPD versus microsatellite DNA markers in population studies of the *Massasauga rattlesnake*. *Journal of Heredity* 91, 458–463.
- Mantziafou A. and Lascaratos A.** (2004) An eddy resolving numerical study of the general circulation and deep-water formation in the Adriatic Sea. *Deep-Sea Research Part I* 51, 921–952.
- Marin A., Fujimoto T. and Arai K.** (2013) Genetic structure of the Peruvian scallop *Argopecten purpuratus* inferred from mitochondrial and nuclear DNA variation. *Marine Genomics* 9, 1–8.
- Martinez G., Mettifofo L., Perez M.A. and Callejas C.** (2007) A method to eliminate self-fertilization in a simultaneous hermaphrodite scallop. 1. Effects on growth and survival of larvae and juveniles. *Aquaculture* 273, 459–469.
- Meirmans P.G. and Van Tienderen P.H.** (2004) GENOTYPE and GENODIVE: two programs for the analysis of genetic diversity of asexual organisms. *Molecular Ecology Notes* 4, 792–794.
- Myrand B., Tremblay R., Sevigny J.M.** (2009) Decreases in multi-locus heterozygosity in suspension-cultured mussels (*Mytilus edulis*) through loss of the more heterozygous individuals. *Aquaculture* 295, 188–194.
- Nei M.** (1978) Estimation of average heterozygosity and genetic distance from small number of individuals. *Genetics* 89, 583–590.
- Nikula R. and Väinölä R.** (2003) Phylogeography of *Cerastoderma glaucum* (Bivalvia: Cardiidae) across Europe: a major break in the Eastern Mediterranean. *Marine Biology* 143, 339–350.
- Olson D.B., Kourafalou V.H., Johns W.E., Samuels G. and Veneziani M.** (2007) Aegean surface circulation from a Satellite-Tracked Drifter Array. *Journal of Physical Oceanography* 37, 1898–1917.
- Oraic D. and Zrncic S.** (2005) An overview of Health Control in Croatian Aquaculture. *Veterinary Research Communication* 29 (Supplement 2), 139–142.
- Ortiz-Zarragoitia M. and Cajarville M.P.** (2010) Intersex and oocyte atresia in a mussel population from the Biosphere's Reserve of Urdaibai (Bay of Biscay). *Ecotoxicology and Environmental Safety* 73, 693–701.
- Page R.D.M.** (1996) TreeView: an application to display phylogenetic trees on personal computers. *Computer Applications in Biosciences* 12, 357–358.
- Panucci-Papadopoulou M.A., Raitso D.E. and Corsini-Foka M.** (2012) Biological invasions and climatic warming: implications for south-eastern Aegean ecosystem functioning. *Journal of the Marine Biological Association of the United Kingdom* 92, 777–789.
- Parisi G., Centoducati G., Gasco L., Gatta P.P., Moretti V.M., Piccolo G., Roncarati A., Terova G. and Pais A.** (2012) Molluscs and echinoderms aquaculture: biological aspects, current status, technical progress and future perspectives for the most promising species in Italy. *Italian Journal of Animal Science* 11, 397–413.
- Peakall R.O.D. and Smouse P.E.** (2006) GENALEX 6: genetic analysis in Excel. Population genetic software for teaching and research. *Molecular Ecology Notes* 6, 288–295.
- Piry S., Alapetite A., Cornuet J.M., Paetkau D., Baudouin L. and Estoup A.** (2004) GeneClass2: a software for genetic assignment and first-generation migrant detection. *Journal of Heredity* 95, 536–539.
- Pompanon F., Bonin A., Bellemain E. and Taberlet P.** (2005) Genotyping errors: causes, consequences and solutions. *Nature Reviews Genetics* 6, 847–859.
- Presá P., Pérez M. and Diz A.P.** (2002) Polymorphic microsatellite markers for blue mussels (*Mytilus* spp.). *Conservation Genetics* 3, 441–443.
- Radović J., Čivić K., Topić R. and Vukelić VP** (2009) *Biodiversity of Croatia*. 2nd revised edition. Zagreb: State Institute for Nature Protection, Ministry of Culture, Republic of Croatia.
- Raymond M. and Rousset F.** (1995) GENEPOP (version 1.2): population genetics software for exact tests and ecumenism. *Journal of Heredity* 86, 248–249.
- Reynolds J., Weir B.S. and Cockerham C.C.** (1983) Estimation of the coancestry coefficient: basis for a short term genetic distance. *Genetics* 105, 767–779.
- Rhode C., Vervalle J., Bester-van der Merwe A.E. and Roodt-Wilding R.** (2013) Detection of molecular signatures of selection at microsatellite loci in the South African abalone (*Haliotis midae*) using a population genomic approach. *Marine Genomics* 10, 27–36.
- Riginos C. and Cunningham C.W.** (2005) Local adaptation and species segregation in two mussel (*Mytilus edulis* × *Mytilus trossulus*) hybrid zones. *Molecular Ecology* 14, 381–400.
- Schuelke M.** (2000) An economic method for the fluorescent labelling of PCR fragments. *Nature Biotechnology* 18, 233–234.
- Ryman N. and Palm S.** (2006) POWSIM: a computer program for assessing statistical power when testing for genetic differentiation. *Molecular Ecology Notes* 6, 600–602.
- Sanna D., Cossu P., Dedola G.L., Scarpa F., Maltagliati F., Castelli A., Franzoi P., Lai T., Cristo B., Curini-Galletti M., Francalacci P. and Casu M.** (2013) Mitochondrial DNA reveals genetic structuring of *Pinna nobilis* across the Mediterranean Sea. *PLoS ONE* 8, e67372. doi:10.1371/journal.pone.0067372.
- Sekino M., Hara M. and Taniguchi N.** (2002) Loss of microsatellite and mitochondrial DNA variation in hatchery strains of Japanese flounder *Paralichthys olivaceus*. *Aquaculture* 213, 101–122.
- Smietanka B., Zbawicka M., Sanko T., Wenne R. and Burzynski A.** (2013) Molecular population genetics of male and female mitochondrial genomes in subarctic *Mytilus trossulus*. *Marine Biology* 160, 1709–1721.
- Stambuk A., Srut M., Satovic Z., Tkalec M. and Klobucar G.I.V.** (2013) Gene flow vs. pollution pressure: genetic diversity of *Mytilus galloprovincialis* in eastern Adriatic. *Aquatic Toxicology* 136–137, 22–31.
- Teixeira de Sousa J., Joaquim S., Matias D., Ben-Hamadou R. and Leitao A.** (2012) Evidence of non-random chromosome loss in bivalves: differential chromosomal susceptibility in aneuploid metaphases of *Crassostrea angulata* (Ostreidae) and *Ruditapes decussatus* (Veneridae). *Aquaculture* 344, 239–241.
- Theodorou J.A., Viaene J., Sorgeloos P. and Tzovenis I.** (2011) Production and marketing trends of the cultured Mediterranean mussel *Mytilus galloprovincialis* Lamarck 1819, in Greece. *Journal of Shellfish Research* 30, 859–874.
- Toro J.E., Ojeda J.A. and Vergara A.M.** (2004) The genetic structure of *Mytilus chilensis* (Hupe 1854) populations along the Chilean coast based on RAPDs analysis. *Aquaculture Research* 35, 1466–1471.

- Triantafyllidis A., Apostolidis A.P., Katsares V., Kelly E., Mercer J., Hughes M., Jorstad K.E., Tsolou A., Hynes R. and Triantaphyllidis C.** (2005) Mitochondrial DNA variation in the European lobster (*Hommarus gammarus*) throughout the range. *Marine Biology* 146, 223–235.
- Van Oosterhout C., Hutchinson W.F., Willis D.P.M. and Shipley P.F.** (2004) MICROCHECKER: software for identifying and correcting genotyping errors in microsatellite data. *Molecular Ecology Notes* 4, 535–538.
- Varela M.A., Gonzáles-Tizón A., Marinas L. and Martínez-Lage A.** (2007) Genetic divergence detected by ISSR markers and characterization of microsatellite regions in *Mytilus* mussels. *Biochemical Genetics* 45, 565–578.
- Venier P., Tallandini L. and Bisol P.M.** (2003) Characterization of coastal sites by applying genetic and genotoxicity markers in *Mytilus galloprovincialis* and *Tapes philippinarum*. *Journal of Chemical Ecology* 19, 113–128.
- Verhoeven K.J.F., Simonsen K.L. and McIntyre L.M.** (2005) Implementing false discovery rate control: increasing your power. *Oikos* 108, 643–647.
- Vidal R., Penalzoa C., Urzua R. and Toro J.E.** (2009) Screening of ESTs from *Mytilus* for the detection of SSR markers in *Mytilus californianus*. *Molecular Ecology Resources* 9, 1409–1411.
- Wei K., Wood A.R. and Gardner J.P.A.** (2013) Population genetic variation in the New Zealand greenshell mussel: locus-dependent conflicting signals of weak structure and high gene flow balanced against pronounced structure and high self-recruitment. *Marine Biology* 160, 931–949.
- Westfall K.M. and Gardner J.P.A.** (2010) Genetic diversity of southern hemisphere blue mussels (Bivalvia: Mytilidae) and the identification of non-indigenous taxa. *Biological Journal of the Linnean Society* 101, 898–909.
- White T.A., Fotherby H.A., Stephens P.A. and Hoelzel A.R.** (2011) Genetic panmixia and demographic dependence across the North Atlantic in the deep-sea fish, blue hake (*Antimora rostrata*). *Heredity* 106, 690–699.
- and
- Yu H. and Li Q.** (2007) Development of EST-SSRs in the Mediterranean blue mussel, *Mytilus galloprovincialis*. *Molecular Ecology Notes* 7, 1308–1310.

**Correspondence should be addressed to:**

A.P. Apostolidis  
 Laboratory of Ichthyology and Fisheries, Department of  
 Animal Production, Faculty of Agriculture, Aristotle  
 University of Thessaloniki, Thessaloniki, 54124, Greece  
 email: [apaposto@agro.auth.gr](mailto:apaposto@agro.auth.gr)